RESEARCH ARTICLE

Maternal Caffeine Consumption during Pregnancy and Risk of Low Birth Weight: A Dose-Response Meta-Analysis of Observational Studies

Jongeun Rhee1*, Rockli Kim2, Yongjoo Kim2, Melanie Tam3, Yizhen Lai4, NaNa Keum5,6, Catherine Elizabeth Oldenburg6

1 Department of Environmental Health, Harvard T.H. Chan School of Public Health, Boston, Massachusetts, United States of America, 2 Department of Social and Behavioral Sciences, Harvard T.H. Chan School of Public Health, Boston, Massachusetts, United States of America, 3 Department of Global Health and Population, Harvard T.H. Chan School of Public Health, Boston, Massachusetts, United States of America, 4 Department of Health Policy and Management, Harvard T.H. Chan School of Public Health, Boston, Massachusetts, United States of America, 5 Department of Nutrition, Harvard T.H. Chan School of Public Health, Boston, Massachusetts, United States of America, 6 Department of Epidemiology, Harvard T.H. Chan School of Public Health, Boston, Massachusetts, United States of America

* jerhee@mail.harvard.edu

Abstract

Epidemiologic studies have shown inconsistent conclusions about the effect of caffeine intake during pregnancy on the risk of low birth weight (LBW). We performed a meta-analysis and linear-dose response analysis examining the association between caffeine consumption during pregnancy and risk of LBW. PubMed and EMBASE were searched for relevant articles published up to March 2014. Eight cohort and four case-control studies met all inclusion criteria. Using a random-effects model of the twelve studies, the pooled odds ratio (OR) for the risk of LBW comparing the highest versus lowest level of caffeine intake during pregnancy was 1.38 (95% CI: 1.10, 1.73). Linear dose-response analysis showed that every additional 100 mg of caffeine intake (1 cup of coffee or 2 cups of tea) per day during pregnancy was associated with a 3.0% increase in OR for LBW. There was a moderate level of overall heterogeneity with an I-squared value of 55% (95% CI: 13, 76%), and no evidence of publication bias based on Egger's test (P = 0.20) and the funnel plot. Thus, high caffeine intake during pregnancy is associated with a significant increase in the risk of LBW, and this risk appears to increase linearly as caffeine intake increases.

Introduction

A recommendation to limit caffeine intake during pregnancy was issued by the United States Food and Drug Administration in 1980 [1]. More recently, the American Congress of Obstetricians and Gynecologists reported that moderate caffeine consumption (<200 mg/day) during
pregnancy does not seem to be a major risk factor of miscarriage or preterm birth; however it was noted that the association between maternal caffeine intake and infant growth restriction remains undetermined [2].

Pregnant women have slower caffeine metabolism, with 1.5 to 3.5 times longer half-life needed to eliminate caffeine, compared to non-pregnant woman [3]. Caffeine has been detected in the amniotic fluid, umbilical cord, urine, and plasma of fetuses, which suggests that caffeine is easily transmitted across the placenta [4, 5]. The immaturity of a fetus’ liver produces a low level of enzymes necessary for caffeine metabolism, and it leaves neonates at risk of adverse outcomes including low birth weight (LBW) [6]. Infant LBW, defined as a birth weight smaller than 2,500g, is a well-established risk factor associated with several adult diseases, such as hypertension and diabetes mellitus, and obesity [7].

Epidemiologic studies have reported inconsistent conclusions about the effects of caffeine intake during pregnancy on LBW. Larroque et al [8] and Fortier et al [9] reported no association between caffeine consumption and birth weight. However, a meta-analysis of seven studies in 1998 [10] found a significant increase in the risk of LBW associated with caffeine consumption. Recently, Greenwood et al [11] found that consuming an increment of 100 mg/day of caffeine was associated with a 7% increase in the risk of LBW in a dose-response meta-analysis. These publications necessitate a revisit to the risk of maternal caffeine consumption on LBW. Hence, the purpose of our study is to systematically review the literature and perform a meta-analysis, including a dose-response analysis, on maternal caffeine consumption during pregnancy and associated risk of LBW.

Materials and Methods

Search strategy

A literature search was performed in PubMed and Embase for all studies published up to March 2014 using the Medical Subject Headings (MeSH) terms or key words pregnancy, pregnancy outcome, perinatal, maternal, caffeine, coffee, birth weight, infant, and low birth weight (S2 File). Five blinded investigators (JR, RK, YK, MT, and YL) independently reviewed the titles, abstracts and full texts using pre-specified eligibility criteria. Additionally, the reference lists of all retrieved articles and previous relevant meta-analysis/review articles were checked to identify additional studies. We followed the guidelines of Meta-analysis Of Observational Studies in Epidemiology (MOOSE) throughout the design, conduct, analysis, and reporting of this meta-analysis [12].

Study selection

We included cohort and case-control studies that examined the association between maternal caffeine intake and LBW. All sources of caffeine exposure, such as coffee, tea, cocoa/chocolate and soda drinks, were included. The outcome, LBW, was defined as birth weight smaller than 2,500 grams. We also restricted to studies providing measures of association, relative risk (RR) or odds ratio (OR) estimate and 95% confidence interval (CI), to perform meta-analysis (Fig 1). Studies presenting more than three categories of caffeine consumption, as well as category-specific number of cases and non-cases, the estimates of ORs or RRs, and associated 95% CI, were further eligible to conduct the dose-response meta-analysis. Non-English articles, literature reviews, abstracts, posters, case reports, animal studies, unpublished results, and studies examining birth complications other than LBW (preterm birth, short for gestational age, etc.) were excluded (S1 Table). Study selection was completed through discussion among the five authors.
Data extraction

From each selected study we extracted the following information: authors, publication year, study design (cohort or case-control), study type (retrospective or prospective), source of caffeine (beverage type), exposure timing (trimester), number of cases and non-cases, sample size, most fully adjusted estimates of association and corresponding 95% CIs, study region, percentage of women age \( \geq 35 \), measurement of caffeine consumption (validated or not),
measurement of outcomes (medical record or not), adjustment for potential confounders (Tables 1 and 2 and S3 Table). For articles that reported cases and non-cases only, we manually calculated the crude effect estimates and 95% CIs [13–15]. For the dose-response meta-analysis, we extracted category-specific doses of caffeine consumption (range, median), the most fully adjusted ORs and their 95% CI (S2 Table). All extracted data were cross-checked at least twice by the authors.

Statistical analysis

Firstly, meta-analysis for highest vs lowest exposure was performed. The summary ORs and 95% CIs were calculated by using random effects model, which allows for between-study variation. If consumption was reported as cups/day, we used a conversion equation of 100mg/day of caffeine for 1 cup of coffee and 50mg/day for 1 cup of tea [16]. The forest plot of the overall highest vs lowest meta-analysis is presented (Fig 2 and S3 Table) as well as the plots of subgroup analysis by cohort and case-control study designs (Figs 3 and 4). Every point in the forest plots (Figs 2, 3 and 4) is each study’s odds ratio of LBW comparing highest versus lowest maternal caffeine consumption.

A linear dose-response meta-analysis was conducted based on the assumption that there is a linear relationship between maternal caffeine consumption and LBW (Fig 5 and S2 Table). We used the Greenland and Longnecker method [17] to calculate study-specific and overall linear dose-response slopes (ORs) and their 95% CIs. The reference group was the lowest category of caffeine consumption. We calculated the median value of caffeine intakes in each exposure category and matched it to the corresponding OR. For categories with undefined upper boundary (ie. more than 5 cups of coffee per day), we calculated the median value assuming the length of the category to be no different from other defined intervals in the study. We present a spaghetti plot, developed by Ding (E Ding, Harvard School of Public Health, 2014), to summarize the linear dose-response relationship between caffeine consumption and risk of LBW (Fig 5). We used STATA version 13 (StataCorp, College Station, TX) for all statistical analysis.

Quality assessment

Potential publication bias was evaluated by the asymmetrical shape of a funnel plot (Fig 6) and by the p-value from Egger’s test [18]. As sensitivity analysis, we examined the combined risk of LBW excluding studies for which we manually calculated the effect estimates [13–15]. Similarly, we repeated our analysis after omitting studies with very discordant results from the pooled estimates.

Between-study heterogeneity was assessed using Cochran Q-statistic and quantified by Higgins I² statistic and associated 95% CI [19]. According to the Cochrane review standard, a p-value greater than 0.05 for the Q-statistic indicates the absence of heterogeneity and an I² statistic of less than 60% represents moderate or no heterogeneity. Sources of heterogeneity were explored by performing meta-regression and subgroup analyses with respect to study design (case-control, cohort), study region (North America, other regions), exposure measurement assessment (validated or not), sources of caffeine (coffee only, tea only, all types), proportion of women aged 35 (<10%, = 10%), exposure timing (first trimester, across the entire pregnancy or other trimesters), publication year (1980s, 1990s, 2000s), and the level of confounder adjustment (smoking, maternal weight, alcohol consumption, SES, race/ethnicity) (Table 3).
<table>
<thead>
<tr>
<th>Author, publication year, country</th>
<th>N cases, total N</th>
<th>Intake measurement</th>
<th>Trimester considered</th>
<th>Comparisons made by authors</th>
<th>Risk estimate</th>
<th>95% CI</th>
<th>Comparisons made in meta-analysis</th>
<th>Dose-response analysis</th>
<th>Variables adjusted for</th>
</tr>
</thead>
<tbody>
<tr>
<td>Linn et al, 1982, USA</td>
<td>927, 12205</td>
<td>Coffee, tea †</td>
<td>First</td>
<td>0 cup/d</td>
<td>1</td>
<td>&gt; 400 mg/d vs 0 mg/d</td>
<td></td>
<td>age&gt;35, parity, race, college education, smoking at delivery, alcohol in the first trimester, on welfare, previous stillbirth, previous induced abortion, previous spontaneous abortion, ponderal index</td>
<td></td>
</tr>
<tr>
<td></td>
<td>70, 3654</td>
<td>Coffee, tea, colas, drugs</td>
<td>Throughout pregnancy</td>
<td>0 mg/d</td>
<td>1</td>
<td>&gt; 300 mg/d vs 0 mg/d ✓</td>
<td></td>
<td>age, marriage, ethnicity, education, cigarettes, alcohol, marijuana use, parity, previous spontaneous abortion, previous induced abortion, previous stillbirth, weight gain, BMI</td>
<td></td>
</tr>
<tr>
<td>Martin et al, 1987, USA</td>
<td>1742, 30,445</td>
<td>Coffee †</td>
<td>Throughout pregnancy</td>
<td>0 cup/d</td>
<td>1</td>
<td>1,000 mg/d vs ✓</td>
<td></td>
<td>age, number of prior pregnancies, previous spontaneous abortion, previous LBW infant, prepregnancy weight, ethnic group, education, employment at start of pregnancy, smoking, alcohol</td>
<td></td>
</tr>
<tr>
<td>McDonald et al, 1992, Canada</td>
<td>307, 7855</td>
<td>Caffeinated coffee, tea, cola</td>
<td>Second</td>
<td>N/A</td>
<td>1</td>
<td>Highest vs Lowest</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>391, 11591</td>
<td>Coffee</td>
<td>First to Second</td>
<td>0–3 cups/d</td>
<td>1</td>
<td>&gt; 800 mg/d vs ✓</td>
<td>✓</td>
<td>smoking, social group, parity, alcohol intake</td>
<td></td>
</tr>
<tr>
<td>Olsen et al, 1991, Denmark</td>
<td>(1 cup = 100 mg caffeine)</td>
<td></td>
<td>(Continued)</td>
<td>3–7 cups/d</td>
<td>1.4</td>
<td>1.10, 1.70</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
Table 1. (Continued)

<table>
<thead>
<tr>
<th>Author, publication year, country</th>
<th>N cases, total N</th>
<th>Intake measurement</th>
<th>Trimester considered</th>
<th>Comparisons made by authors</th>
<th>Risk estimate</th>
<th>95% CI</th>
<th>Comparisons made in meta-analysis</th>
<th>Dose-response analysis*</th>
<th>Variables adjusted for</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bracken et al, 2003, USA</td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>&gt; 300 mg/d vs 0 mg/d</td>
<td>✓</td>
<td>age, parity, #prior pregnancies, marital status, race, education, height, smoking during third trimester, prepregnancy weight</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>&lt; 149 mg/d</td>
<td>1.45</td>
<td>0.89, 2.35</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>150–299 mg/d</td>
<td>1.59</td>
<td>0.70, 3.60</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>&gt; 300 mg/d</td>
<td>1.32</td>
<td>0.46, 3.78</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fortier et al, 1993, Canada</td>
<td></td>
<td></td>
<td>Throughout pregnancy</td>
<td>1</td>
<td>&gt; 300 mg/d vs 5 mg/d</td>
<td>✓</td>
<td>cigarette consumption, # of previous low birth weight newborns, family income, and parity</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>0–10 mg/d</td>
<td>1</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>&gt; 300 mg/d vs 5 mg/d</td>
<td>✓</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Bakker et al, 2010, Netherlands</td>
<td></td>
<td></td>
<td>Third</td>
<td>1</td>
<td>&lt; 2 unit/d</td>
<td>0.99</td>
<td>0.52, 1.87</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>2–3.9 unit/d</td>
<td>1.08</td>
<td>0.84 to 1.40</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>4–5.9 unit/d</td>
<td>1.19</td>
<td>0.73 to 1.95</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>&gt; 6 unit/d</td>
<td>2.58</td>
<td>1.26 to 5.30</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mills et al, 1993, USA</td>
<td></td>
<td></td>
<td>Throughout pregnancy</td>
<td>0 mg/d</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>1–99mg/d</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>100–199mg/d</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>200–299mg/d</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>1</td>
<td>&gt; 300mg/d</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Abbreviations: Confidence interval (CI)
* Check mark (✓) indicates studies that are included in the dose-response analysis
† Caffeine consumption was originally reported in cups and was converted as1 cup = 100 mg caffeine

doi:10.1371/journal.pone.0132334.t001
Results

Literature search and study selection

The searching scheme resulted in a total of 336 studies. A total of 13 studies, nine cohort [6, 9, 13–16, 20–22] and four case-control studies [23–26], were included with a total of 4,919 cases (Fig 1). Of these, three cohort [9, 13, 15] and one case-control [26] studies had not been incorporated in the recent meta-analysis [11]. Mill et al. [15] was not included for meta-analysis.
since it did not have desired outcomes as defined by our inclusion criteria. For dose-response analysis, ten studies, seven cohort [6, 9, 14, 16, 20–22] and three case-control studies [23–25] with a total of 4,499 cases, met the eligibility criteria.

**Overall effect estimate (highest versus lowest) & dose-response analysis**

The pooled estimate of the twelve studies presented that maternal caffeine consumption during pregnancy was associated with increased risk of LBW (pooled OR: 1.38, 95% CI: 1.10, 1.73). The linear dose-response analysis showed that one additional cup of coffee or two additional cups of tea per day during pregnancy was associated with a 3.0% increase in the OR for LBW (pooled OR: 1.03, 95% CI: 1.01, 1.05).
### Test of heterogeneity and subgroup analysis

The test of heterogeneity resulted in a moderate level of heterogeneity \((P = 0.01, I^2 = 55\%, 95\% \text{ CI: 13, 76\%})\). The meta-regression tests showed that none of the study characteristics, including study design, proportion of women aged 35 and older, region, exposure assessment, sources of caffeine, timing of the exposure, publication year, and confounders adjustment status, significantly modified the pooled estimate for the effect of maternal caffeine intake on the risk of LBW. The results of the meta-regression tests for heterogeneity and subgroup meta-analysis are displayed in Table 3.

### Influence test & publication bias

The sensitivity analysis showed that there were no changes in directionality and significance of the pooled ORs of high vs low meta-analysis after excluding studies with manually calculated effect estimate (Eskenazi et al. [13] \((\text{pooled OR} = 1.41, 95\% \text{ CI: 1.07, 1.85})\) and Linn et al. [14] \((\text{pooled OR} = 1.43, 95\% \text{ CI: 1.10, 1.87})\)), and most influential study (Martin et al. [20] \((\text{pooled...})))
ORA R = 1.27, 95% CI: 1.05, 1.52). Similarly, removing Linn et al. [14] did not change the directionality and significance of the OR for the linear dose-response meta-analysis (OR = 1.03 95% CI: 1.02, 1.05). However, after removing Martin et al. [20], the OR was no longer significant (OR = 1.07 95% CI = 0.18, 6.19).

Visual inspection of the funnel plot suggested that the studies were nearly symmetrically distributed around the log of the pooled estimate, and Egger’s tests formally presented that there was no significant evidence of publication bias (P = 0.20).

Discussion
This meta-analysis of twelve studies identified an overall 37.8% increase (Fig 2) in the odds of LBW among women in the highest caffeine intake group compared to those in the lowest group. A dose-response analysis based on ten studies found a 3.0% increase (Fig 5) in the odds of LBW for every 100 mg of caffeine consumed per day during pregnancy, which is equivalent to about one cup of coffee or two cups of tea. The effect size of our high vs low meta-analysis is relatively small compared to well recognized risk factors of LBW, such as active maternal smoking [27–30]. Jaddoe et al. [27] and Horta et al. [28] found active maternal smoking during pregnancy increased the risk of LBW incidence by 75% and 59%, respectively. However, the
OR of LBW among pregnant women exposed to environmental tobacco smoke (ETS) is similar to our result [29, 31, 32]. Salmasi et al. [31] and Leonardi-Bee [32] reported ETS exposure increased the risk of LBW births by 16% and 22%, respectively. Although the effect size is small in our dose-response analysis, the result is more precisely estimated (Fig 5), compared to conventional analysis (Fig 2), by using pooling data from ten studies and indicated that there is a significant association even with very small dose of exposure. Since there is no evidence of threshold effect or plateau in the linear dose-response curve (Fig 5), recommendations to pregnant women regarding caffeine intake should consider the absolute risk of increasing maternal caffeine consumption.

Our findings are in agreement with a meta-analysis conducted by Fernandes et al. [10] in 1998, which reported an increased risk of LBW among pregnant women who consumed more than 150 mg of caffeine per day. Also, Sengpiel et al. [33] reported a 21–28 g decrease in birth weight for each additional 100 mg of caffeine consumed per day. Similar results have been reported in several recent studies [34, 35], while others have not found a significant association to exist [36, 37].

---

**Fig 5.** Linear dose-response analysis on the effect of caffeine intake during pregnancy on low birth weight.

doi:10.1371/journal.pone.0132334.g005
The effect of caffeine consumption during pregnancy is of public health concern because caffeine passes through placental barriers [38]. The cytochrome P450 1A2 enzyme (CYP1A2) predominantly metabolizes caffeine [39]. Tsutsumi et al. [40] reported that CYP1A2 activity in early (8–16 weeks) and late (32–39 weeks) pregnancy is reduced by 35% and 52%, respectively. During pregnancy, the half-life of caffeine increases, which causes caffeine to be retained in the body longer [41, 42]. Caffeine can then cross the placenta and be present in the plasma of newborns [5]. Since the levels of CYP1A2 are believed to be low in the placenta and fetus [43], the fetus can be exposed to caffeine for a long period of time. The pharmacological effects of caffeine related to fetal growth are the blockade of adenosine receptors and the inhibition of cyclic nucleotide phosphodiesterase (PDE) [44]. When caffeine acts as an antagonist of adenosine receptor, adenosine is unable to regulate the local blood flow during hypoxia [9]. The acute maternal hypoxia can negatively impact the fetal cardiovascular function and fetal growth [45]. Also, when PDE is inhibited by caffeine, the levels of cyclic adenosine monophosphate (cAMP) will be increased because PDE degrades cAMP, which may interfere with fetal growth. For example, Bistoletti et al [46] explained that fetal asphyxia is associated with higher cyclic AMP levels.

Fig 6. Funnel plot of meta-analysis on the effect of caffeine intake during pregnancy on low birth weight.

doi:10.1371/journal.pone.0132334.g006
Table 3. Subgroup meta-analyses of maternal caffeine intake during pregnancy and LBW.

<table>
<thead>
<tr>
<th>Study characteristics</th>
<th># of Studies</th>
<th>N of Cases</th>
<th>Total N</th>
<th>Pooled OR</th>
<th>95% CI</th>
<th>P for Heterogeneity</th>
<th>I²</th>
<th>95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Study design (P for interaction = 0.54)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cohort</td>
<td>8</td>
<td>3,887</td>
<td>74,885</td>
<td>1.43</td>
<td>1.14, 1.79</td>
<td>0.06</td>
<td>49%</td>
<td>0, 77%</td>
</tr>
<tr>
<td>Case-control</td>
<td>4</td>
<td>703</td>
<td>2,156</td>
<td>1.3</td>
<td>0.66, 2.60</td>
<td>0.05</td>
<td>62.00%</td>
<td>0, 87%</td>
</tr>
<tr>
<td>Region (P for interaction = 0.37)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>North America (US, Canada)</td>
<td>8</td>
<td>3,713</td>
<td>63,402</td>
<td>1.48</td>
<td>1.15, 1.91</td>
<td>0.08</td>
<td>44%</td>
<td>0, 75%</td>
</tr>
<tr>
<td>Other regions</td>
<td>4</td>
<td>877</td>
<td>16,086</td>
<td>1.19</td>
<td>0.72, 1.95</td>
<td>0.03</td>
<td>67.70%</td>
<td>0, 89%</td>
</tr>
<tr>
<td>Aged&gt; = 35 composition (P for interaction = 0.84)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;10%</td>
<td>6</td>
<td>893</td>
<td>23,816</td>
<td>1.52</td>
<td>1.04, 2.22</td>
<td>0.03</td>
<td>58.60%</td>
<td>0, 83%</td>
</tr>
<tr>
<td>&gt;10%</td>
<td>4</td>
<td>3,303</td>
<td>53,280</td>
<td>1.32</td>
<td>1.08, 1.62</td>
<td>0.73</td>
<td>0%</td>
<td>0, 85%</td>
</tr>
<tr>
<td>Exposure validation (P for interaction = 0.86)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>validated</td>
<td>3</td>
<td>516</td>
<td>6,684</td>
<td>1.29</td>
<td>0.81, 2.04</td>
<td>0.42</td>
<td>0%</td>
<td>0, 90%</td>
</tr>
<tr>
<td>not validated</td>
<td>9</td>
<td>4,074</td>
<td>72,804</td>
<td>1.4</td>
<td>1.08, 1.83</td>
<td>0</td>
<td>64.50%</td>
<td>27, 83%</td>
</tr>
<tr>
<td>Sources of caffeine (P for interaction = 0.55)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>All</td>
<td>9</td>
<td>2,365</td>
<td>35,349</td>
<td>1.49</td>
<td>1.07, 2.06</td>
<td>0</td>
<td>66.10%</td>
<td>31, 83%</td>
</tr>
<tr>
<td>Coffee only</td>
<td>2</td>
<td>2,133</td>
<td>42,036</td>
<td>1.31</td>
<td>1.03, 1.67</td>
<td>0.48</td>
<td>-</td>
<td></td>
</tr>
<tr>
<td>Tea only</td>
<td>1</td>
<td>92</td>
<td>2,103</td>
<td>0.99</td>
<td>0.32, 3.06</td>
<td>-</td>
<td>-</td>
<td></td>
</tr>
<tr>
<td>Timing of the exposure (P for interaction = 0.70)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First trimester</td>
<td>3</td>
<td>3,338</td>
<td>60,785</td>
<td>1.55</td>
<td>0.92, 2.62</td>
<td>0.2</td>
<td>38.80%</td>
<td>0, 81%</td>
</tr>
<tr>
<td>All or other trimesters</td>
<td>8</td>
<td>1,252</td>
<td>18,703</td>
<td>1.38</td>
<td>1.03, 1.85</td>
<td>0</td>
<td>66.30%</td>
<td>29, 84%</td>
</tr>
<tr>
<td>Publication Year (P for interaction = 0.78)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1980s</td>
<td>2</td>
<td>997</td>
<td>15,859</td>
<td>2.19</td>
<td>0.58, 8.36</td>
<td>0.003</td>
<td>89%</td>
<td>-</td>
</tr>
<tr>
<td>1990s</td>
<td>7</td>
<td>3,393</td>
<td>59,234</td>
<td>1.23</td>
<td>0.97, 1.55</td>
<td>0.1</td>
<td>44.50%</td>
<td>0, 77%</td>
</tr>
<tr>
<td>2000s</td>
<td>3</td>
<td>200</td>
<td>4,395</td>
<td>1.72</td>
<td>0.96, 3.09</td>
<td>0.31</td>
<td>15.80%</td>
<td>0, 91%</td>
</tr>
<tr>
<td>Adjustment for alcohol drinking (P for interaction = 0.08)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>7</td>
<td>3,368</td>
<td>62,532</td>
<td>1.7</td>
<td>1.24, 2.34</td>
<td>0.02</td>
<td>59.40%</td>
<td>7, 82%</td>
</tr>
<tr>
<td>No</td>
<td>5</td>
<td>1,222</td>
<td>16,956</td>
<td>1.05</td>
<td>0.78, 1.42</td>
<td>0.21</td>
<td>32.10%</td>
<td>0, 74%</td>
</tr>
<tr>
<td>Adjustment for smoking (P for interaction = 0.88)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

(Continued)
Of the studies included in this meta-analysis, four [6, 13, 20, 21] reported a significant positive association between caffeine intake and LBW while eight found no significant association [9, 14, 16, 22–26]. In case-control studies, mothers with an LBW outcome are more likely to report their caffeine consumption to be less than true amount, which may attenuate the effect estimate towards the null or bias it towards inverse association [47]. Attenuated linear dose-response line was shown (Fig 5) rather than strong linear line when mothers who consumed high levels of caffeine under-reported their actual consumption. This could explain why we observed different results for case-control and cohort studies in our high vs low meta-analysis. Among the cohort studies, there was a significant increase in risk of LBW among mothers who consumed higher amounts of caffeine. Among the case-control studies, the effect was no longer significant, although the directionality of the association was consistent.

Overall, there was a moderate level of heterogeneity between the studies, but we were not able to identify significant sources of heterogeneity after performing meta-regression tests based on different study characteristics. However, removing the most influential study, Martin et al [20], reduced the level of heterogeneity from I-squared of 55% to 34%. After reviewing its study design and adjusted confounders, we were not able to identify characteristics of this study [20] that may have caused it to be particularly influential. Removing another study that did not provide an adjusted OR, Eskenazi et al. [13] Linn et al. [14], did not change the directionality or significance of the results.

This meta-analysis also considered possible confounders of the relationship between caffeine intake and LBW by using adjusted ORs reported by the studies (Tables 1 and 2). The most common confounders adjusted for were smoking status, parity, alcohol use, and ethnicity. Maternal age, weight, and trimester were also adjusted for in many studies. Levels of smoking,

### Table 3. (Continued)

<table>
<thead>
<tr>
<th>Study characteristics</th>
<th># of Studies</th>
<th>N of Cases</th>
<th>Total N</th>
<th>Pooled OR</th>
<th>95% CI</th>
<th>P for Heterogeneity</th>
<th>$\bar{b}$</th>
<th>95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Yes</td>
<td>11</td>
<td>307</td>
<td>7,855</td>
<td>1.41</td>
<td>1.07, 1.85</td>
<td>0.01</td>
<td>58.70%</td>
<td>20, 79%</td>
</tr>
<tr>
<td>No</td>
<td>1</td>
<td>4,283</td>
<td>71,633</td>
<td>1.34</td>
<td>1.02, 1.76</td>
<td>-</td>
<td>-</td>
<td></td>
</tr>
<tr>
<td><strong>Adjustment for maternal weight (P for interaction = 0.39)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>5</td>
<td>2,487</td>
<td>39,791</td>
<td>1.19</td>
<td>0.76, 1.85</td>
<td>0.08</td>
<td>52.90%</td>
<td>0, 83%</td>
</tr>
<tr>
<td>No</td>
<td>7</td>
<td>2,103</td>
<td>39,697</td>
<td>1.5</td>
<td>1.13, 2.00</td>
<td>0.02</td>
<td>58.70%</td>
<td>5, 82%</td>
</tr>
<tr>
<td><strong>Adjustment for Ethnicity (P for interaction = 0.32)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>8</td>
<td>3,479</td>
<td>55,625</td>
<td>1.62</td>
<td>1.11, 2.36</td>
<td>0</td>
<td>69.60%</td>
<td>37, 85%</td>
</tr>
<tr>
<td>No</td>
<td>4</td>
<td>1,111</td>
<td>23,863</td>
<td>1.24</td>
<td>1.02, 1.52</td>
<td>0.81</td>
<td>0.00%</td>
<td>0, 85%</td>
</tr>
<tr>
<td><strong>Adjustment for SES (P for interaction = 0.52)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>7</td>
<td>3,602</td>
<td>61,381</td>
<td>1.28</td>
<td>1.08, 1.53</td>
<td>0.5</td>
<td>0.00%</td>
<td>0, 71%</td>
</tr>
<tr>
<td>No</td>
<td>5</td>
<td>988</td>
<td>18,107</td>
<td>1.71</td>
<td>0.96, 3.05</td>
<td>0.3</td>
<td>78.90%</td>
<td>50, 91%</td>
</tr>
</tbody>
</table>

Abbreviations: Confidence interval (CI)

doi:10.1371/journal.pone.0132334.t003
alcohol use, and maternal age have previously been found to have a positive correlation with
levels of caffeine consumption [48].

This meta-analysis has a few limitations. Frequency, quantity, and sources of caffeine intake
during pregnancy were self-reported by the mothers or expectant mothers in all accepted stud-
ies. Thus, recall bias might have occurred in terms of exposure assessment, especially for the
seven retrospective studies [9, 14, 21, 23–26]. Most of the studies did not conduct exposure
measurement validation. However, Fenster et al. [24] and Bracken et al. [16] reported that
retrospective data was reliable. Fenster et al. [24] found that 77% of the respondents could
reproduce their caffeine consumption record from 6 months earlier within one cup of coffee.
Bracken et al. [16] observed that prospective data collected at 28 and 36 weeks of pregnancy
were similar to respective data gathered retrospectively. It is not possible to perform individual
exposure quality assessment, but misclassification of exposure would have likely occurred ran-
domly, causing non-differential bias. Exposure validation was not found to be a source of het-
erogeneity in our pooled estimates.

Another limitation is that studies differed in the units used to measure caffeine intake. The
amount of caffeine consumption was measured in cups per day by four studies [14, 21, 22, 26],
while the other studies reported milligrams per day. To make these estimates comparable for
dose-response analysis, we used conversion factors of 100mg/day for 1 cup of coffee and 50mg/
day for 1 cup of tea. Since each study had different conversion factors, exposure levels may
have been slightly under- or overestimated. For example, Fenster et al. [24] assumed a caffeine
content of 107mg/cup of coffee and 34mg/cup of tea, and Bakker et al. [6] estimated that 1 cup
of caffeinated coffee contains 90mg caffeine However, we believe that an approximate differ-
ence of 10mg for estimating one cup of coffee or tea would not lead to clinically relevant differ-
ences in the results.

A strength of this meta-analysis is the extensiveness of the literature search; all studies ever
conducted prior to March 2014 were considered for inclusion. Our findings not only reaffirm
the recent meta-analysis results [11, 49] that identified an inverse relationship between caffeine
consumption and birth weight, but also induce a comprehensive conclusion by including more
studies. The quality of outcome measurement in all studies was reliable as they were collected
by birth certificate, from hospitals, or were confirmed by physicians. Furthermore, evaluation
of the funnel plot and Egger’s test found no evidence of publication bias.

The results of our meta-analysis suggest that consuming 100 mg of caffeine per day may
lead to a small, but significant, increase in the risk of LBW. It is important that women are
aware of the effects that caffeine may have on their infants’ health and make appropriate
adjustments to their levels of consumption.

**Supporting Information**

S1 File. PRISMA 2009 Checklist.

(DOC)

S2 File. Medical Subject Heading (MeSH) terms and title/abstract (tiab) keywords in
PubMed and Embase search strategy.

(DOCX)

S1 Table. The reasons of excluding 27 articles.

(TIF)

S2 Table. Data used for dose-response analysis.

(TIF)
S3 Table. Data used for meta-analysis (highest vs lowest).

(TIF)

Acknowledgments

We would like to acknowledge and thank the contribution of Dr. C. Chung Hsieh, Dr. Julie Goodman, and Dr. David Christiani from Harvard T.H. Chan School of Public Health for their support and comments on revising the manuscript.

Author Contributions

Conceived and designed the experiments: JR RK YK MT YL. Analyzed the data: JR YK. Contributed reagents/materials/analysis tools: JR RK NK CEO. Wrote the paper: JR RK YK MT YL.

References


