Average Spectra of Massive Galaxies in the Sloan Digital Sky Survey

Citation

Eisenstein, Daniel J., David W. Hogg, Masataka Fukugita, Osamu Nakamura, Mariangela Bernardi, Douglas P. Finkbeiner, David J. Schlegel, et al. 2003. "Average Spectra of Massive Galaxies in the Sloan Digital Sky Survey." *The Astrophysical Journal* 585 (2) (March 10): 694–713. doi:10.1086/346233.

Published Version

doi:10.1086/346233

Permanent link

<http://nrs.harvard.edu/urn-3:HUL.InstRepos:33462894>

Terms of Use

This article was downloaded from Harvard University's DASH repository, and is made available under the terms and conditions applicable to Other Posted Material, as set forth at <http://nrs.harvard.edu/urn-3:HUL.InstRepos:dash.current.terms-of-use#LAA>

Share Your Story

The Harvard community has made this article openly available.
Please share how this access benefits you. [Submit a story](#).

[Accessibility](#)

AVERAGE SPECTRA OF MASSIVE GALAXIES IN THE SLOAN DIGITAL SKY SURVEY

DANIEL J. EISENSTEIN,^{1,2,3,4} DAVID W. HOGG,^{5,6} MASATAKA FUKUGITA,⁷ OSAMU NAKAMURA,⁷ MARIANGELA BERNARDI,^{2,8,9}
 DOUGLAS P. FINKBEINER,¹⁰ DAVID J. SCHLEGEL,¹⁰ J. BRINKMANN,¹¹ ANDREW J. CONNOLLY,¹² ISTVÁN CSABAI,^{13,14}
 JAMES E. GUNN,¹⁰ ŽELJKO IVEZIĆ,¹⁰ DON Q. LAMB,² JON LOVEDAY,¹⁵ JEFFREY A. MUNN,¹⁶ ROBERT C. NICHOL,⁸
 DONALD P. SCHNEIDER,¹⁷ MICHAEL A. STRAUSS,¹⁰ ALEX SZALAY,¹⁴ AND DON G. YORK²

Received 2002 August 18; accepted 2002 November 20

ABSTRACT

We combine Sloan Digital Sky Survey spectra of 22,000 luminous, red, bulge-dominated galaxies to get high signal-to-noise ratio average spectra in the rest-frame optical and ultraviolet (2600–7000 Å). The average spectra of these massive, quiescent galaxies are early type with weak emission lines and with absorption lines indicating an apparent excess of α -elements over solar abundance ratios. We make average spectra of subsamples selected by luminosity, environment, and redshift. The average spectra are remarkable in their similarity. What variations do exist in the average spectra as a function of luminosity and environment are found to form a nearly one-parameter family in spectrum space. We present a high signal-to-noise ratio spectrum of the variation. We measure the properties of the variation with a modified version of the Lick index system and compare to model spectra from stellar population syntheses. The variation may be a combination of age and chemical abundance differences, but the conservative conclusion is that the quality of the data considerably exceeds the current state of the models.

Subject headings: cosmology: observations — galaxies: abundances — galaxies: clusters: general — galaxies: elliptical and lenticular, cD — galaxies: evolution — methods: statistical

On-line material: color figures, data files

1. INTRODUCTION

Luminous elliptical and bulge-dominated galaxies are the most massive galaxies in the universe. These objects show little rotation, have smooth radial profiles, are massive and kinematically hot, have a narrow range of stellar mass-to-light ratios, show high metallicities, and reside preferentially in the universe's denser environments (e.g., Kormendy & Djorgovski 1989; Roberts & Haynes 1994; Bender & Saglia 1999). These galaxies are also the most intrinsically luminous galaxies in the universe (e.g., Tammann, Yahil, &

Sandage 1979; Blanton et al. 2001). That the most optically luminous galaxies are red is remarkable, given that unextincted red stellar populations generally have higher stellar mass than blue populations at constant luminosity; it means that they are considerably more massive than their bluer, fainter spiral neighbors.

A common paradigm for the formation of bulge-dominated systems is hierarchical merging of smaller, star-forming progenitors (Barnes & Hernquist 1992; Kauffmann 1996; Baugh, Cole, & Frenk 1996). In this model, the systems appear old and metal-rich because the merging is more common at high redshift and because the resulting starburst consumes or expels the remaining gas, thereby ending star formation (Kauffmann & Charlot 1998). This dependence upon cosmological merger rates and feedback suggests that the stellar populations of bulges should display subtle but important dependencies on mass and environment.

In this paper, we study the aggregate stellar populations of luminous, red, bulge-dominated galaxies using averaged spectra from the Sloan Digital Sky Survey (SDSS; York et al 2000). We select these galaxies because we know that they dominate the stellar mass density of the universe (Fukugita, Hogan, & Peebles 1998; Hogg et al. 2002) and because they show great regularities in their properties (e.g., Oke & Sandage 1968; Faber 1973; Visvanathan & Sandage 1977; Djorgovski & Davis 1987; Dressler et al. 1987; Kormendy & Djorgovski 1989; Bower, Lucey, & Ellis 1992; Roberts & Haynes 1994; Bernardi et al. 2003a, 2003b, 2003c, 2003d). In addition, the great age of their stellar populations eases some, but not all, of the difficulties in understanding superposed younger populations (Faber 1972; O'Connell 1976; Gunn, Stryker, & Tinsley 1981; Charlot, Worthey, & Bresnan 1996).

Since the luminosity-weighted average spectrum does measure the total stellar population of an entire sample of galaxies, it is a well-defined mean property that can be

¹ Steward Observatory, University of Arizona, 933 North Cherry Avenue, Tucson, AZ 85121.

² University of Chicago, Astronomy and Astrophysics Center, 5640 South Ellis Avenue, Chicago, IL 60637.

³ Hubble Fellow.

⁴ Alfred P. Sloan Fellow.

⁵ Center for Cosmology and Particle Physics, Department of Physics, New York University, 4 Washington Place, New York, NY 10003.

⁶ Institute for Advanced Study, Princeton, NJ 08540.

⁷ Institute for Cosmic Ray Research, University of Tokyo, Kashiwa 2778582, Japan.

⁸ Department of Physics, Carnegie Mellon University, Pittsburgh, PA 15213.

⁹ Fermilab National Accelerator Laboratory, P.O. Box 500, Batavia, IL 60510.

¹⁰ Princeton University Observatory, Peyton Hall, Princeton, NJ 08544.

¹¹ Apache Point Observatory, P.O. Box 59, Sunspot, NM 88349.

¹² Department of Physics and Astronomy, University of Pittsburgh, Pittsburgh, PA 15217.

¹³ Department of Physics of Complex Systems, Eötvös University, Pázmány Péter sétány 1, H-1518 Budapest, Hungary.

¹⁴ Department of Physics and Astronomy, Johns Hopkins University, 3701 San Martin Drive, Baltimore, MD 21218.

¹⁵ Astronomy Centre, University of Sussex, Falmer, Brighton BN1 9QJ, UK.

¹⁶ United States Naval Observatory, Flagstaff Station, P.O. Box 1149, Flagstaff, AZ 86002.

¹⁷ Department of Astronomy and Astrophysics, Pennsylvania State University, University Park, PA 16802.

compared to theoretical models of galaxy formation (e.g., Baldry et al. 2002). We focus not on the interpretations of the average spectrum itself but rather on the variations of this average with luminosity, environment, and redshift. These variations should be able to test predictions that more luminous (or massive) galaxies have higher metallicity and galaxies in clusters have older stellar populations. Previous observational studies (e.g., Bower et al. 1990; Guzmán et al. 1992; Rose et al. 1994; Tantaló, Chiosi, & Bressan 1998; James & Mobasher 1999; Terlevich et al. 1999; Poggianti et al. 2001) comparing spectra of individual galaxies have found evidence to support these hypotheses. Of course, one would also expect to find that galaxies at high redshift should have younger stellar populations than similar galaxies at low redshift, but even this trend can be altered if massive, bulge-dominated galaxies are still forming at low redshift (Franx 1995; van Dokkum & Franx 2001).

Because we are interested in comparing the average spectra of samples, it is critical that the samples be selected so that the variation between the samples is as controlled as possible. With SDSS data, we can select uniform samples according to rest-frame photometry and morphological properties across a wide range of environments and redshifts, allowing trends in the data to be attributed to astrophysics rather than selection biases.

The average spectra are of extremely high signal-to-noise ratio, thereby allowing analyses that would be impractical on individual SDSS spectra. Over 100 absorption lines are apparent in the composite. Indeed, the resulting spectra exceed the ability of the best spectral models of stellar populations to interpret, especially given the nonsolar ratio of α -element to iron-peak abundances (Worthey, Faber, & González 1992; González 1993; Davies, Sadler, & Peletier 1993; Paquet 1994; Worthey 1998; Proctor & Sansom 2002; Bernardi et al. 2003d) apparent in the data.

The outline of this paper is as follows: In § 2 we describe the relevant aspects of SDSS data and introduce the methods for our analysis, including principal component analysis (PCA) and a modification of the Lick absorption-line index system. In § 3 we describe our sample selection and environment estimation. We present a general analysis of the average spectra in § 4 and analysis of the variations with luminosity and environment in § 5. Variations with redshift are presented in § 6. In § 7 we use the highest redshift sample to determine the average spectrum in the mid-ultraviolet. We conclude in § 8.

Except where noted otherwise, we adopt a conventional cosmological world model with $H_0 = 100 h \text{ km s}^{-1} \text{ Mpc}^{-1}$ and $(\Omega_M, \Omega_\Lambda) = (\frac{1}{3}, \frac{2}{3})$.

2. DATA AND PROCEDURES

2.1. SDSS Imaging

The SDSS is obtaining u , g , r , i , and z -band drift-scan images of the northern Galactic cap and spectra of roughly 10^6 galaxies in that region (York et al. 2000). The photometric system and imaging hardware are described in detail elsewhere (Fukugita et al. 1996; Gunn et al. 1998; Hogg et al. 2001; Smith et al. 2002). An automated image-processing system detects astronomical sources and measures their photometric and astrometric properties (Lupton et al. 2001; Stoughton et al. 2002; Pier et al. 2002; R. H. Lupton et al. 2003, in preparation). Finally, galaxies are selected for

spectroscopy by two algorithms (Strauss et al. 2002; Eisenstein et al. 2001) described further in § 3.

The photometric parameters of interest for this study are as follows: The Petrosian (1976) radius $\theta_{r,\text{Pet}}$ is the angular radius at which the mean surface brightness of the source in the SDSS r -band image inside that radius is 5 times higher than the mean surface brightness in a narrow annulus centered on that radius. The Petrosian flux (with corresponding magnitude m_{Pet}) is the total flux (in any of the five SDSS bandpass images) within a circular aperture of radius $2\theta_{r,\text{Pet}}$ (twice the Petrosian radius in the r -band). The half-light and 90% radii θ_{50} and θ_{90} are the angular radii within which 50% and 90% of the r -band Petrosian flux is found. The concentration c is defined to be the ratio θ_{90}/θ_{50} . The de Vaucouleurs likelihood L_{deV} is the likelihood that the galaxy's r -band two-dimensional image was generated by a true de Vaucouleurs elliptical profile convolved with the seeing. The exponential likelihood L_{exp} is a similar likelihood of an exponential profile. The model magnitude m_{model} is the total magnitude corresponding to the best-fit profile, which is the best-fit de Vaucouleurs profile for all galaxies in this study. All magnitudes are corrected for extinction using the Schlegel, Finkbeiner, & Davis (1998) predictions.

Two samples of galaxies are selected from the SDSS imaging for spectroscopy. The larger sample, with about 88% of the allotment, is a flux-limited sample of galaxies that extends to $r_{\text{Pet}} < 17.77$ (Strauss et al. 2002). We refer to this as the MAIN sample. The smaller portion, with the remaining 12% of the allotment, uses two color-magnitude cuts to seek luminous, early-type galaxies that are fainter than the MAIN limit. This sample is known as the luminous, red galaxy (LRG) sample. The LRG sample is roughly volume-limited to $z \approx 0.4$ and contains additional galaxies to $z \approx 0.55$ with a flux limit of $r_{\text{Pet}} = 19.5$. Further details are found in Eisenstein et al. (2001).

2.2. SDSS Spectroscopy

The SDSS uses two fiber-fed double spectrographs to measure spectra of objects from 3800 to 9200 Å with a resolution of about 1800. Each plug plate holds 640 fibers, yielding 608 spectra of galaxies, quasars, and stars and 32 sky spectra per pointing (Blanton et al. 2003). Fibers cover 3" diameter circular apertures on the sky. Each spectrograph handles 320 fibers (which we call a half-plate), and the two halves are reduced independently.

In detail, the instrumental resolution and pixel scale are close to constant in logarithmic wavelength rather than wavelength. The instrumental resolution is about 170 km s^{-1} FWHM. The “idlSpec2d” software pipeline (D. J. Schlegel et al. 2003, in preparation) combines multiple exposures of a given object and resamples the total spectrum onto a grid of wavelengths that is logarithmically spaced by 69.1 km s^{-1} .

The flux calibration procedure is summarized in Stoughton et al. (2002, §§ 3.3 and 4.10.1) and will be fully detailed in D. J. Schlegel et al. (2003, in preparation). The flux calibration is imposed on each plate by a set of eight spectrophotometric standard stars, chosen by color to be F subdwarf stars. The SDSS does not use an atmospheric refraction corrector, so the effective fiber position on the sky shifts slightly as a function of wavelength. In the presence of brightness gradients, this creates a fluxing error. This is corrected by referring the broadband spectrophotometry to

5" $\times$ 8" aperture "smear" exposures. Since early-type galaxies have small color gradients, we expect that this smear correction will remove the effects of atmospheric refraction.

The spectrophotometric calibration appears to be correct in the average to better than 10%. Since each average spectrum involves galaxies from a moderate range of redshifts, systematic wavelength-dependent spectrophotometry biases in the average spectra will be strongly suppressed on scales below a few hundred angstroms. Comparing our average spectra in different redshift bins (see § 3.2) shows differences of under 10% (4% rms), some of which may be due to actual evolution. Since this involves comparison of spectrophotometry at different observed wavelengths, we infer that the mean spectrophotometry is accurate at this level, up to perhaps an overall tilt. We return to this point in § 5.2. Recent efforts (C. A. Tremonti & D. J. Schlegel 2002, private communication) suggest that the current calibrations are slightly too blue (roughly 10% in flux over the spectral range) with additional residuals blueward of 4300 Å (observed frame).

Redshifts are found by comparing to stellar templates (J. A. Frieman et al. 2003, in preparation; D. J. Schlegel et al. 2003, in preparation). Repeat observations show that the errors in the redshift estimates are less than 30 km s⁻¹ for MAIN sample galaxies and degrade to about 100 km s⁻¹ for the highest redshift LRGs.

2.3. Spectroscopic Sample

We use spectra from 261 separate plates, a total sample of about 120,000 galaxies. Certain plates have been observed more than once, and a small fraction of objects appear on more than one plate (for quality assurance). We use only one spectrum of each object, chosen by the best median signal-to-noise ratio on each candidate half-plate. Only plates that meet survey quality standards are used.

A few spectra were found to contain sufficiently serious reduction failures as to dominate the principal component analysis; we rejected the entire plate in this case (so as not to skew the average by some undetected pattern among the corrupted objects). This eliminated 12 plates, although some of these had other exposures that could have been used.

We select samples of red, bulge-dominated galaxies from these plates; the selection is described in § 3.

2.4. Principal Component Analysis

We are interested in average spectra; however, we choose to go further and perform PCAs on each set of spectra. This isolates a set of eigenspectra ranked by the amount of the total variance in the input set that each explains. PCA is useful for identifying variations in the spectra beyond the mean, and we find that emission-line components are isolated by this technique.

We can recover the average spectra from the PCA decomposition by finding the mean of each PCA coefficient and summing the eigenspectra weighted by these means. Were we to include all the eigenspectra, the resulting linear combination would be exactly the average of the spectra used as input to the PCA. However, we include only the first 20 eigenspectra. We have verified that this procedure makes no difference relative to the true average; the means of the PCA coefficients converge rapidly to zero. One can see that the truncation does not

constitute a smoothing of the spectra by noting that permuting the order of the pixels within the definition of the spectral space would not affect the PCA results.

We begin with spectra in F_λ units, but we use the redshifts to rescale each source to its intrinsic luminosity spectrum L_λ . This does not fully reproduce the mean spectrum of a volume-limited set of galaxies, but in practice the luminosity range of our samples is sufficiently narrow—0.5 mag—that the corrections to a volume-limited mean are negligible. To close approximation, each average spectrum will represent the volume-averaged stellar population of galaxies with the chosen intrinsic properties.

Before performing the PCA step, we precondition the spectra. We shift each spectrum to the rest frame according to its redshift. We do not include shifts smaller than a unit spacing of our wavelength grid (69.1 km s⁻¹). This is equivalent to adding a small additional velocity dispersion to the sample; however, the effect is only 20 km s⁻¹ [69/(12)^{1/2}, the standard deviation of a boxcar], which is small compared to the actual velocity dispersions of the galaxies. We then select a particular range of rest-frame wavelengths, e.g., 3650–7000 Å, to ensure that the full range of wavelengths is covered by SDSS spectra for the full range of redshifts included in a particular sample.

We iteratively interpolate over pixels for which "id1Spec2d" has set warning mask flags. We begin by linear-interpolating over all masked pixels and computing the mean spectrum. We then return to the masked pixels, interpolate with the mean spectrum times a linear function, and recompute the mean. We repeat this step four times to ensure convergence. We increase the masked region around the sky features at 5577, 5890, 5896, 6300, 6364, 7245, and 7283 Å, typically to 10 Å, so as to avoid rare problems of poor sky subtraction. Since the samples involve a wide range of redshift, masking at a fixed observed wavelength enters at different rest-frame wavelengths and therefore has little effect on the average spectrum. We remove entirely any objects that have more than 20 pixels masked for missing data (not counting pixels in the first and last 50 pixels of a given spectrum).

We subtract the continuum of each individual spectrum by removing the first N_F sinusoidal components. We pick $N_F = 8$ for our 3650–7000 Å stacking, $N_F = 6$ for 3500–6000 Å stacking, and $N_F = 5$ for 2600–4400 Å stacking. We have tried other choices with little change. For example, we have tried extracting narrow-wavelength regions with no continuum subtraction and get the same answers.

Because we subtract the continuum from our spectra before performing the PCA, our eigenspectra lack a continuum. We restore this continuum after the PCA step by adding back the average of the subtracted continua. This procedure exactly reproduces the mean spectrum but protects the PCA from being dominated by broadband variations. When comparing spectra from different samples, we rescale to a common overall amplitude but do not otherwise adjust the continua.

Before performing the PCA, we renormalize the different wavelengths by the inverse of the average error in that pixel. This does not affect the average spectrum but does affect subsequent components. In practice, the weighting is mild. After performing the PCA, or more exactly, the diagonalization of the product of the matrix of spectra with its transpose, we restore the original physical normalization of the wavelengths in each eigenspectrum.

In practice, we find that a single component dominates the spectrum of eigenvalues. This eigenspectrum shows a familiar early-type absorption-line spectrum. Because the ratio of this eigenvalue to the others is so large, the first component can deviate only slightly from the mean of the data set.

Regarding different samples, it is important to stress that the entire PCA process has been applied to the two sets of galaxies completely independently. In particular, the comparison of average spectra is not constrained to differ only in a low (20) dimensional subspace of retained eigenspectra.

2.5. Error Estimation

To construct an estimate of the error on the average spectrum, we divide each sample into two, find the average spectrum of each half, and analyze the rms differences between the two halves. Before differencing, we remove 20 low-pass sinusoidal modes (from over 2000 pixels) to be sure that the continuum is not affecting our estimate of the small-scale noise. There is essentially no structure apparent in the differenced spectra. Because there are small correlations between neighboring pixels (e.g., because of the subpixel wavelength shifts between exposures), we smooth the difference with a boxcar filter of 5 pixel width and multiply the result by $\sqrt{5}$. Were the pixels independent, this would not matter, but for correlated pixels this smoothing creates the residual that should apply to multipixel averages such as are found in spectral indexes. We form the square of the smoothed difference spectrum and then smooth with a 100 pixel boxcar. This effectively creates the variance of the differenced spectrum as a function of wavelength.

The galaxies are split into the two halves according to whether their plate number is even or odd. This means that reduction errors that might affect an entire plate are included in the noise estimate, because all of the affected galaxies will go into one half or the other. Conversely, any secular changes in target selection algorithm or imaging properties are isolated from the error analysis. If we had divided the sample into early plates and late plates, small differences in target selection (perhaps simply caused by different photometric calibrations in different parts of the sky) could have resulted in artificial differences between the mean spectra.

If one has two spectra of equal signal-to-noise ratio, then the noise in the difference is twice the noise in the average. Hence, since our even and odd spectra should have quite similar signal-to-noise ratios, we estimate the variance in the average spectra as being one-quarter of that in the smoothed square difference.

Although the noise in a single spectrum can be a complicated and spiky function of wavelength due to the presence of sky lines and detector flaws, the noise in the combined spectrum should be a smooth function of wavelength. Noise features at a narrow observed wavelength range are smeared out by the superposition of many different redshifts, while detector issues such as bad CCD columns are eliminated by the superposition of spectra from many different positions in the detector. Only the broad variations in response across wavelength (e.g., the drop in sensitivity shortward of 4000 Å) should survive to affect the noise in the stacked spectrum.

The error at certain rest-frame wavelengths is dominated not by photon noise but rather by the shot noise of the presence of rare galaxies in the sample. In particular, a small fraction of our selected galaxies have emission lines, and the presence or absence of a few strong emission-line galaxies can affect the results at the line wavelengths. We could treat this by bootstrap resampling of the input samples, but the line emissions are sufficiently weak that we focus only on the isolation and subtraction of the lines and not on the error in that subtraction.

The signal-to-noise ratios of the average spectra are very high, typically a few hundred per pixel (see Table 1). Because our error estimation is based on internal differences, it does not account for errors that are common to the two half-samples. Generally, systematic effects are functions of observed wavelength, which means that their impact will be smeared out by the combination of objects of different redshifts. This suppresses such error modes on scales less than a few hundred angstroms. We have few ways at present to assess the errors on larger scales, but the differencing of samples at very different redshifts does suggest that the errors are small (4% rms). We are not aware of any systematic effects that would be functions of rest-frame wavelength. We present another validation of the quoted errors in § 5.2.

2.6. Comparison between Spectra and to Models

The galaxies in the samples to be defined in § 3 differ in their velocity dispersions. In order to make fair comparisons among the average spectra, we therefore smooth all the average spectra to a common velocity dispersion of $\sigma_v = 325 \text{ km s}^{-1}$, which is the largest velocity dispersion of the individual samples.

In order to interpret our average spectra in terms of age, metallicity, or other variations, we need to compare our spectra to model spectra of stellar population synthesis models. We use the models of Vazdekis (1999, hereafter V99), who provides 2 Å resolution spectra of single-age, single-metallicity, solar-element ratio stellar populations. Because the resolution of these models is better than the resolution achieved in galaxies with large velocity dispersion, we can treat these model spectra identically to our data and thereby do fair comparisons between the two. We use single-power-law initial mass function V99 models throughout.

2.7. *sLick* Indexes

We measure absorption-line strengths by a revision of the most recent version of the “Lick” system (Faber et al. 1985; Worthey et al. 1994; Trager et al. 1998). In addition to the standard Lick indexes, we have included four indexes for measuring the H γ and H δ absorption lines (Worthey & Ottaviani 1997). The Lick system involves measuring line fluxes in narrow wavelength ranges relative to a “pseudo-continuum” found by linearly interpolating between two spectral averages on either side of the lines.

We cannot match the fluxing and calibration of the Lick system but have attempted an intermediate step of matching the wavelength-dependent resolution of the Lick IDS instrument (using a linear interpolation of the table in Worthey & Ottaviani 1997). The highest velocity dispersion of any of our spectra is 325 km s^{-1} at SDSS resolution. This corresponds to 275 km s^{-1} at 6000 Å as

TABLE 1
SUMMARY OF SAMPLE PROPERTIES

Magnitude	Environment/Redshift	Number	Used	S/N	$\langle z \rangle$	$\langle M_r \rangle$	σ Applied	Color Reject
$-22.5 < M_r < -22.0$	All	2568	2394	487	0.193	-22.19	166	72
	$0 \leq N \leq 3$	356	326	159	0.197	-22.17	170	21
	$4 \leq N \leq 9$	822	778	323	0.195	-22.19	156	11
	$N \geq 10$	531	494	260	0.186	-22.22	134	6
$-22.0 < M_r < -21.5$	All	5830	5412	464	0.170	-21.73	181	276
	$0 \leq N \leq 3$	1253	1148	303	0.171	-21.72	198	77
	$4 \leq N \leq 9$	1905	1773	400	0.170	-21.73	192	77
	$N \geq 10$	732	686	256	0.166	-21.75	166	18
$-21.5 < M_r < -21.0$	All	6991	6477	697	0.144	-21.25	200	414
	$0 \leq N \leq 3$	1784	1633	385	0.143	-21.25	223	116
	$4 \leq N \leq 9$	2166	2009	414	0.144	-21.25	210	128
	$N \geq 10$	648	603	264	0.142	-21.25	199	22
$-21.0 < M_r < -20.5$	All	5111	4749	626	0.126	-20.78	231	424
	$0 \leq N \leq 3$	1425	1319	350	0.125	-20.77	237	130
	$4 \leq N \leq 9$	1446	1343	395	0.126	-20.78	223	112
	$N \geq 10$	488	463	205	0.125	-20.77	228	25
$M_g < -21.8$	$0.15 < z < 0.25$	297	276	184	0.210	-21.98 ($\langle M_g \rangle$)	90	
	$0.25 < z < 0.35$	685	641	184	0.305	-21.99 ($\langle M_g \rangle$)	51	
	$0.35 < z < 0.45$	1286	1200	169	0.403	-22.00 ($\langle M_g \rangle$)	71	
	$0.45 < z < 0.50$	762	700	87	0.475	-22.01 ($\langle M_g \rangle$)	139	

NOTES.—The column titled “Number” lists the number of objects that pass our morphology, luminosity, density, and color cuts. The column titled “Color Reject” lists the number that passed the morphology, luminosity, and density cuts but failed the color cut. The column titled “Used” lists the number actually used; some objects fail our spectral quality cuts on the amount of missing or masked data. The column titled “S/N” lists the average signal-to-noise ratio per pixel of spectrum (roughly 1 Å) in the stacked spectrum. The average redshift and absolute magnitude in each sample are given, as is the dispersion of the Gaussian required to smooth the average spectrum to a velocity dispersion of 325 km s⁻¹. The sums of the three environment samples at a given luminosity do not equal the “All” environment sample because we exclude galaxies from the environment sample that fall less than 0.5 h^{-1} Mpc from the edge of the survey imaging.

smoothed to Lick IDS resolution. Hence, when computing indexes, we smooth our spectra to match, at each wavelength, an intrinsic dispersion of 275 km s⁻¹ at the Lick IDS resolution. For example, this corresponds to a total dispersion of 450 km s⁻¹ at 4000 Å. The result is the spectrum of a galaxy with 275 km s⁻¹ dispersion as it would be observed by the Lick IDS instrument. Because, strictly speaking, the Lick system is defined at zero velocity dispersion (at Lick IDS spectral resolution), we refer to the indexes used in this work, normalized to 275 km s⁻¹, as smoothed-Lick or “sLick” indexes.

We introduce the sLick indexes rather than work on the true Lick system because (1) the Lick IDS spectra were referenced to a tungsten blackbody for overall calibration; (2) the indexes are, strictly speaking, calibrated to a defined set of standard stars, which have not been observed with the SDSS telescope; and (3) the Lick system is defined at zero velocity dispersion, at which no galaxies can ever be observed, although some empirical corrections for velocity dispersion have been defined (Trager et al. 1998). For all these reasons, direct comparison of our sLick index measurements (in Tables 2A–2C) to measurements in the Lick system found in the literature must be made with care.

On the other hand, we have treated the V99 model spectra in an identical way to the observed spectra, so the internal consistency of comparing our observed indexes with the models is reliable. For the same reason, our measurements of differences or derivatives of sLick indexes among our average spectra given in Tables 2A–2C are accurate.

3. MASSIVE GALAXY SAMPLES IN THE SDSS

3.1. Morphology and Color Selection in the MAIN Sample

The MAIN sample contains galaxies of all types, but we would like to select massive, red, bulge-dominated galaxies. We do this selection primarily on the basis of morphology but cannot avoid the need to remove color outliers.

We consider four luminosity bins within the MAIN sample of galaxies: $-20.5 < M_r < -21$, $-21 < M_r < -21.5$, $-21.5 < M_r < -22$, and $-22 < M_r < -22.5$. These absolute magnitudes are $k + e$ corrected, dereddened, r -band Petrosian magnitudes. We use the bandpass and evolution ($k + e$) correction of a passively evolving old stellar population (Eisenstein et al. 2001) to adjust the galaxy photometry to $z = 0$. Since the redshift range of each bin is small and since the evolution correction is only 1 mag per unit redshift, the assumed evolution correction is not important for the results but merely serves to compare the galaxies to those at $z = 0$. Note that $M_r^* \approx -20.8$ from the SDSS luminosity function (Blanton et al. 2001), so the least luminous bin has galaxies whose luminosity is roughly L^* , while the most luminous bin has galaxies 4 times more luminous.

We make two purely morphological cuts on the sample: (1) we require that the inverse concentration $1/c$ be less than 0.37, and (2) we require that the de Vaucouleurs model fit be 20 times more likely than the exponential fit (excluding galaxies for which both likelihoods underflow to zero). See § 2.1 for the definitions of these statistics. Both of these cuts favor the traditional surface brightness

FIG. 1.—Rest-frame $g-r$ color vs. absolute r -band magnitude (both corrected for passive evolution) of MAIN galaxies. The effects of the morphological radial profile cuts are shown in the bottom two panels. Galaxies failing the morphology cut are in the top right panel. The bottom right panel shows the color cuts and luminosity bins. The dashed lines show the locus of color and magnitude of an old population with a fading 20% young single-age burst (Bruzual & Charlot 1993 [2001 version]). We stress that the color cut is excluding many fewer galaxies at fixed stellar mass than at fixed luminosity. [See the electronic edition of the *Journal* for a color version of this figure.]

profiles of elliptical galaxies and bulges. We do not perform cuts on the asymmetries or residual lumpiness of the galaxy.

We find that these radial profile cuts do significantly alter the distribution of galaxies in a rest-frame color versus absolute magnitude plot. The red sequence of early-type galaxies is strongly favored. Based on similar samples in Shimasaku et al. (2001), Strateva et al. (2001), and Hogg et al. (2002), we expect that more than two-thirds of the sample are of E/S0 morphology, with the remainder being early-type spiral galaxies.

Figure 1 shows the distribution of galaxies in rest-frame color versus absolute magnitude with and without the morphological cuts. A small fraction of the radial-profile-selected galaxies lie well below the red sequence in $g-r$. If one considers slices at constant r -band luminosity, this fraction of blue galaxies is $\sim 5\%$. It should be stressed that the fraction of blue galaxies in slices of constant stellar mass would be considerably lower. Blue galaxies fade as they age and redden. Hence, one should compare the number of lower luminosity galaxies on the red sequence to the number of higher luminosity blue galaxies. In Figure 1, we display the loci of magnitude and color populated by a burst of 20% new stars on top of an old population, using the 2001 version of the models from Bruzual & Charlot (1993). Following these curves rather than lines of constant luminosity, we would say that very few massive galaxies have been excluded by our color cut. In this model, 20% bursts require about 2 Gyr to redden enough to pass our color cut.

Since we are interested in massive quiescent galaxies, we decide to cut these blue outliers using a color cut. We fit the rest-frame color-magnitude relation to a line of $(g-r)_0 = 0.74 - 0.014(M_r + 20)$ and keep only those galaxies that fall within 0.1 mag in rest-frame color (the region shown in Fig. 1). This excludes 3% of the galaxies in the highest luminosity bin and 8% in the lowest (see Table 1).

Our color selection obviously excludes vigorously star-forming galaxies from the sample. Hence, we cannot claim that the average spectrum that we find represents the average light from *all* bulge-dominated systems. The average age of the sample is biased toward older ages (although of course the dominant problem in the interpretation of optical spectra is the opposite, namely, that a few young stars will bias the inferred age low).

We prefer to compare galaxies at equal redshifts, and therefore we restrict our sample to redshifts between 0.1 and 0.25. However, our lower luminosity subsamples are drawn from a characteristically lower portion of this range, so there may be a small amount of passive evolution mixed into the spectral comparison across luminosity. Histograms of the redshift distribution are shown in Figure 2.

Table 1 lists the sample size, mean redshift, mean luminosity, selection properties, and average spectrum signal-to-noise ratio for the four luminosity samples (see the rows with environment “All”), as well as for the environment and redshift samples discussed in the next two sections.

3.2. Luminous LRG Selection

Eisenstein et al. (2001) show that the LRG spectroscopic sample contains luminous galaxies with red rest-frame colors and early-type spectra out to $z \approx 0.55$. The sample is roughly volume-limited at $z < 0.4$; however, the selection does create an oblique dependence on luminosity and rest-frame color. Intrinsically bluer galaxies must be more luminous than their redder cousins to enter the sample.

Luminous galaxies drawn from the MAIN sample at $0.15 < z < 0.2$, where there is no color selection, show an obvious ridge in color-magnitude space from the prevalence of early-type galaxies. At the highest luminosities, the red sequence is particularly obvious because the frequency of

FIG. 2.—Redshift histograms for the MAIN galaxy luminosity-binned samples.

galaxies drops significantly blueward of the sequence. We would like to select all of the galaxies on the red sequence.

This red sequence is quite narrow, but not infinitely so. Hence, we have the problem that for some range of luminosities, galaxies on the red side of the red sequence would qualify for the LRG sample while those on the blue side would not. To gather a sample for which the entire red sequence will be included, we must use an even higher luminosity cut than the normal LRG selection criteria.

We therefore require $M_g < -21.8$ within the LRG sample. This luminosity refers to the passively evolved, rest-frame g -band luminosity k -corrected from the r -band Petrosian magnitude, as described in Eisenstein et al. (2001). For comparison to our samples from MAIN, we note that this cut corresponds to $M_r \lesssim -22.55$. As shown in Figure 3, this cut allows the LRG sample to include the full red sequence. In addition, the higher luminosity cut means that the sample, which would otherwise be subject to flux limitations at $z \gtrsim 0.40$, retains its volume-limited nature until $z \approx 0.5$.

We limit the LRG selection to $0.15 < z < 0.50$ and divide this into three redshift bins 0.1 in width plus a fourth bin at $0.45 < z < 0.50$. The limit at low redshift is primarily due to the small number of sufficiently luminous galaxies in the small closer volume. Similarly, there are not many galaxies in the LRG sample at $z > 0.50$, and these would not be volume-limited to $M_g = -21.8$.

3.3. Local Density Estimation

We wish to study the spectral properties of our galaxies as a function of their environment. Since our spectra are of luminous galaxies, we do not have full spectroscopic cover-

age of the fainter neighbors. We therefore resort to counting nearby galaxies from the SDSS imaging data. This has the advantage that it is independent of the SDSS spectroscopic coverage, which of course is only partially complete, and that the same luminosity cuts can be applied to all galaxies. To include a poor man's photometric redshift scheme, we restrict the count to galaxies consistent with early-type colors at the appropriate redshift.

Given the spectroscopic redshift of a particular luminous galaxy, we calculate the angular scale corresponding to a transverse distance of $0.5 h^{-1}$ Mpc, the r -band luminosity of an L^* passively evolved early-type galaxy, and the $g-i$ color of that galaxy. The last of these is based on the average color of the two evolving models in Eisenstein et al. (2001). We then count the number of galaxies that fall within 0.15 mag in $g-i$ model color, within 1 mag brightward and 2 mag faintward in r_{Pet} flux, and between 0.03 and $0.5 h^{-1}$ Mpc in transverse distance. Primary galaxies that fall within $0.5 h^{-1}$ Mpc of the edge of the imaging sample are excluded from the environmental studies. This means that the samples for environment variations are smaller than those for luminosity variations (see Table 1).

This count of nearby early-type galaxies gives us an estimate of the local environment of our candidate. However, the statistic is noisy. With random positions inside the survey region, we find an average of 3.6 neighbors, nearly independent of redshift. The average number of neighbors around the actual galaxies increases with luminosity: 5.8 , 5.7 , 6.4 , and 8.95 for luminosity bins of $-21.0 < M_r < -20.5$, $-21.5 < M_r < -21.0$, $-22.0 < M_r < -21.5$, and $-22.5 < M_r < -22.0$, respectively. Hence, while we have clear evidence that galaxies are clustered, the Poisson error on the background is not small relative to the average signal. We cannot make fine distinctions about the local density but instead can construct generously large bins that should be monotonic in their typical density.

For each luminosity bin in the MAIN sample, we divide the sample into three density bins—those with 10 or more neighbors, those with 4 to 9 neighbors, and those with 3 or fewer neighbors—for a total of 12 samples. While the dividing line between the samples will be blurred, with equivalent galaxies jumping the boundary because of background Poisson fluctuations, bins this large should be reasonably distinct as regards the actual (nonprojected) density of the included galaxies.

4. THE AVERAGE SPECTRA

Figure 4 gives an example of the type of average spectrum that we are analyzing. We have used 365 spectra of luminous early-type galaxies in the redshift range $0.30 < z < 0.35$ to produce this stacked spectrum. The wavelength range of 3000 – 6800 Å rest frame is shown; this corresponds roughly to 4000 – 9000 Å in observed wavelength. Careful inspection of the blue and red ends reveals some noise on the pixel scale. In the blue, this is caused by the lower sensitivity of the spectrographs and of course the intrinsic faintness of early-type galaxies in the rest-frame UV. In the red, imperfect sky subtraction limits the performance. In between, the error per pixel is smaller than the thickness of the line. The velocity dispersion of the ensemble creates a visually obvious scale of the width of astrophysical features in the spectrum; comparing that width to the pixel scale revealed by the noise should convince the viewer that

FIG. 3.—Rest-frame color vs. magnitude (both corrected for passive evolution) distribution of all galaxies at $0.15 < z < 0.25$ (top left) and LRGs in different redshift slices (other panels). The horizontal dot-dashed line is a fit to the color-magnitude locus of the red sequence. The full LRG sample would result in a bias toward galaxies on the red side of the red sequence. By imposing an additional magnitude cut (vertical dashed line), we can include the whole red sequence. The sample is volume-limited to $z \approx 0.5$. The diagonal dashed lines show fading bursts as in Fig. 1. [See the electronic edition of the Journal for a color version of this figure.]

FIG. 4.—Average spectrum of 365 luminous, red galaxies, selected as described in § 3.2, in the redshift range $0.3 < z < 0.35$ over the rest-frame wavelength range of 3000–6800 Å. The noise and pixel scale are visible on the blue and red ends as a fine-scale roughness; in between, the noise per pixel is generally smaller than the thickness of the line. The vertical normalization is arbitrary.

essentially all of the features in the spectrum are real. Nearly all of the average spectra that we consider have higher signal-to-noise ratios than this one.

4.1. Absorption-Line Strengths

All of the average spectra in this paper are typical early-type galaxy spectra; they are dominated by the absorption-line features found in the spectra of old stellar populations. This can be made quantitative with comparison to the V99 models via the sLick indexes.

Broadly, the sLick index measurements (given in Tables 2A–2C) are consistent with the conclusion that our sample is composed primarily of an old, metal-rich stellar population. However, there is a large scatter in the inferred ages or metallicities among different sLick indexes; i.e., none of the V99 models is a good match to the total spectrum.

Figure 5 shows this graphically. In each panel, a sLick index based on H δ is plotted against one of various metal-line indexes. The measurements from 16 average spectra, our luminosity and environment samples, are shown as squares. The V99 models are shown as a grid of metallicity and age. Three metallicities (–0.4, 0, and +0.2 dex relative to solar) and six ages (4, 5, 6.3, 8, 10, 12, and 15 Gyr) are shown. We defer discussion of the trends in luminosity and environment until § 5; here our point is simply that no single V99 model matches all indexes for a given spectrum. The α -elements (e.g., Mg *b*) tend to show higher metallicities, while iron-peak elements (e.g., Fe5270) show lower metallicities. This can be explained, at least in part, by a mismatch in chemical abundances. In particular, the Vazdekis models have solar abundance ratios, and it appears that the average spectra show enhanced abundances for α -process elements relative to iron-peak elements. This α -enhancement has been noted previously (Worthey et al. 1992; V99; Davies et al. 2001; Trager et al. 2000a; Proctor & Sansom 2002) and is expected for a stellar population that converts its gas entirely into stars rapidly, before Type Ia supernovae have

time to significantly pollute the interstellar medium with iron-peak elements.

The hydrogen Balmer lines have been widely used to try to disentangle metallicity and age estimation in old stellar systems (e.g., O’Connell 1976; Rose 1985; Worthey et al. 1994; Worthey & Ottaviani 1997; Vazdekis & Arimoto 1999). We study the H β , H γ , and H δ lines. For galaxies with velocity dispersions of 275 km s^{–1}, the Vazdekis models predict that the H β sLick index is the least sensitive to metal content, while the H γ and H δ sLick indexes still carry significant metal dependence (in the sense that more metal-rich systems would be inferred to be younger). However, as we show below, H β has some contamination from interstellar emission lines, and so we focus on the H δ and H γ indexes, where the effect is negligible.

For H δ and H γ , comparing the V99 models at solar metallicity suggests a relatively young age, roughly 5–6 Gyr, for a single-age burst (see the H δ_A versus Fe4383 panel of Fig. 5). All four indexes (H δ_A , H δ_F , H γ_A , and H γ_F) give similar ages. As usual, a higher metallicity implies a younger age. Such age inferences are famously biased by the presence of young A and F stars. We therefore consider histories in which some fraction of the stars are indistinguishably old (we adopt 12.5 Gyr) while the rest are formed at a constant rate between 1 and 12.5 Gyr (we exclude recent formation on the grounds that our color cut would exclude blue star-forming galaxies), assuming solar metallicity throughout (a poor assumption; see Maraston & Thomas 2000). To match H γ and H δ , one would need to form $\sim 50\%$ of the stars in the continuous phase, which gives an average age of ~ 9 Gyr.

H β , on the other hand, implies an old age, above 12 Gyr for the solar-metallicity V99 model. The emission in the average spectrum (§ 4.2) undoubtedly makes this an upper limit. When correcting for nonzero velocity dispersions and line emission by the prescriptions of Trager et al. (1998, 2000a), we find H β , Mg *b*, Fe5270, and Fe5335 indexes that fall comfortably within the $z = 0$ locus found by Trager

TABLE 2A
SLICK INDEXES ON AVERAGE SPECTRA

Magnitude	Environment/Redshift	H δ_A (Å)	H δ_F (Å)	CN ₂ (mag)	Ca4227 (Å)	H γ_A (Å)	H γ_F (Å)	G4300 (Å)
$-22.5 < M_r < -22.0$	All	-1.347 (19)	0.301 (13)	0.0970 (5)	0.860 (10)	-5.508 (17)	-1.371 (11)	5.041 (22)
$-22.0 < M_r < -21.5$	All	-1.089 (14)	0.437 (10)	0.0824 (5)	0.862 (8)	-5.224 (12)	-1.233 (8)	4.920 (18)
$-21.5 < M_r < -21.0$	All	-0.772 (13)	0.575 (10)	0.0690 (5)	0.848 (6)	-4.915 (13)	-1.079 (8)	4.818 (14)
$-21.0 < M_r < -20.5$	All	-0.544 (19)	0.650 (14)	0.0566 (4)	0.831 (8)	-4.729 (17)	-0.989 (12)	4.735 (19)
Slope per mag.....		-0.552	-0.242	+0.0267	+0.021	-0.536	-0.264	+0.203
Error		0.016	0.012	0.0004	0.008	0.015	0.010	0.018
$-22.5 < M_r < -22.0$	$N \geq 10$	-1.649 (37)	0.189 (26)	0.1109 (13)	0.884 (24)	-5.818 (32)	-1.571 (20)	5.163 (39)
	$4 \leq N \leq 9$	-1.259 (23)	0.334 (16)	0.0925 (11)	0.870 (21)	-5.447 (29)	-1.349 (19)	5.011 (39)
	$0 \leq N \leq 3$	-1.082 (47)	0.408 (34)	0.0865 (11)	0.852 (26)	-5.170 (44)	-1.191 (29)	4.916 (57)
	N slope	-0.599	-0.231	+0.0236	+0.033	-0.662	-0.389	+0.256
	Error	0.059	0.042	0.0017	0.036	0.054	0.035	0.068
$-22.0 < M_r < -21.5$	$N \geq 10$	-1.337 (32)	0.329 (23)	0.0944 (11)	0.900 (20)	-5.503 (29)	-1.389 (20)	5.078 (40)
	$4 \leq N \leq 9$	-1.069 (24)	0.424 (18)	0.0809 (7)	0.840 (9)	-5.264 (20)	-1.258 (13)	4.946 (22)
	$0 \leq N \leq 3$	-0.897 (32)	0.527 (23)	0.0733 (11)	0.827 (20)	-4.982 (23)	-1.099 (15)	4.754 (31)
	N slope	-0.440	-0.198	+0.0208	+0.071	-0.525	-0.293	+0.333
	Error	0.046	0.033	0.0016	0.029	0.038	0.025	0.051
$-21.5 < M_r < -21.0$	$N \geq 10$	-0.995 (37)	0.478 (27)	0.0776 (11)	0.871 (17)	-5.176 (29)	-1.181 (18)	4.925 (40)
	$4 \leq N \leq 9$	-0.818 (22)	0.546 (15)	0.0697 (7)	0.840 (11)	-4.962 (21)	-1.101 (14)	4.824 (24)
	$0 \leq N \leq 3$	-0.662 (25)	0.594 (18)	0.0639 (5)	0.806 (9)	-4.813 (20)	-1.029 (14)	4.747 (26)
	N slope	-0.328	-0.112	+0.0132	+0.065	-0.354	-0.152	+0.173
	Error	0.044	0.032	0.0012	0.019	0.035	0.023	0.046
$-21.0 < M_r < -20.5$	$N \geq 10$	-0.887 (51)	0.478 (37)	0.0686 (14)	0.856 (21)	-5.069 (45)	-1.152 (29)	4.910 (54)
	$4 \leq N \leq 9$	-0.585 (28)	0.625 (20)	0.0574 (7)	0.839 (13)	-4.758 (23)	-1.010 (15)	4.750 (28)
	$0 \leq N \leq 3$	-0.429 (33)	0.685 (24)	0.0529 (9)	0.830 (12)	-4.560 (28)	-0.906 (19)	4.688 (31)
	N slope	-0.427	-0.190	+0.0141	+0.024	-0.482	-0.237	+0.194
	Error	0.059	0.043	0.0016	0.024	0.051	0.033	0.058
$M_g < -21.8$	$0.15 < z < 0.25$	-1.508 (47)	0.228 (34)	0.1066 (15)	0.876 (32)	-5.590 (43)	-1.457 (28)	5.095 (49)
	$0.25 < z < 0.35$	-1.315 (41)	0.301 (30)	0.1040 (13)	0.833 (27)	-5.332 (56)	-1.277 (37)	5.030 (75)
	$0.35 < z < 0.45$	-0.864 (56)	0.587 (42)	0.0936 (19)	0.786 (28)	-4.978 (51)	-1.110 (34)	4.940 (77)
	$0.45 < z < 0.50$	-0.588 (81)	0.644 (60)	0.0817 (23)	0.715 (47)	-4.315 (78)	-0.795 (53)	4.558 (106)
	Slope per $\Delta z = 0.1$	+0.319	+0.158	-0.0080	-0.051	+0.378	+0.201	-0.138
	Error	0.027	0.020	0.0008	0.017	0.025	0.017	0.033
Model: 8 Gyr.....	$[Z/H] = 0.0$	-1.849	0.074	0.0353	0.981	-6.272	-1.471	5.338
Model: 12 – 8 Gyr	$[Z/H] = 0.0$	-1.183	-0.480	0.0261	0.216	-1.179	-0.562	0.245
Model: 8 Gyr.....	$[Z/H] = 0.2 - \text{solar}$	-0.941	-0.296	0.0502	0.125	-0.906	-0.362	0.194

NOTES.—Stacked spectra and models (V99) are smoothed until the measured velocity dispersion is $\sigma = 325 \text{ km s}^{-1}$. The smoothed spectra are then further convolved with a Gaussian to achieve the resolution that a $\sigma = 275 \text{ km s}^{-1}$ galaxy would have with the Lick/IDS system. However, unlike the official Lick system, we have kept the spectra in their approximately flux-calibrated form. No attempt to correct the indexes to zero velocity dispersion has been made. The focus is on comparison between the different samples. The 1 σ error in each measurement is given in parentheses.

The top set compares different luminosities within the MAIN set. The middle four sets compare different environments within a given luminosity bin. The final set compares different redshifts within the LRG set. Each set is fitted to a linear regression, and the slope and its error are quoted. The slopes vs. environment are computed from the best-fit line considering the three bins to be given abscissa values of 0, 0.5, and 1 from least dense to most dense. Italic numbers indicate 3σ detections of a nonzero slope.

Model spectra from (V99) are measured in the same fashion. Results for the 8 Gyr, solar metallicity model are quoted. Also quoted are the difference between the 12 and 8 Gyr models, both at solar metallicity, and the $[Z/H] = 0.2$ and solar metallicity models, both at 8 Gyr.

et al. (2000a), although our H β lines are slightly deeper (i.e., younger) than their median value.

Above all, however, we feel that the obvious failings of solar abundance ratio models to simultaneously fit multiple metal indexes caution against any quantitative conclusion on the age of the stellar populations. The Balmer-line indexes are affected by metal absorption, and it may be that neglecting the α -to-Fe abundance ratio differences skews the age inferences or disturbs the agreement between H δ and H β .

4.2. Emission-Line Components

In addition to generating the average spectrum, the PCA algorithm provides us with all of the spectral components that contribute significantly to the diversity (variance) of the spectra in our samples. There is a clear emission-line component that affects roughly 10% of the spectra. We isolate all variations from the “typical” spectrum by seeking outliers in the locus of PCA coefficients. In particular, noting that the PCA coefficients of most galaxies fall in a tight

TABLE 2B
sLICK INDEXES ON AVERAGE SPECTRA

Magnitude	Environment/Redshift	Fe4383 (Å)	Ca4455 (Å)	Fe4531 (Å)	Fe4668 (Å)	H β (Å)	Mg ₂ (mag)	Mg <i>b</i> (Å)
$-22.5 < M_r < -22.0$	All	4.158 (25)	1.017 (12)	2.975 (19)	6.377 (31)	1.583 (13)	0.2570 (2)	3.781 (14)
$-22.0 < M_r < -21.5$	All	4.039 (18)	1.005 (8)	2.959 (11)	6.121 (17)	1.575 (9)	0.2441 (2)	3.684 (12)
$-21.5 < M_r < -21.0$	All	3.959 (18)	0.976 (8)	2.918 (11)	5.774 (20)	1.497 (8)	0.2345 (2)	3.543 (12)
$-21.0 < M_r < -20.5$	All	3.798 (25)	0.933 (7)	2.856 (14)	5.372 (17)	1.472 (7)	0.2265 (2)	3.382 (13)
Slope per mag.....		+0.226	+0.061	+0.085	+0.704	+0.086	+0.0202	+0.270
Error		0.022	0.008	0.014	0.020	0.009	0.0003	0.012
$-22.5 < M_r < -22.0$	$N \geq 10$	4.251 (46)	1.099 (22)	3.002 (40)	6.450 (63)	1.512 (26)	0.2685 (7)	4.006 (28)
	$4 \leq N \leq 9$	4.133 (40)	1.020 (18)	2.985 (31)	6.324 (41)	1.578 (21)	0.2552 (5)	3.770 (22)
	$0 \leq N \leq 3$	4.062 (70)	0.977 (32)	2.980 (53)	6.311 (99)	1.674 (38)	0.2477 (10)	3.678 (42)
	N slope	+0.198	+0.128	+0.024	+0.171	-0.156	+0.0220	+0.358
	Error	0.081	0.038	0.066	0.111	0.045	0.0012	0.049
$-22.0 < M_r < -21.5$	$N \geq 10$	4.119 (35)	1.030 (13)	3.010 (24)	6.284 (46)	1.521 (25)	0.2559 (7)	3.801 (31)
	$4 \leq N \leq 9$	3.988 (29)	0.981 (11)	2.932 (20)	6.072 (34)	1.537 (13)	0.2441 (2)	3.655 (12)
	$0 \leq N \leq 3$	3.976 (44)	0.992 (16)	2.898 (30)	5.952 (39)	1.601 (23)	0.2351 (5)	3.529 (21)
	N slope	+0.158	+0.043	+0.116	+0.325	-0.083	+0.0202	+0.268
	Error	0.056	0.021	0.039	0.061	0.034	0.0008	0.036
$-21.5 < M_r < -21.0$	$N \geq 10$	4.006 (41)	0.986 (18)	2.932 (26)	5.881 (43)	1.542 (25)	0.2411 (8)	3.619 (32)
	$4 \leq N \leq 9$	3.937 (37)	0.959 (14)	2.891 (20)	5.761 (34)	1.529 (16)	0.2368 (5)	3.556 (18)
	$0 \leq N \leq 3$	3.872 (30)	0.956 (15)	2.874 (20)	5.682 (36)	1.493 (12)	0.2313 (4)	3.437 (15)
	N slope	+0.134	+0.029	+0.055	+0.196	+0.054	+0.0101	+0.199
	Error	0.051	0.024	0.033	0.057	0.026	0.0008	0.032
$-21.0 < M_r < -20.5$	$N \geq 10$	3.889 (60)	0.957 (28)	2.903 (42)	5.535 (61)	1.519 (27)	0.2348 (11)	3.490 (51)
	$4 \leq N \leq 9$	3.827 (34)	0.966 (10)	2.879 (25)	5.322 (36)	1.452 (16)	0.2282 (5)	3.420 (22)
	$0 \leq N \leq 3$	3.719 (41)	0.907 (16)	2.834 (23)	5.335 (43)	1.474 (19)	0.2233 (5)	3.313 (22)
	N slope	+0.179	+0.073	+0.074	+0.160	+0.030	+0.0107	+0.192
	Error	0.071	0.029	0.045	0.073	0.033	0.0011	0.048
$M_g < -21.8$	$0.15 < z < 0.25$	4.149 (70)	1.041 (29)	3.016 (49)	6.663 (82)	1.607 (40)	0.2670 (8)	3.934 (34)
	$0.25 < z < 0.35$	4.157 (78)	1.050 (38)	2.952 (51)	6.334 (81)	1.595 (28)	0.2622 (8)	3.863 (35)
	$0.35 < z < 0.45$	4.004 (73)	0.991 (32)	2.891 (42)	6.045 (66)	1.713 (22)	0.2423 (11)	3.608 (43)
	$0.45 < z < 0.50$	3.634 (91)	0.911 (45)	2.802 (77)	5.954 (109)	1.810 (46)	0.2287 (21)	3.418 (86)
	Slope per $\Delta z = 0.1$	-0.152	-0.039	-0.068	-0.259	+0.076	-0.0123	-0.166
	Error	0.036	0.016	0.026	0.040	0.018	0.0006	0.023
Model: 8 Gyr.....	[Z/H] = 0.0	4.373	...	...	...	1.804	0.2189	3.134
Model: 12 – 8 Gyr	[Z/H] = 0.0	0.672	...	...	...	-0.212	0.0314	0.364
Model: 8 Gyr.....	[Z/H] = 0.2 – solar	0.777	...	...	...	-0.007	0.0359	0.381

NOTE.—See Table 2A notes.

distribution near zero (save for the first component, which has a tight distribution with a nonzero mean), we form the sample of galaxies that have all coefficients (excluding the first and as usual any beyond the 20th component) within some tolerance of zero. The numerical value of this threshold is meaningless without discussion of the normalization of the PCA eigenspectra, but visually it imposes a tight envelope around the cloud of points containing the bulk of the galaxies. It is likely that much of the scatter in this cloud is observational scatter. We call the selected galaxies the “median” sample. We then form the average spectrum of this sample and subtract that from the average spectrum of the full sample.

An example of the resulting spectra is shown in Figure 6. Here we have taken our MAIN sample bulge-dominated galaxies in the highest luminosity bin, $-22.5 < M_r < -22$. Only 5% of the galaxies fell in the outlier sample, although this fraction rises at lower luminosities (see Table 3). The

residual spectrum has obvious emission lines, as well as reduced absorption at Na D and Mg *b*. There is a hint of additional absorption in the higher Balmer lines as well as an overall blue tilt. It is worth noting, however, that the differences between the average spectrum of all galaxies and the median sample are small, as one can see by comparing the vertical axes in Figures 6 and 8 (see also Willis et al. 2002). The effects on the sLick indexes are negligible, save for the stronger emission lines (e.g., [O III] and H β).

We do not claim that this residual spectrum contains all of the line emission; clearly, a baseline emission rate would not be included. However, if emission is the exception rather than the rule for these galaxies, then this residual spectrum would contain most of the emission. It is our impression that a few objects dominate the total emission strength. Note that the method will select all the objects with strong H α , [O II], and/or [O III] emission; the weak emission at H γ and H δ need not be detected in individual spectra for it to

TABLE 2C
SLICK INDEXES ON AVERAGE SPECTRA

Magnitude	Environment/Redshift	Fe5270 (Å)	Fe5335 (Å)	Fe5406 (Å)	Fe5709 (Å)	Fe5782 (Å)	Na D (Å)	TiO ₂ (mag)
$-22.5 < M_r < -22.0$	All	2.340 (18)	1.873 (19)	1.246 (11)	0.744 (7)	0.594 (8)	3.687 (13)	0.0745 (2)
$-22.0 < M_r < -21.5$	All	2.380 (11)	1.945 (11)	1.273 (6)	0.765 (5)	0.597 (4)	3.493 (7)	0.0723 (2)
$-21.5 < M_r < -21.0$	All	2.348 (10)	1.919 (8)	1.263 (5)	0.771 (5)	0.587 (3)	3.295 (5)	0.0709 (2)
$-21.0 < M_r < -20.5$	All	2.292 (13)	1.856 (14)	1.228 (9)	0.769 (6)	0.573 (4)	3.096 (7)	0.0701 (1)
Slope per mag.....		+0.050	+0.031	+0.017	-0.014	+0.018	+0.395	+0.0027
Error		0.013	0.014	0.008	0.006	0.005	0.008	0.0001
$-22.5 < M_r < -22.0$	$N \geq 10$	2.361 (25)	1.924 (26)	1.274 (18)	0.755 (15)	0.625 (14)	3.934 (32)	0.0778 (4)
	$4 \leq N \leq 9$	2.341 (25)	1.911 (27)	1.248 (20)	0.757 (9)	0.600 (13)	3.708 (18)	0.0744 (4)
	$0 \leq N \leq 3$	2.374 (37)	1.889 (39)	1.261 (34)	0.733 (15)	0.594 (16)	3.583 (24)	0.0729 (8)
	N slope	-0.005	+0.034	+0.023	+0.022	+0.033	+0.336	+0.0055
	Error	0.043	0.046	0.036	0.022	0.022	0.040	0.0008
$-22.0 < M_r < -21.5$	$N \geq 10$	2.398 (27)	1.950 (28)	1.284 (24)	0.746 (10)	0.588 (11)	3.678 (21)	0.0759 (4)
	$4 \leq N \leq 9$	2.334 (14)	1.913 (15)	1.244 (12)	0.763 (7)	0.593 (8)	3.449 (11)	0.0723 (2)
	$0 \leq N \leq 3$	2.323 (19)	1.930 (26)	1.223 (16)	0.761 (10)	0.583 (10)	3.263 (17)	0.0699 (2)
	N slope	+0.064	+0.017	+0.057	-0.015	+0.006	+0.300	+0.0056
	Error	0.032	0.038	0.028	0.014	0.015	0.027	0.0004
$-21.5 < M_r < -21.0$	$N \geq 10$	2.390 (36)	1.914 (31)	1.269 (21)	0.755 (15)	0.585 (12)	3.373 (16)	0.0728 (4)
	$4 \leq N \leq 9$	2.315 (15)	1.875 (20)	1.235 (15)	0.760 (9)	0.577 (6)	3.283 (10)	0.0709 (2)
	$0 \leq N \leq 3$	2.305 (16)	1.869 (18)	1.217 (13)	0.755 (11)	0.567 (6)	3.251 (12)	0.0701 (2)
	N slope	+0.057	+0.038	+0.048	+0.001	+0.018	+0.114	+0.0026
	Error	0.034	0.035	0.025	0.019	0.013	0.021	0.0005
$-21.0 < M_r < -20.5$	$N \geq 10$	2.318 (46)	1.846 (30)	1.235 (24)	0.751 (20)	0.566 (15)	3.170 (25)	0.0729 (4)
	$4 \leq N \leq 9$	2.315 (21)	1.859 (22)	1.243 (16)	0.773 (11)	0.569 (9)	3.122 (10)	0.0699 (2)
	$0 \leq N \leq 3$	2.226 (28)	1.838 (30)	1.225 (17)	0.758 (14)	0.562 (9)	3.046 (14)	0.0694 (4)
	N slope	+0.117	+0.007	+0.014	-0.001	+0.006	+0.133	+0.0035
	Error	0.051	0.044	0.029	0.025	0.017	0.027	0.0006
$M_g < -21.8$	$0.15 < z < 0.25$	2.386 (31)	1.967 (36)	1.265 (26)	0.739 (24)	0.576 (22)	3.952 (33)	...
	$0.25 < z < 0.35$	2.353 (30)	1.890 (33)	1.254 (29)	0.709 (22)	0.628 (24)	3.872 (40)	...
	$0.35 < z < 0.45$	2.325 (46)	1.834 (49)	1.191 (32)	0.627 (25)	0.555 (26)	3.614 (60)	...
	$0.45 < z < 0.50$	2.127 (105)	1.762 (107)	1.153 (59)	0.764 (94)	0.579 (77)	3.405 (142)	...
	Slope per $\Delta z = 0.1$	-0.048	-0.067	-0.037	-0.044	-0.007	-0.160	...
	Error	0.024	0.026	0.017	0.017	0.016	0.029	...
Model: 8 Gyr.....	$[Z/H] = 0.0$	2.598	1.907	1.295	...	...	...	...
Model: 12 – 8 Gyr	$[Z/H] = 0.0$	0.259	0.220	0.157	...	...	...	...
Model: 8 Gyr.....	$[Z/H] = 0.2 - \text{solar}$	0.335	0.293	0.217	...	...	...	...

NOTE.—See Table 2A notes.

appear in the residual spectrum. Finally, it is important to note that the residual spectrum contains all deviations from the median, but it is not required that these deviations occur along a single “direction” in the space of PCA coefficients. While we detect emission and absorption, it is possible that different outlier galaxies contribute these in different proportions.

In detail, the emission-line spectrum satisfies most of the definitions of a LINER (Heckman 1980; Veilleux & Osterbrock 1987). [N II] $\lambda 6583$, [O I] $\lambda 6300$, and [S II] $\lambda \lambda 6716, 6731$ are strong compared to H α , while [O III] $\lambda 5007$ is weak compared to H β . Of course, it is not straightforward to give a physical interpretation to this average spectrum. It is likely that the galaxies show a diversity of star formation and AGN emission mechanisms, and we do not claim that the average spectrum proves that star formation is negligible. Future work will explore the details of the emission-line behavior of these early-type galaxies.

Lower luminosity bulge-dominated galaxies show stronger emission-line components as well as a more obvious Balmer absorption component. This is discussed in § 5.3.

5. VARIATIONS WITH LUMINOSITY AND ENVIRONMENT

We now proceed to our primary task, the analysis of the average spectra of luminous bulge-dominated galaxies as a function of luminosity and environment. Recall that, as described in §3, we break each of our four luminosity bins into three environment bins. We also analyze each of the four luminosity bins in its entirety.

There is a striking level of similarity among the average spectra of these subsamples. Dividing by redshift, luminosity, and environment always yields the same qualitative spectrum. However, we do find small but highly significant differences among the average spectra. In broad terms, the

FIG. 5.—Each panel shows a plot of two sLick indexes, with the $H\delta_A$ index plotted against various metal indexes. Points show the values measured from different samples. Squares indicate the environment subsamples. Black squares are the $N > 10$, high-density sample. White squares are the $N \leq 3$, low-density sample. Gray squares are the $4 \leq N \leq 9$, intermediate-density sample. The size of the square indicates the luminosity of the sample, with the largest square being $-22.5 < M_r < -22$, and the smallest square being $-21 < M_r < -20.5$. Circles indicate the LRG sample divided by redshift, with the darker color being higher redshift. Error bars in both indexes are shown but are often smaller than the points. The line grid shows the predictions from the V99 model, as described in § 2.6. Three metallicities are shown (-0.4 , 0 , and $+0.2$ dex relative to solar) and six different ages (4, 5, 6.3, 8, 10, 12, and 15 Gyr). The thicker line marks the 10 Gyr locus. The letters YS, YR, and OP indicate the “young, solar,” “young, metal-rich,” and “old, metal-poor” boundaries of the grid. [See the electronic edition of the *Journal* for a color version of this figure.]

FIG. 6.—Difference spectrum of outliers from the locus of luminous bulge-dominated galaxies in the luminosity range $-22.5 < M_r < -22$. The 5% of the galaxies whose PCA coefficients fall outside of the primary distribution of points are labeled as outliers. This spectrum is the change to the average spectrum from including or excluding these objects. The resulting spectrum shows a number of emission lines, with relative strengths characteristic of a LINER, as well as some alterations to some absorption lines, including the higher order Balmer lines.

trends are similar to those found in previous work (e.g., Dressler et al. 1987; Bower et al. 1990; Guzmán et al. 1992; Bender, Burstein, & Faber 1993; Rose et al. 1994; Trager et al. 2000b; Bernardi et al. 2003d).

5.1. Variations in the sLick Indexes

Figures 5 and 7 show the spectral variations as a function of luminosity and environment as measured by the sLick indexes. Lower luminosity galaxies or those in poorer environments have indexes that would be interpreted as younger or more metal-poor. In some cases, notably the iron indexes, the variations would be interpreted primarily as age. However, we are concerned that the poor match between the data and the models in other indexes (e.g., Mg *b*) may cast doubt on such detailed interpretation.

The quantitative values of the sLick indexes are presented in Tables 2A–2C, along with their errors and the V99 model predictions (where available). We also present the slopes of linear regressions versus luminosity, environment, and redshift within the sets of comparison samples. Slopes that are significant at 3σ are shown in *italic*.

5.2. A One-dimensional Spectral Space

A striking feature of Figures 5 and 7 is that the locus of values found in the luminosity and environment samples are nearly one-dimensional. In other words, with two different parameters at play, one might have expected the sLick indexes to populate two-dimensional regions in these plots. Instead, the points fall (primarily) on tight curves, even for pairs of indexes for which the models predict a separation of the effects of age and metallicity. This suggests that luminosity and environment control the same parameter in the average spectra. The H β index is an exception, which we discuss below.

TABLE 3
EMISSION AND ABSORPTION LINE FLUXES

MAGNITUDE	ENVIRONMENT	F_{em}	EMISSION								ABSORPTION			
			H α	H β	H γ	H δ	[S II]	[N II]	[O III]	[O II]	H α	H β	H γ	H δ
$-22.5 < M_r < -22$	All	0.05	0.31	0.073	0.028	0.017	0.21	0.40	0.17	0.28	...	0.005	0.065	0.074
	$N \geq 10$	0.07	0.17	0.053	0.017	0.010	0.15	0.32	0.11	0.19	...	0.005	0.019	0.012
	$4 \leq N \leq 9$	0.07	0.34	0.086	0.033	0.017	0.22	0.40	0.11	0.27	...	0.052	0.098	0.106
	$0 \leq N \leq 3$	0.06	0.22	0.067	0.024	0.010	0.16	0.35	0.35	0.25	...	...	0.039	0.293
$-22 < M_r < -21.5$	All	0.08	0.64	0.133	0.047	0.026	0.33	0.62	0.33	0.36	...	0.022	0.137	0.128
	$N \geq 10$	0.07	0.37	0.073	0.021	0.015	0.20	0.30	0.13	0.20	...	0.039	0.120	0.117
	$4 \leq N \leq 9$	0.07	0.62	0.124	0.049	0.031	0.33	0.57	0.39	0.39	...	0.009	0.133	0.140
	$0 \leq N \leq 3$	0.08	0.65	0.125	0.048	0.019	0.30	0.58	0.26	0.32	0.061	0.032	0.187	0.139
$-21.5 < M_r < -21$	All	0.10	1.02	0.207	0.073	0.035	0.42	0.86	0.41	0.41	...	0.052	0.239	0.225
	$N \geq 10$	0.11	0.73	0.143	0.060	0.030	0.33	0.66	0.33	0.39	...	...	0.158	0.153
	$4 \leq N \leq 9$	0.10	0.82	0.172	0.062	0.022	0.35	0.68	0.27	0.35	...	0.045	0.254	0.198
	$0 \leq N \leq 3$	0.11	1.10	0.216	0.078	0.037	0.43	0.92	0.43	0.41	...	0.069	0.255	0.246
$-21 < M_r < -20.5$	All	0.12	1.34	0.289	0.106	0.052	0.53	1.03	0.58	0.51	...	0.086	0.350	0.334
	$N \geq 10$	0.10	0.85	0.185	0.078	0.046	0.43	0.79	0.51	0.49	...	0.009	0.180	0.173
	$4 \leq N \leq 9$	0.22	1.26	0.270	0.096	0.046	0.47	0.96	0.40	0.42	...	0.114	0.328	0.294
	$0 \leq N \leq 3$	0.20	1.26	0.264	0.097	0.045	0.50	0.93	0.47	0.46	...	0.076	0.343	0.325
Continuum F_λ ($\text{\AA}^{-1}$)			1.20	1.14	0.92	0.75	1.15	1.2	1.10	0.46	1.20	1.14	0.92	0.75

NOTES.—Emission-line strengths are computed from the residual spectrum formed by the difference of the average spectrum and the “median” spectrum. The quantity F_{em} is the fraction of spectra that were declared as outliers and therefore omitted from the “median” spectrum. The higher-order Balmer lines are also seen to be in absorption; we fitted a broad absorption Gaussian and a narrow emission Gaussian. The results are quoted as line fluxes in arbitrary units; the F_λ of the continuum is also given (in units of flux per angstrom). Equivalent widths can be found by dividing the line flux by the continuum F_λ . [S II] is the sum of the 6731 and 6716 Å lines, [N II] is the sum of the 6548 and 6583 Å lines, [O III] is only the 5007 Å line, and [O II] is the sum of the 3726 and 3729 Å lines.

FIG. 7.—As in Fig. 5, each panel shows a plot of two sLick indexes. The top two panels show the H β index against the H δ_A and Mg b indexes. These show that H β does not follow a simple one-dimensional locus; it is unclear whether this is the result of emission-line contamination. The lower four panels show the H δ_A index against the H γ_F index and three metal-line indexes plotted against the H δ_A index. The grids of V99 models are absent in the three panels where the index wavelength falls outside the wavelength range of the models. [See the electronic edition of the *Journal* for a color version of this figure.]

To investigate this regularity, we took the 16 average spectra (12 from the luminosity/environment bins and four from the pure luminosity bins), subtracted the mean, and performed a PCA analysis on the residuals. No points were masked and no continua were subtracted, but as usual the spectra were smoothed to 325 km s^{-1} before performing the PCA. Eighty-nine percent of the variance in the 16 spectra was contained in the first PCA component; i.e., removing this component drops the rms variation from 1% to 0.3%. The mean spectrum and the first component of variation are shown in Figure 8. The normalization of the first variation component relative to the mean is chosen so that the rms amplitude of the variation is unity. In other words, the relative normalization in Figure 8 is the typical spread in the 16 spectra, which is 1% rms of the mean.

The first component of variation is clearly a mix of emission and absorption lines. Emission at $H\alpha$, $H\beta$, $[O \text{ II}]$, $[O \text{ III}]$, $[N \text{ II}]$, and $[S \text{ II}]$ is visible. Higher order Balmer lines are seen in absorption, and the strong Mg and Na absorption lines at 5175 and 5900 Å are varying in strength. The variation in the Mg line is slightly displaced to the red relative to the deepest part of the line in the mean spectrum; this indicates that blending of the lines can be important. Finally, there is a broadband tilt to the spectrum, in the sense that variations in the direction of extra nebular emission and stronger Balmer absorption correlate with a bluer spectrum, as one might expect.

The behavior at $H\delta$ and $H\gamma$ is of concern: although the variation spectrum appears to have extra absorption at these frequencies, the absorption troughs are wider than

FIG. 8.—Average of the average spectra from all of the environmental and luminosity samples (note that this double-counts some galaxies), along with the first component of variation between these 16 sets. This component accounts for 89% of the variance in the 16 spectra. The normalization of the first component of variation relative to the mean spectrum is set by the rms amplitude of this component. Higher components have very small amplitudes (~ 0.002) on this scale.

FIG. 9.—Amplitude of the first component of variation relative to the mean (these two spectra are shown in Fig. 8) as a function of luminosity for three different environment bins. The normalization of the spectral component is chosen so that the rms value of this amplitude is unity; as shown in Fig. 8, this corresponds to $\sim 1\%$ variations in the spectra.

expected (compare the line shapes to those of the higher order Balmer lines near 3800 Å). In other words, we worry that other nearby metal lines may be impacting the H δ and H γ sLick indexes.

The second component of variation in the PCA analysis is smaller by a factor of 4. It is primarily broadband fluctuations, with a small amount of H α emission.

Figure 9 shows the amplitudes of the first component required by our 16 different samples. As expected, lower luminosity galaxies and those in poorer environments have more emission and Balmer absorption.

We stress that the information in Figure 8 goes well beyond the sLick indexes. We have found the empirical primary direction of variation with respect to luminosity and environment for the ensemble of luminous, early-type galaxies over most of the optical range. While we are limited in the resolution available from hot stellar systems, there are hundreds of spectral elements measured to high signal-to-noise ratio in this spectrum. Matching these variations would be a stringent test for stellar population synthesis. We have not attempted more detailed comparisons to the models because it is clear that the available solar element abundance ratio models are a poor match to the mean spectrum.

As an aside, the fact that the first component of spectral variation contains 89% of the variance between the 16 spectra while being only 1% variation rms is a striking confirmation of the high signal-to-noise ratio of the average spectra. If these spectra had a signal-to-noise ratio of only 100 pixel $^{-1}$, then the PCA eigenspectra corresponding to this noise would have been comparable to the astrophysical variation, so that the first component would have explained only $\sim 1/16$ of the variance. Hence, the signal-to-noise ratio of the spectra must considerably exceed 100 for errors that are not common to all the spectra. This is particularly remarkable given that no broadband modes were removed prior to this PCA analysis; indeed, much of the first component's variation is broadband, since it contains only 0.5% rms of variation when four broadband modes are removed.

5.3. Variations in Emission-Line Components

In § 4.2 we isolated the spectral variations of outlier galaxies by using the PCA coefficients to identify the outliers and then considering the difference between the average spectrum of all the galaxies and the average spectrum of those in the median sample. We now consider these residual spectra as functions of luminosity and environment. We generally find that galaxies of lower luminosity and poorer environment have stronger emission-line components.

In the case of H γ and H δ , we clearly see small emission lines sitting in deeper absorption lines (Fig. 6, *inset*), indicating that the residual spectrum contains both Balmer absorption from young stars and Balmer emission from young stars and active nuclei. We stress again that this superposition is a property of the composite of outliers and may reflect multiple populations of outliers from the “typical” old stellar spectrum.

To quantify the strength of these lines, we fit separate Gaussians to the emission and absorption lines, assuming that all emission lines in the spectrum have the same velocity width. The resulting line fluxes are shown in Table 3. Converting these to equivalent widths shows that the contribution to the H γ and H δ indexes is very small: a 15% effect on the inferred age in the worst case. We therefore infer that the H γ and H δ indexes are negligibly affected by emission in the full sample.

H β , on the other hand, does seem to have significant emission, roughly enough to cancel out what must be non-negligible absorption. The two contributions are more difficult to disentangle, and it seems possible that the age inferred from H β would be significantly affected. The fact that H δ and H γ are well correlated with metal-line indexes while H β has a lot of scatter may also be suggestive of emission in H β . However, it might indicate that the higher lines have significant contamination from metallicity variations.

H α has even stronger emission, and the presence of strong [N II] makes it impossible to detect an absorption component. The line ratio of [N II] to H α displays an increasing trend with luminosity and density, which may indicate a decreasing role of star formation in driving the emission (Veilleux & Osterbrock 1987).

In general, the Balmer decrement is larger than expected from conventional predictions of the recombination spectrum. This may be caused by dust obscuration of the emission regions.

5.4. Selection Biases

The interpretation of the spectral differences is subject to considerations of biases in the sample selection and analysis. Fortunately, the comparison between samples from different environments is substantially immune to redshift-dependent effects, such as variations in the sample selection or change in physical spectroscopic aperture across redshift, because the redshift distributions of the three environment bins are nearly identical. Target selection and all aspects of the observations were done without consideration of the environment of the galaxies. The only significant effect is that galaxies in clusters tend to be more luminous. However, the 0.5 mag bin size in luminosity limits the shift in the mean luminosity between environment subsamples to less than 2% in the lower three luminosity samples and 4% in the most luminous sample (Table 1).

The comparison across luminosity, on the other hand, may include some biases in sample selection. First, there is a small shift in redshift (listed in Table 1 and shown in Fig. 2) between the luminosity bins. If all samples were passively evolving, the most luminous galaxies would appear ~ 1 Gyr younger than the least luminous relative to a fixed redshift comparison. Second, galaxies of different luminosities have different physical sizes, so that the spectroscopic fiber will include a different fraction of, e.g., the effective radius of the de Vaucouleurs model. Radial population gradients will therefore bias our results (McClure 1969; Munn 1992 and references therein). This effect is partially offset by the change in angular diameter distance with redshift. Finally, and most important, the lower luminosity bins almost certainly have more contamination from objects that would not be classified by eye as red, elliptical galaxies. Our morphology cut is relatively crude and does not include obvious mainstays of quantitative morphology, such as bulge-disk decomposition and smoothness. It is not clear that equivalent samples are being selected at all luminosities or even that equivalent samples can be selected in principle. The difficulty of homogenizing stellar disk contributions could also bias the environmental subsamples because of the density-morphology relation.

5.5. Interpretation of Absorption Lines

Figures 5, 7, and 8 and Tables 2A–2C all demonstrate that there are statistically significant differences between the luminosity and environment subsamples. Tables 2A–2C display the slopes and errors resulting from least-squares fits to each sLick index across luminosity and environment. Indexes with more than 3σ detections of variation are indicated by asterisks. In particular, the $H\delta$, CN_2 , $H\gamma$, G4300, C₂4668, Mg *b*, Na D, and TiO₂ indexes all show strong variations.

In principle, such well-detected variations should be able not only to determine age and mean metallicity shifts but also to probe for variations in abundance ratios. Trippico & Bell (1995, hereafter TB95) calculate the variation in Lick indexes induced by changes in the abundances of individual elements for three specific model stars. This grid of abundances and stars is too sparse to interpret the observed differences in our data. Nevertheless, if one adopts the overly simple model in which the TB95 cool giant and hot dwarf models are combined in arbitrary ratio with abundance variations allowed in groups of C and N, O, α -elements, and iron-peak elements and then resolves our observed index variations against these five vectors of model variations, then one finds that the variations among our spectra are best described by variations in the α -element abundance. Of course, age variations are available to this model only insofar as they could be mapped to changes in the giant-to-dwarf ratio. As an additional caveat (were there not enough already!), we note that the TB95 Lick indexes are quoted at zero velocity dispersion, while our variations are at 275 km s^{-1} .

The indication that abundance ratio variations may be the controlling parameter in the spectral variation with luminosity and environment is intriguing. We believe that this hint should offer strong encouragement for the construction of spectral models of stellar populations with non-solar abundance ratios (e.g., Thomas, Maraston, & Bender 2002).

6. AVERAGE SPECTRUM VERSUS COSMIC TIME

Whereas our comparison of MAIN samples was all at low redshift, our LRG sample spans a large range in redshift, $0.15 < z < 0.50$, with galaxies that are similar in luminosity and rest-frame color. We can therefore look for the evolution of the average spectrum across cosmic time. We divide the galaxies into four separate redshift bins: $0.15 < z < 0.25$, $0.25 < z < 0.35$, $0.35 < z < 0.45$, and $0.45 < z < 0.50$.

While the selected galaxies are luminous early types at all redshifts, there are subtle redshift dependences in the selection. For example, the fine details of the rest-frame luminosity-color region imposed at each redshift by the LRG selection have mild redshift dependences (Fig. 3). At this point, we have not assessed these because of uncertainties in the appropriate *k*-corrections and evolution corrections as well as the need to model the effect of the *noise* in the observed colors and magnitudes on the sample selection. Detailed modeling of these selection effects is now under way.

To first order, the galaxies in each bin should be similar, in the sense that a passively evolving old galaxy would be in all the samples. However, the physical aperture of our $3''$ diameter spectroscopic fiber is increasing with redshift, so that we are including more of the light of the higher redshift galaxies. If galaxies have older or more metal-rich stellar populations near their centers, then the higher redshift objects will be biased to look younger or more metal-poor. Of course, the size of such effects is bounded by the weakness of color gradients in giant elliptical galaxies (Franx, Illingworth, & Heckman 1989; Peletier et al. 1990; Michard 2000) and by the fact that we observe the luminosity-weighted spectrum within the physical aperture.

The sLick indexes measured for these average spectra are given in Tables 2A–2C and plotted as circles in Figures 5 and 7. The results show a clear dependence on redshift. Were this difference due only to age, the higher redshift galaxies would be younger by ~ 3 Gyr, with some variation with index. This is consistent with the concordance cosmology prediction of ~ 2.5 Gyr. In the figures, one should note that the results do not in all cases lie along the locus defined by the low-redshift luminosity and environment variations. If we view the redshift sample as defining the true time derivative of a metal-rich passively evolving population, then the luminosity and environment dependences may require abundance variations.

We next consider the variations of the full spectrum across redshift. We take the four average spectra, subtract the mean, and compute the least-squares slope for each pixel as a function of the central redshift of each spectrum. The vector of slopes is the spectrum of ordered variations within this set of four spectra. The highest redshift average spectra have noise roughly twice as large as the others, so we down-weight them by a factor of 4 in the least-squares fitting. We normalize the slope to a difference of 0.1 in redshift. The PCA analysis of § 5.2 yields a similar variation spectrum to this least-squares method.

The resulting mean and variation spectra are shown in Figure 10. The variation spectrum shows more noise than the equivalent spectrum in the environment/luminosity comparison. This is partially because the highest redshift average spectra are noisier, but it is also because the LRG sample spectra required less smoothing to match to

FIG. 10.—*Top panel of each pair:* Average of the average spectra from all of the redshift samples. *Bottom panel of each pair:* Least-squares variation per 0.1 in redshift. The sign of the variation is such that one should add the variation spectrum to the mean when considering higher redshifts. The first component variation spectrum from Fig. 8 is shown for comparison (*dotted line*). The relative normalization of the two variation spectra is arbitrary. The broadband offsets of the redshift variation spectrum should not be trusted because of spectrophotometry uncertainties; comparing fixed rest-frame wavelengths across redshift requires accurate spectrophotometry across a significant range of observed wavelengths, whereas the comparisons across environment and luminosity were less demanding of this. The redshift variation spectrum has more small-scale noise in part because the more luminous galaxies required less smoothing to reach a velocity dispersion of 325 km s^{-1} , thereby including more photon noise. Simulation of the noise indicates that, for example, the three largest features between 4225 and 4325 Å are significant at roughly 5σ confidence, while the features between 4150 and 4200 Å are only marginally significant. [See the electronic edition of the Journal for a color version of this figure.]

325 km s^{-1} than the MAIN sample spectra. Despite this, many features in the variation spectrum are well detected. We assess the significance by creating realizations of the noise of the original spectra, smoothing, and applying the least-squares analysis. The smaller features in the variation spectrum are not significant, but features such as the three peaks between 4225 and 4325 Å are clearly significant. Further SDSS data should make interesting improvements in the signal-to-noise ratio of these spectra.

The variation spectrum shows strong alterations to absorption features at H δ , H γ , Mg b , and Na D, as well as

weaker changes to other lines. Mg b and Na D are shallower at higher redshift, while the two Balmer lines are deeper. Little hint of emission lines is found, save for [O II], probably in part because the galaxies in the LRG sample are considerably more luminous than those in the MAIN sample. Overlaying the component of variation from the environment and luminosity analysis shows matching trends in the strongest features but imperfect agreement in the amplitudes of variations in different lines. Hence, the two variation spectra (in this section and § 5.2) represent two different combinations of age and metallicity; one might

FIG. 11.—Average UV spectrum of 726 luminous, red galaxies at $0.47 < z < 0.55$. The inset shows an expansion of the bluest region of the spectrum. The 2900 and 2640 Å spectral breaks are clearly visible, as are the interstellar Mg lines at 2796 and 2852 Å. The vertical normalization is arbitrary.

speculate that the redshift variation is dominantly due to age.

7. ULTRAVIOLET AVERAGE SPECTRUM

In Figure 11 we present the composite UV spectrum from our highest redshift galaxies, 726 luminous, red galaxies at $0.47 < z < 0.55$. Many absorption lines are detected, although the signal-to-noise ratio is poor at the bluest wavelengths because of the dropping efficiency of the SDSS spectrographs shortward of 4000 Å. The 2900 and 2640 Å spectral breaks are clearly visible, as are the strong Mg lines at 2796 and 2852 Å. Using the definitions of Spinrad et al. (1997), we measure the strengths of these breaks to be $B(2640) = 1.33 \pm 0.06$ and $B(2900) = 1.41 \pm 0.04$.

The strengths of the UV breaks suggest that the UV spectrum has contributions from F-type stars (Spinrad et al. 1997; Fioc & Rocca-Volmerange 1997). Such stars can be present in the spectrum because either (1) there has been some star formation within the previous few Gyr or (2) there is enough metallicity range among stars in the composite spectrum that there are old, F-type stars, either dwarfs or giants. Both explanations are plausible; we expect some ongoing star formation even among these old stellar populations, and we expect a wide metallicity range.

8. SUMMARY AND CONCLUSIONS

We have defined samples of luminous, red, bulge-dominated galaxies from the SDSS that represent controlled variations in luminosity, environment, and redshift. We have focused entirely on luminous galaxies, roughly L^* and above. The samples are a nearly complete representation of the most massive galaxies in the local universe.

We have computed luminosity-weighted average spectra so as to probe the aggregate stellar populations of these galaxies. The spectra indicate fairly old, metal-rich stellar

populations that are enhanced in α -elements relative to iron-peak elements. Our primary goal has been to compare the relative differences in the populations as a function of luminosity, environment, and redshift. We find that the spectral variations across luminosity and environment are nearly a one-dimensional set. With a velocity-broadened form of the Lick indexes, most index-index comparisons track a one-dimensional locus despite having two independent parameters (luminosity and environment). This is true even for pairs of indexes that are predicted to split the familiar age-metallicity degeneracy. We find some indication that the controlling parameter of the variation may be the α -to-Fe abundance ratio. Our results include not simply a vector of sLick index offsets, but a full-resolution spectrum of the variation that includes over a hundred clearly visible lines.

We have compared the average spectra directly to models from V99, analyzing these model spectra identically to our data, and find that the variations are plausible even if the mean values of the indexes may be shifted. More luminous galaxies, lower redshift galaxies, or those in denser environments would be interpreted as more metal-rich or older. However, we stress that the obviously nonsolar α -to-iron-peak ratio in our data probably invalidates any detailed interpretation of the differences with the solar-abundance ratio models of V99.

We hope that the detail available in these spectra and their derivatives with luminosity, environment, and redshift will be a stimulus to detailed, full-resolution modeling of α -enhanced, supersolar stellar populations. The spectra are available as a gzipped tar file in the electronic edition of the *Journal*.

We thank Jim Peebles, Scott Trager, and Ann Zabludoff for useful conversations and the referee for helpful comments. D. J. E. thanks New York University for hospitality while a portion of this research was completed. D. J. E. was supported by National Science Foundation (NSF) grant AST 00-98577, by Hubble Fellowship grant HF-01118.01-99A from the Space Telescope Science Institute, which is operated by the Association of Universities for Research in Astronomy, Inc., under NASA contract NAS5-26555, and by an Alfred P. Sloan Research Fellowship. D. W. H. was partially supported by NSF grant 0101738 and NASA contract NAG5-11669. This research made use of the NASA Astrophysics Data System.

Funding for the creation and distribution of the SDSS archive has been provided by the Alfred P. Sloan Foundation, the Participating Institutions, NASA, NSF, the US Department of Energy, the Japanese Monbukagakusho, and the Max Planck Society. The SDSS Web site is <http://www.sdss.org>.

The SDSS is managed by the Astrophysical Research Consortium for the Participating Institutions. The Participating Institutions are the University of Chicago, Fermilab, the Institute for Advanced Study, the Japan Participation Group, The Johns Hopkins University, Los Alamos National Laboratory, the Max-Planck-Institute for Astronomy, the Max-Planck-Institute for Astrophysics, New Mexico State University, the University of Pittsburgh, Princeton University, the United States Naval Observatory, and the University of Washington.

REFERENCES

- Baldry, I. K., et al. 2002, *ApJ*, 569, 582
 Barnes, J. E., & Hernquist, L. 1992, *ARA&A*, 30, 705
 Baugh, C., Cole, S., & Frenk, C. S. 1996, *MNRAS*, 283, 1361
 Bender, R., Burstein, D., & Faber, S. M. 1993, *ApJ*, 411, 153
 Bender, R., & Saglia, R. P. 1999, in *ASP Conf. Ser.* 182, *Galaxy Dynamics*, ed. D. R. Merritt, M. Valluri, & J. A. Sellwood (San Francisco: ASP), 113
 Bernardi, M., et al. 2003a, *AJ*, in press
 ———. 2003b, *AJ*, in press
 ———. 2003c, *AJ*, in press
 ———. 2003d, *AJ*, in press
 Blanton, M. R., Lupton, R. H., Maley, F. M., Young, N., Zehavi, I., & Loveday, J. 2003, *AJ*, in press
 Blanton, M. R., et al. 2001, *AJ*, 121, 2358
 Bower, R. G., Ellis, R. S., Rose, J. A., & Sharples, R. M. 1990, *AJ*, 99, 530
 Bower, R. G., Lucey, J. R., & Ellis, R. S. 1992, *MNRAS*, 254, 601
 Bruzual A., G., & Charlot, S. 1993, *ApJ*, 405, 538
 Charlot, S., Worthey, G., & Bressan, A. 1996, *ApJ*, 457, 625
 Davies, R. L., Sadler, E. M., & Peletier, R. F. 1993, *MNRAS*, 262, 650
 Davies, R. L., et al. 2001, *ApJ*, 548, L33
 Djorgovski, S., & Davis, M. 1987, *ApJ*, 313, 59
 Dressler, A., Lynden-Bell, D., Burstein, D., Davies, R. L., Faber, S. M., Terlevich, R. J., & Wegner, G. 1987, *ApJ*, 313, 42
 Eisenstein, D. J., et al. 2001, *AJ*, 122, 2267
 Faber, S. M. 1972, *A&A*, 20, 361
 ———. 1973, *ApJ*, 179, 731
 Faber, S. M., Friel, E. D., Burstein, D., & Gaskell, C. M. 1985, *ApJS*, 57, 711
 Fioc, M., & Rocca-Volmerange, B. 1997, *A&A*, 326, 950
 Franx, M. 1995, in *IAU Symp.* 164, *Stellar Populations*, ed. P. C. van der Kruit & G. Gilmore (Dordrecht: Kluwer), 269
 Franx, M., Illingworth, G., & Heckman, T. 1989, *AJ*, 98, 538
 Fukugita, M., Hogan, C. J., & Peebles, P. J. E. 1998, *ApJ*, 503, 518
 Fukugita, M., Ichikawa, T., Gunn, J. E., Doi, M., Shimasaku, K., & Schneider, D. P. 1996, *AJ*, 111, 1748
 González, J. 1993, Ph.D. thesis, Univ. California, Santa Cruz
 Gunn, J. E., Stryker, L. L., & Tinsley, B. M. 1981, *ApJ*, 249, 48
 Gunn, J. E., et al. 1998, *AJ*, 116, 3040
 Guzmán, R., Lucey, J. R., Carter, D., & Terlevich, R. J. 1992, *MNRAS*, 257, 187
 Heckman, T. M. 1980, *A&A*, 87, 152
 Hogg, D. W., Finkbeiner, D. P., Schlegel, D. J., & Gunn, J. E. 2001, *AJ*, 122, 2129
 Hogg, D. W., et al. 2002, *AJ*, 124, 646
 James, P. A., & Mobasher, B. 1999, *MNRAS*, 306, 199
 Kauffmann, G. 1996, *MNRAS*, 281, 487
 Kauffmann, G., & Charlot, S. 1998, *MNRAS*, 294, 705
 Kormendy, J., & Djorgovski, S. 1989, *ARA&A*, 27, 235
 Lupton, R., Gunn, J. E., Ivezić, Z., Knapp, G. R., Kent, S., & Yasuda, N. 2001, in *ASP Conf. Ser.* 238, *Astronomical Data Analysis Software and Systems X*, ed. F. R. Harnden, Jr., F. A. Primini, & H. E. Payne (San Francisco: ASP), 269
 Maraston, C., & Thomas, D. 2000, *ApJ*, 541, 126
 McClure, R. D. 1969, *AJ*, 74, 50
 Michard, R. 2000, *A&A*, 360, 85
 Munn, J. A. 1992, *ApJ*, 399, 444
 O'Connell, R. W. 1976, *ApJ*, 206, 370
 Oke, J. B., & Sandage, A. 1968, *ApJ*, 154, 21
 Paquet, A. 1994, Ph.D. thesis, Univ. Heidelberg
 Peletier, R. F., Davies, R. L., Illingworth, G. D., Daview, L. E., & Cawson, M. 1990, *AJ*, 100, 1091
 Petrosian, V. 1976, *ApJ*, 209, L1
 Pier, J. R., et al. 2002, *AJ*, submitted
 Poggianti, B. M., et al. 2001, *ApJ*, 562, 689
 Proctor, R. N., & Sansom, A. E. 2002, *MNRAS*, 333, 517
 Roberts, M. S., & Haynes, M. P. 1994, *ARA&A*, 32, 115
 Rose, J. A. 1985, *AJ*, 90, 1927
 Rose, J. A., Bower, R. G., Caldwell, N., Ellis, R. S., Sharples, R. M., & Teague, P. 1994, *AJ*, 108, 2054
 Schlegel, D. J., Finkbeiner, D. P., & Davis, M. 1998, *ApJ*, 500, 525
 Shimasaku, K., et al. 2001, *AJ*, 122, 1238
 Smith, J. A., et al. 2002, *AJ*, 123, 2121
 Spinrad, H., Dey, A., Stern, D., Dunlop, J., Peacock, J., Jimenez, R., & Windhorst, R. 1997, *ApJ*, 484, 581
 Stoughton, C., et al. 2002, *AJ*, 123, 485
 Strateva, I., et al. 2001, *AJ*, 122, 1861
 Strauss, M. A., et al. 2002, *AJ*, 124, 1810
 Tammann, G. A., Yahil, A., & Sandage, A. 1979, *ApJ*, 234, 775
 Tantalo, R., Chiosi, C., & Bressan, A. 1998, *A&A*, 333, 419
 Terlevich, A. I., Kuntschner, H., Bower, R. G., Caldwell, N., & Sharples, R. M. 1999, *MNRAS*, 310, 445
 Thomas, D., Maraston, C., & Bender, R. 2002, preprint (astro-ph/0202166)
 Trager, S. C., Faber, S. M., Worthey, G., & González, J. J. 2000a, *AJ*, 119, 1645
 ———. 2000b, *AJ*, 120, 165
 Trager, S. C., Worthey, G., Faber, S. M., Burstein, D., & González, J. J. 1998, *ApJS*, 116, 1
 Trippico, M. J., & Bell, R. A. 1995, *AJ*, 110, 3035 (TB95)
 van Dokkum, P. G., & Franx, M. 2001, *ApJ*, 553, 90
 Vazdekis, A. 1999, *ApJ*, 513, 224 (V99)
 Vazdekis, A., & Arimoto, N. 1999, *ApJ*, 525, 144
 Vveilleux, S., & Osterbrock, D. E. 1987, *ApJS*, 63, 295
 Visvanathan, N., & Sandage, A. 1977, *ApJ*, 216, 214
 Willis, J. P., Hewett, P. C., Warren, S. J., & Lewis, G. F. 2002, *MNRAS*, 337, 953
 Worthey, G. 1998, *PASP*, 110, 888
 Worthey, G., Faber, S. M., & González, J. J. 1992, *ApJ*, 398, 69
 Worthey, G., Faber, S. M., González, J. J., & Burstein, D. 1994, *ApJS*, 94, 687
 Worthey, G., & Ottaviani, D. L. 1997, *ApJS*, 111, 377
 York, D. G., et al. 2000, *AJ*, 120, 1579