Real world effectiveness of primary implantable cardioverter defibrillators implanted during hospital admissions for exacerbation of heart failure or other acute co-morbidities: cohort study of older patients with heart failure

Citation

Published Version
doi:10.1136/bmj.h3529

Permanent link
http://nrs.harvard.edu/urn-3:HUL.InstRepos:17820687

Terms of Use
This article was downloaded from Harvard University’s DASH repository, and is made available under the terms and conditions applicable to Other Posted Material, as set forth at http://nrs.harvard.edu/urn-3:HUL.InstRepos:dash.current.terms-of-use#LAA

Share Your Story
The Harvard community has made this article openly available. Please share how this access benefits you. Submit a story.

Accessibility
Real world effectiveness of primary implantable cardioverter defibrillators implanted during hospital admissions for exacerbation of heart failure or other acute co-morbidities: cohort study of older patients with heart failure

Chih-Ying Chen,1 Lynne Warner Stevenson,2 Garrick C Stewart,2 Deepak L Bhatt,2 Manisha Desai,3 John D Seeger,1 Lauren Williams,1 Jessica J Jalbert,1 Soko Setoguchi4

ABSTRACT

OBJECTIVES
To examine the effectiveness of primary implantable cardioverter defibrillators (ICDs) in elderly patients receiving the device during a hospital admission for exacerbation of heart failure or other acute co-morbidities, with an emphasis on adjustment for early mortality and other factors reflecting healthy candidate bias rather than the effect of the ICD.

DESIGN
Retrospective cohort study.

SETTING
Linked data from the Centers for Medicare and Medicaid Services and American College of Cardiology-National Cardiovascular Data Registry ICD registry, nationwide heart failure registry, and Medicare claims data 2004-09.

POPULATION
23 111 patients aged ≥66 who were admitted to hospital for exacerbation of heart failure or other acute co-morbidities and eligible for primary ICDs.

MAIN OUTCOME MEASURES
All cause mortality and sudden cardiac death. Latency analyses with Cox regression were used to derive crude hazard ratios and hazard ratios adjusted for high dimension propensity score for outcomes after 180 days from index implantation or discharge.

RESULTS
Patients who received an ICD during a hospital admission had lower crude mortality risk than patients who did not receive an ICD (40% v 60% at three years); however, with conditioning on 180 day survival and with adjustment for high dimension propensity score, the apparent benefit with ICD was no longer evident for sudden cardiac death (adjusted hazard ratio 0.95, 95% confidence interval 0.78 to 1.17) and had a diminished impact on total mortality (0.91, 0.82 to 1.00). There were trends towards a benefit with ICD in reducing mortality or sudden cardiac death in patients who had had a myocardial infarction more than 40 days previously, left bundle branch block, or low serum B type natriuretic peptide; however, these trends did not reach significance.

CONCLUSION
After adjustment for healthy candidate bias and confounding, the benefits of primary ICD therapy seen in pivotal trials were not apparent in patients aged 66 or over who received ICDs during a hospital admission for exacerbation of heart failure or other acute co-morbidities. Future research is warranted to further identify subgroups of elderly patients who are more likely to benefit from ICDs. Recognition of those patients whose dominant risk factors are from decompensated heart failure and non-cardiac co-morbidities will allow better focus on ICDs in those patients for whom the device offers the most benefit and provides meaningful prolonging of life.
The benefit of primary ICD therapy in landmark trials was shown in patients with heart failure in their 60s. The results from these trials, however, might not directly apply to older populations. Real world recipients of ICDs generally have more non-cardiac co-morbidities. Furthermore, primary ICD trials were conducted in outpatients with symptoms of stable mild-to-moderate heart failure. About a third of older recipients of ICDs have undergone an implantation during a hospital admission for exacerbation of heart failure or other acute co-morbidities. In patients with chronic heart failure, the early post-discharge period after an acute admission is associated with a high risk of mortality, during which progressive heart failure is the most likely cause of death. It is therefore unclear how the impact of primary ICDs on the prevention of sudden cardiac death translates to overall survival benefits among elderly patients who received the devices during acute admissions.

We examined the effectiveness of ICDs outside the previously studied populations in ICD trials. Our target population was elderly patients who received the device during acute admissions for exacerbation of heart failure or other acute co-morbidities. Assessment of the potential impact on survival with real world data requires caution because of healthy candidate bias. This type of selection bias could arise when patients at high risk of complications or deemed to be too sick to benefit are not selected for an ICD, and when patients are less interested in preventing sudden death because of the existing burden of other chronic illness. Evaluation of the clinical effectiveness of ICDs without consideration of the healthy candidate effect could overestimate its benefit. We therefore used specific design and analytic approaches to account for this bias when assessing the clinical effectiveness of ICDs implanted during admission for exacerbation of heart failure or other acute causes among elderly patients and potential differences in the effectiveness of ICDs by demographic and clinical characteristics. This included adjustment for mortality during the first 180 days, a period during which previous trials have shown no benefit of ICDs.

Methods
Data sources
We conducted a retrospective cohort study using the ICD registry of the Centers for Medicare and Medicaid Services (CMS) (2005-08); ICD registry of the American College of Cardiology-National Cardiovascular Data Registry (ACC-NCDR) (2005-08); a nationwide heart failure registry aggregated from several quality improvement and accreditation programs, including the American Heart Association’s Get With the Guidelines program (2005-08); and Medicare institutional and non-institutional claims (2004-09).

ACC-NCDR and CMS-ICD registry
The Medicare/Medicaid ICD registry is a subset of the American College of Cardiology-National Cardiovascular Data Registry’s ICD registry, which is the sole repository for data on ICD implantation for Medicare beneficiaries. Hospital personnel enter data on the registry under routine quality control review. The registry includes information on patient history, clinical characteristics, drugs, facility information, provider information, indications for ICD, device information, and inpatient complications.

Heart failure registry
The data from the national clinical registries for patients with heart failure were aggregated from several quality improvement and accreditation programs managed by Outcome Sciences using common data elements, data clarification procedures, and quality assurance practices. The aggregate database includes data from over 800 US hospitals in 50 states with close to 300,000 patients with a primary diagnosis of heart failure. The dataset has been successfully used to assess quality of care outcomes in patients with heart failure.

Medicare institutional and non-institutional files
Medicare is the national health insurance program administered by the US government. Most of its beneficiaries are aged 65 or older. The Medicare institutional and non-institutional files contain data on final claims submitted by healthcare providers. Main information contained in those files includes diagnosis and procedures, dates of service, reimbursement amount, provider identifiers, and demographic information on beneficiaries. Appendix 1 provides more details of these data sources.

Data linkage
We linked the combined ICD registry and the heart failure registry to Medicare claims data using four non-unique identifiers: date of birth, sex, admission date for implantation of the ICD, and provider identifiers, which is described in detail elsewhere and in appendix 1. Briefly, we validated this linkage among the subset of 196,923 patients who had a unique identifier in the ICD registry. Our linkage using non-unique identifiers yielded 98% specificity, 95% sensitivity, and a 98% positive predictive value compared with the linkage method using both non-unique and unique identifiers.

Study population
Our study population consisted of elderly patients with heart failure with and without ICDs who had acute hospital admissions for heart failure or any co-morbidities and were considered eligible for ICD therapy for primary prevention. The primary cohort consisted of older patients who were covered by Medicare and who could be linked to either the ICD or heart
failure registry. The secondary cohort consisted of older patients who were nested within the heart failure registry linked to Medicare data. To ensure their eligibility for primary ICD therapy, all study patients were required to have an ejection fraction of ≤35% at the time of admission. We excluded patients with cardiac arrest or sustained ventricular tachycardia, for whom ICDs would be secondary prevention. To ensure our study patients had a sufficient look-back period for assessing pertinent covariates, we required that patients had health insurance coverage for one year before the index procedure or admission. We also required patients to be aged ≥66 to ensure everyone had a one year look-back period. The exposure status (ICD implantation) was defined as a patient having a record of ICD implantation in the ICD registry. Lastly, we excluded patients who received cardiac resynchronization therapy with a defibrillator (CRT-D) as these patients met additional criteria for this indication.19-21

Outcomes

Our primary outcome was all cause mortality. The date of death was obtained from the Medicare beneficiary summary file. The secondary outcome was sudden cardiac death, defined by using a previously validated algorithm (positive predictive value 87%).22 The designation of sudden cardiac death was made if the patient was not staying at a terminal institution (that is, hospital or nursing home) on the date of death, their code for underlying cause of death was consistent with sudden cardiac death (see appendix 2), and they did not have a “terminal procedure inconsistent with unresuscitated cardiac arrest, such as radiology, thrombolysis or general anesthetic.”22 The cause of death was obtained from the National Death Index23 (appendix 3). We obtained index data on all the patients with an ICD and a randomly selected sample (70%) of those without in the primary cohort.

Latency analyses

We used latency analysis to adjust for potential healthy candidate bias, analogous to the healthy worker effect.24-26 Follow-up began after a prespecified latent period after the index date. In the current study, the index date was the date of implantation for the ICD group or the date of discharge for those without an ICD. As previous trials have shown that survival benefits of ICDs are not apparent until 1-1.5 years after implantation,6 7 we used latency periods of 180 days (primary latency period) and 365 days (secondary latency period). All patients were followed until the occurrence of an outcome event (death or sudden cardiac death) or the end of the study period (31 December 2009).

Statistical analyses

Patient characteristics were described as percentages for categorical variables by ICD exposure status. Medians and interquartile ranges or means and standard deviations were used for continuous variables. We plotted observed mortality by ICD exposure status using Kaplan-Meier estimates. Cox regression was used to derive crude hazard ratios and hazard ratios adjusted for high dimensional propensity score for outcomes using the groups without an ICD (that is, older people admitted with equivalent indication for a primary ICD but who did not receive an ICD) as a reference.

We used the high dimensional propensity score methods to adjust for surrogates of unmeasured factors to overcome residual confounding. The algorithm was used to thoroughly screen Medicare claims data to identify covariates that could collectively be surrogates for unobserved factors influencing the patient selection for ICDs.27 This allows for maximum control of potential confounders given the available information in our data sources. For example, although we did not have information on NYHA class for heart failure, we can adjust for this based on other available proxies for severity of heart failure in the dataset, including numbers of previous hospital admissions for heart failure, ejection fraction, blood pressure, B type natriuretic peptide, or co-morbidities. To calculate high dimensional propensity scores, we used data from the year before the index date. The 200 most common codes in each data dimension were identified, from which 500 likely confounders were selected based on their prevalence and potential for confounding in the study population. The scores at the index date were derived from predicted probabilities from logistic regression models containing all of the empirically identified covariates and several predefined variables: demographic characteristics; cause of index admission, admission source, admission type, and diagnostic/laboratory test results for ejection fraction, systolic blood pressure, serum sodium, serum B type natriuretic peptide, and estimated glomerular filtration rate.28

We handled variables with missing values (such as systolic blood pressure, serum sodium, B type natriuretic peptide, and creatinine) by multiple imputation29 and assumed an underlying multivariate normal distribution. Our analysis was based on five imputed datasets in which the imputation model included all variables in the outcome model (ICD use outcomes and potential confounders) as well as variables related to missingness (see appendix 4 for potential predictors of missing values included in the imputation model). We repeated all analyses in the secondary cohort.

Subgroup analyses

We assessed the heterogeneity of the effects of ICD by demographic characteristics (age, sex, and race) and three clinical characteristics (history of myocardial infarction,30 31 presence of left bundle branch block,32 33 and type natriuretic peptide concentration at index admission) using separate Cox models in each subgroup.

Age was categorized in four groups in five year increments. We used a previously validated claim
based definition to identify myocardial infarction and classified this population into three groups: recent myocardial infarction (one or more myocardial infarctions within the 40 days before the index date), old myocardial infarction (one or more myocardial infarctions 41-365 days before the index date), and no myocardial infarction within 365 days before the index admission. We also used a claim based definition to identify patients with left bundle branch block within 365 days before the index date (ICD-9 (international classification of diseases, ninth revision, clinical modification), 426.2x or 426.3x). We required the diagnosis of left bundle branch block to be made during a hospital admission. We classified our patients into two groups according to B type natriuretic peptide concentration (low v high) using a cut-off value of 800 ng/L.36-37

We adjusted for potential confounding using the high dimensional propensity score estimated from the entire cohort.38

Sensitivity analyses
We repeated all analyses in a subset of the population with complete laboratory data and a subset of the population matched on high dimensional propensity score (that is, the high dimensional propensity score matched analyses). This was done to assess the impact of the missing data assumption on our findings and the robustness of the adjustment using the high dimensional propensity score.

All analyses were conducted with SAS 9.2 (SAS Institute, Cary, NC).

Results
Study population and characteristics
We identified a cohort of 23111 patients with heart failure (5258 with an ICD and 17 853 without) who met eligibility criteria (figs 1 and 2). For over 90% of the patients, the diagnosis that led to the index admission was heart failure or other cardiac causes. Patients with ICDs were younger and were more likely to be men than patients without ICDs. They also had a lower ejection fraction, more previous admissions for cardiac diseases, and more physician visits, and their heart failure was more likely to have an ischemic cause. Patients with ICDs also had a higher prevalence of non-cardiac admissions, chronic kidney disease, metastatic cancer, lower estimated glomerular filtration rate, and higher B type natriuretic peptide (table1). These findings were similar in the secondary cohort (appendix 5).
Crude mortality risks and Kaplan-Meier curves

During follow-up (average 2.8 years, range 1 day-5 years), 12,293 (53%) patients died. The crude mortality risk among our Medicare population admitted to the hospital was 34% (95% confidence interval 33% to 35%) at one year and 56% (55% to 57%) at three years. The mortality curves for the patients with and without ICDs (fig 3) began to diverge immediately after ICD implantation (2.4% (2.0% to 2.8%) v 12.7% (12% to 13%) at 30 days). Crude mortality at one year was lower for ICD recipients than for eligible patients without an ICD (18% (17% to 19%) v 39% (38% to 40%) at one year and 40% (38% to 41%) v 60% (60% to 61%) at three years). However, the crude mortality in these hospitalized Medicare patients with an ICD at one year was similar to the mortality seen at three years in trials of ICDs in ambulatory recipients.67
non-significant 26% reduction in sudden cardiac death (0.74, 0.40 to 1.35).

Effectiveness by presence of left bundle branch block
ICD use was associated with a lower risk of mortality (hazard ratio 0.64, 95% confidence interval 0.34 to 1.17) and sudden cardiac death (0.51, 0.16 to 1.61) among patients with left bundle branch block (table 5); although the confidence intervals were wide.

Effectiveness by serum B type natriuretic peptide status
A total of 7134 patients had their admission serum B type natriuretic peptide value documented (2091 (40%) patients with ICD v 5043 (28%) without). Risks of mortality and sudden cardiac death associated with ICD therapy were numerically lower among patients with low serum B type natriuretic peptide (hazard ratios 0.86 (95% confidence interval 0.67 to 1.10) for mortality and 1.11 (0.69 to 1.77) for sudden cardiac death) than those with high type natriuretic peptide (0.94 (0.78 to 1.13) and 1.20 (0.85 to 1.69), respectively); however, these risk estimates were not significant (table 5).

Discussion
Main findings
The benefits of primary ICD therapy that had been previously shown in ambulatory patients with heart failure do not seem to translate to elderly patients who receive the device during acute hospital admissions for exacerbation of heart failure or other acute co-morbidities. Adjustment for potential confounding and healthy candidate bias11 in this specific population reduced the apparent impact of ICD therapy to a 5% reduction in sudden cardiac death and a 9% reduction in all cause mortality, which were not significant. This trend remained similar among subgroups of patients aged under 80 and across the sexes and racial subgroups. There was, however, a trend towards benefit of ICDs implanted during acute admissions to hospital in reducing mortality or sudden cardiac death in patients who had non-recent myocardial infarction more than 40 days prior to implantation, left bundle branch block, or lower serum B type natriuretic peptide, although these also did not reach significance.

Strength and limitations
Our study is the first to use latency analysis to account for the healthy candidate bias in assessing effectiveness

---

Table 2 | Number of events and incidence rates for death and sudden cardiac death in patients with and without implantable cardioverter defibrillator (ICD) in latency 180 day* and latency 365 day* analyses

<table>
<thead>
<tr>
<th>Group</th>
<th>Latency 180 day</th>
<th></th>
<th>Latency 365 day</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Death</td>
<td>Sudden cardiac death</td>
<td></td>
</tr>
<tr>
<td>Event/IR†</td>
<td>HR (95% CI)</td>
<td>Crude</td>
<td>Adjusted‡</td>
</tr>
<tr>
<td>Primary cohort</td>
<td>ICD (n=5258)</td>
<td>1307/159</td>
<td>0.69 (0.65 to 0.74)</td>
</tr>
<tr>
<td></td>
<td>No ICD (n=17 853)</td>
<td>5299/241</td>
<td>Reference</td>
</tr>
<tr>
<td>Secondary cohort</td>
<td>ICD (n=412)</td>
<td>125/201</td>
<td>0.90 (0.71 to 1.14)</td>
</tr>
<tr>
<td></td>
<td>No ICD (n=17 853)</td>
<td>5299/241</td>
<td>Reference</td>
</tr>
</tbody>
</table>

---

Table 3 | Primary and sensitivity analyses of ICD effectiveness. Figures are hazard ratios (95% CI) adjusted for high dimension propensity score (hdPS)*Complete case analyses not conducted in secondary cohort because of smaller size of cohort.

<table>
<thead>
<tr>
<th>Primary cohort</th>
<th>Secondary cohort</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>ICD/no ICD</td>
</tr>
<tr>
<td>Primary analyses</td>
<td>5258/17 853</td>
</tr>
<tr>
<td>hdPS matched analyses</td>
<td>2254/2254</td>
</tr>
<tr>
<td>Complete case analyses</td>
<td>1801/3261</td>
</tr>
<tr>
<td>Primary cohort</td>
<td>Adjusted</td>
</tr>
<tr>
<td>Secondary cohort</td>
<td>0.95 (0.78 to 1.17)</td>
</tr>
<tr>
<td>Sudden cardiac death</td>
<td>0.5 0.8 1 2</td>
</tr>
<tr>
<td>Primary cohort</td>
<td>Adjusted</td>
</tr>
<tr>
<td>Secondary cohort</td>
<td>0.95 (0.78 to 1.17)</td>
</tr>
</tbody>
</table>
### Table 4 | Effectiveness of implantable cardioverter defibrillators (ICDs) in demographic subgroups in latency 180 day* and latency 365 day* analyses

<table>
<thead>
<tr>
<th>Outcome</th>
<th>Sample size (ICD/no ICD)</th>
<th>% of eligible patients received ICD</th>
<th>Event (ICD/no ICD)</th>
<th>IR† (ICD/no ICD)</th>
<th>Adjusted HR (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>Latency 180 day</td>
</tr>
<tr>
<td>Death</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age (years):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>66-70</td>
<td>1322/2534</td>
<td>34</td>
<td>285/568</td>
<td>127/137</td>
<td>0.89 (0.70 to 1.13)</td>
</tr>
<tr>
<td>71-75</td>
<td>1347/2950</td>
<td>31</td>
<td>301/755</td>
<td>136/168</td>
<td>1.11 (0.88 to 1.40)</td>
</tr>
<tr>
<td>76-80</td>
<td>1394/3670</td>
<td>28</td>
<td>362/1068</td>
<td>172/216</td>
<td>0.92 (0.75 to 1.12)</td>
</tr>
<tr>
<td>≥81</td>
<td>1195/8699</td>
<td>12</td>
<td>359/2908</td>
<td>213/345</td>
<td>0.76§ (0.65 to 0.93)</td>
</tr>
<tr>
<td>Sex:</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Men</td>
<td>3763/8532</td>
<td>14</td>
<td>362/2520</td>
<td>152/237</td>
<td>0.90 (0.75 to 1.07)</td>
</tr>
<tr>
<td>Women</td>
<td>1495/8532</td>
<td></td>
<td>326/2520</td>
<td>152/237</td>
<td>0.90 (0.75 to 1.07)</td>
</tr>
<tr>
<td>Race:</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>White</td>
<td>4392/15068</td>
<td>23</td>
<td>1086/4402</td>
<td>562/241</td>
<td>0.90 (0.81 to 1.03)</td>
</tr>
<tr>
<td>Black</td>
<td>613/1790</td>
<td>26</td>
<td>154/594</td>
<td>172/240</td>
<td>0.93 (0.70 to 1.24)</td>
</tr>
<tr>
<td>Other</td>
<td>253/995</td>
<td>20</td>
<td>65/303</td>
<td>181/240</td>
<td>0.94 (0.60 to 1.47)</td>
</tr>
<tr>
<td>Sudden cardiac death</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age (years):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>66-70</td>
<td>1322/2534</td>
<td>34</td>
<td>73/125</td>
<td>33/30</td>
<td>1.04 (0.63 to 1.72)</td>
</tr>
<tr>
<td>71-75</td>
<td>1347/2950</td>
<td>31</td>
<td>59/176</td>
<td>27/39</td>
<td>0.81 (0.48 to 1.35)</td>
</tr>
<tr>
<td>76-80</td>
<td>1394/3670</td>
<td>28</td>
<td>99/215</td>
<td>47/44</td>
<td>1.43 (0.95 to 2.14)</td>
</tr>
<tr>
<td>≥81</td>
<td>1195/8699</td>
<td>12</td>
<td>99/810</td>
<td>59/96</td>
<td>0.74 (0.52 to 1.04)</td>
</tr>
<tr>
<td>Sex:</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Men</td>
<td>3763/9321</td>
<td>29</td>
<td>248/732</td>
<td>42/65</td>
<td>1.00 (0.78 to 1.28)</td>
</tr>
<tr>
<td>Women</td>
<td>1495/8532</td>
<td>14</td>
<td>82/594</td>
<td>35/56</td>
<td>0.86 (0.59 to 1.26)</td>
</tr>
<tr>
<td>Race:</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>White</td>
<td>4392/15068</td>
<td>23</td>
<td>288/1121</td>
<td>41/61</td>
<td>0.98 (0.79 to 1.23)</td>
</tr>
<tr>
<td>Black</td>
<td>613/1790</td>
<td>26</td>
<td>31/135</td>
<td>35/54</td>
<td>0.77 (0.40 to 1.47)</td>
</tr>
<tr>
<td>Other</td>
<td>253/995</td>
<td>20</td>
<td>11/70</td>
<td>31/55</td>
<td>0.75 (0.24 to 2.33)</td>
</tr>
</tbody>
</table>

*Latency 180 day analyses: starting follow-up from 180 days after index time; latency 365 day analyses: starting follow-up from 365 days after index time
†IR = incidence rate per 1000 person years.
‡Adjusted for high dimension propensity score.
§P < 0.05.

Fig 5 | Hazard ratios (adjusted for high dimension propensity score) for death and sudden cardiac death among demographic subgroups in latency 180 day analyses

of ICDs with observational data. Several lines of evidence suggest the existence of healthy candidate bias in the observational studies of ICD. Although the mortality curves of previous randomized trials show no ICD benefit until 1.5 years, we observed an immediate separation of the mortality curves in both our current and previous study, which is probably caused by healthy candidate bias due to selection of patients for ICD rather than immediate ICD benefit. The existence of healthy candidate bias among patients in hospital has been supported by evidence that ICD recipients had a 40-50% lower risk of adverse events, such as non-traumatic hip fracture and admission to a skilled nursing facility, than similar patients who did not receive an ICD. This bias in patient and physician selection for ICD implantation cannot be completely eliminated by adjustment for known risk factors and can lead to overestimation of the net benefit of ICDs. The disparity between groups is further suggested by a greater difference in total mortality (including non-cardiac death) than in sudden cardiac death, which is the only event that is expected to be reduced by ICDs.

Among strategies that have been developed to account for healthy candidate bias, latency analysis has been shown to be useful and is suitable to evaluate ICD effectiveness because of the delayed benefit seen in trials. This method allows a less biased evaluation by focusing on a time period more likely to reflect true ICD effectiveness.
effectiveness rather than preferential selection of ICDs for and by healthier patients. A few caveats should be noted. Healthy candidate bias can continue to influence outcomes beyond the initial chosen latency period, which is likely given our conservative 180 day latency period. Therefore, our latency analysis might still overestimate ICD survival benefit. Additionally, latency analyses could underestimate the survival benefit of ICDs if lifesaving events occurred more frequently during the prespecified latency period. Nevertheless, trial data have indicated such an underestimation is likely minimal.6 7

Several limitations need to be considered when our findings are interpreted. First, and foremost, our findings are not applicable to elderly patients who would undergo ICD implantation electively as outpatients. Our study population was limited to elderly patients who received ICDs during acute admissions to hospital for reasons other than ICD implantation. We selected this population for analysis of effectiveness, as in a previous study by Hernandez and colleagues,39 because data are available for comparison of similar patients admitted with heart failure who did not receive ICDs. Many elderly patients, however, receive the device as an elective procedure. The effectiveness of ICD among these healthier patients is more likely to be comparable with that of the trials on which the guideline recommendations are based. Additionally, our findings cannot be generalized to patients who received cardiac resynchronization therapy with their ICD (CRT-D) as this is likely to decrease heart failure events. We also did not include patients who received the ICD as secondary prevention, for which lifesaving ICD therapies are more likely to occur.

We could not identify all patients with a history of myocardial infarction, left bundle branch block, or low serum B type natriuretic peptide values because we used a claims based definition and because of missing data on B type natriuretic peptide. Therefore, we did not have a sufficient sample size to confirm the trends seen for those subgroups thought to derive more benefit from ICDs.30-33 35 Neither did we have complete information on all the recognized risk factors for death from heart failure, such as the New York Heart Association class and duration of QRS. Residual confounding is possible. Regarding the general problem of missing values in registry variables, our results were robust to the missingness assumption, as the trends seen for those subgroups thought to derive more benefit from ICDs.30-33 35 Neither did we have complete information on all the recognized risk factors for death from heart failure, such as the New York Heart Association class and duration of QRS. Residual confounding is possible. Regarding the general problem of missing values in registry variables, our results were robust to the missingness assumption, as the results based on imputed datasets were similar to complete case analyses.40

<table>
<thead>
<tr>
<th>Table 5</th>
<th>Effectiveness of implantable cardioverter defibrillators (ICDs) in clinical subgroups in latency 180 day* and latency 365 day* analyses</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Outcome=death</strong></td>
<td><strong>Sample size (ICD/no ICD)</strong></td>
</tr>
<tr>
<td><strong>Latency 180 day</strong></td>
<td></td>
</tr>
<tr>
<td><strong>Myocardial infarction:</strong></td>
<td>Recent</td>
</tr>
<tr>
<td></td>
<td>Old</td>
</tr>
<tr>
<td></td>
<td>No</td>
</tr>
<tr>
<td><strong>Left bundle branch block:</strong></td>
<td>No</td>
</tr>
<tr>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td><strong>B type natriuretic peptide:</strong></td>
<td>≤800</td>
</tr>
<tr>
<td></td>
<td>&gt;800</td>
</tr>
<tr>
<td><strong>Outcome=sudden cardiac death</strong></td>
<td><strong>Sample size (ICD/no ICD)</strong></td>
</tr>
<tr>
<td><strong>Latency 180 day</strong></td>
<td></td>
</tr>
<tr>
<td><strong>Myocardial infarction:</strong></td>
<td>Recent</td>
</tr>
<tr>
<td></td>
<td>Old</td>
</tr>
<tr>
<td></td>
<td>No</td>
</tr>
<tr>
<td><strong>Left bundle branch block:</strong></td>
<td>No</td>
</tr>
<tr>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td><strong>B type natriuretic peptide:</strong></td>
<td>≤800</td>
</tr>
<tr>
<td></td>
<td>&gt;800</td>
</tr>
</tbody>
</table>

*Latency 180 day analyses: starting follow-up from 180 days after index time; latency 365 day analyses: starting follow-up from 365 days after index time.
†IR=incidence rate per 1000 person years.
‡Adjusted for high dimension propensity score.
§P<0.05.
in randomized trials.41 More importantly, previous trials were conducted among ambulatory patients with symptoms of stable mild-to-moderate heart failure, many of whom had not previously been admitted to hospital with heart failure. Our study focused on a population of patients admitted for exacerbation of heart failure or other acute causes. This particular subset of patients has a higher baseline burden of heart failure and other co-morbidities than trial populations. Older age,42-44 advanced heart failure,42 45 46 and non-cardiac co-morbidities46 increase the likelihood of mortality that will not be prevented by an ICD, and thus present competing risks for prolonged survival with an ICD.

Four previous studies reported that primary ICD implantation in routine clinical practice was associated with a sizable survival benefit comparable with those seen in major trials,39 47-49 but only one39 included the subset of elderly patients who underwent device implantation during an acute hospital admission. In addition, two of these studies enrolled participants from outpatient cardiology clinics, where the ICD was generally implanted as an elective procedure.48 49 The previous observational studies39 48-49 also found immediate separation of survival curves at a point in time unlikely to be strongly influenced by ICDs.

Implications for practice and future research

Our results extend the understanding of the clinical effectiveness of ICDs in elderly patients admitted to hospital, who are not typical of patients in trials. Patients admitted for acute exacerbation of heart failure or other co-morbidities might be at greater risk both during ICD implantation and for non-arrhythmic events after discharge. For early risk, it is possible that there is some similarity between patients admitted for heart failure and patients early after myocardial infarction, who also showed no benefit from ICD and an increased risk of non-sudden cardiac death.30 31 50 For patients non-electively admitted to hospital for heart failure or their comorbidities, it might be appropriate to delay the decision to implant a primary ICD until they have been discharged and can be re-evaluated in the outpatient setting.

Our findings provide no reason to restrict access to ICDs for older patients with heart failure who otherwise seem similar to patients in pivotal ICD trials. The subset of elderly patients who received primary prevention ICDs in trials in outpatient settings were previously shown to derive benefit.41 While an ideal study would be a randomized trial in elderly patients, this is not likely to be performed in time to inform imminent clinical decision making. We failed to show a benefit only for those older patients receiving ICDs during an urgent admission for exacerbation of heart failure or other acute causes. Future research is warranted to identify other groups of older patients who have a high or low likelihood of benefit from ICDs to maximize the lifesaving potential of their use.

The efficacy of ICDs in women has also been questioned, in large part because they were under-represented in previous trials.6 7 51 In our study of 10 027 eligible women, we observed no heterogeneity in effectiveness of ICDs between the sexes; however, it has been shown that female candidates for primary ICD are likely have a lower risk of sudden cardiac death than male candidates.52 Women also more often experience complications with ICDs.53 54 Thus, the benefit-risk equation of ICDs among women might require further investigation.

The higher survival associated with ICDs in patients aged over 81 further emphasizes the likelihood of residual bias that cannot be adjusted for using currently reported patient characteristics. The oldest group of patients in our study comprised the lowest proportion of ICD recipients out of eligible recipients. This probably indicates a particularly rigorous selection of healthy ICD recipients in this age group, excluding patients with obvious co-morbidities and more general frailty. Therefore, despite the use of latency analysis and adjustment for high dimension propensity scores, accounting for patient and physician selection is still challenging in assessing real world clinical effectiveness.

Shared decision making regarding primary prevention ICD has been recommended to involve explicit consideration of patient preferences and the likelihood of competing risks for mortality in all patients.55 These decisions require particular scrutiny for patients admitted to hospital for exacerbation of heart failure or other acute causes. Recognition of those patients whose dominant risks are from decompensated heart failure and non-cardiac co-morbidities will allow for focused ICD therapy in those patients for whom the device offers the most benefit to provide meaningful prolonging of life.

We thank Jeptha Curtis of Yale School of Medicine and Yale-New Haven Hospital, Sherri Dodd of Medtronic, Kenneth Ellenbogen of Virginia Commonwealth University School of Medicine, Marcel E Salive of the National Institute of Aging at National Institutes of Health, and Lynette Voshage-Stahl of Boston Scientific for the guidance and expertise they contributed to this project by serving on the technical expert panel. We also thank Drew Pratt of the National Cancer Institute at the National Institutes of Health for his review of the manuscript. The views expressed in this article are those of the authors and do not represent opinions of the panel members.

This article was presented in part at the AHA Scientific Sessions 2013 on 18 November 2013 in Dallas, TX, and at the 2014 International Conference of Pharmacoepidemiology on 25 October 2014 in Taipei, Taiwan.

The views expressed in this article are those of the authors and do not represent policies of the AHRQ, CMS or the US OHHS. This manuscript was prepared while CYC was employed at Brigham and Women’s Hospital/Brigham Medical School. CYC is now a visiting scientist at the Division of Epidemiology II, Office of Pharmacovigilance and Epidemiology, Office of Surveillance and Epidemiology, Center for Drug Evaluation and Research, US Food and Drug Administration. The opinions expressed in this work are the author’s own and do not reflect the view of the Food and Drug Administration, the Department of Health and Human Services, or the US government.

Contributors: SS and LWS contributed to the acquisition of the data. CYC, SS and LWS developed the study design. CYC, SS, and LWS drafted the manuscript. All authors contributed to the interpretation of the data.
and the revision of the work, and all approved the final version to be published. CYC and SS are the guarantors. The authors of this report are responsible for its content. Statements in the report should not be construed as endorsement by the Agency for Healthcare Research and Quality or the US Department of Health and Human Services.

Funding: This project is funded by contract No HHSN290-2005-0016-I TO3 from the Agency for Healthcare Research and Quality (AHRQ), US Department of Health and Human Services (DHHS) as part of the Developing Evidence to Inform Decisions about Effectiveness (DEcIDE) program, IAA Contract 500-2010-000011 TO6 and CEA Contract 500-2010-000011 TO2 from the Centers for Medicare and Medicaid Services (CMS), US DHHS. The funding agency had no role in the design and conduct of the study and in the collection, analysis, and interpretation of the data. The manuscript was based on a report done under contract to AHRQ; AHRQ had the draft report reviewed by independent peer reviewers before acceptance of the final report.

Competing interests: All authors have completed the ICJME uniform disclosure form at http://www.icmje.org/coiDisclosure.pdf (available on request from the corresponding author) and declare no financial relationships with any organizations that might have an interest in the submitted work in the previous 3 years, no other relationships or activities that could appear to have influenced the submitted work. SS is supported by a mid-career development award grant K02-HD017731 from the AHRQ, US DHHS. She also reported receiving research support from Johnson and Johnson and receiving personal income for consulting from Sanofi-Aventis. SS has made, available online a detailed listing of financial disclosures (http://www.dcri.duke.edu/about-us/conflict-of-interest/). JDS is a paid consultant to Optum Insight and WHSCON. DBL discloses the following relationship—advisory board: Cardax, Elsevier Practice Update Cardiology, Medscape Cardiology, Regado Biosciences, board of directors: Boston VA Research Institute, Society of Cardiovascular Patient Care; chair: American Heart Association Get With The Guidelines Steering Committee; data monitoring committees: Duke Clinical Research Institute, Harvard Clinical Research Institute, Mayo Clinic, Population Health Research Institute; honorariums: American College of Cardiology (senior associate editor, Clinical Trials and News, ACC.org), Belvior Publications (editor in chief, Harvard Heart Letter), Duke Clinical Research Institute (clinical trial steering committees), Harvard Clinical Research Institute (clinical trial steering committee), HMP Communications (editor in chief, Journal of Invasive Cardiology), Journal of the American College of Cardiology (associate editor), Population Health Research Institute (clinical trial steering committee), Slack Publications (Chief Medical Editor, Cardiology Today’s Intervention), WebMD (CMC steering committees), other: Clinical Cardiology (deputy editor, research funding: Amarin, AstraZeneca, Biotronik, Bristol-Myers Squibb, Eisai, Ethicon, Forest Laboratories, Ischemix, Medtronic, Pfizer, Roche, Sanofi Aventis, St. Jude Medical, The Medicines Company; trustee: American College of Cardiology, unfunded research: FlowCo, Plexa Pharma, Takeda).

Ethical approval: This study was approved by the institutional review board of Brigham and Women’s Hospital (IRB No 20090902819).

Transparency: The lead authors, CYC and SS, affirm that this manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned and as reported in this manuscript are acceptable of the final report.

Data sharing: No additional data available.

This is an Open Access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 4.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited and the use is non-commercial. See: http://creativecommons.org/licenses/by-nc/4.0/.