Surgical Site Infections: Volume-Outcome Relationship and Year-to-Year Stability of Performance Rankings

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Background: Surgical site infection (SSI) rates are publicly reported as quality metrics and increasingly used to determine financial reimbursement.

Objective: To evaluate the volume-outcome relationship as well as the year-to-year stability of performance rankings following coronary artery bypass graft (CABG) surgery and hip arthroplasty.

Research Design: We performed a retrospective cohort study of Medicare beneficiaries who underwent CABG surgery or hip arthroplasty at US hospitals from 2005 to 2011, with outcomes analyzed through March 2012. Nationally validated claims-based surveillance methods were used to assess for SSI within 90 days of surgery. The relationship between procedure volume and SSI rate was assessed using logistic regression and generalized additive modeling. Year-to-year stability of SSI rates was evaluated using logistic regression to assess hospitals’ movement in and out of performance rankings linked to financial penalties.

Results: Case-mix adjusted SSI risk based on claims was highest in hospitals performing <50 CABG/year and <200 hip arthroplasty/year compared with hospitals performing ≥200 procedures/year.

At that same time, hospitals in the worst quartile in a given year based on claims had a low probability of remaining in that quartile the following year. This probability increased with volume, and when using 2 years’ experience, but the highest probabilities were only 0.59 for CABG (95% confidence interval, 0.52–0.66) and 0.48 for hip arthroplasty (95% confidence interval, 0.42–0.55).

Conclusions: Aggregate SSI risk is highest in hospitals with low annual procedure volumes, yet these hospitals are currently excluded from quality reporting. Even for higher volume hospitals, year-to-year random variation makes past experience an unreliable estimator of current performance.

Key Words: hospital quality, public reporting, performance measurement, health care-associated infection, Medicare, health care reform, hospital reimbursement, policy

(Med Care 2017;55: 79–85)
Study Population

We used 2005–2011 Medicare Provider Analysis and Review Research Identifiable Files to identify all short-stay acute care US hospitals performing CABG and primary hip arthroplasty on fee-for-service Medicare patients between January 1, 2005 and December 31, 2011, based on Medicare Part A inpatient claims data. We refer to these below as Medicare patients and Medicare procedures.

We identified CABG cases using International Classification of Diseases, Ninth Revision (ICD-9) codes 36.10–36.17, 36.19, and 36.2, and primary hip arthroplasty cases using ICD-9 codes 81.51 and 81.52, with multiple codes for the same procedure type on the same day counted as a single case. We then assessed claims within 90 days of the surgical procedure for ICD-9 codes suggestive of a deep and organ/space SSI, including ICD-9 diagnosis codes 513.1, 682.2, 730.08, 996.61, 996.62, 998.31, 998.32, 998.51, and 998.59 following CABG, and 996.60, 996.66, 996.67, 996.69, 998.51, and 998.59 following primary hip arthroplasty. Prior national studies showed that these codes accurately rank hospitals when compared with expert adjudication of full text medical records, often identifying cases missed by hospital surveillance. Therefore, we feel this is an acceptable methodology for comparing SSI outcomes across hospitals. In fact, similar methodology is currently being used by the CMS Hospital Inpatient Quality Reporting Program to validate hospital SSI reporting.

To focus on surgical complications rather than preexisting infections, we excluded all codes listed as present on admission codes during any readmissions to a hospital within 90 days of the surgical procedure. For patients who underwent another major surgery in the 90-day postoperative surveillance window, we censored our surveillance at the time of the subsequent surgery. We used 2012 Medicare Provider Analysis and Review Research Identifiable Files data to capture coding and readmissions for procedures performed in the last 90 days of 2011.

For each patient in our CABG and primary hip arthroplasty cohorts, we collected data on age, sex, and co-morbidities at the time of the surgical admission for individual-level risk adjustment. Comorbidities were assessed using publicly available comorbidity software from the Agency for Healthcare Research and Quality’s Healthcare Cost and Utilization Project, based on methods previously described by Elixhauser et al.

SSI Risk by Procedure Volume

Medicare procedure volume was categorized into 1–24, 25–49, 50–99, 100–199, or 200+ surgical cases on Medicare patients each year, and hospitals could change volume categories across years. These categories correspond to hospitals that would be excluded from public reporting based on expected SSI rates of 4%, 2%, 1%, 0.05%, respectively, using the current CMS methodology of excluding hospitals with <1 expected SSI in a given year based on procedure volume.

We used logistic regression to calculate the odds of having an SSI code by the annual Medicare surgical volume. Generalized estimating equations were used to control for repeated measures across years within individual hospitals, and the model was adjusted for the age, sex, and co-morbidities of each Medicare patient. We also examined the relationship between continuous procedure volume and the probability of a patient having an SSI code graphically using generalized additive models.

A more in depth discussion of the statistical methods and models is included in a supplementary appendix, Supplemental Digital Content 1 (http://links.lww.com/MLR/B247).

SSI Risk Over Time

Within each year, we fit a logistic regression mixed effects model with random intercepts for each hospital, including each individual’s age, sex, and coded comorbidities as fixed effects. The predicted random intercepts from the models for each year indicate case mix-adjusted relative performance. Hospitals were ranked based on these predicted random intercepts and divided into case mix-adjusted quartiles of performance. We then used logistic regression, with generalized estimating equations to account for repeated measures across years within hospital, to model the outcome of being in the worst quartile next year based on: (1) worst, middle 2, and best quartile status this year, for all hospitals combined and by volume category; (2) worst, middle 2, and best quartile status this year and the year prior, by volume category. Our modeling of quartile status was based on the fact that the CMS HAC Reduction Program uses quartiles to determine financial penalties.
We also plotted the performance rank of randomly selected individual hospitals from 2005 through 2011. This allowed us to examine the stability of SSI performance visually.

**RESULTS**

**Study Population**

From 2005 through 2011, an average of 1165 US hospitals performed CABG procedures and 3117 US hospitals performed hip arthroplasty procedures for fee-for-service Medicare patients each year. An average of 135,882 CABG procedures and 217,223 hip arthroplasty procedures were performed each year. The median number of annual fee-for-service Medicare surgical cases per hospital was 89 for CABG (interquartile range, 46–150) and 46 for hip arthroplasty (interquartile range, 19–95). There was a 17% decline in the number of CABG procedures and a 17% increase in the number of hip arthroplasty procedures performed per year from 2005 through 2011.

**SSI Risk by Procedure Volume**

Table 1 shows SSI coding rates in US hospitals by the annual volume of surgical cases performed on Medicare patients. The table also shows the unadjusted and adjusted odds of SSI for each group compared with the largest volume hospitals. For both CABG and hip arthroplasty, coded SSI rates were highest at hospitals that performed the fewest procedures per year.

For CABG, patients at hospitals performing <50 Medicare procedures per year had significantly higher odds of an SSI code: 30% higher for patients in hospitals performing fewer than 25 procedures and 21% higher in hospitals performing 25–49 procedures. The risk in hospitals that performed 50–199 procedures per year was not different from those performing ≥200 procedures. Hospitals performing fewer than 50 procedures per year accounted for more than a quarter of all US hospitals performing CABG procedures from 2005 to 2011, while only accounting for 7%–8% of the surgical volume.

For hip arthroplasty, there was an inverse correlation between procedure volume and patients’ SSI risk, with patients at hospitals in each of the 4 lower volume categories having significantly higher risk than patients at hospitals performing ≥200 procedures per year. The excess odds of an SSI code ranged from 58% in hospitals performing fewer than 25 procedures to 14% in hospitals performing 100–199 procedures. Over 90% of US hospitals performing hip arthroplasty procedures from 2005 to 2011 had SSI risks that were significantly higher than those in the largest volume hospitals.

Figure 1 plots the probability of a patient having an SSI code by volume of Medicare procedures. The data shown is for 2011, but the trends were similar for each year. For CABG, coded SSI rates decreased with increasing volume up to approximately 100 Medicare fee-for-service procedures per year, with similar risk above 100 procedures per year. For hip arthroplasty, SSI risk continued to decline as hospital volume increased.

**Stability of SSI Performance over Time**

Table 2 shows the probability that a hospital ranked in the worst quartile in 1 year remained in that quartile the following year based on claims data suggestive of SSI.

For CABG, hospitals ranked in the worst quartile in the current year had a 44% chance of remaining in that quartile the following year. For comparison, the chance of remaining in the worst quartile would be 25% if quartile rankings were solely due to random chance. The probability of remaining in the worst quartile was directly related to a hospital’s annual procedure volume. For example, the probability of remaining in the worst quartile for hospitals performing fewer than 25 CABG procedures annually on Medicare patients was 0.20, and statistically no different from chance [95% confidence interval (CI), 0.14–0.29], while this probability was 0.52 in the highest volume hospitals (95% CI, 0.45–0.60).

### TABLE 1. Procedure Volume and Rates of Assignment of Surgical Site Infection Codes in Medicare Fee-for Service Patients

<table>
<thead>
<tr>
<th>Annual Procedure Volume (Medicare)</th>
<th>Hospitals Performing Procedurea [N (%)]</th>
<th>Annual No. Surgical Casesb [N (%)]</th>
<th>Annual Coded Infectionsc</th>
<th>Coded SSI Rate (%)</th>
<th>Crude OR (95% Confidence Interval)</th>
<th>Adjusted OR (95% Confidence Interval)</th>
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<tr>
<td>1–24</td>
<td>121 (10)</td>
<td>1581 (1.2)</td>
<td>120</td>
<td>7.57</td>
<td>1.29 (1.16–1.43)</td>
<td>1.30 (1.17–1.44)</td>
</tr>
<tr>
<td>25–49</td>
<td>202 (17)</td>
<td>7522 (5.5)</td>
<td>529</td>
<td>7.04</td>
<td>1.19 (1.11–1.28)</td>
<td>1.21 (1.12–1.30)</td>
</tr>
<tr>
<td>50–99</td>
<td>323 (28)</td>
<td>23,714 (17.5)</td>
<td>1448</td>
<td>6.11</td>
<td>1.02 (0.96–1.09)</td>
<td>1.04 (0.98–1.11)</td>
</tr>
<tr>
<td>100–199</td>
<td>343 (29)</td>
<td>48,013 (35.3)</td>
<td>2811</td>
<td>5.85</td>
<td>0.98 (0.92–1.04)</td>
<td>1.01 (0.95–1.07)</td>
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<tr>
<td>200+</td>
<td>176 (15)</td>
<td>55,052 (40.5)</td>
<td>3290</td>
<td>5.98</td>
<td>Ref.</td>
<td>Ref.</td>
</tr>
<tr>
<td>1–24</td>
<td>984 (32)</td>
<td>11,250 (5.2)</td>
<td>348</td>
<td>3.10</td>
<td>1.63 (1.51–1.75)</td>
<td>1.58 (1.47–1.69)</td>
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<tr>
<td>25–49</td>
<td>657 (21)</td>
<td>23,809 (11.0)</td>
<td>628</td>
<td>2.64</td>
<td>1.38 (1.29–1.48)</td>
<td>1.34 (1.26–1.44)</td>
</tr>
<tr>
<td>50–99</td>
<td>744 (24)</td>
<td>53,234 (24.5)</td>
<td>1274</td>
<td>2.39</td>
<td>1.25 (1.17–1.33)</td>
<td>1.22 (1.15–1.30)</td>
</tr>
<tr>
<td>100–199</td>
<td>535 (17)</td>
<td>73,201 (33.7)</td>
<td>1631</td>
<td>2.23</td>
<td>1.16 (1.09–1.23)</td>
<td>1.14 (1.07–1.21)</td>
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<tr>
<td>200+</td>
<td>195 (6)</td>
<td>55,723 (25.7)</td>
<td>1074</td>
<td>1.93</td>
<td>Ref.</td>
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*Adjusted odds ratios in bold text were significantly higher than the null value of 1 (P<0.05).

*Logistic regression model with generalized estimating equations to account for clustering within hospitals, robust standard errors used for confidence interval.

*Logistic regression model adjusted for age, sex, and coded co-morbidities, with generalized estimating equations to account for clustering within hospitals, robust standard errors used for confidence interval.

*CABG indicates coronary artery bypass graft; OR, odds ratio; Ref., reference; SSI, surgical site infection.
For hip arthroplasty, hospitals ranked in the worst quartile in the current year had a 34% chance of remaining in that quartile the following year. Similar to CABG, the probability of remaining in the worst quartile for hospitals performing fewer than 25 hip arthroplasty procedures annually on Medicare patients was no different from chance (0.26, 95% CI, 0.23–0.29). In the highest volume hospitals, the probability was 0.44 (95% CI, 0.37–0.50).

The modeled probability of moving from the best quartile to the worst quartile for CABG was 0.11 and for hip was 0.18. In other words, 1 of 9 hospitals ranked in the best quartile in the current year based on coding suggestive of SSI following CABG and 1 of 6 hospitals ranked in the best quartile in the current year based on coding suggestive of SSI following hip arthroplasty would be expected to be ranked in the worst quartile the following year. This movement in annual performance ranking is demonstrated graphically in Figure 2, which plots the yearly case mix-adjusted ranks for 5 randomly selected hospitals in each volume category. Regardless of annual procedure volume, there was great variability in the year-to-year rank for each hospital from 2005 to 2011. In fact, we observed consecutive year movement from the worst quartile to the best quartile and then back to the worst quartile in a number of hospitals for both procedures.

The effect of using 2 years of data to predict the next year’s performance is shown in Table 3. For both CABG and hip arthroplasty, hospitals ranked in the worst quartile 2 years in a row had a slightly higher probability of being ranked in the worst quartile the following year. However, even for the largest hospitals, the probabilities were still only 0.59 for CABG (95% CI, 0.52–0.66) and 0.48 for hip arthroplasty (95% CI, 0.42–0.55).

**CONCLUSIONS**

Public reporting of SSI rates and their role in determining hospital reimbursement have led to an increasing national focus on these potentially preventable health care–associated infections.\(^{26}\) However, data from hospitals with low procedure volume are often ignored due to well founded concerns about instability in their individual rates. This exclusion is important, as our data suggest that procedure volume is a strong predictor of outcome following CABG and hip arthroplasty, with the highest SSI risk found in those hospitals currently excluded from public scrutiny. In addition, while SSI rates have declined nationally over time,\(^{27}\) our modeling of longitudinal data from individual hospitals suggests that annual SSI performance rankings may be highly unstable from year-to-year, even for hospitals performing many procedures. This raises important concerns about the use of SSI data for both quality reporting and reimbursement programs.

Our finding that hospitals that perform a lower volume of surgical procedures have a higher SSI risk adds to prior literature showing volume-outcome associations for post-operative complications, costs, and mortality.\(^{7–11}\) In one study, lower surgical procedure volume was associated with longer operative durations, a risk factor for SSI.\(^{28}\) It is also possible that infection prevention practices are more standardized at hospitals with higher surgical volumes.\(^{3}\) Whatever the reason,
the association is robust, and procedure volume should be considered a strong determinant of SSI risk. Instead, current policies ignore hospitals with lower procedure volume.

As for why coded SSI rates declined with increasing surgical volume only up to around 100 annual CABG cases on Medicare fee-for-service patients, while the volume-outcome relationship continued beyond this surgical case volume for hip arthroplasty, we do not know of existing literature highlighting this finding. It is not clear why this should differ between CABG and hip arthroplasty, and it is worth further investigation.

Two recent papers assessed the hospitals being penalized by the HAC Reduction Program and the Hospital Value-Based Purchasing Program. These programs appear to disproportionately penalize larger volume hospitals, and major teaching hospitals in particular. Exclusion of hospitals with low surgical volume is an important contributor to this phenomenon.

Even among the included hospitals, there is a high degree of variability in the year-to-year performance ranking of individual hospitals. It is not clear that outcome from a single year, or even 2 years, are truly indicative of better or worse performance. Similar instability was previously reported for metrics comparing postoperative mortality. This prior work showed a failure to predict future outcomes based on past performance, with an argument that this was due to an insufficient number of procedures to adequately assess outcomes in relation to other hospitals. For this reason, some have advocated for changes in sampling strategies to capture additional cases in an effort to improve reliability.

In exploring the impact of a broader sampling strategy, we did find a modest improvement in prediction by using quartile ranking over 2 consecutive years to predict whether a hospital would remain in the worst quartile. Even with this extra year of data, however, prediction remained poor. This was true even in the highest volume group, where 41% of hospitals performing CABG and 52% of hospitals performing hip arthroplasty were no longer in the worst quartile in the subsequent year despite having been in the worst quartile for 2 consecutive years. In addition, broader sampling strategies may compromise interpretability by combining data from different periods and the possibility for direct improvement based on early indicators of worse outcomes.

One interesting article suggested a statistic to measure rankability. Unfortunately, since the statistic depends on fitting a fixed effects model, we were unable to estimate it in our full sample of hospitals, due to the relative rareness of the outcome and small number of procedures in some hospitals. Instead, our analyses used empirical Bayes methods in an effort to improve reliability when analyzing year-to-year performance ranking. These methods have been advocated by others comparing quality outcomes between hospitals. Although there are benefits to this approach, these methods pull hospitals with low procedure volumes toward the middle of the distribution, meaning they
are less likely to appear in the extreme quartiles. In seeking greater certainty, the outcomes in many small volume hospitals are ignored through these empirical Bayes methods, despite the fact that in the aggregate these hospitals have the highest complication rates. At the same time, the variability that we noted in year-to-year performance ranking was based on hospitals with the most robust data.

Our study does have some limitations. First, our analyses used claims data to identify SSI outcomes rather than data from hospital-based prospective surveillance systems. For both CABG and hip arthroplasty, we previously published national validation work showing that billing codes that are suggestive of SSI successfully identify reportable infections with similar confirmation rates among patients with an SSI code in hospitals with high versus low SSI coding. This suggests that hospitals with higher rates of SSI coding also have higher rates of chart confirmed SSIs.18–21 We believe claims-based ranking is both more complete because it includes the substantial number of SSIs that become manifest after discharge and do not return to the hospital that performed the surgery, and also superior in that this method minimizes variable implementation of surveillance between hospitals.

Second, our analyses were limited to Medicare fee-for-service patients. In general, though, Medicare surgical volume tracks well with overall surgical volume when looking at data from State Inpatient Databases.35 In addition, there is no reason to think that the selected diagnosis codes would be used differently in Medicare fee-for-service patients versus other patients in a way that would influence our results.

Third, we only included inpatient claims data, and did not look at coding for potential SSIs that may have been treated in the outpatient setting. In prior work, however, we showed that our SSI codes identified 87% of deep incisional primary and organ/space SSIs following CABG and 92% following hip arthroplasty when limiting to Part A inpatient claims data.18

Fourth, it is possible that surgical procedures with higher SSI rates might have different year-to-year variability. An example is colon surgery that has national SSI rates of 5.6%, compared with 1.3% for primary hip replacement and 2.8% for CABG.36

Fifth, we only used age, sex, and comorbidity data coded in inpatient claims for case-mix adjustment. It is possible that other factors not available in claims could further improve our case-mix adjustment, but a hospital’s case-mix typically does not change significantly from year-to-year, so we do not believe that this significantly impacts our study findings.

Although we recognize that any grading system is subject to year-to-year reclassification, it is important to consider when this variation is so great that the process being used to compare hospital performance loses its value. The high degree of year-to-year variation in the performance rank of individual hospitals, shown in our data (Fig. 2), raises concerns about the value of a hospital’s rank in any given year. Instead, there should be a greater focus on sustained performance trends which are more likely to be a better measure of quality. One alternative would be to financially reward hospitals with sustained longitudinal improvement, rather than focusing on performance in a single year or 2 years.

This is important, as hospital SSI rates following colon surgery and abdominal hysterectomy are now being publicly reported and used to determine CMS reimbursement.14 If the problems we identified for CABG and hip arthroplasty apply for these other 2 procedures, hospitals may incur financial penalties and patients may make choices on the basis of information that is a poor measure of the hospital’s current performance.

SSI risk for CABG and hip arthroplasty is greatest in the large number of US hospitals performing too few procedures to be included in public reporting. At the same time, for those hospitals that are included in public reports, past performance is a poor predictor of future performance. Assigning financial penalties and providing guidance to the public based on current SSI performance metrics risks ignoring the large group of low volume hospitals that have the poorest outcomes, and targeting hospitals whose outlier status in a given year may be due to random variation.
REFERENCES


