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Traditional Risk Factors Versus Biomarkers for Prediction of Secondary Events in Patients With Stable Coronary Heart Disease: From the Heart and Soul Study

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Background—Patients with stable coronary heart disease (CHD) have widely varying prognoses and treatment options. Validated models for risk stratification of patients with CHD are needed. We sought to evaluate traditional and novel risk factors as predictors of secondary cardiovascular (CV) events, and to develop a prediction model that could be used to risk stratify patients with stable CHD.

Methods and Results—We used independent derivation (912 participants in the Heart and Soul Study) and validation (2876 participants in the PEACE trial) cohorts of patients with stable CHD to develop a risk prediction model using Cox proportional hazards models. The outcome was CV events, defined as myocardial infarction, stroke, or CV death. The annual rate of CV events was 3.4% in the derivation cohort and 2.2% in the validation cohort. With the exception of smoking, traditional risk factors (including age, sex, body mass index, hypertension, dyslipidemia, and diabetes) did not emerge as the top predictors of secondary CV events. The top 4 predictors of secondary events were the following: N-terminal pro-type brain natriuretic peptide, high-sensitivity cardiac troponin T, urinary albumin:creatinine ratio, and current smoking. The 5-year C-index for this 4-predictor model was 0.73 in the derivation cohort and 0.65 in the validation cohort. As compared with variables in the Framingham secondary events model, the Heart and Soul risk model resulted in net reclassification improvement of 0.47 (95% CI 0.25 to 0.73) in the derivation cohort and 0.18 (95% CI 0.01 to 0.40) in the validation cohort.

Conclusions—Novel risk factors are superior to traditional risk factors for predicting 5-year risk of secondary events in patients with stable CHD. (J Am Heart Assoc. 2015;4:e001646 doi: 10.1161/JAHA.114.001646)

Key Words: coronary disease • epidemiology • prevention • risk prediction

An estimated 15.5 million adults in the United States live with coronary heart disease (CHD).1 With advances in the treatment of acute coronary syndromes and aggressive risk factor management, patients now live longer with chronic CHD, and secondary prevention has become a major focus.2 To date, predicting risk of secondary events has received little attention because all patients with stable CHD are recommended to receive similar treatment with lipid-lowering medications, antiplatelet agents, β-blockers, angiotensin inhibitors, smoking cessation, and glycemic control, regardless
of disease severity. However, given the widely varying prognoses\textsuperscript{3} and expanding range of therapeutic options for patients with chronic CHD, such as novel antiplatelet agents and revascularization procedures, it has become increasingly necessary to define distinct risk groups for whom different treatment strategies may be preferred.\textsuperscript{4}

Most cardiovascular (CV) risk models have focused on predicting incident CHD\textsuperscript{5,6} or outcomes after acute coronary syndromes.\textsuperscript{7} In these models, traditional risk factors, such as age, gender, smoking, hypertension, cholesterol, and diabetes, have remained the cornerstone of risk stratification.\textsuperscript{8} However, it is increasingly recognized that risk factors for incident CHD may not predict secondary events in patients with prevalent CHD, in part because better risk factor profiles may reflect more aggressive control of sicker patients.\textsuperscript{9} Once clinical CHD exists, markers of end-organ damage may be more important than risk factors for incident disease. Thus, new approaches for risk prediction in patients with stable CHD are needed.

Many novel risk factors have been shown to provide incremental prognostic information in patients with stable CHD.\textsuperscript{10–19} Likewise, combining traditional risk factors with symptom severity, ejection fraction, and standard laboratory test results can improve prediction of secondary events.\textsuperscript{20–27} However, unlike the Framingham risk score for developing incident CHD, the few existing secondary prevention models have not become widely accepted for use in clinical practice,\textsuperscript{19,20,22,23,27} in part because of model complexity and lack of external validation. Simple integrated prediction models that capture risk of secondary events are lacking.

Our objectives were to evaluate traditional risk factors and novel biomarkers as predictors of 5-year risk of secondary events, and to develop a prediction model that could be used to risk stratify patients with chronic stable CHD.

Methods

Derivation Cohort

The Heart and Soul Study is a prospective cohort study that was originally designed to investigate the effect of psychosocial factors on prognosis in patients with stable CHD. Methods have been previously described.\textsuperscript{28} Subjects were eligible if they met one of the following criteria: (1) history of myocardial infarction, (2) history of coronary revascularization, (3) \( \geq 50\% \) angiographic stenosis in at least 1 coronary artery, or (4) exercise-induced ischemia by treadmill ECG or nuclear perfusion imaging. Exclusion criteria were the following: (1) history of myocardial infarction within the past 6 months, (2) inability to walk 1 block, or (3) intention to move out of the local area within 3 years. The protocol was approved by the appropriate institutional review boards, and all participants provided written informed consent. Between September 2000 and December 2002, 1024 participants enrolled and completed a baseline study visit that included an interview, questionnaire, 12-hour fasting blood draw, and echocardiogram. Serum was stored at \(-70^\circ\text{C}\). The current analysis was restricted to the 912 participants with complete baseline data. The last date of follow-up was June 14, 2011.

Model Covariates

Eighteen candidate predictors were chosen based on clinical judgment, prior literature, and reproducibility of tests (Table 1). Age, sex, current smoking, medication use, history of myocardial infarction, and history of heart failure were collected by self-report questionnaire. Diabetes was defined by self-report, fasting glucose \( \geq 126 \text{ mg/dL} \), or taking diabetic medications. Hypertension was defined by self-report, systolic blood pressure \( \geq 140 \text{ mm Hg} \), or diastolic blood pressure \( \geq 90 \text{ mm Hg} \). Blood pressure was measured by trained study personnel using calibrated sphygmomanometers. Body mass index was calculated as weight in kilograms divided by height in meters squared. Left ventricular ejection fraction was measured by echocardiography. Physical activity and medication adherence were assessed by self-report.\textsuperscript{28,29}

Estimated glomerular filtration rate was calculated using the combined (serum creatinine and cystatin C) CKD-Epi equation.\textsuperscript{30} Urinary albumin and creatinine were measured from 24-hour urine collections by nephelometry and the rate Jaffe method, respectively, and urinary albumin to creatinine ratio (uACR) was calculated as milligrams of albumin divided by grams of creatinine. Fasting total and high-density lipoprotein cholesterol were measured, and low-density lipoprotein cholesterol was calculated by the Friedewald equation. High-sensitivity C-reactive protein was measured by the Roche Integra assay or the Beckman Extended Range analyzer. N-terminal pro-type brain natriuretic peptide (NT-proBNP) was measured with the Roche Elecsys NT-proBNP assay.

Outcome Variables

The composite outcome variable was time to first nonfatal myocardial infarction, stroke, or CV death. Annual telephone interviews with participants or their proxies were conducted regarding recent emergency room visits, hospitalizations, or death. For any reported event, 2 independent and blinded
adjudicators reviewed medical records, ECGs, death certificates, and coroner’s reports. If the adjudicators agreed on the outcome classification, their classification was binding. If they disagreed, a third blinded adjudicator reviewed the event and determined the outcome classification. Nonfatal myocardial infarction and stroke were defined using standard criteria. CV death was defined as due to myocardial infarction, sudden death, heart failure, coronary revascularization, stroke, or peripheral vascular disease.

**External Validation Cohort**

We are not aware of any other prospective observational cohort study of patients with stable CHD that has measured all 3 biomarkers (uACR, troponin, and NT-proBNP) that were selected in our derivation sample. However, the PEACE trial was a randomized trial of trandolapril versus placebo in 8290 patients with stable CHD and left ventricular ejection fraction >40% who were enrolled from November 1996 to June 2000 and were followed for up to 7 years (median 4.8 years) (last date of follow-up December 31, 2003). Of these, 2876 participants had measurements of smoking status, NT-proBNP, hs-cTnT, and uACR, as well as detailed adjudication of CV events (urine samples for uACR were available for only 2977 participants. Additionally, 80 were excluded due to lack of serum biomarker measurements, 3 due to lack of smoking status, and 8 due to lack of follow-up).

**Statistical Analysis**

Cox proportional hazards models were developed for risk stratification. At the outset, we omitted the 4 weakest predictors—low-density lipoprotein, high-density lipoprotein (HDL), prior history of myocardial infarction, and hypertension (all \( P \)-values >0.5). We then exhaustively screened candidate models with 4 to 11 predictors and potential

### Table 1. Characteristics of Patients With or Without Subsequent MI, Stroke, or CV Death

<table>
<thead>
<tr>
<th>Candidate Predictor Variables</th>
<th>Derivation Cohort (n=912)</th>
<th>Validation Cohort (n=2876)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Case</td>
<td>Non-Case</td>
<td>Case</td>
</tr>
<tr>
<td>Age, y</td>
<td>70±12</td>
<td>66±11</td>
</tr>
<tr>
<td>Male</td>
<td>178 (88%)</td>
<td>572 (81%)</td>
</tr>
<tr>
<td>BMI, kg/m²</td>
<td>28.2±5.4</td>
<td>28.5±5.4</td>
</tr>
<tr>
<td>Current smoking</td>
<td>42 (21%)</td>
<td>137 (19%)</td>
</tr>
<tr>
<td>Diabetes</td>
<td>78 (39%)</td>
<td>212 (30%)</td>
</tr>
<tr>
<td>Hypertension</td>
<td>168 (83%)</td>
<td>538 (76%)</td>
</tr>
<tr>
<td>History of myocardial infarction</td>
<td>123 (61%)</td>
<td>367 (52%)</td>
</tr>
<tr>
<td>History of congestive heart failure†</td>
<td>51 (25%)</td>
<td>108 (15%)</td>
</tr>
<tr>
<td>Medication nonadherence</td>
<td>20 (10%)</td>
<td>52 (7%)</td>
</tr>
<tr>
<td>Physical inactivity</td>
<td>90 (45%)</td>
<td>241 (34%)</td>
</tr>
<tr>
<td>High-sensitivity troponin T, pg/mL‡</td>
<td>16.5 (9.6 to 26.0)</td>
<td>8.6 (5.3 to 14.0)</td>
</tr>
<tr>
<td>LVEF &lt;50%</td>
<td>40 (20%)</td>
<td>68 (10%)</td>
</tr>
<tr>
<td>LDL-C, mg/dL*</td>
<td>103±34</td>
<td>105±34</td>
</tr>
<tr>
<td>HDL-C, mg/dL*</td>
<td>43±14</td>
<td>46±14</td>
</tr>
<tr>
<td>C-reactive protein, mg/L†</td>
<td>2.7 (1.4 to 6.3)</td>
<td>2.0 (0.8 to 4.5)</td>
</tr>
<tr>
<td>NT-proBNP, pg/mL‡</td>
<td>400 (147 to 1087)</td>
<td>141 (64 to 324)</td>
</tr>
<tr>
<td>BNP, pg/mL†,§</td>
<td>275 (107 to 887)</td>
<td>123 (50 to 296)</td>
</tr>
<tr>
<td>eGFR, mL/min per 1.73 m²‡</td>
<td>62±23</td>
<td>74±20</td>
</tr>
<tr>
<td>Urine albumin:creatinine ratio, mg/g‡</td>
<td>14.0 (7.8 to 56.3)</td>
<td>7.9 (4.7 to 14.6)</td>
</tr>
</tbody>
</table>

BMI indicates body mass index; CV, cardiovascular; eGFR, estimated glomerular filtration rate; HDL-C, high-density lipoprotein cholesterol; LDL-C, low-density lipoprotein cholesterol; LVEF, left ventricular ejection fraction; MI, myocardial infarction; NT-proBNP, N-terminal pro-type brain natriuretic peptide.

*Mean±standard deviation.
†History of congestive heart failure was an exclusion criterion in the validation cohort.
‡Median (interquartile range).
§BNP was not among the 18 candidate predictors; in a sensitivity analysis, BNP replaced NT-proBNP.
interactions using 20 repetitions of 10-fold cross-validation of Harrell’s C-index, a measure of discrimination.\textsuperscript{37} Internal cross-validation was used to minimize overfitting. Non-normal continuous variables were log-transformed to improve linearity.

After the model selection algorithm was run, we selected the top 4-predictor model based on cross-validated C-index (CVCI). Although they had slightly better discrimination (higher C-indices), models incorporating more than 4 predictors included predictors that were not independently associated with CV events in multivariate models. Including additional variables did not meaningfully improve C-index or calibration. Calibration was assessed by visual comparison of observed and model-based risk by decile of model-based risk and informal comparison of the average model-based 5-year event risk with nonparametric Kaplan–Meier estimates. We also assessed model fit with a likelihood ratio test for addition of deciles of model-based risk to the continuous model. In addition, we evaluated the cross-validated calibration slope\textsuperscript{38} as a measure of overfitting.

The Heart and Soul model was externally validated in 2876 participants from the PEACE trial. We calculated the C-index for the validation cohort and fitted 5-year risks using Cox model coefficients and 5-year baseline survival estimates from the derivation cohort.

To compare the relative strength of individual variables in the top model, we calculated standardized hazard ratios (per standard deviation increase in log-transformed variable), cross-validated C-indices for each variable alone, and net reclassification of each variable added to a base model containing the other 3 variables.

Category-free net reclassification improvement\textsuperscript{39} was used to compare our model to the Framingham secondary events model for stable CHD.\textsuperscript{19} To optimize performance of the Framingham model, we used a model in which coefficients for the Framingham predictor variables were refit to the derivation cohort. In the validation cohort, we substituted ln(total cholesterol) for ln(total cholesterol/HDL) because HDL was not measured, and re-estimated the coefficients of both models. We also compared the Framingham secondary events model and the Heart and Soul model C-indices in both derivation and validation cohorts, and estimated bootstrap confidence intervals using 500 repetitions. All statistical analyses were performed using STATA 11.0, 12.1, or 13.1 (Stata Corp, College Station, TX).

<table>
<thead>
<tr>
<th>Event</th>
<th>Derivation Cohort (N=912)</th>
<th>Validation Cohort (N=2876)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Myocardial infarction</td>
<td>109 1.8</td>
<td>164 1.2</td>
</tr>
<tr>
<td>Stroke</td>
<td>37 0.6</td>
<td>58 0.4</td>
</tr>
<tr>
<td>Cardiovascular death</td>
<td>106 1.7</td>
<td>98 0.7</td>
</tr>
<tr>
<td>Any event</td>
<td>202 3.4</td>
<td>304 2.2</td>
</tr>
</tbody>
</table>

Discrimination

After screening for CVCI, the final model selected from the derivation cohort, the Heart and Soul risk model, contained, in order of importance: NT-proBNP, hs-cTnT, uACR, and current smoking. Larger models of up to 7 predictors only marginally improved discrimination, and even larger models beyond 7 predictors actually performed worse than the more parsimonious models. The 4-predictor Heart and Soul model had a CVCI of 0.73 (95% CI 0.69 to 0.76). Notably, traditional CV risk factors such as age, hypertension, low-density lipoprotein, diabetes, and obesity did not appear in the top predictive models. Including interactions between sex and other predictors did not increase the CVCI. Replacing NT-proBNP with BNP resulted in minimally changed CVCI of 0.73 (95% CI 0.69 to 0.76). We also performed a sensitivity analysis adding baseline medications (aspirin, \(\beta\)-blockers, statins, angiotensin-converting enzyme inhibitors, angiotensin receptor blockers, and diuretics) to the model, and did not find any improvement.

Calibration

Observed estimates of 5-year event incidence were similar to predicted risk (Figure – Panels A and B) in the derivation cohort. The likelihood-ratio goodness-of-fit \(P\)-value was 0.07, and the cross-validated calibration slope was 0.96, indicating acceptable fit.

External Validation

In the validation cohort, C-index was 0.65 (95% CI 0.61 to 0.68), the goodness-of-fit \(P\)-value was 0.13, and the

Results

The 912 participants in the Heart and Soul derivation cohort contributed 5979 person-years with a mean (SD) of 6.6 (2.8) years follow-up. The 2876 participants in the PEACE validation cohort contributed 13,806 person-years with a mean (SD) of 4.8 (1.4) years follow-up. Compared to the derivation cohort, the validation cohort was significantly healthier at baseline (Table 1) and had lower annual rates of CV events (Table 2).
calibration slope was 0.85. Calibration is visually demonstrated in Figure – Panels C and D.

Comparison of Individual Variables

Standardized hazard ratios for the 4 individual predictors in the Heart and Soul model showed that NT-proBNP was the strongest biomarker (Table 3), followed by hs-cTnT and then uACR. The top model with 5 predictors included the 4 predictors in the Heart and Soul model with the addition of high-sensitivity C-reactive protein, but high-sensitivity C-reactive protein was not significantly associated with CV events after multivariable adjustment for NT-proBNP, hs-cTnT, uACR, and smoking (hazard ratio 1.11 per SD increase, 95% CI 0.97 to 1.27; *P*=0.15). The hazard ratios of all the biomarkers were weaker in the validation cohort compared to the derivation cohort. C-indices of each individual variable for prediction of CV events and net reclassification of each variable also showed consistent results in terms of relative strength of variables as well as weaker performance of the biomarkers in the validation cohort (Table 4).

Comparison of Heart and Soul Model to Framingham (Traditional) Secondary Events Model

The Framingham secondary events model contains age, diabetes, and ln(total cholesterol/HDL) for men, plus an additional 2 variables for women: ln(SBP) and smoking. There was a statistically significant improvement in C-index between the Framingham secondary events model and the Heart and Soul risk model in both cohorts. In the derivation cohort, the
CVCI improved from 0.61 (95% CI 0.55 to 0.64) to 0.73 (95% CI 0.69 to 0.76), an absolute improvement of 0.12 (95% CI 0.08 to 0.19; \(P < 0.001\)). In the validation cohort, the C-index improved from 0.57 (95% CI 0.52 to 0.61) to 0.65 (95% CI 0.61 to 0.68), an absolute improvement of 0.08 (95% CI 0.04 to 0.13; \(P = 0.001\)). When compared with the Framingham secondary events model, the Heart and Soul model resulted in significant reclassification improvement, and categorized patients into a wide range of risk groups, highlighting the heterogeneity of disease course in patients with stable CHD. Traditional risk factors for incident CHD, most notably age, did not emerge as the top predictors of secondary events in patients with prevalent CHD.

Previously developed risk prediction models have incorporated an array of risk factors with varying abilities to discriminate risk. The Framingham secondary events model relies upon traditional risk factors, including age, gender, blood pressure, cholesterol, smoking, and diabetes, but results in poor risk discrimination for CV events.\(^{19}\) Other risk models of up to 16 risk factors, including patient-reported symptoms, left ventricular ejection fraction, and standard laboratory measurements have been developed.\(^{20–27}\) However, most of these models are too complex for routine use, and only 1 has been externally validated.\(^{27}\) Since prediction

### Table 3. Multivariable-Adjusted HR for MI, Stroke, or CV Death in the Heart and Soul Risk Model

<table>
<thead>
<tr>
<th>Variable</th>
<th>Derivation Cohort</th>
<th>Validation Cohort</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Adjusted* HR (95% CI)</td>
<td>(P) Value</td>
</tr>
<tr>
<td>NT-proBNP(^{†})</td>
<td>1.40 (1.24, 1.58)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>hs-cTnT(^{†})</td>
<td>1.65 (1.32, 2.06)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>uACR(^{†})</td>
<td>1.18 (1.08, 1.28)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Current smoking</td>
<td>1.57 (1.11, 2.22)</td>
<td>0.01</td>
</tr>
</tbody>
</table>

CV indicates cardiovascular; HR, hazard ratios; hs-cTnT, high-sensitivity cardiac troponin T; MI, myocardial infarction; NT-proBNP, N-terminal pro-type brain natriuretic peptide; uACR, urine albumin to creatinine ratio.

*Adjusted for other variables in the model.

\(^{†}\)Per standard-deviation increase in log variable.

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### Table 4. C-Index and Net Reclassification for Individual Variables Predicting MI, Stroke, or CV Death in Derivation and Validation Cohorts

<table>
<thead>
<tr>
<th>Variable</th>
<th>Derivation Cohort</th>
<th>Validation Cohort</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>C-Index* (95% CI)</td>
<td>Net Reclassification Improvement(^{†})</td>
</tr>
<tr>
<td>NT-proBNP</td>
<td>0.69 (0.65 to 0.73)</td>
<td>0.32 (0.16 to 0.49)</td>
</tr>
<tr>
<td>hs-cTnT</td>
<td>0.69 (0.65 to 0.73)</td>
<td>0.21 (0.04 to 0.39)</td>
</tr>
<tr>
<td>uACR</td>
<td>0.66 (0.62 to 0.70)</td>
<td>0.14 (−0.08 to 0.28)</td>
</tr>
<tr>
<td>Current smoking</td>
<td>0.49 (0.49 to 0.50)</td>
<td>0.05 (−0.08 to 0.18)</td>
</tr>
</tbody>
</table>

CV indicates cardiovascular; hs-cTnT, high-sensitivity cardiac troponin T; MI, myocardial infarction; NT-proBNP, N-terminal pro-type brain natriuretic peptide; uACR, urine albumin to creatinine ratio.

*C-index calculated for each individual variable alone.

\(^{†}\)Nested category-free net reclassification was calculated by adding each variable to a base model containing the other 3 variables.
model creation tends to result in optimistic estimates, external validation provides valuable information about the performance of a model.40 Though we observed a lower C-index in the validation population with our model, an externally validated C-index of 0.65 is well within the range of other commonly used CV prediction models, such as the CHADS-2 model for predicting stroke in patients with atrial fibrillation (validated C-index 0.63 to 0.70)41,42 and the Thrombolysis in Myocardial Infarction risk score for predicting death and ischemic events in patients with unstable angina or non-ST elevation myocardial infarction (validated C-index 0.59 to 0.65).7,43

Furthermore, even when we ran the selection algorithm on the PEACE cohort as if it were the derivation cohort, the highest CVCI was still <0.7. This suggests that the decrease in C-index in the validation cohort is less a reflection of overfitting, but rather that the general health and homogeneity of a clinical trial population makes it harder to risk stratify. Therefore, the decrease in discrimination is largely a reflection of PEACE being a healthier cohort of clinical trial subjects.

The inclusion of multiple novel risk factors in a risk model for secondary CV events represents a significant departure from previous risk models. Using both traditional and novel risk factors as candidate predictors, we found that the top model identified using a prespecified algorithm included only 1 traditional risk factor (smoking) and 3 novel risk factors (NT-proBNP, hs-cTnT, uACR). It has been observed that the risk factors that predict a first event are often not the same factors that predict secondary events, possibly due to more aggressive treatment of traditional risk factors in diseased patients.9 The participants included in this study were individuals with stable CHD, most of whom were receiving treatment for traditional risk factors, which may explain why traditional risk factors did not emerge as strong predictors of subsequent CV events. Biomarkers may provide more information about the burden of disease than traditional risk factors, including age. Notably, the novel risk factors identified as top predictors may reflect risk related to ongoing hemodynamic stress (NT-proBNP), myocardial damage (hs-cTnT), and renal dysfunction (uACR). In addition to improving overall risk estimation, measurement of novel risk factors could also play a role in directing intensification of therapies.

In current practice, measurement of these novel risk factors is not routine, but there is interest in determining whether their use can improve clinical outcomes. Routine measurement of natriuretic peptides is gaining acceptance in the management of heart failure.44 High-sensitivity troponin assays are not commercially available in the United States, but evidence of their utility in the setting of acute chest pain and risk prediction is strong.12,18,45 Measurement of uACR is a recommended method for detecting chronic kidney disease and a Class IIa guideline recommendation in patients with CV disease because of their increased risk of concomitant kidney disease.46,47 Further study is needed to determine whether using a risk model based on multiple novel risk factors can improve clinical outcomes.

Several limitations must be considered in the interpretation of our results. First, both the derivation and validation cohorts included 82% men, so future studies are required to test this model in cohorts with larger numbers of women. Second, HDL was not available in the validation cohort, and is a variable in the Framingham secondary events model;
therefore we could not export the published coefficients directly from the Framingham model, and perform a true validation by category-free net reclassification improvement. Nevertheless, by refitting both models to the validation cohort, we were able to make a fair comparison of the performance of the variables themselves, confirming that novel markers were stronger than traditional risk factors in the validation cohort as well. Third, the drop in C-index from 0.73 in the derivation cohort to 0.65 in the validation cohort likely reflects differences in the cohorts, but could also reflect some degree of model overfitting. Fourth, this model was created in the context of the candidate predictors listed in Table 1 and does not account for any variation in risk that could occur due to other predictors not included in this analysis. Finally, like most CV risk prediction models, this model does not take into account treatment effect during the course of follow-up. However, this model is meant to predict subsequent risk for stable and already well-treated CHD patients.

In summary, we developed and validated a simple 4-predictor model for 5-year risk of CV events in patients with stable treated CHD. A risk model including multiple novel risk factors may be useful for estimating prognosis in patients with stable CHD to inform treatment decisions or for use as a risk stratification tool in future research.

Acknowledgments

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Disclosures

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