Surveillance of medication use: early identification of poor adherence

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ABSTRACT

Background We sought to measure population-level adherence to antihyperlipidemics, antihypertensives, and oral hypoglycemics, and to develop a model for early identification of subjects at high risk of long-term poor adherence.

Methods Prescription-filling data for 2 million subjects derived from a payor’s insurance claims were used to evaluate adherence to three chronic drugs over 1 year. We relied on patterns of prescription fills, including the length of gaps in medication possession, to measure adherence among subjects and to build models for predicting poor long-term adherence.

Results All prescription fills for a specific drug were sequenced chronologically into drug eras. 61.3% to 66.5% of the prescription patterns contained medication gaps >30 days during the first year of drug use. These interrupted drug eras include long-term discontinuations, where the subject never again filled a prescription for any drug in that category in the dataset, which represent 23.7% to 29.1% of all drug eras. Among the prescription-filling patterns without large medication gaps, 0.8% to 1.3% exhibited long-term poor adherence. Our models identified these subjects as early as 60 days after the first prescription fill, with an area under the curve (AUC) of 0.81. Model performance improved as the predictions were made at later time-points, with AUC values increasing to 0.93 at the 120-day time-point.

Conclusions Dispensed medication histories (widely available in real time) are useful for alerting providers about poorly adherent patients and those who will be non-adherent several months later. Efforts to use these data in point of care and decision support facilitating patient care are warranted.

INTRODUCTION

Although medications may be the single most important healthcare intervention for chronic disease in the developed world,1 2 their impact is modulated by patient adherence. Even for conditions where using a single medication just once a day is known to improve outcomes, only about 50% of patients adhere to the correct regimen.3–5 Studies of patients taking antihypertensives, antidepressives, hypoglycemics, antipsychotics, and antihyperlipidemics show that poor adherence is predictive of poor outcomes and high costs, including future hospitalizations and treatment discontinuation.6–18 Conversely, consistent adherence to statins, antihypertensives, antidiabetic agents, and antiviral therapies is associated with positive health outcomes.3 16 17 19

Causes of non-adherence are multifactorial, including tolerability, side effects, complexity of prescribing and filling procedures, level of understanding of the importance of taking the medication, cultural norms, out of pocket cost of the medication, and the possible lack of discernable effects of the medication.1 20–24 Healthcare providers tend to overestimate medication adherence.25–30 Surveillance of adherence can be achieved by monitoring the records of dispensed prescription medications.31 32 Prescription-filling data provide an indirect measure of a patient’s adherence to prescription medication. Although filled prescription medications are not always taken, it is generally true that patients who do not fill a prescription are non-adherent. Early identification of patients at high risk of poor adherence would enable early interventions by clinicians that could improve health outcomes.30 33–35 Attempts to predict poor adherence using only patient characteristics (such as socio-economic status and comorbidities) have found those variables to be weak predictors.24 36 However, these studies did not consider the predictive potential of early prescription-filling patterns. We sought to measure the adherence of patients filling prescriptions for three classes of chronic drugs: antihyperlipidemics, antihypertensives, and oral hypoglycemics. Furthermore, we sought to develop models for early identification of patients who will likely be long-term poor adherers to these classes of drugs.

METHODS

This study is a retrospective analysis of temporally sequenced dispensed medication data. We use a two-phase prediction model for early identification of patients at risk of poor adherence to chronic medications. The first phase is a straightforward approach to identifying prolonged gaps in prescription-filling patterns. The second phase uses a regression model to isolate more complex and subtle patterns predictive of future poor adherence.

Data

The subjects for this study were identified from more than 8.5 million beneficiaries of a large private insurance plan covered during the period of January 1999 to December 2006, who were taking at least one of three classes of drugs: antihyperlipidemics, antihypertensives, and oral hypoglycemics. The three drug classes were selected because for these chronic medications, improved health outcomes require consistent adherence over time. The individual drugs included for each category are given in online supplementary appendix table A1. The prescription-filling data included the date, National...
Drug Code (NDC), quantity dispensed (in days), and information about whether the prescription was filled through a mail order or retail mechanism in addition to other administrative information. The age and gender of the subjects were also available, as well as information about hospitalizations and dates of membership of the insurance plan.

For each subject, all instances of prescription fills for a specific drug were sequenced chronologically into drug eras. Each drug era represented the prescription-filling history of a particular subject for a specific drug during the time period included in the dataset.

**Study period and cohort construction**

Drug eras coinciding with a gap in membership of the insurance program during the first year of prescriptions were excluded (figure 1). Left and right censoring was performed by excluding the drug eras beginning within the first 6 months of the study period, as well as those beginning in the last 15 months of the study period. Drug eras coinciding with hospitalizations during the first year of prescriptions were excluded. A final exclusion was performed to remove drug eras with unrealistic or unusual data including ages greater than 200 years, quantities greater than 200 supply days, and unlisted genders.

**Adherence analysis**

The drug eras were first evaluated based on medication gaps during the first year of prescription fills (figure 2). Those containing at least one medication gap greater than 50 days were flagged as poor adherers. This did not include drug eras where the subject switched to another drug in the same category within 30 days of running out of their last prescription. A switch to a new drug must have been initiated no sooner than 90 days before the last prescription fill.

The drug eras that were flagged by the large-gap detector were further subdivided into three categories: (1) those that switched to another drug in the same category within 90 days of the flagging date; (2) long-term discontinuers who never filled another prescription for that particular drug in this dataset; and (3) short-term discontinuers who filled at least one more prescription for the same drug after the flagging date.

Drug eras that were neither flagged by the large-gap detector nor switched to another drug in the same category (‘No gap, no switch’ category in figure 2) exhibited consistent filling of prescriptions for at least 1 year. For this group of drug eras, we developed models to identify those exhibiting poor adherence by consistently filling prescriptions late.

**Model development and evaluation**

The outcome variable was adherence as measured by the medication possession ratio (MPR), a standard measure of possession of filled prescription medication over time. The study period considered was the first year of the drug era, measured from the first observed prescription fill. An MPR of 0.80 is the accepted standard indicator of good adherence. The goal of our modeling was to identify the subjects who did not meet this criterion of good adherence at the 1-year time point (‘Poor adherence’ category in figure 2). Only drug eras that did not contain medication gaps greater than 50 days during the first year were considered (‘No gap, no switch’ category in figure 2).

The drug eras were used to build and test a logistic regression model for predicting the outcome variable using two independent measures: age and the latest calculated MPR value. Gender and drug names were also initially considered as variables but were found not to significantly contribute to the models, and were excluded. Three models were developed to make predictions at three time points early in the drug era (days 60, 90, and 120). Each of the models was trained with one third of the drug eras to identify those that would not meet the criterion for good adherence 1 year after the first filled prescription. To demonstrate the value of the fill data patterns, comparison models for each drug category were also developed for performance comparison using only the gender and age of the subject. The models were tested on the remaining two thirds of the dataset and evaluated using the area under the receiver operator characteristic curve (AUC). Additional metrics for performance included the specificity, accuracy, positive predictive value, and negative predictive value, with the sensitivity of the models set to 90.0%.

**RESULTS**

Over 2 million subjects were identified (table 1). After the exclusion criteria were applied, the antihypertensives class had the most subjects (1,294,521) followed by antihyperlipidemics (790,883) and oral hypoglycemics (278,029). The antihypertensives class contained more than four times as many drugs (74) as the two other classes (17 and 16 for antihyperlipidemics and oral hypoglycemics, respectively) (online supplementary appendix table A1). In all three classes, between 50% and 60%
of subjects filled prescriptions for more than one drug in the same class, resulting in 1.4 to 1.9 times as many drug eras as unique subjects in each class. The final cohort had slightly more females than males in all three classes, and the average age on the date of the first prescription was between 62 and 67 years (table 1).

A two-phase adherence monitor was applied. First, a large-gap detector, identifying periods of more than 30 days between expected fills, was applied. For antihyperlipidemics, antihypertensives, and oral hypoglycemics, respectively, 66.5%, 64.5%, and 61.3% of drug eras contained medication gaps greater than 30 days. These include long-term discontinuations, where the subject never again filled a prescription for that drug in the dataset, which represent 29.0%, 29.1%, and 23.7% of all the drug eras and 59.5%, 68.6%, and 51.8% of the drug eras containing large medication gaps. Among the long-term discontinuations, 50.4%, 50.3%, and 36.6% had only one prescription fill. On average, long-term discontinuation drug eras had 1.7, 1.6, and 2.2 prescription fills, and 71.3, 59.6, and 59.5 dispensed supply days. The drug eras defined as short-term discontinuations—containing large gaps followed by prescription fills (short-term discontinuations)—had average gap lengths of 105.6, 120.0, and 107.6 days for antihyperlipidemics, antihypertensives, and oral hypoglycemic, respectively.

After applying the large-gap detector, we applied an algorithm to detect non-adherence defined as an MPR of less than 0.80 at 1 year. The MPR, a standard measure of adherence, is defined as the days supply of medication divided by the days between refills. Among drug eras not containing large gaps between fills (‘No gap, no switch’) category in figure 2), those which had MPRs less than 0.80 at the 1-year outcome date (‘Poor adherence’ category in figure 2) met our definition of long-term poor adherence. For antihyperlipidemics, antihypertensives, and oral hypoglycemic, respectively, these drug eras represented 1.0%, 0.8%, and 1.3% of all the drug eras, and 3.1%, 2.6%, and 3.5% of the drug eras in the ‘No gap, no switch’ category. The distributions of the 1-year MPR values for the drug eras in this group are similar for the three drug categories (figure 3). These drug eras were used to build logistic regression models for predicting long-term poor adherence. Models were built for making predictions at three early time-points in the drug era: 60, 90, and 120 days after the first prescription fill.

The logistic regression models are in the form:

\[
\text{risk of poor adherence} = \frac{1}{1 + e^{\beta_0 + \beta_1 x_1 + \beta_2 x_2}}
\]

where \( x_1 \) is the latest available MPR value and \( x_2 \) is the age of the subject at the beginning of the drug era. The intercepts (\( \beta_0 \)) and coefficients (\( \beta_1 \) and \( \beta_2 \)) for the nine models are given in table 2. The performance of the prediction models developed for each drug category and each time point to detect drug eras with poor adherence is shown in figure 4.

The models performed similarly...
for the three drug classes, and performance improved across the three time points of prediction when starting at a later point. For antihyperlipidemics, antihypertensives, and oral hypoglycemics, respectively, for predictions made at 120 days after the first fill, compared with 60 days, the areas under the curve (AUC) improved from 0.85, 0.81, and 0.85 to 0.95, 0.92, and 0.95 (figure 4).

With sensitivity held constant at 90%, table 3 shows the performance statistics for each model, which like the AUC, improved when making predictions at later time points. Specificity, accuracy, positive predictive value, and negative predictive value all increased from below 50% to above 80% between days 60 and 120. The negative predictive value remained constant at 99% for all three drug categories and the three time points.

**DISCUSSION**

A high rate of non-adherence is evident in claims data. Similar data are available in real time from pharmacy benefit managers (PBMs) and pharmacies, suggesting a ready approach to real time adherence surveillance. Beyond simply detecting overt non-adherence, we have also shown that poor adherence is predictable very early based on patterns of medication fills.

Among the patients who continuously fill their prescriptions over the course of at least 1 year, the small percentage of long-term poor adherers can be identified as early as 60 days after the first prescription fill. Poor adherers in this class are harder to catch than those with significant medication gaps, and identifying them early in the treatment could greatly reduce the risks associated with long-term poor adherence.

**Table 2 Coefficients for the logistical regression models where the risk of poor adherence is calculated using the latest available MPR value and the age of the subject at the beginning of the drug era**

<table>
<thead>
<tr>
<th></th>
<th>Day 60 model</th>
<th>Day 90 model</th>
<th>Day 120 model</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Antihyperlipidemics</strong></td>
<td></td>
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<td></td>
</tr>
<tr>
<td>$\beta_0$ (intercept)</td>
<td>7.195</td>
<td>11.791</td>
<td>16.462</td>
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<td>$\beta_1$ (MPR)</td>
<td>-8.850</td>
<td>-14.481</td>
<td>-20.350</td>
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<tr>
<td>$\beta_2$ (age)</td>
<td>-0.039</td>
<td>-0.034</td>
<td>-0.031</td>
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<tr>
<td><strong>Antihypertensives</strong></td>
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<td></td>
<td></td>
</tr>
<tr>
<td>$\beta_0$ (intercept)</td>
<td>5.911</td>
<td>10.430</td>
<td>14.397</td>
</tr>
<tr>
<td>$\beta_1$ (MPR)</td>
<td>-8.140</td>
<td>-13.420</td>
<td>-18.260</td>
</tr>
<tr>
<td>$\beta_2$ (age)</td>
<td>-0.031</td>
<td>-0.029</td>
<td>-0.027</td>
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<tr>
<td><strong>Oral hypoglycemics</strong></td>
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<td></td>
</tr>
<tr>
<td>$\beta_0$ (intercept)</td>
<td>6.406</td>
<td>10.759</td>
<td>14.524</td>
</tr>
<tr>
<td>$\beta_1$ (MPR)</td>
<td>-8.438</td>
<td>-13.595</td>
<td>-18.260</td>
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<tr>
<td>$\beta_2$ (age)</td>
<td>-0.031</td>
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MPR, medication possession ratio.

Our analysis and modeling were conducted on claims data that included both mail order and retail mechanisms for acquiring medications. The main difference between the two mechanisms is that mail order prescriptions were generally for 90-day supplies, while retail prescriptions were for fewer supply days, most commonly 30 days. The mechanism for mail order still required that the beneficiary actively request each new medication course. We expect that in contrast, mail order mechanisms requiring minimal or no action on the part of the subject would make prescription-filling data less predictive of future adherence to chronic medications, potentially disguising non-adherence.

The MPR metric used in this model, although widely accepted, is still only an indirect metric of adherence to taking medication, as patients may fill the prescription but not take the medication. It is certainly reasonable to assume that in most cases, if the subject did not fill the prescriptions, then the subject was not taking sufficient medication. However, it is possible that the subjects were receiving supplies of the medication through methods not captured in the data available for this study, such as out of pocket payments, particularly when the copay would be higher than the advertised rate. An additional concern is pill splitting, which would be observed as poor adherence in the dataset, but might well be a physician-approved regimen. Samples provided by the physician at visit intervals may also have disrupted the perceived fill cycle. Further, although this could be corrected by including e-prescribing data in the model, without it, the method misses patients who are prescribed a medication and never fill.

The definition of poor adherence as an MPR value of less than 80%, although widely used in studies of adherence measure, is arbitrary and may not be appropriate in all cases. In the context of this study, the threshold of 80% adherence labeled a very small portion of the overall population as poor adherers. Although a single threshold was used in this study to define poor adherence, similar models can be built and evaluated with different threshold values.

The models presented here incorporate information widely available through pharmacies, pharmacy benefit managers, and insurers. Bringing these data and their interpretation to the point of care presents an opportunity for a clinician to quantify a patient’s adherence to a specific medication regime and to intervene where appropriate. In practice, an automated system that uses both the large-gap detector and the regression model could identify the vast majority of patients at risk of poor outcomes due to poor adherence. By running the algorithm on a prescription-filling dataset on a daily basis, subjects would be identified as soon as they met the models’ criteria for alert, rather than waiting until a future interaction with a clinician.
Once a patient is identified as being at risk, a clinician could evaluate whether the subject is indeed non-adherent and initiate an interaction with the subject to attempt to improve adherence.

Dispensed medication data for chronic medications could be useful in alerting providers about patients currently poorly adherent as well as those who will be non-adherent several months afterwards. Even moderate non-adherence early in the treatment course can be an indicator of future poor adherence. However, there are practical considerations around integrating an automated adherence screening tool into clinical practice including integrating a point of care application into clinician workflow and devising approaches to notifying physicians without inducing ‘alert fatigue.’ False alerts could be reduced if e-prescribing were incorporated into the algorithm, eliminating detection of intended medication discontinuations.

Studies of interventions have shown mixed results in improving adherence, and new approaches are being developed that good adherers are systematically more health-seeking than non-adherers,

which suggests that identifying patients at high risk of poor adherence may also help identify patients in need of intervention to improve not only adherence to chronic medications, but also other aspects of health.

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Competing interests None.

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REFERENCES