Consumption of Dairy Products and Colorectal Cancer in the European Prospective Investigation into Cancer and Nutrition (EPIC)

The Harvard community has made this article openly available. Please share how this access benefits you. Your story matters.

Citation

Published Version
doi:10.1371/journal.pone.0072715

Citable link
http://nrs.harvard.edu/urn-3:HUL.InstRepos:11877029

Terms of Use
This article was downloaded from Harvard University’s DASH repository, and is made available under the terms and conditions applicable to Other Posted Material, as set forth at http://nrs.harvard.edu/urn-3:HUL.InstRepos:dash.current.terms-of-use#LAA
Consumption of Dairy Products and Colorectal Cancer in the European Prospective Investigation into Cancer and Nutrition (EPIC)

Neil Murphy1, Teresa Norat14, Pietro Ferrari2, Mazda Jenab2, Bas Bueno-de-Mesquita3,4, Guri Skeie5, Anja Olsen6, Anne Tjønneland6, Christina C Dahm7,8, Kim Overvad8, Marie Christine Boutron-Ruault9,10, Françoise Clavel-Chapelon9,10, Laura Nailler9,10, Rudolf Kaaks11, Birgit Teucher11, Heiner Boeing12, Manuela M Bergmann12, Antonia Trichopoulou13,14, Pagona Lagiou13,15,16, Dimitrios Trichopoulos14,15,16, Domenico Palli17, Valeria Pala18, Rosario Tumino19, Paolo Vineis1,20, Salvatore Panico21, Petra H. M. Peeters22, Vincent K. Dik4, Elisabete Weiderpass2, Eiliiv Lund2, Jose Ramon Quiros Garcia23, Raul Zamora-Ros24, Maria José Sánchez Pérez25,26, Miren Dorronsoro27, Carmen Navarro26,28,29, Eva Ardanaz26,30, Jonas Manjer31, Martin Almquist32, Ingegerd Johansson33, Richard Palmqvist34, Kay-Tee Khaw35, Nick Wareham36, Timothy J. Key37, Francesca L. Crowe37, Veronika Fedirko2, Marc J. Gunter1, Elio Riboli1

1 Department of Epidemiology and Biostatistics, School of Public Health, Imperial College London, London, United Kingdom, 2 International Agency for Research on Cancer, Lyon, France, 3 The National Institute for Public Health and the Environment, Bilthoven, The Netherlands, 4 Department of Gastroenterology and Hepatology, University Medical Centre, Utrecht, The Netherlands, 5 Institute of Community Medicine, University of Tromsø, Tromsø, Norway, 6 Danish Cancer Society, Institute of Cancer Epidemiology, Copenhagen, Denmark, 7 Department of Cardiology, Aarhus University Hospital, Aalborg, Denmark, 8 Department of Epidemiology, School of Public Health, Aarhus University, Aarhus, Denmark, 9 Inserm, Centre for Research in Epidemiology and Population Health, Institut Gustave Roussy, Villejuif, France, 10 Paris South University, UMRS 1018, F-94805, Villejuif, France, 11 German Cancer Research Center, Division of Cancer Epidemiology, Heidelberg, Germany, 12 Department of Epidemiology, German Institute of Human Nutrition, Potsdam-Rehbrücke, Germany, 13 World Health Organization Collaborating Center for Food and Nutrition Policies, Department of Hygiene, Epidemiology and Medical Statistics, University of Athens Medical School, Athens, Greece, 14 Hellenic Health Foundation, Athens, Greece, 15 Department of Epidemiology, Harvard School of Public Health, Boston, Massachusetts, United States of America, 16 Bureau of Epidemiologic Research, Academy of Athens, Athens, Greece, 17 Molecular and Nutritional Epidemiology Unit, Cancer Research and Prevention Institute – ISPO, Florence, Italy, 18 Nutritional Epidemiology Unit, Fondazione IRCCS Istituto Nazionale Tumori, Milan, Italy, 19 Cancer Registry and Histopathology Unit, “Civile – M.P.Arezzo” Hospita, ASP Ragusa, Italy, 20 HuGe Foundation, Torino, Italy, 21 Department of Clinical and Experimental Medicine, Federico II University, Naples, Italy, 22 Julius Centre, University Medical Centre Utrecht, Utrecht, The Netherlands, 23 Public Health Directorate, Asturias, Spain, 24 Unit of Nutrition, Environment and Cancer Catalan Institute of Oncology, Barcelona, Spain, 25 Andalusian School of Public Health, Granada, Spain, 26 CIBER Epidemiology and Public Health CIBERESP, Madrid, Spain, 27 Public Health Division of Gipuzkoa, Basque Regional Health Department and Ciberesp-Biodonostia, San Sebastian, Spain, 28 Department of Epidemiology, Murcia Regional Health Council, Murcia, Spain, 29 Sociosanitary Sciences Department, Universidad de Murcia, Murcia, Spain, 30 Navarre Public Health Institute, Pamplona, Spain, 31 Department of Plastic Surgery, Lund University, Malmo, Sweden, 32 Department of Surgery, Lund University, Malmo, Sweden, 33 Department of Odontology, Umeå University, Umeå, Sweden, 34 Department of Medical Biosciences, Pathology, Umeå University, Umeå, Sweden, 35 University of Cambridge, Cambridge, United Kingdom, 36 MRC Epidemiology Unit, Cambridge, United Kingdom, 37 Cancer Epidemiology Unit, Nuffield Department of Clinical Medicine, University of Oxford, Oxford, United Kingdom

Abstract

Background: Prospective studies have consistently reported lower colorectal cancer risks associated with higher intakes of total dairy products, total milk and dietary calcium. However, less is known about whether the inverse associations vary for individual dairy products with differing fat contents.

Materials and Methods: In the European Prospective Investigation into Cancer and Nutrition (EPIC), we investigated the associations between intakes of total milk and milk subtypes (whole-fat, semi-skimmed and skimmed), yoghurt, cheese, and dietary calcium with colorectal cancer risk amongst 477,122 men and women. Dietary questionnaires were administered at baseline. Multivariable hazard ratios (HRs) and 95% confidence intervals (CIs) were estimated using Cox proportional hazards models, adjusted for relevant confounding variables.

Results: During the mean 11 years of follow-up, 4,513 incident cases of colorectal cancer occurred. After multivariable adjustments, total milk consumption was inversely associated with colorectal cancer risk (HR per 200 g/day 0.93, 95% CI: 0.89–0.98). Similar inverse associations were observed for whole-fat (HR per 200 g/day 0.90, 95% CI: 0.82–0.99) and skimmed milk (HR per 200 g/day 0.90, 95% CI: 0.79–1.02) in the multivariable models. Inverse associations were observed for cheese and yoghurt in the categorical models; although in the linear models, these associations were non-significant. Dietary calcium was inversely associated with colorectal cancer risk (HR per 200 mg/day 0.95, 95% CI: 0.91–0.99); this association was limited to dairy sources of calcium only (HR per 200 mg/day 0.95, 95% CI: 0.91–0.99), with no association observed for non-dairy calcium sources (HR per 200 mg/day 1.00, 95% CI: 0.81–1.24).

Conclusions: Our results strengthen the evidence for a possible protective role of dairy products on colorectal cancer risk. The inverse associations we observed did not differ by the fat content of the dairy products considered.
Introduction

Colorectal cancer is the third most common cancer worldwide, with over 1.2 million new diagnoses estimated to have occurred in 2008 [1]. Variation in international incidence rates [2,3] coupled with findings from migrant studies [4,5] suggests that colorectal cancer etiology is influenced by modifiable lifestyle factors, such as diet. In the recent WCRF/AICR Continuous Update Project, alcoholic drinks and red and processed meat were judged to be “convincing” factors associated with increased colorectal cancer risk; whilst foods containing dietary fibre were similarly judged but associated with reduced risk [6]. For total dairy products, an updated meta-analysis (the WCRF Continuous Update Project) recently reported a 17% lower colorectal cancer risk per 400 g/day increased intake, [7] but indicated that evidence for individual products was lacking and/or uncertain.

Although an inverse association between consumption of total milk with colorectal cancer risk has been consistently observed, [7,8] whether the fat content of milk offsets a potential anti-carcinogenic role is unclear. Animal models have shown that high-fat consumption results in bile acid production, which in turn promotes colorectal cancer, [9] but associations between milk subtypes, with different fat contents, and colorectal cancer have rarely been examined in prospective studies [10]. Similarly, how other high-fat dairy products, such as cheese and yoghurt, are associated with colorectal cancer risk is unclear, as mixed results have been reported from the handful of previous prospective studies. For cheese consumption, four prospective studies reported null associations [8,11–13] and one study reported an inverse association [14]. For yoghurt, three cohort studies have not found any association, [8,11,12] but a recent analysis within the European Prospective Investigation into Cancer and Nutrition (EPIC)-Italy cohorts reported reduced risks amongst those with higher consumption, even after adjustment for calcium intake [15].

The principal anti-carcinogenic component contained within dairy products is believed to be calcium. Most, [8,11,12,16,17] but not all [18] cohort studies that have investigated calcium intake in relation to colorectal cancer have reported inverse associations. Previously within EPIC, a nested case-control study based on 1,248 colorectal cancer cases reported higher intakes of dietary calcium were associated with lower colorectal cancer risk [19]. Although, whether this association differed according to dairy and non-dairy sources of calcium was not investigated, nor was a potential non-linear relationship that has been observed in other cohorts [8,11].

In this present analysis, we investigated how intakes of milk with different fat content (total, whole-fat, semi-skimmed, and skimmed), cheese, yoghurt, and dietary calcium (total, dairy and non-dairy sources) relate to colorectal cancer risk in the EPIC study. The EPIC is a large prospective cohort from 10 European countries with a wide range of dietary intakes. The large number of participants and colorectal cancer cases accrued provided high statistical power to investigate relationships according to individual dairy products and across cancer sub-sites.

Methods

Outline

EPIC is an on-going multicentre prospective cohort study designed to investigate the associations between diet, lifestyle, genetic and environmental factors and various types of cancer. A detailed description of the methods has previously been published [20,21]. In summary, 521,448 participants (~70% women) mostly aged 35 years or above were recruited between 1992 and 2000. Participants were recruited from 23 study centres in ten European countries: Denmark, France, Germany, Greece, Italy, the Netherlands, Norway, Spain, Sweden, and United Kingdom (UK). Participants were recruited from the general population, with the following exceptions: the French cohort were teacher health insurance programme members; the Italian and Spanish cohorts included members of blood donor associations and the general population; the Utrecht (the Netherlands) and Florence (Italy) cohorts contained participants from mammographic screening programs; the Oxford (UK) cohort included a large proportion of vegetarians, vegans, and low meat eaters; finally, only women participated in the cohorts of France, Norway, Naples (Italy) and Utrecht (the Netherlands). Written informed consent was provided by all study participants. Ethical approval for the EPIC study was obtained from the review boards of the International Agency for Research on Cancer (IARC) and local participating centres. Exclusions prior to the onset of the analyses included: participants with prevalent cancer at enrolment (n = 28,285); participants with missing dietary, lifestyle, and anthropometric data (n = 6,253); participants in the highest and lowest 1% of the distribution for the ratio between energy intake to estimated energy requirement (n = 9,600); and finally participants with extreme total dairy
intakes above 2000 g/day (n = 190). Our study therefore included 477,122 participants (334,901 women and 142,141 men).

Diet and lifestyle questionnaires

Dietary information over the previous 12 months was obtained at study baseline using validated country/centre specific dietary questionnaires. In Malmo (Sweden), a dietary questionnaire was combined with a 7-day food registration and interview. In Greece, two Italian centres, and Spain, interviewers administered the dietary questionnaires. In all other centres/countries, the questionnaires were self-administered. In Spain, France, and Ragusa (Italy) questions were structured by meals, while in other countries the structure was by food groups. Also at baseline, standardized computer-based single 24-hour dietary recalls (24-hdr) were collected from 36,994 study participants. This additional dietary assessment was used to calibrate for differences in questionnaires across countries [22]. Individual dairy products were categorized as milk, cheeses, and yoghurts. Due to relatively low intakes and incomplete measurements across centres, other individual dairy products such as ice cream, cream desserts and milk-based puddings, milk beverages, dairy creams and creams for milk and coffee were not analysed individually. Total milk was assessed as the sum of all types of milk consumed (whole-fat, skimmed, semi skimmed, and not specified). Semi-skimmed milk was defined as milk containing 0.5–2.5% fat, and skimmed milk was defined as having <0.5% fat content. Milk subtype information was unavailable in Norway, and only partially available in Germany, Greece (both whole-fat milk only), and three Italian centres (Florence, Varese, Turin; whole-fat and semi-skimmed milks only). Cheese included all kinds of fresh, fermented, and matured cheese. Yoghurt included natural and flavoured in all cohorts, and additionally fermented milk in Sweden, Norway, and Denmark. Intakes of calcium were obtained from the EPIC Nutrient Data Base (ENDB); in which the nutritional composition of foods across the different countries has been standardized [23].

Lifestyle questionnaires were used to obtain information on education (used as a proxy for socioeconomic status), smoking status and intensity, alcohol consumption, and physical activity levels. Height and weight were measured at the baseline examination in all centres apart from part of Oxford, and all of the Norway and France subcohorts, where measurements were self-reported via the lifestyle questionnaire [20].

Ascertainment of colorectal cancer incidence

Population cancer registries were used in Denmark, Italy, the Netherlands, Norway, Spain, Sweden and the United Kingdom to identify incident cancer diagnoses. In France, Germany and Greece cancer cases during follow-up were identified by a combination of methods including: health insurance records, cancer and pathology registries, and by active follow-up directly through study participants or through next-of-kin. Complete follow-up censoring dates varied amongst centres, ranging between 2005 and 2010.

Cancer incidence data were coded using the 10th Revision of the International Classification of Diseases (ICD-10) and the second revision of the International Classification of Disease for Oncology (ICDO-2). Proximal colon cancer included those within the cecum, appendix, ascending colon, hepatic flexure, transverse colon, and splenic flexure (C18.0–18.5). Distal colon cancer included those within the descending (C18.6) and sigmoid (C18.7) colon. Overlapping (C18.8) and unspecified (C18.9) lesions of the colon were grouped among colon cancers only. Cancer of the rectum included cancer occurring at the recto sigmoid junction (C19) and rectum (C20).

Statistical analysis

Hazard ratios (HRs) and 95% confidence intervals (CIs) were estimated using Cox proportional hazards models. Age was the primary time variable in all models. Time at entry was age at recruitment. Exit time was age at whichever of the following came first: colorectal cancer diagnosis, death, or the last date at which follow-up was considered complete in each centre. To control for differing follow-up procedures, questionnaire design, and other differences across centres, models were stratified by study centre. Models were also stratified by sex and age at recruitment in 1-year categories. Possible non-proportionality was assessed using an analysis of Schoenfeld residuals, [24] with no evidence of non-proportionality being detected.

Dietary intakes were modelled using either quintiles defined across cohort participants (total milk, total dairy and calcium), pre-defined categories (whole-fat, semi-skimmed, and skimmed milks: non consumers, <100, 100–199, 200–299, ≥300 g/day); and a predefined low intake reference category and quartiles defined across the remaining participants (cheese reference category = <5 g/day; yoghurt reference category = non-consumers). Intakes were also modelled as continuous variables, with HR expressed per increments of: 200 g/day for milk; 100 g/day for yoghurt; 50 g/day for cheese; 400 g/day for total dairy intake, and 200 mg/day for calcium. Trend tests across intake categories were calculated by assigning the median value of each intake quintile/category and modelling as continuous terms into Cox regression models.

Analyses for colorectal, colon, proximal colon, distal colon, and rectal cancers were conducted for both sexes combined as no interactions by sex were observed for intakes of total dairy products (P = 0.26), milk (P = 0.28), cheese (P = 0.58), yoghurt (P = 0.51), and dietary calcium (P = 0.11). The results by sex are in Tables S1, S2, S3, and S4 in File S1. All models were adjusted for total energy intake, using the standard model, to obtain isocaloric risk estimates and partly control for measurement error of dairy products and calcium intake estimates. All models were additionally adjusted for: body mass index (BMI; kg/m²; continuous); physical activity (inactive, moderately inactive, moderately active, active, or missing); smoking status and intensity (never, current, 1–15 cigarettes per day; current, 16–25 cigarettes per day; current, 25+ cigarettes per day; former, quit ≤10 years; former, quit 11–20 years; former, quit 20+ years; current, pipe/cigar/occasional; current/former, missing; or unknown); education level (none/primary school completed, technical/professional school, secondary school, longer education – including university, or unknown); menopausal status (premenopausal, postmenopausal, perimenopausal/unknown menopausal status, or surgical postmenopausal); ever use of oral contraceptive (yes, no, or unknown); ever use of menopausal hormone therapy (yes, no, or unknown); and intakes of alcohol (yes or no; continuous, g/day), red and processed meats, and fibre (both continuous, g/day). Finer adjustment for body shape was attempted by also controlling for waist circumference in a subset of the cohort for which measurements were available. When included in the multivariable models, instead of, or with BMI, the risk estimates were virtually unchanged; and accordingly, we adjusted solely for BMI. In the analyses for whole-fat, semi-skimmed, and skimmed milk, the models included the covariates as detailed above, plus additional adjustment for the other milk subtypes. Similarly, the dairy and non-dairy calcium analyses were mutually adjusted for one another.

To determine whether the dietary calcium-colorectal cancer association differed according to anthropometric, lifestyle, and dietary characteristics, we included interaction terms (multiplicative scale) in separate models. The statistical significance of
the cross-product terms were evaluated using the likelihood ratio test.

Cox proportional hazard restricted cubic spline models were used to explore possible deviation from a non-linear calcium-colorectal cancer relationship, with five knots specified at the median of each quintile of intake [23]. Heterogeneity of associations across anatomical cancer sub-sites was assessed by calculating $\chi^2$ statistics. The heterogeneity across countries was explored by taking a meta-analytic approach [26]. To evaluate possible reverse causality, cases diagnosed within the first 2 and 5 years of follow-up were excluded from the analyses.

To improve comparability of data across study centres and to partially correct the relative risk estimates for the measurement error of dietary intakes, a linear regression calibration model was used utilizing the 24-hdr taken at baseline from a subset of the cohort (n = 34,426 in this analysis) [27,28]. The 24-hdr were regressed on dietary questionnaire values, with adjustment for the same list of covariates detailed above, and further control for the week day and season of recall measurements. Country and sex-specific calibration models were used to obtain individual calibrated values of dietary exposure for all participants. Cox proportional hazards regression models were then applied using the calibrated values for each participant on a continuous scale. The standard error of the de-attenuated coefficients was corrected through bootstrap sampling. The P-value for the trend of the de-attenuated coefficients was calculated by dividing the de-attenuated coefficient by the bootstrap-derived standard error and approximating the standardized normal distribution. (29).

Statistical tests used in the analysis were all two-sided and a $P$-value of $<0.05$ was considered statistically significant. Analyses were conducted using SAS version 9.1 and Stata version 11.0.

Results

After a mean (SD) follow-up of 11.0 (2.8) years, 4,513 colorectal cancer cases were documented amongst the 477,122 participants. Of the 4,513 colorectal cancer cases, 2,868 were colon tumours (1,298 proximal; 1,266 distal and 304 overlapping or unspecified), and 1,645 were rectal tumours. The total person-years and cancer cases were documented amongst the 477,122 participants. Although this difference was not significant ($P$ = 0.83; distal colon vs. proximal colon $P$ Heterogeneity = 0.76) (Table 3). In calibrated models, colorectal cancer risk was 7% lower for each 200 g/day higher intake of total milk. Over 17% of participants reported consuming more than one milk subtype. The linear inverse associations for colorectal, colon, and rectal cancers were of similar strength for whole-fat and skimmed milk, but there were no significant associations for semi-skimmed milk (Table 4).

However, in sensitivity analyses, when the models included only sole consumers of each milk subtype, identical inverse colorectal cancer risk estimates were observed for whole-fat (HR per 200 g/day 0.87, 95% CI: 0.79–0.95), semi-skimmed (HR per 200 g/day 0.87, 95% CI: 0.78–0.97) and skimmed milks (HR per 200 g/day 0.87, 95% CI: 0.76–0.99) (data not tabulated).

Total milk and milk subtypes by fat content

Total milk was similarly inversely related to the cancer risk across all locations of the bowel (colon vs. rectal $P$ Heterogeneity = 0.63; distal colon vs. proximal colon $P$ Heterogeneity = 0.76) (Table 3). In calibrated models, colorectal cancer risk was 7% lower for each 200 g/day higher intake of total milk. Over 17% of participants reported consuming more than one milk subtype. The linear inverse associations for colorectal, colon, and rectal cancers were of similar strength for whole-fat and skimmed milk, but there were no significant associations for semi-skimmed milk (Table 4).

However, in sensitivity analyses, when the models included only sole consumers of each milk subtype, identical inverse colorectal cancer risk estimates were observed for whole-fat (HR per 200 g/day 0.87, 95% CI: 0.79–0.95), semi-skimmed (HR per 200 g/day 0.87, 95% CI: 0.78–0.97) and skimmed milks (HR per 200 g/day 0.87, 95% CI: 0.76–0.99) (data not tabulated).

Cheese

Cheese consumption was inversely associated with colorectal cancer in the categorical model (Table 3). The association was significant for colon ($\geq$56 g/day vs. <5 g/day HR, 0.83, 95% CI: 0.71–0.97; $P$-trend = 0.047) but not rectal cancer, although this difference was not significant ($P$ Heterogeneity = 0.39). In the linear calibrated models, non-significant inverse associations were observed for colorectal, colon and rectal cancers of similar strength for whole-fat and skimmed milk, but there were no significant associations for semi-skimmed milk (Table 4).

Table 1. Descriptive information of the European Prospective Investigation into Cancer and Nutrition participant countries.

<table>
<thead>
<tr>
<th>Country</th>
<th>Number of participants</th>
<th>Total person-years</th>
<th>Number of colorectal cancer cases</th>
<th>Total dairy products intake (g/day) *</th>
<th>Dietary calcium intake (mg/day) *</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Men</td>
<td>Women</td>
<td>Men</td>
<td>Women</td>
<td>Men</td>
</tr>
<tr>
<td>Denmark</td>
<td>26,266</td>
<td>28,699</td>
<td>284,431</td>
<td>316,511</td>
<td>474</td>
</tr>
<tr>
<td>France</td>
<td>–</td>
<td>67,372</td>
<td>–</td>
<td>699,221</td>
<td>–</td>
</tr>
<tr>
<td>Germany</td>
<td>21,135</td>
<td>27,386</td>
<td>208,164</td>
<td>271,857</td>
<td>263</td>
</tr>
<tr>
<td>Greece</td>
<td>10,807</td>
<td>15,225</td>
<td>99,108</td>
<td>148,604</td>
<td>61</td>
</tr>
<tr>
<td>Italy</td>
<td>14,029</td>
<td>30,510</td>
<td>158,917</td>
<td>341,469</td>
<td>173</td>
</tr>
<tr>
<td>Spain</td>
<td>15,147</td>
<td>24,849</td>
<td>182,950</td>
<td>299,557</td>
<td>185</td>
</tr>
<tr>
<td>Sweden</td>
<td>22,287</td>
<td>26,374</td>
<td>289,320</td>
<td>349,295</td>
<td>339</td>
</tr>
<tr>
<td>The Netherlands</td>
<td>9,618</td>
<td>26,854</td>
<td>115,334</td>
<td>315,529</td>
<td>81</td>
</tr>
<tr>
<td>United Kingdom</td>
<td>22,852</td>
<td>52,543</td>
<td>252,096</td>
<td>586,301</td>
<td>324</td>
</tr>
<tr>
<td>All EPIC</td>
<td>142,141</td>
<td>334,981</td>
<td>1,590,320</td>
<td>3,670,621</td>
<td>1,900</td>
</tr>
</tbody>
</table>

*Data are median intake information collected from 24-hour dietary recalls (n = 34,426 participants).

doi:10.1371/journal.pone.0072715.t001
cancers. For proximal colon cancer, the highest consumers (>56 g/day) had a 27% (95% CI: 0.58–0.93) reduced risk compared to those consuming <5 g/day, but in the calibrated model, this association was not significant. No association was observed for tumours in the distal region of the colon, and the heterogeneity in association by colonic region was not statistically significant ($P_{\text{Heterogeneity}} = 0.82$).

### Table 2. Baseline characteristics of study participants by categories of total dairy intake.

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Quintile of total dairy intake</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Q1</td>
</tr>
<tr>
<td><strong>Dairy intake range (g/day)</strong></td>
<td></td>
</tr>
<tr>
<td>Men</td>
<td></td>
</tr>
<tr>
<td>$N$</td>
<td>33,251</td>
</tr>
<tr>
<td>Colorectal cancer cases</td>
<td>484</td>
</tr>
<tr>
<td>Age at recruitment (years) ‡</td>
<td>52.1</td>
</tr>
<tr>
<td>Body mass index (kg/m2)</td>
<td>26.8</td>
</tr>
<tr>
<td>Education ‡</td>
<td>24.7</td>
</tr>
<tr>
<td>Smoking status and intensity ‡</td>
<td>34.0</td>
</tr>
<tr>
<td>Physical activity ‡</td>
<td>22.7</td>
</tr>
<tr>
<td>Total energy intake (kcal/day)</td>
<td>2233</td>
</tr>
<tr>
<td>Red and processed meat intake (g/day)</td>
<td>100.5</td>
</tr>
<tr>
<td>Calcium intake (mg/day)</td>
<td>685</td>
</tr>
<tr>
<td>Fibre intake (g/day)</td>
<td>23.2</td>
</tr>
<tr>
<td>Alcohol intake (g/day)</td>
<td>26.7</td>
</tr>
<tr>
<td>Women</td>
<td></td>
</tr>
<tr>
<td>$N$</td>
<td>62,174</td>
</tr>
<tr>
<td>Colorectal cancer cases</td>
<td>494</td>
</tr>
<tr>
<td>Age at recruitment (years) ‡</td>
<td>50.4</td>
</tr>
<tr>
<td>Body mass index (kg/m2)</td>
<td>25.1</td>
</tr>
<tr>
<td>Education ‡</td>
<td>19.6</td>
</tr>
<tr>
<td>Smoking status and intensity ‡</td>
<td>24.5</td>
</tr>
<tr>
<td>Physical activity ‡</td>
<td>10.7</td>
</tr>
<tr>
<td>Ever use of contraceptive pill ‡</td>
<td>56.1</td>
</tr>
<tr>
<td>Ever use of menopausal hormone therapy ‡</td>
<td>22.9</td>
</tr>
<tr>
<td>Menopausal status ‡</td>
<td>41.8</td>
</tr>
<tr>
<td>Postmenopausal (%)</td>
<td></td>
</tr>
<tr>
<td>Total energy intake (kcal/day)</td>
<td>1694</td>
</tr>
<tr>
<td>Red and processed meat intake (g/day)</td>
<td>62.8</td>
</tr>
<tr>
<td>Calcium intake (mg/day)</td>
<td>610</td>
</tr>
<tr>
<td>Fibre intake (g/day)</td>
<td>20.7</td>
</tr>
<tr>
<td>Alcohol intake (g/day)</td>
<td>8.7</td>
</tr>
</tbody>
</table>

Mean and standard deviation unless stated otherwise.

doi:10.1371/journal.pone.0072715.t002
Table 3. Multivariable hazard ratios (95% confidence intervals) of colorectal cancer risk by dairy product consumption categories.

<table>
<thead>
<tr>
<th>Food group (g/day)</th>
<th>CRC cases (n) *</th>
<th>Person-years</th>
<th>Colorectal cancer †</th>
<th>Colorectal cancer ‡</th>
<th>All colon †</th>
<th>Proximal †</th>
<th>Distal †</th>
<th>Rectal cancer ‡</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(n = 4,513)</td>
<td></td>
<td>(n = 4,513)</td>
<td>(n = 2,868)</td>
<td>(n = 1,298)</td>
<td>(n = 1,266)</td>
<td>(n = 1,645)</td>
<td></td>
</tr>
<tr>
<td>Total milk</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1</td>
<td>&lt; 9</td>
<td>808</td>
<td>1,013,915</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2</td>
<td>9–89</td>
<td>935</td>
<td>1,044,757</td>
<td>1.00 (0.88–1.07)</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Q3</td>
<td>90–187</td>
<td>836</td>
<td>1,037,641</td>
<td>0.92 (0.83–1.02)</td>
<td>0.92</td>
<td>0.95</td>
<td>0.95</td>
<td>0.88</td>
</tr>
<tr>
<td>Q4</td>
<td>188–324</td>
<td>988</td>
<td>1,079,981</td>
<td>0.89 (0.80–0.98)</td>
<td>0.90</td>
<td>0.90</td>
<td>0.88</td>
<td>0.94</td>
</tr>
<tr>
<td>Q5</td>
<td>≥ 325</td>
<td>946</td>
<td>1,084,647</td>
<td>0.80 (0.72–0.89)</td>
<td>0.81</td>
<td>0.80</td>
<td>0.84</td>
<td>0.78</td>
</tr>
<tr>
<td>P-trend</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>&lt;0.001</td>
<td>0.01</td>
<td>0.09</td>
<td>0.017</td>
</tr>
<tr>
<td>Per 200 g/day – uncalibrated</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.94 (0.91–0.97)</td>
<td>0.93 (0.90–0.97)</td>
<td>0.95 (0.89–1.01)</td>
<td>0.94 (0.88–0.99)</td>
</tr>
<tr>
<td>Per 200 g/day – calibrated</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.93 (0.89–0.98)</td>
<td>0.93 (0.88–0.98)</td>
<td>0.93 (0.87–0.99)</td>
<td>0.94 (0.87–1.02)</td>
</tr>
<tr>
<td>Cheese</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1</td>
<td>&lt; 5</td>
<td>495</td>
<td>506,354</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2</td>
<td>5–18</td>
<td>1073</td>
<td>1,218,699</td>
<td>0.93 (0.83–1.04)</td>
<td>0.92</td>
<td>0.87</td>
<td>0.76</td>
<td>1.06</td>
</tr>
<tr>
<td>Q3</td>
<td>19–32</td>
<td>1114</td>
<td>1,173,780</td>
<td>0.99 (0.88–1.11)</td>
<td>0.99</td>
<td>0.94</td>
<td>0.84</td>
<td>1.02</td>
</tr>
<tr>
<td>Q4</td>
<td>33–55</td>
<td>980</td>
<td>1,185,851</td>
<td>0.88 (0.78–0.99)</td>
<td>0.89</td>
<td>0.86</td>
<td>0.70</td>
<td>1.01</td>
</tr>
<tr>
<td>Q5</td>
<td>≥ 56</td>
<td>851</td>
<td>1,176,256</td>
<td>0.86 (0.76–0.98)</td>
<td>0.87</td>
<td>0.83</td>
<td>0.73</td>
<td>0.91</td>
</tr>
<tr>
<td>P-trend</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.09</td>
<td>0.02</td>
<td>0.04</td>
<td>0.02</td>
</tr>
<tr>
<td>Per 50 g/day – uncalibrated</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.95 (0.90–1.00)</td>
<td>0.94 (0.88–1.01)</td>
<td>0.90 (0.81–1.01)</td>
<td>0.95 (0.85–1.06)</td>
</tr>
<tr>
<td>Per 50 g/day – calibrated</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.92 (0.80–1.06)</td>
<td>0.88 (0.75–1.04)</td>
<td>0.85 (0.68–1.05)</td>
<td>0.82 (0.66–1.03)</td>
</tr>
<tr>
<td>Yoghurt</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1</td>
<td>&lt; 17.8</td>
<td>958</td>
<td>1,051,433</td>
<td>0.96 (0.87–1.07)</td>
<td>0.97</td>
<td>0.98</td>
<td>1.02</td>
<td>0.93</td>
</tr>
<tr>
<td>Q2</td>
<td>17.9–53</td>
<td>768</td>
<td>1,031,493</td>
<td>0.91 (0.82–1.00)</td>
<td>0.93</td>
<td>0.98</td>
<td>0.96</td>
<td>0.97</td>
</tr>
<tr>
<td>Q3</td>
<td>54–108</td>
<td>824</td>
<td>1,047,993</td>
<td>0.88 (0.80–0.98)</td>
<td>0.92</td>
<td>0.98</td>
<td>1.00</td>
<td>0.95</td>
</tr>
<tr>
<td>Q4</td>
<td>≥ 109</td>
<td>889</td>
<td>1,069,512</td>
<td>0.86 (0.78–0.95)</td>
<td>0.90</td>
<td>0.88</td>
<td>0.94</td>
<td>0.93</td>
</tr>
<tr>
<td>P-trend</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.02</td>
<td>0.04</td>
<td>0.03</td>
<td>0.09</td>
</tr>
<tr>
<td>Per 100 g/day – calibrated</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.99 (0.95–1.03)</td>
<td>0.99 (0.94–1.04)</td>
<td>1.00 (0.94–1.07)</td>
<td>0.98 (0.91–1.05)</td>
</tr>
<tr>
<td>Per 100 g/day – uncalibrated</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.97 (0.90–1.04)</td>
<td>0.98 (0.90–1.06)</td>
<td>1.02 (0.92–1.13)</td>
<td>0.93 (0.82–1.06)</td>
</tr>
<tr>
<td>Total dairy</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1</td>
<td>&lt; 134</td>
<td>978</td>
<td>1,028,047</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Q2</td>
<td>134–228</td>
<td>850</td>
<td>1,031,665</td>
<td>0.89 (0.81–0.97)</td>
<td>0.90</td>
<td>0.86</td>
<td>0.75</td>
<td>0.91</td>
</tr>
<tr>
<td>Q3</td>
<td>229–332</td>
<td>871</td>
<td>1,053,198</td>
<td>0.83 (0.75–0.91)</td>
<td>0.85</td>
<td>0.86</td>
<td>0.81</td>
<td>0.86</td>
</tr>
</tbody>
</table>
Table 3. Cont.

<table>
<thead>
<tr>
<th>BOTH SEXES</th>
<th>Basic model</th>
<th>Multivariable models</th>
</tr>
</thead>
<tbody>
<tr>
<td>Food group (g/day)</td>
<td>CRC cases (n) *</td>
<td>Person-years</td>
</tr>
<tr>
<td>n</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(n = 4,513)</td>
<td>(n = 4,513)</td>
</tr>
<tr>
<td>Q4</td>
<td>333–489</td>
<td>853</td>
</tr>
<tr>
<td>Q5</td>
<td>≥490</td>
<td>961</td>
</tr>
<tr>
<td>P-trend</td>
<td>&lt;0.001</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Per 400 g/day – uncalibrated</td>
<td>0.88 (0.83–0.93)</td>
<td>0.87 (0.80–0.93)</td>
</tr>
<tr>
<td>Per 400 g/day – calibrated</td>
<td>0.86 (0.79–0.94)</td>
<td>0.85 (0.76–0.95)</td>
</tr>
</tbody>
</table>

*Basic model – Cox regression using total energy intake (continuous), and stratified by age (1-year categories), sex, and centre.
†Multivariable model – Cox regression using total energy intake (continuous), body mass index (continuous), physical activity index (inactive, moderately inactive, moderately active, active, or missing), smoking status and intensity (never; current, 1–15 cigarettes per day; current, 16–25 cigarettes per day; current, 26+ cigarettes per day; former, quit ≤10 years; former, quit 11–20 years; former, quit 20+ years; current, pipe/cigar/occasional; current/former, missing; unknown), education status (none, primary school completed, technical/professional school, secondary school, longer education including university, or not specified), ever use of contraceptive pill (yes, no, or unknown), ever use of menopausal hormone therapy (yes, no, or unknown), menopausal status (premenopausal, postmenopausal, perimenopausal/unknown menopausal status, or surgical menopause), alcohol consumption (yes or no; and continuous), and intakes of red and processed meat and fibre (both continuous), and stratified by age (1-year categories), sex, and centre. 

Yoghurt intake was significantly inversely related to colorectal cancer risk in categorical models (≥109 g/day vs. non-consumers, HR 0.90, 95% CI: 0.81–0.99; P-trend = 0.043) (Table 3). The inverse association was restricted to the colon and not observed for tumours in the rectum, although the difference was not statistically significant (P Heterogeneity = 0.79). Within the colon, the difference in association across the distal and proximal regions was non-significant (P Heterogeneity = 0.29). No associations were observed in the linear calibrated models for cancers across all bowel locations. After adjustment for dietary calcium intake the inverse association for colorectal cancer using the categorical model was no longer significant (≥109 g/day vs. non-consumers, HR 0.94, 95% CI: 0.83–1.04; P-trend = 0.33; data not tabulated).

Total dairy intake

Total dairy intake was significantly inversely associated to colorectal cancer risk (≥490 g/day vs. <134 g/day, HR 0.77, 95% CI: 0.70–0.86; P-trend <0.001) (Table 3). In calibrated models, each 400 g/day higher intake of total dairy products was associated with a 14% lower risk. The inverse association was of similar magnitude for colon and rectal cancer (P Heterogeneity = 0.72); and within the colon, there was no evidence of heterogeneity across distal and proximal regions (P Heterogeneity = 0.66).

Dietary calcium

For dietary calcium, similar strength inverse associations were observed across all locations of the colorectum (colon vs. rectal P Heterogeneity = 0.36; distal colon vs. proximal colon P Heterogeneity = 1.00) (Table 5). There was no deviation from linearity for the relationship between dietary calcium and colorectal cancer in the restricted cubic spline model (P = 0.43) (data not shown). Calcium intake from dairy foods was inversely associated to cancer risk across all locations of the bowel. When calcium and milk were included in the same models, the inverse associations for milk weakened and became non-significant, but the significant inverse associations for calcium remained (data not shown). Dietary calcium from non-dairy sources was not inversely associated with colorectal cancer risk. The association between dietary calcium intake and risk of colorectal cancer did not differ by BMI (P = 0.56), waist circumference (men P = 0.74; women P = 0.64), physical activity (P = 0.26), smoking status (P = 0.37), alcohol consumption (P = 0.75), and intakes of red and processed meat (P = 0.50), and fibre (P = 0.65) (data not tabulated).

Between country heterogeneity and inclusion of preclinical disease

There was evidence of significant heterogeneity by country for total dairy products (P = 0.034) (Figure S1 in File S1); although risk estimates ≤1 were observed in most countries. No associations were observed in the Sweden and Denmark cohorts. Non-significant between country heterogeneity was observed for intakes of dietary calcium (P = 0.60; Figure S2 in File S1), total milk (P = 0.13), cheese (P = 0.64), and yoghurt (P = 0.12).

Excluding the participants with less than 2 and 5 years of follow-up (including 502 and 1,483 colorectal cancer cases respectively) from the total dairy, total milk, cheese, yoghurt, and calcium intake analyses resulted in negligible differences in the colorectal cancer associations (data not shown).

Discussion

In this analysis of the EPIC cohort, after a mean follow-up of 11 years where 4,513 cases accrued, higher intakes of all subtypes of milk, cheese, yoghurt, total dairy products and dietary calcium from dairy sources were associated with reduced colorectal cancer risk. Overall, our results provided no evidence for divergent...
relationships for high and low-fat dairy products with colorectal cancer risk.

The inverse association we observed for total milk consumption was similar to what was reported by both the Pooling Project of cohort studies, and a recent systematic review [7,8]. Few prospective studies have previously investigated the associations for milk by fat content. In the Adventist Health Study, a stronger inverse association was reported for non-fat milk consumers compared to consumers of milks containing higher fat [10]. In our larger analysis, similar strength inverse associations were observed for all milk subtypes, refuting the notion that the milk-colorectal cancer association differs according to fat content, at least in the range of intakes recorded within EPIC.

The inverse cheese-colorectal cancer association observed in the smaller analysis, similar strength inverse associations were observed for all milk subtypes, refuting the notion that the milk-colorectal cancer association differs according to fat content, at least in the range of intakes recorded within EPIC.

The inverse cheese-colorectal cancer association observed in the smaller analysis, similar strength inverse associations were observed for all milk subtypes, refuting the notion that the milk-colorectal cancer association differs according to fat content, at least in the range of intakes recorded within EPIC.
Table 5. Multivariable hazard ratios (95% confidence intervals) of colorectal cancer risk by dietary calcium intake categories.

<table>
<thead>
<tr>
<th>BOTH SEXES</th>
<th>Calcium (mg/day)</th>
<th>Dairy calcium (mg/day)</th>
<th>Non-dairy calcium (mg/day)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>CRC cases (n) *</td>
<td>Person-years (n=4,513)</td>
<td>Person-years (n=4,513)</td>
</tr>
<tr>
<td><strong>Basic model</strong></td>
<td><strong>Multivariable model</strong></td>
<td><strong>Multivariable models</strong></td>
<td><strong>Colon cancer</strong></td>
</tr>
<tr>
<td>Q1</td>
<td>&lt;661</td>
<td>943</td>
<td>1,034,125</td>
</tr>
<tr>
<td>Q2</td>
<td>662–845</td>
<td>885</td>
<td>1,048,337</td>
</tr>
<tr>
<td>Q3</td>
<td>846–1030</td>
<td>921</td>
<td>1,054,640</td>
</tr>
<tr>
<td>Q4</td>
<td>1031–1279</td>
<td>891</td>
<td>1,059,887</td>
</tr>
<tr>
<td>Q5</td>
<td>≥1280</td>
<td>873</td>
<td>1,063,951</td>
</tr>
<tr>
<td>$P$-trend</td>
<td>&lt;0.001</td>
<td>&lt;0.001</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Dairy calcium (mg/day)</th>
<th>Per 200 mg/day – uncalibrated</th>
<th>Per 200 mg/day – calibrated</th>
</tr>
</thead>
<tbody>
<tr>
<td>Q1</td>
<td>&lt;308</td>
<td>998</td>
</tr>
<tr>
<td>Q2</td>
<td>309–462</td>
<td>869</td>
</tr>
<tr>
<td>Q3</td>
<td>463–621</td>
<td>866</td>
</tr>
<tr>
<td>Q4</td>
<td>622–838</td>
<td>866</td>
</tr>
<tr>
<td>Q5</td>
<td>≥839</td>
<td>909</td>
</tr>
<tr>
<td>$P$-trend</td>
<td>&lt;0.001</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Non-dairy calcium (mg/day)</th>
<th>Per 200 mg/day – uncalibrated</th>
<th>Per 200 mg/day – calibrated</th>
</tr>
</thead>
<tbody>
<tr>
<td>Q1</td>
<td>&lt;276</td>
<td>816</td>
</tr>
<tr>
<td>Q2</td>
<td>277–344</td>
<td>971</td>
</tr>
<tr>
<td>Q3</td>
<td>345–410</td>
<td>975</td>
</tr>
<tr>
<td>Q4</td>
<td>411–501</td>
<td>891</td>
</tr>
<tr>
<td>Q5</td>
<td>≥502</td>
<td>855</td>
</tr>
<tr>
<td>$P$-trend</td>
<td>0.244</td>
<td>0.87</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Per 200 mg/day</th>
<th>Per 200 mg/day – calibrated</th>
</tr>
</thead>
<tbody>
<tr>
<td>Q1</td>
<td>1.02 (0.94–1.10)</td>
</tr>
<tr>
<td>Q2</td>
<td>1.06 (0.96–1.17)</td>
</tr>
<tr>
<td>Q3</td>
<td>1.05 (0.91–1.21)</td>
</tr>
<tr>
<td>Q4</td>
<td>1.07 (0.92–1.25)</td>
</tr>
<tr>
<td>Q5</td>
<td>0.95 (0.83–1.08)</td>
</tr>
</tbody>
</table>

*Basic model – Cox regression using total energy intake (continuous), and stratified by age (1-year categories), sex, and centre.
Multivariable model – Cox regression using total energy intake (continuous), body mass index (continuous), physical activity index (inactive, moderately inactive, moderately active, active, or missing), smoking status and intensity (never, current, 1–15 cigarettes per day; current, 16–25 cigarettes per day; current, 16+ cigarettes per day; former, quit ≤10 years; former, quit 11–20 years; former, quit 20+ years; current, pipe/cigar/occasional; current/former, missing; unknown), education status (none, primary school completed, technical/professional school, secondary school, longer education including university, or not specified), ever use of contraceptive pill (yes, no, or unknown), ever use of menopausal hormone therapy (yes, no, or unknown), menopausal status (premenopausal, postmenopausal, perimenopausal, unknown)
For cheese consumption, results from the limited previous research have usually reported null results [11–13]. For yoghurt, an inverse colorectal cancer association in the categorical model was also not replicated in the linear calibrated model. Some evidence suggests that lactic acid bacteria contained within yoghurt products may protect against colorectal cancer [29]. Recently, an analysis of the EPIC-Italy cohorts reported a 35% reduced colorectal cancer risk - after adjustment for calcium intake – amongst participants who consumed more than 25 g/day of yoghurt compared to non-consumption (less than 1 g/day) [15]. When we additionally adjusted for calcium intake, the inverse colorectal association in the categorical model disappeared. However, our results do not rule out the lactic acid hypothesis, as the types of yoghurt consumed across EPIC countries may differ in lactic acid content, and this information may not have been captured within our study.

We observed consistent inverse associations across all cancer sub-sites for dietary calcium intake, in line with the majority of published cohort studies [8,11,12,16,17]. Some studies have reported a threshold level for calcium intake (~1,000 to ~1,400 mg/day), above which reductions in colorectal cancer risk are not observed [8,11]. In our analysis we did not observe a threshold association at any level of intake, or any departure from linearity. Our inverse associations were limited to dairy sources of calcium, as we observed either null or weak non-statistically significant associations in the non-dairy calcium models. Other prospective studies have reported no association [18,30] or increased risk [16] for non-dairy calcium intake with colorectal and colon cancers. A possible explanation for the non-inverse associations for non-dairy calcium could be that plant sources of calcium – the main contributors to non-dairy calcium intake amongst EPIC participants – contain oxalate and phytate which have been shown to inhibit calcium absorption [31]. Across all EPIC centres, milk contributes most to the consumption of total dairy products [32]. Lactose and casein in milk may increase the bioavailability of calcium, which could also explain the inverse associations we observed for dairy calcium [33]. The primary anticarcinogenic component contained within dairy foods is believed to be calcium [29]. Laboratory studies have shown that calcium can induce apoptosis in colonic epithelium cells, [34] and alter colonic Kras gene mutations [35]. Animal and human intervention studies have shown that calcium impacts upon colonic cell differentiation: indirectly, by binding to available bile acids and fatty acids, suppressing their ability to modify colonic cell proliferation [36,37]; and directly, by suppressing colonic epithelial cell proliferation and inducing terminal differentiation [38]. Evidence from clinical trials suggests that calcium supplementation reduces the recurrence of colorectal adenoma [39]. Beyond the calcium content of dairy products, other constituents contained within these products may explain the inverse associations observed. For instance lactoferrin, vitamin D in fortified dairy products, and certain fatty acids, such as butyric acid, have been linked with having possible beneficial roles against colorectal cancer. [29] However, isolating the influence of individual food components present simultaneously in the same foods is difficult.

The public health implications of our results are complicated by the contrasting associations between calcium intake and prostate cancer. Dietary calcium has been consistently associated with increased prostate cancer risk, and the WCRF/AICR 2007 Expert report judged it a “probable” cause of the disease [40]. Within EPIC, a 300 mg/day intake of dietary calcium was previously associated with a 9% increased risk of prostate cancer [41]. In our analysis, an equivalent daily intake amongst men and women would be associated with 7% statistically significant reduced colorectal cancer risks. At present, the available evidence for the divergent associations between cancer sites has not been considered convincing enough to justify potential sex-specific calcium and dairy product intake recommendations.

Strengths of our study include its large-scale prospective design, the large number of colorectal cancer cases and the possibility of controlling for the main potential confounders. However, information on past bowel cancer screening and previous endoscopy procedures were unknown; although previous studies have observed unchanged inverse calcium-colon cancer relationships when the multivariable models were additionally adjusted for endoscopy history [30]. A further limitation was that intake of calcium supplements could not be included in our analysis; although other large cohort studies have observed only minor differences in associations between total calcium intakes (supplements plus diet) compared solely to dietary sources [8,17]. In our study, diet was assessed through dietary questionnaires, which are subject to measurement error. Random misclassification may have thus caused an attenuation of the estimates of the diet-disease association; however, we partially corrected our estimates through regression calibration using 24-hdr data [23]. Another study limitation was that changes in diet during follow-up could not be taken into account; however, this does not appear to have influenced our conclusions since our results are consistent with those of other cohort studies, some of which used cumulative estimates of diet over time [8,30].

In conclusion, our study supports potential beneficial roles for dietary intakes of dairy products and calcium on colorectal cancer prevention. Inverse associations were observed for low-fat and high-fat dairy products, indicating that the fat content contained within dairy products does not influence this relationship.

Ethical approval review board information for local EPIC centres

The National Committee on Health Research Ethics (Denmark); Comité de Protection des Personnes (France); Ethics Committee of the Heidelberg University Medical School (Heidelberg, Germany); Ethikkommission der Landesarztekammer Brandenburg Cottbus (Potsdam, Germany); University of Athens Medical School (Greece) Comitato Etico Indipendente, Fondazione IRCCS Istituto Nazionale dei Tumori, Milano (Italy); Human Genetics Foundation Torino: Ethics Committee (Turin, Italy); The Medical Ethical Committee (METC= Medisch Ethische Toetsingscommissie) of the University Medical Center Utrecht (UMCU), Utrecht, the Netherlands (The Netherlands); Regional ethical committee for Northern Norway and the Norwegian Data Inspectorate (Norway); CEIC Comité de Ética de Investigación Clinica (Spain); Ethics Committee of Lund University (Malmo, Sweden); Umea Regional Ethical Review Board (Umea, Sweden); Norwich District Ethics Committee (Cambridge, UK); Scotland A Research Ethics Committee (Oxford); and the Imperial College Research Ethics Committee [ICREC] (UK).
Supporting Information

File S1 Supporting information. Figure S1. Multivariable hazard ratios and 95% confidence intervals of colorectal cancer risk by country, per 400 g/day increase in total dairy intake. Hazard ratios estimated by Cox proportional hazards models adjusting for total energy intake (continuous), body mass index (continuous), physical activity index (inactive, moderately inactive, moderately active, active, or missing), smoking status and intensity (never, current, 1–15 cigarettes per day; current, 16–25 cigarettes per day; current, 16+ cigarettes per day; former, quit ≤10 years; former, quit 11–20 years; former, quit 20+ years; current, pipe/cigar/occasional; current/former, missing; unknown), education status (none, primary school completed, technical/professional school, secondary school, longer education including university, or not specified), ever use of contraceptive pill (yes, no, or unknown), ever use of menopausal hormone therapy (yes, no, or unknown), menopausal status (premenopausal, postmenopausal, perimenopausal/unknown menopausal status, or surgical postmenopausal), alcohol consumption (yes or no; and continuous) and intakes of red and processed meat and fibre (both continuous), and stratified by age (1-year categories), sex, and centre. Figure S2. Multivariable hazard ratios and 95% confidence intervals of colorectal cancer risk by country, per 200 mg/day increase in total dietary calcium (B). Hazard ratios estimated by Cox proportional hazards models adjusting for total energy intake (continuous), body mass index (continuous), physical activity index (inactive, moderately inactive, moderately active, active, or missing), smoking status and intensity (never, current, 1–15 cigarettes per day; current, 16–25 cigarettes per day; current, 16+ cigarettes per day; former, quit ≤10 years; former, quit 11–20 years; former, quit 20+ years; current, pipe/cigar/occasional; current/former, missing; unknown), education status (none, primary school completed, technical/professional school, secondary school, longer education including university, or not specified), ever use of contraceptive pill (yes, no, or unknown), ever use of menopausal hormone therapy (yes, no, or unknown), menopausal status (premenopausal, postmenopausal, perimenopausal/unknown menopausal status, or surgical postmenopausal), alcohol consumption (yes or no; and continuous) and intakes of red and processed meat and fibre (both continuous), and stratified by age (1-year categories), sex, and centre. Table S1. Multivariable hazard ratios (95% confidence intervals) of colorectal cancer risk in men by dietary calcium intake categories. Table S2. Multivariable hazard ratios (95% confidence intervals) of colorectal cancer risk in women by dietary calcium intake categories. Table S3. Multivariable hazard ratios (95% confidence intervals) of colorectal cancer risk in men by dietary calcium intake categories. Table S4. Multivariable hazard ratios (95% confidence intervals) of colorectal cancer risk in women by dietary calcium intake categories. (DOCX)

Author Contributions

Analysis of the data: NM TN. Wrote the paper: NM TN. Analysed the data: NM TN. Wrote the paper: PF MJ BBM GS ER. Interpretation of results: NM TN PF MJ BBM GS ER. Commented on the analysis and interpretation of the findings: AO A. Tjønneland CCD KO MCBR FCC EN RR BK HB MMB A. Trichopoulou PL DT DP VF RT PV SP PHMP VKD EW EL JRQG RZR MJSP MD CN EA JM MA JJ RP KTK NW TJK FLG VF MJG.

References

Dairy Products and Colorectal Cancer

PLOS ONE | www.plosone.org 12 September 2013 | Volume 8 | Issue 9 | e72715


