Risk prediction by genetic risk scores for coronary heart disease is independent of self-reported family history

The Harvard community has made this article openly available. Please share how this access benefits you. Your story matters

Citation

Published Version
doi:10.1093/eurheartj/ehv462

Citable link
http://nrs.harvard.edu/urn-3:HUL.InstRepos:25658417

Terms of Use
This article was downloaded from Harvard University’s DASH repository, and is made available under the terms and conditions applicable to Other Posted Material, as set forth at http://nrs.harvard.edu/urn-3:HUL.InstRepos:dash.current.terms-of-use#LAA
Risk prediction by genetic risk scores for coronary heart disease is independent of self-reported family history

Hayato Tada1,2,3, Olle Melander4,5*, Judy Z. Louie6, Joseph J. Catanese6, Charles M. Rowland6, James J. Devlin6, Sekar Kathiresan1,2,3*, and Dov Shiffman6*

1Center for Human Genetic Research and Cardiovascular Research Center, Massachusetts General Hospital, Boston, MA, USA; 2Program in Medical and Population Genetics, Broad Institute, Cambridge, MA, USA; 3Department of Medicine, Harvard Medical School, Boston, MA, USA; 4Department of Clinical Sciences, Lund University, Malmö, Sweden; 5Department of Internal Medicine, Skåne University Hospital, Malmö, Sweden; and 6Quest Diagnostics, Alameda, CA, USA

Received 26 February 2015; revised 27 July 2015; accepted 19 August 2015; online publish-ahead-of-print 20 September 2015

Aims

Genetic risk scores (GRSs) have been associated with coronary heart disease (CHD) in large studies. We asked whether expanding an established 27-variant GRS (GRS27) to a 50-variant GRS (GRS50) improved CHD prediction and whether GRSs are independent of self-reported family history of CHD.

Methods and results

The association between GRSs and incident CHD was assessed in Cox models adjusting for established risk factors in 23,595 participants of the Malmö Diet and Cancer study—a prospective, population-based study. During a median follow-up of 14.4 years, 2213 participants experienced a first CHD event. After adjustment for established risk factors, both GRS27 and GRS50 were associated with incident CHD [hazard ratio (HR) = 1.70 for high (top quintile) vs. low (bottom quintile) of GRS27; 95% confidence interval (CI): 1.48–1.94; P trend = 1.6 × 10⁻¹⁵ and HR = 1.92 for GRS50; 95% CI: 1.67–2.20; P trend = 6.2 × 10⁻²²]. Adding 23 single nucleotide polymorphisms (SNPs) to GRS27 improved risk prediction (P = 3 × 10⁻⁶). Further adjustment for self-reported family history did not appreciably change the risk estimates of either GRS27 (HR = 1.65; 95% CI: 1.45–1.89) or GRS50 (HR = 1.87; 95% CI: 1.63–2.14). The addition of GRS50 to established risk factors, including self-reported family history, improved discrimination (P < 0.0001) and reclassification (continuous net reclassification improvement index = 0.17, P < 0.0001). In young participants (below median age), those with high GRS50 had 2.4-fold greater risk (95% CI: 1.85–3.12) than those with low GRS50.

Conclusion

The addition of 23 SNPs to an existing GRS27 improved CHD risk prediction and was independent of self-reported family history. Coronary heart disease risk assessment by GRS could be particularly useful in young individuals.

Keywords

Coronary heart disease risk • Genetic risk scores

Clinical perspective

Assessing risk of coronary events is integral to the prevention and treatment of cardiovascular disease. However, current risk assessment algorithms do not explicitly incorporate information about a patient’s genetic risk. This large, population-based, prospective study of middle-aged Europeans, found that genetic risk—as measured by a genetic risk score comprising dozens of single nucleotide polymorphisms—is independent of traditional risk factors, including family history of cardiovascular disease.
Introduction

Coronary heart disease (CHD) is the leading cause of death in the European Union as well as in the USA—an estimated ~1.8 million Europeans and ~400 000 Americans die of CHD annually, at an estimated annual cost of €60 billion in the European Union and $108.9 billion in the USA. Therefore, improving CHD risk prediction in order to effectively direct CHD risk prevention resources is an important public health goal. The association of established risk factors with CHD in large prospective studies has been used to develop a number of CHD risk prediction models, most recently by the European Association for Cardiovascular Prevention & Rehabilitation and US American College of Cardiology/American Heart Association task force on practice guidelines. Typically, these models incorporate information about age, sex, hypertension, blood cholesterol, smoking history, and history of diabetes to calculate the probability of CHD event in the short term (10 years) or long term. The heritability of CHD is well documented, which motivated the inclusion of family history of CHD as an option for risk assessment in patients with borderline risk. However, despite the identification of numerous genetic variants that are associated with CHD, a direct assessment of a patient’s genetic risk is not typically included in accepted CHD risk prediction models.

Genetic risk scores (GRSs)—based on single nucleotide polymorphisms (SNPs) associated with CHD—have been shown to be associated with future CHD events in large case-control and prospective studies. However, it has not been fully assessed whether increasing the number of genome-wide significant (GWS) SNPs in GRS-based CHD risk prediction continue to improve risk prediction, whether GRS could improve CHD risk assessment beyond self-reported family history, and whether GRS risk differs between younger and older individuals. We investigated these questions in a large population-based prospective study of middle-aged men and women.

Methods

Study participants

The Malmö Diet and Cancer (MDC) study is a community-based, prospective observational study of 30 447 participants drawn from ~230 000 residents of Malmö, Sweden. Men aged 46–73 years and women aged 45–73 years were invited to participate and were enrolled between 1991 and 1996. Details of the MDC design have been previously reported. The primary endpoints of the study were time to first occurrence of CHD (composite endpoint of coronary event, cardiovascular death, and revascularizations). After exclusions, 23 595 participants were included in the current study. A detailed description of the inclusion and exclusion criteria in the current study, baseline assessment, and endpoint determination are provided in the Supplemental material online.

Modelling of genetic risk score

We assessed the association of CHD with each of two GRSs (Supplementary material online, Table S1). One GRS is the 27-SNP GRS (GRS27) described by Mega et al. A second GRS comprised the GRS27 SNPs and 23 additional SNPs, for a total of 50 SNPs (GRS50). Each of these 23 additional SNPs has been shown to be associated with CHD at a GWS level. The GRS of each individual in the current study was calculated as follows: the previously reported risk estimate for each SNP was multiplied by the product of the probability of each risk allele at a SNP occurring in each individual (Supplementary material online, Table S1) was natural log transformed and multiplied by one (for heterozygotes) or two (for homozygotes); these products were then summed. The mean and standard deviation (0.34 for GRS27 and 0.43 for GRS50) of the study population were used to standardize each GRS to have a mean of 0 and unit variance. Genetic risk was assessed based on the standardized GRS as well as by comparing those with high GRS (Quintile 5), those with intermediate risk score (Quintiles 2 to 4), and those with low GRS (Quintile 1).

Table 1 Baseline characteristics according to coronary heart disease event status

<table>
<thead>
<tr>
<th>Baseline characteristics</th>
<th>Events (n = 2213)</th>
<th>Non-events (n = 21 382)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (years)</td>
<td>61.5 ± 6.9</td>
<td>57.7 ± 7.7</td>
</tr>
<tr>
<td>Men, n (%)</td>
<td>1420 (64.2)</td>
<td>7553 (35.3)</td>
</tr>
<tr>
<td>Body mass index (kg/m²)</td>
<td>26.7 ± 4.0</td>
<td>25.6 ± 4.0</td>
</tr>
<tr>
<td>Current smoker, n (%)</td>
<td>790 (35.7)</td>
<td>5832 (27.3)</td>
</tr>
<tr>
<td>Systolic blood pressure (mmHg)</td>
<td>150.3 ± 20.2</td>
<td>140.1 ± 19.9</td>
</tr>
<tr>
<td>Diastolic blood pressure (mmHg)</td>
<td>88.7 ± 9.8</td>
<td>85.2 ± 10.0</td>
</tr>
<tr>
<td>Use of anti-hypertensives, n (%)</td>
<td>646 (29.2)</td>
<td>3407 (15.9)</td>
</tr>
<tr>
<td>Prevalent diabetes mellitus, n (%)</td>
<td>242 (10.9)</td>
<td>688 (3.2)</td>
</tr>
<tr>
<td>Apolipoprotein A-I (g/L)</td>
<td>1.47 ± 0.26</td>
<td>1.58 ± 0.28</td>
</tr>
<tr>
<td>Apolipoprotein B (g/L)</td>
<td>1.18 ± 0.26</td>
<td>1.06 ± 0.26</td>
</tr>
<tr>
<td>Self-reported family history of CHD, n (%)</td>
<td>998 (0.45)</td>
<td>7790 (0.36)</td>
</tr>
<tr>
<td>GRS27</td>
<td>0.18 ± 1.01</td>
<td>−0.02 ± 1.00</td>
</tr>
<tr>
<td>GRS50</td>
<td>0.20 ± 1.00</td>
<td>−0.02 ± 1.00</td>
</tr>
</tbody>
</table>

Data are presented as mean ± standard deviation unless indicated.

GRS27, 27-variant genetic risk score; GRS50, 50-variant genetic risk score.
Results

Study population and genetic risk scores

The baseline characteristics of the 23,595 MDC participants in this study are provided in Table 1 (stratified by incident CHD event status) and Supplementary material online, Table S2 (stratified by self-reported family history). During a median follow-up of 14.4 years, 2,213 participants experienced a first CHD event.

We genotyped 50 SNPs reported to be associated with CHD at a GWAS level (Supplementary material online, Table S1). None of the SNPs were in strong linkage disequilibrium ($r^2 < 0.45$ for any pair of SNPs). We assessed the association of incident CHD with two GRSSs: the previously described GRSS27 as well as GRSS50, an expanded GRS that included all GRS27 SNPs as well as 23 additional SNPs. We calculated a weighted GRS for each participant using the literature risk estimates as weights for the risk allele of each of the SNPs. The standardized GRS means and standard deviations for those with and without events are reported in Table 1.

Genetic risk scores and incident coronary heart disease

Both GRSS27 and GRSS50 were associated with incident CHD [hazard ratio (HR) = 1.20 per SD; 95% confidence interval (CI): 1.15–1.25 and HR = 1.23; 95% CI: 1.18–1.28, respectively] after adjustment for established risk factors including age, sex, systolic blood pressure, hypertension treatment, smoking, apoB, apoA-I, and prevalent diabetes (Table 2). Those with high GRSS50 had 1.92-fold greater risk of CHD than those with low genetic risk (95% CI: 1.67–2.20, $P = 7.5 \times 10^{-21}$). For GRSS27, those with high genetic risk had 1.7-fold greater risk of CHD than those with low genetic risk (95% CI: 1.48–1.94). When both GRSS27 and a GRS comprising the 23 SNPs present in GRSS50 but not in GRSS27 were included in a model that also adjusted for established risk factors, both GRSS27 and GRSS23 were associated with incident CHD events ($P = 2 \times 10^{-13}$ and $3 \times 10^{-16}$, respectively). That is, CHD risk prediction by a model that included GRSS27 was improved by adding 23 additional SNPs.

Discrimination and reclassification

27-Variant genetic risk score and GRSS50 each improved the discrimination of a model that included established risk factors even after adding self-reported family history ($P \leq 2 \times 10^{-15}$) although the magnitude of the improvement of c-statistic was modest (Supplementary material online, Table S3). Self-reported family history improved discrimination of an established risk factors model. Risk classification by a model that included established risk factors and self-reported family history was improved by both GRSS27 [continuous net reclassification improvement index (cNRI) = 0.15, $P < 0.0001$, 7% in those without events and 8% in those with events, Supplementary material online, Table S4] and GRSS50 (cNRI = 0.17, $P < 0.0001$, 10% in those without events and 7% in those with events). Risk classification of an established risk factor model was also improved by self-reported family history, although among patients with events, more patients were reclassified in the wrong direction (lower risk). Risk classification was not improved by GRSS27 and GRSS50 in a categorical analysis (above and below 7.5% 10-year risk categories $^{4}$; Supplementary material online, Table S5).

Genetic risk scores and self-reported family history

27-Variant genetic risk score and GRSS50 were associated with incident CHD events in participants with and without self-reported family history. In a stratified analysis that adjusted for established risk factors, the HR for CHD for a high compared with a low GRSS50 was 1.75 (95% CI: 1.43–2.15) among those with family history and was 1.96 (95% CI: 1.63–2.35) among those without family history (Table 3). The association between GRSS50 and CHD events did not differ according to self-reported family history status: in an

---

Table 2 Genetic risk scores and incident coronary heart disease

<table>
<thead>
<tr>
<th>GRS risk</th>
<th>Low</th>
<th>Intermediate</th>
<th>High</th>
</tr>
</thead>
<tbody>
<tr>
<td>N (event)</td>
<td>4719 (343)</td>
<td>14,157 (1294)</td>
<td>4719 (576)</td>
</tr>
<tr>
<td>Event rate (95% CI)</td>
<td>5.21 (4.67–5.79)</td>
<td>6.62 (6.27–7.01)</td>
<td>8.97 (8.25–9.74)</td>
</tr>
<tr>
<td>HR (95% CI)</td>
<td>Reference</td>
<td>1.26 (1.12–1.42)</td>
<td>1.70 (1.48–1.94)</td>
</tr>
<tr>
<td>P value (vs. low)</td>
<td>$1.1 \times 10^{-7}$</td>
<td>$9.2 \times 10^{-7}$ (Perr = $1.6 \times 10^{-15}$)</td>
<td>$5.0 \times 10^{-18}$</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>GRS50</th>
<th>Low</th>
<th>Intermediate</th>
<th>High</th>
</tr>
</thead>
<tbody>
<tr>
<td>N (event)</td>
<td>4719 (318)</td>
<td>14,157 (1303)</td>
<td>4719 (592)</td>
</tr>
<tr>
<td>Event rate (95% CI)</td>
<td>4.82 (4.31–5.39)</td>
<td>6.67 (6.31–7.04)</td>
<td>9.25 (8.52–10.03)</td>
</tr>
<tr>
<td>HR (95% CI)</td>
<td>Reference</td>
<td>1.39 (1.23–1.57)</td>
<td>1.92 (1.67–2.20)</td>
</tr>
<tr>
<td>P value (vs. low)</td>
<td>$1.4 \times 10^{-7}$</td>
<td>$7.5 \times 10^{-21}$ (Perr = $6.2 \times 10^{-22}$)</td>
<td>$6.8 \times 10^{-23}$</td>
</tr>
</tbody>
</table>

Risk estimates were adjusted for age, sex, systolic blood pressure, hypertension treatment, smoking, apoB, apoA-I, and prevalent diabetes. GRSS27 risk boundaries (SD): low ¼ 0.8547; intermediate ¼ 0.8360; high ¼ 0.8360. GRSS50 risk boundaries (SD): low ¼ 0.8547; intermediate ¼ 0.8517; high ¼ 0.8360. Event rates are per 1000 person-years. HR, hazard ratio; CI, confidence interval; GRS, genetic risk score; GRSS27, 27-variant genetic risk score; GRSS50, 50-variant genetic risk score.
Table 3  Genetic risk scores and incident coronary heart disease according to self-reported family history

<table>
<thead>
<tr>
<th>Self-reported family history</th>
<th>GRS</th>
<th>Intermediate riska</th>
<th></th>
<th>High riskb</th>
<th></th>
<th>Pinteraction</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>HR (95% CI)</td>
<td>P value</td>
<td>HR (95% CI)</td>
<td>P value</td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>GRS50</td>
<td>1.29 (1.07–1.56)</td>
<td>0.007</td>
<td>1.75 (1.43–2.15)</td>
<td>7.7 × 10⁻⁸</td>
<td>0.33</td>
</tr>
<tr>
<td>No</td>
<td>GRS50</td>
<td>1.43 (1.21–1.68)</td>
<td>1.9 × 10⁻⁵</td>
<td>1.96 (1.63–2.35)</td>
<td>7.4 × 10⁻¹³</td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>GRS27</td>
<td>1.26 (1.05–1.52)</td>
<td>0.013</td>
<td>1.64 (1.34–2.01)</td>
<td>2.1 × 10⁻⁶</td>
<td>0.38</td>
</tr>
<tr>
<td>No</td>
<td>GRS27</td>
<td>1.23 (1.05–1.44)</td>
<td>0.009</td>
<td>1.67 (1.39–1.99)</td>
<td>2.1 × 10⁻⁸</td>
<td></td>
</tr>
</tbody>
</table>

Risk estimates were adjusted for age, sex, systolic blood pressure, hypertension treatment, smoking, apoB, apoA-I, and prevalent diabetes. GRS50 risk boundaries (SD): low ≤ −0.8547; intermediate > −0.8547 and ≤ 0.8236; high > 0.8236. GRS50 risk boundaries (SD): low ≤ −0.8517; intermediate > −0.8517 and ≤ 0.8360; high > 0.8360. Pinteraction, between GRS as a continuous variable and self-reported family history status for the incident CHD outcome.

HR, hazard ratio; CI, confidence interval; GRS, genetic risk score; GRS27, 27-variant genetic risk score; GRS50, 50-variant genetic risk score.

aIntermediate risk: Quintiles 2, 3, and 4 compared with low risk (Quintile 1).
bHigh risk: Quintile 5 compared with low risk (Quintile 1).

Table 4  Genetic risk scores and coronary heart disease risk in young and old

<table>
<thead>
<tr>
<th>&lt;Median agea</th>
<th>Intermediate riskc</th>
<th>High riskd</th>
<th>&gt;Median agea</th>
<th>Intermediate riskc</th>
<th>High riskd</th>
<th>Pinteraction</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>HR (95% CI)</td>
<td>P value</td>
<td></td>
<td>HR (95% CI)</td>
<td>P value</td>
<td></td>
</tr>
<tr>
<td>Model 1e</td>
<td></td>
<td></td>
<td>Model 2f</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>GRS50</td>
<td>1.55 (1.22–1.98)</td>
<td>0.0004</td>
<td>GRS50</td>
<td>1.53 (1.20–1.94)</td>
<td>0.0005</td>
<td></td>
</tr>
<tr>
<td>GRS27</td>
<td>1.53 (1.20–1.94)</td>
<td>0.0005</td>
<td>GRS27</td>
<td>1.53 (1.20–1.96)</td>
<td>0.0006</td>
<td></td>
</tr>
<tr>
<td></td>
<td>2.40 (1.84–3.12)</td>
<td>7.5 × 10⁻¹¹</td>
<td>2.35 (1.81–3.06)</td>
<td>2.0 × 10⁻¹⁰</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>1.33 (1.15–1.53)</td>
<td>7.9 × 10⁻³</td>
<td>1.32 (1.14–1.52)</td>
<td>0.0002</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>1.75 (1.49–2.06)</td>
<td>6.5 × 10⁻¹²</td>
<td>1.71 (1.45–2.01)</td>
<td>6.5 × 10⁻¹¹</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>1.18 (1.03–1.35)</td>
<td>0.02</td>
<td>1.17 (1.02–1.34)</td>
<td>0.03</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>1.52 (1.30–1.78)</td>
<td>1.8 × 10⁻⁷</td>
<td>1.48 (1.27–1.73)</td>
<td>9.6 × 10⁻⁷</td>
<td>0.0003</td>
<td></td>
</tr>
</tbody>
</table>

GRS27 had a mean of 2.49 and standard deviation of 0.34; GRS50 had a mean of 3.82 and standard deviation of 0.43. Risk boundaries (27-SNPs GRS): low GRS ≤ −0.8547; intermediate GRS > −0.8547 and ≤ 0.8236; high GRS > 0.8236. Risk boundaries (50-SNPs GRS): low GRS ≤ −0.8517; intermediate GRS > −0.8517 and ≤ 0.8360; high GRS > 0.8360. Pinteraction, for interaction between continuous GRS and median age status for the incident CHD outcome.

HR, hazard ratio; CI, confidence interval; GRS, genetic risk score; GRS27, 27-variant genetic risk score; GRS50, 50-variant genetic risk score.
aMedian age: 57.6.
bPinteraction for interaction between continuous GRS and median age status for the incident CHD events.
cIntermediate risk: Quintiles 2, 3, and 4 compared with low risk (Quintile 1).
dHigh risk: Quintile 5 compared with low risk (Quintile 1).
eModel 1: adjusted for age, sex, systolic blood pressure, use of antihypertensive medication, smoking, apoB, apoA-I, and prevalent diabetes.
fModel 2: Model 1 and additional adjustment for family history.

Analysis of the combined strata, the P for interaction was 0.33 in a model that adjusted for established risk factors and included a term for the interaction between family history and GRS50. Self-reported family history was associated with incident CHD events after adjustment for established risk factors (HR = 1.43, 95% CI: 1.31–1.56, P = 8 × 10⁻¹¹). Adjustment for self-reported family history in the combined strata did not appreciably change these risk estimates (HR = 1.87; 95% CI: 1.63–2.14 for GRS50 and HR = 1.40; 95% CI: 1.29–1.53 for family history). Similar results were observed for GRS27 (Table 3). Both GRS27 and GRS50 were associated with self-reported family history of CHD (P < 0.0001); the odds of a positive self-reported family history among those with high GRS compared with having low GRS were modest: OR = 1.37 (95% CI: 1.26–1.49) for GRS27 and OR = 1.40 (95% CI: 1.29–1.53) for GRS50. The GRS distribution among participants with and without self-reported family history is shown in Supplementary material online, Figure S2.

Genetic risk score in young and old

Since genetic risk is generally thought to be more important in young individuals, we investigated the interaction between age and GRS and found that the CHD risk associated with GRS varied with age for both GRS27 and GRS50 (Pinteraction = 0.03 for both). We further assessed the GRS in those below and above the median age of this study (57.6) and found that risk associated with either GRS50 or GRS27 differed among those below and above the median age (Pinteraction < 0.003) in models that adjusted for established risk factors (Table 4 and Supplementary material online, Table S6 for event rate by GRS category and age category). Among the young, those with high GRS had more than two-fold greater risk of CHD than those with low GRS (HR = 2.40; 95% CI: 1.84–3.12; P = 7.5 × 10⁻¹¹ for GRS50 and HR = 2.24; 95% CI: 1.72–2.90; P = 1.4 × 10⁻⁹ for GRS27). We examined the event rate in those with and without self-reported family history of CHD according to their GRS status (Figure 1). We found that in the young, the
found that CHD risk assessed by both GRS is independent of self-reported family history, and for both the 27-SNP and the 50-SNP GRS, the associated CHD risk estimates were higher among younger individuals than among older individuals.

Genetic risk scores comprising SNPs that are individually associated with CHD at a GWAS level have been investigated previously. In 2010, Ripatti et al. reported that GRS comprising 13 SNPs was associated with incident CHD in several cohorts; however, this 13-SNP GRS did not improve net reclassification when added to established traditional risk factors. Following the publication of the largest to date meta-analysis genome-wide association in 2013, Andrea et al. investigated a 46-SNP GRS in six Swedish studies comprising ~10,000 individuals and found that this 46-SNP GRS improved net reclassification when added to established risk factors. More recently, Mega et al. investigated a 27-SNP GRS in randomized, placebo-controlled studies of statin therapy and found that the 27-SNP GRS could identify individuals who would benefit from statin therapy.

Since the risk associated with each individual SNP is modest, and since SNPs that have been only identified in recent large studies typically have risk estimates that are lower than SNPs that were identified in smaller, earlier studies, it was unclear whether it would be useful to include these more recently identified SNPs in a GRS. Our study investigated this question directly. We showed a model that includes 23 SNPs in addition to those in the previously reported 27-SNP GRS improves risk prediction. However, adding either GRS27 or GRS50 to established risk factors resulted in similar improvement in discrimination (c-statistic) and reclassification (cNRI).

Our study also found that subjective, self-reported family history of CHD and objectively measured genetic risk are not redundant. Both can contribute to a better assessment of a patient’s CHD risk because a GRS-based genetic risk measure is associated with CHD independent of self-reported family history of CHD, as well as established risk factors, that is, self-reported family history is not a substitute for genetic risk assessment. Since family history reflects both genetic and non-genetic factors, and since the accuracy of patient-reported family history is low, this non-redundancy of self-reported family history and genetic assessment is not surprising.

Since genetics is generally thought to play a more important role in CHD events that occur in younger individuals than in older individuals, we examined the 27-SNP and 50-SNP GRS risk estimates in different age groups and found that those risk estimates were higher in younger individuals than in older individuals. Moreover, younger individuals with no self-reported family history of CHD and with high GRS had greater risk than those with self-reported family history of CHD and a low GRS. This finding suggests that a GRS-based risk assessment could be particularly useful among younger individuals — particularly those with borderline CHD risk by established risk factors — attractive since it could overcome barriers faced by other predictive tools, e.g., 24 h ambulatory blood pressure monitoring.

Our study has several limitations. This study was conducted in Swedish middle-aged individuals; hence, the generalizability to other ethnicities or age groups is uncertain. LDL cholesterol and HDL cholesterol levels were not available for our study population; therefore, we used the available apoA-I and apoB plasma levels as cumulative incidence in the presence of competing risk among those with high GRSS0 without a self-reported family history (0.065 at Year 15) was greater than the cumulative incidence among those with low GRSS0 with a self-reported family history (0.041, P = 0.013 for Gray test; Figure 1A). Similar results were observed in those above the median age (Panel B). Similar results were observed for GRS27 (Supplementary material online, Figure S3). We also examined whether CHD risk prediction by GRS varied by sex and found no evidence for interaction between sex and GRS ($P_{\text{interaction}} > 0.3$).

**Discussion**

In a large community-based prospective study of 23,595 participants, we investigated a 27-SNP and a 50-SNP GRS for CHD and

---

**Figure 1** Cumulative incidence of coronary heart disease events according to self-reported family history of coronary heart disease and 50-variant genetic risk score. Blue and green: those with high 50-variant genetic risk score with (blue) or without (green) a self-reported family history. Red and black: those with low 50-variant genetic risk score with (red) or without (black) a self-reported family history. Inset: those with intermediate 50-variant genetic risk score with (dashed) or without (dotted) a self-reported family history. FH, self-reported family history. Cumulative incidence was estimated while considering non-coronary heart disease death as competing risk. (A) Participants younger than median age ($\leq 57.6$). Median age for this younger group is 51.4 (interquartile range, 48.8–54.2). (B) Participants older than median age ($>57.6$). Median age for this older group is 64.7 (interquartile range, 61.1–67.7).
cotivarites in our established risk factors model. A different definition of self-reported family history might have produced different results. For example, the definition of self-reported family history of CHD in the Framingham Heart Study considers only CHD family member events that occurred before age 55 for men or 65 for women. However, in clinical practice, patients may not know the age at which their family member had a CHD event or may not be asked about the age. For these reasons, the family history question asked at baseline in MDC did not specify any limitations on the age of the family member at the time of event. The fraction of women in MDC is greater than that in other European-based population studies.** However, since we found no interaction between GRS and sex, we believe that our results should be generalizable to populations with different proportion of women. Although our results suggest that genetic risk assessment could be useful in the young, this study population did not include sufficient number of individuals to provide risk estimates in those <45 years of age. Additional studies would be needed to address this question.

In conclusion, a GRS could improve risk assessment for future CHD when added to established risk factor models. We suggest that such genetic assessment could be considered for individuals whose established risk-based treatment decision is uncertain.

**Supplementary material**

Supplementary material is available at European Heart Journal online.

**Funding**

The Malmö Diet and Cancer study was made possible by grants from the Swedish Cancer Society, the Swedish Medical Research Council, the Swedish Dairy Association, and the Albert Pålsson and Gunnar Nilsson Foundations and the Malmö city council. H.T. is supported by a grant from the Japanese Circulation Society to study in the USA. S.K. is supported by a Research Scholar award from the Massachusetts General Hospital (MGH), the Howard Goodman Fellowship from MGH, and the Donovan Family Foundation. O.M. is supported by the European Research Council (StG-282255), the Swedish Heart and Lung Foundation, Swedish Research Council; the Novo Nordisk Foundation, the Skåne University Hospital donation funds; the Medical Faculty, Lund University; the Governmental funding of clinical research within the national health services, the Albert Pålsson Research Foundation, Region Skåne, the King Gustav V and Queen Victoria Foundation and the Marianne and Marcus Wallenberg Foundation. O.M. and S.K. are the recipients of an investigator-initiated grant from Quest Diagnostics. Funding to pay the Open Access publication charges for this article was provided by Quest Diagnostics.

**Conflict of interest:** D.S., J.Z.L., J.J.C., C.M.R., and J.J.D. are employees of Quest Diagnostics. S.K. is a member of a scientific advisory board at Quest Diagnostics. The King Gustav V and Queen Victoria Foundation was provided by Quest Diagnostics.

**References**


---

**CARDIOVASCULAR FLASHLIGHT**

\textit{doi:10.1093/eurheartj/ehv286}

\textit{Online publish-ahead-of-print 3 July 2015}

Radiotherapy-induced vascular damage in mammary arterial graft: correlations between optical coherence tomography and pathology

Nicolas Amabile\(^\star\), Aurélie Veugeois\(^1\), Konstantinos Zannis\(^2\), and Christophe Caussin\(^1\)

\(^1\)Department of Cardiology, Institut Mutualiste Montsouris, 42 Boulevard Jourdan, Paris 75014, France; and \(^2\)Department of Cardiac Surgery, Institut Mutualiste Montsouris, Paris, France

\(^\star\) Corresponding author. Tel: +33 1 56 61 65 58, Email: nicolas.amabile@imm.fr; nicolasamabile@yahoo.fr

A 53-year-old man with a previous history of mediastinal radiotherapy was referred to our institution for coronary artery bypass graft and aortic valve replacement.

Per-operative surgical analysis revealed poor flow within left and right internal mammary arteries (LIMA and RIMA), making their use as pedunculated conduits unsuitable. The proximal end arteries were removed and both vessels were harvested as free grafts. Moreover, aorta was heavily calcified allowing only one proximal anastomosis. A saphenous vein [saphenous venous graft (SVG)] was used to bypass the right coronary artery, then the free LIMA was grafted between the vein and mid-left anterior descending artery; free-RIMA was placed between LIMA and first marginal branch.

The patient experienced a cardiac arrest 2 h after procedure’s end and was successfully resuscitated. An emergency angiography control was decided under extracorporeal membrane oxygenation.

The graft selective angiography revealed a sub-occluded LIMA graft (Panel A; Supplementary material online, Video S1). TIMI 3 flow was restored in LIMA and RIMA following a low-pressure balloon angioplasty (Panel B; Supplementary material online, Video S2). A subsequent LIMA analysis with optical coherence tomography (OCT) was performed (see Supplementary material online, Video S3), revealing post-balloon dissection (Panel C2), infiltrative disease with fibro-lipid plaque atherosclerosis (Panel C3), and marked intimal and medial thickening (Panel C4). A redo emergent surgery was decided: both LIMA and RIMA were explanted and replaced by saphenous grafts.

Pathological analysis of the arterial grafts (Hematoxylin and eosin staining) correlated to OCT images: they showed presence of fibro-lipid plaque (white arrow) with inflammatory cells infiltrate (blue arrowhead), 1 cm above LIMA/SVG anastomosis (Panels D1 and D2) as well as diffuse intima fibrosis (white arrowhead) and media necrosis (white arrows, Panels E1 and E2), suggesting radiotherapy-induced vasculopathy. Subsequent patient evolution was uneventful.

Supplementary material is available at \textit{European Heart Journal} online.

---

Published on behalf of the European Society of Cardiology. All rights reserved. © The Author 2015. For permissions please email: journals.permissions@oup.com.