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The Heritability of Hedonic Capacity and Perceived Stress: A Twin Study Evaluation of
Candidate Depressive Phenotypes

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Running Title: The Heritability of Hedonic Capacity and Perceived Stress

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ABSTRACT

Background. Anhedonia and stress sensitivity have been identified as promising depressive phenotypes. Research suggests that stress-induced anhedonia is a possible mechanism underlying the association between stress and depression. The present proof-of-concept study assessed whether a laboratory-based measure of hedonic capacity and stress perception are heritable and if their genetic and environmental contributions are shared.

Methods. Twenty monozygotic (MZ) and 15 dizygotic (DZ) twin pairs completed a probabilistic reward task that provides an objective behavioral measure of hedonic capacity (reward responsiveness) and completed several questionnaires including the Perceived Stress Scale.

Bivariate Cholesky models were used to investigate whether covariation between 1) depressive symptoms and hedonic capacity, 2) depressive symptoms and perceived stress, and 3) perceived stress and hedonic capacity resulted from shared or residual genetic and environmental factors.

Results. Additive genetic (A) and individual-specific environment (E) factors contributed to 46% and 54% of the variance in hedonic capacity, respectively. For perceived stress, 44% and 56% of the variance was accounted for by A and E factors, respectively. The genetic correlation between depression and hedonic capacity was moderate ($r_a=0.29$), while the correlation between depression and stress perception was large ($r_a=0.67$). Genetic and environmental correlations between hedonic capacity and stress perception were large ($r_a=0.72$ and $r_e=-0.43$).

Conclusions. The present study provides initial feasibility for using a twin approach to investigate genetic contributions of a laboratory-based anhedonic phenotype. Although these preliminary findings indicate that hedonic capacity and perceived stress are heritable with substantial shared additive genetic contributions, replications in larger samples will be needed.

INTRODUCTION

Current mental illness classification systems, such as the Diagnostic and Statistical Manual of Mental Disorders (DSM-IV; APA, 2000), take an atheoretical approach to the etiology and pathophysiology of mental illness by relying upon phenomenological descriptions of symptom clusters and clinical course as diagnostic criteria (Hyman, 2007). One issue stemming from these nosological systems is that they identify categorical illnesses that are inherently heterogeneous. As an illustration, the DSM-IV requires that 5 of 9 symptoms (with at least 1 symptom being depressed mood or anhedonia) must be endorsed to meet criteria for major depressive disorder (MDD); this produces 105 unique symptom combinations. Given that distinct disorder components are likely associated with different pathophysiologies, it is not surprising that this heterogeneity has hindered our ability to identify genetic, neurobiological, and environmental factors contributing to depression (Hasler *et al.* 2004). To overcome these challenges, researchers have suggested focusing on narrowly defined and quantifiable phenotypes, which arguably represent a more direct expression of biological and environmental influences than the overall disorder (e.g., Meyer-Lindenberg & Weinberger, 2006).

Anhedonia, the loss of pleasure or lack of reactivity to pleasurable stimuli, is a promising depressive phenotype; it is a cardinal symptom of depression that has been associated with greater depression severity, poor treatment response, and reduced activity in reward-related brain regions (Hasler *et al.* 2004; Kasch *et al.* 2002; Keedwell *et al.* 2005). Despite theories suggesting that anhedonia is a genetically influenced vulnerability factor for depression (e.g., Meehl, 1975), few studies have investigated the heritability of hedonic capacity. Furthermore, the limited research available has relied exclusively on self-report measures tapping a broad range of hedonic processes. Growing evidence suggests, however, that hedonic capacity is not a

monolithic phenomenon, but can instead be parsed into distinct psychological, neural, and neurochemical subcomponents (e.g., Berridge & Kringelbach, in press). In light of these findings, it is perhaps not surprising that studies based on self-report assessments of anhedonia have yielded wide heritability estimates (from 27% to 82%; Dworkin & Saczynski, 1984; Berenbaum *et al.* 1990; Hay *et al.* 2001; Heath *et al.* 1994; Keller *et al.* 2005; Kendler *et al.* 1991; Linney *et al.* 2003; MacDonald III *et al.* 2001; Ono *et al.* 2002). In the present study, we employed a probabilistic reward task to objectively assess a fundamental aspect of hedonic capacity, reward responsiveness, which can be conceptualized as an individual's ability to modify behavior according to reinforcement history.

Increased stress sensitivity has been identified as a further promising depressive phenotype (Hasler *et al.* 2004). Animal research (e.g., Anisman & Matheson, 2005) supported by limited human findings (e.g., Bogdan & Pizzagalli, 2006) suggests that the depressogenic effects of stress may be partly attributable to stress-induced hedonic deficits. Surprisingly, with the exception of a recent study showing that the heritability of perceived stress ranges from 5 to 45% depending on self-report assessment (Federenko *et al.* 2006), little is known about the heritability of this important depressive phenotype.

The primary goals of the present study were to investigate (1) the feasibility of using a twin approach to assess the genetic contributions of a laboratory-based anhedonic phenotype that was recently shown to characterize MDD subjects (Pizzagalli *et al.* in press); (2) whether this objective measure of reward responsiveness is heritable; and (3) whether genetic and environmental influences are shared between reward responsiveness and perceived stress. A secondary goal was to replicate findings that perceived stress is heritable (Federenko *et al.* 2006).

We hypothesized that both reward responsiveness and perceived stress would be moderately heritable and share genetic and environmental components.

METHOD

Participants

The final sample consisted of 20 monozygotic (MZ) (age: 29.00 ± 10.90 ; 90% female; 95% Caucasian) and 15 dizygotic (DZ) (age: 33.73 ± 13.54 ; 87% female; 87% Caucasian) twin pairs who attended the 30th annual Twins Days Festival in Twinsburg, Ohio.¹ Zygosity groups did not differ in age, education, gender, ethnicity, income, or behavioral task performance ($p > 0.12$). All participants reported normal vision and no current or past psychiatric disorder, neurological illness or learning disorder. Participants received \$5 for their time and “won” \$5 during the probabilistic reward task. All participants provided informed written consent prior to participation. The Committee on the Use of Human Subjects at Harvard University approved the study.

Procedure

Participants completed the probabilistic reward task on a computer in a research booth on festival grounds. The following paper-and-pencil measures were collected: 1) demographic information; 2) two zygosity questionnaires (Kasriel & Eaves, 1976; Ooki *et al.* 1990); 3) the Mood and Anxiety Symptom Questionnaire (MASQ; Watson *et al.* 1995) to assess anxiety-specific symptoms (Anxious Arousal, AA), depression-specific symptoms (Anhedonic Depression, AD), and general distress (General Distress-Anxious Symptoms, GDA; General Distress-Depressive Symptoms, GDD); 4) the Beck Depression Inventory-II (BDI; Beck *et al.* 1996) to assess depressive symptomatology; and 5) the Perceived Stress Scale (PSS; Cohen *et al.*

1983) to assess subjective perception of life stress. In the present sample, Cronbach's alpha reliabilities for all questionnaires were excellent (0.83-0.94).

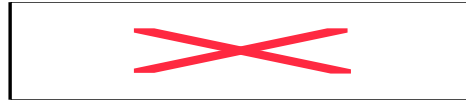
Probabilistic Reward Task

The reward task was adapted from Tripp & Alsop (1999) and has been described in detail and validated in multiple independent samples (e.g., Barr *et al.*, in press; Pizzagalli *et al.* 2005). In addition to standard measures of hit rate and reaction time (RT), this task allows for the computation of *response bias*, which reflects the participant's tendency to select one stimulus regardless of actual stimulus presentation. Unequal frequency of reward following correct identifications of two stimuli produces a systematic preference (response bias) for the response paired more frequently with reward (Macmillan & Creelman, 2005). In the present study, response bias was used to assess how subjects modulated their behavior as a function of prior reinforcement history.

Participants completed 3 blocks of 80 trials in which they decided if a mouth was either long (11 mm) or short (10 mm) by making an appropriate response on a computer keyboard ("v" or "m"; Fig. 1). Importantly, the small size difference between stimuli and the short exposure time made it difficult to ascertain which stimulus was presented. An asymmetric reward schedule between stimulus types was used to induce a response bias. Specifically, in each block, correct identification of one stimulus ("rich stimulus") was rewarded ("Correct!! You won 5 cents") three times more frequently (24 times) than the other ("lean stimulus"; 8 times). Key assignment and stimuli were counterbalanced across pairs. Participants were informed that their goal was to win as much money as possible and that not all correct responses would be rewarded.

Data Reduction

A two-step procedure was used to identify outlier responses (see Bogdan & Pizzagalli, 2006). Next, hit rates [= (number of hits)/(number of hits + number of misses)] and RT scores were calculated for rich and lean stimuli separately. Response bias was computed as follows:



Statistical Analyses

Twin analyses

Pearson correlation analyses provided MZ and DZ twin pair correlations. Model fitting can be used to estimate the extent of additive genetic (A), dominant genetic (D), common environment (C), and nonshared environment/measurement error (E) contributions (Rijsdijk & Sham, 2002; Purcell, 2001). The factor “A” represents the sum of the effects of individual alleles at all loci, whereas “D” captures interactions between alleles. “C” represents environmental influences shared by family members, while “E” captures individual-specific environment influences (and measurement error). In the present study, an ADE model was chosen based on the observations that (1) correlations involving block 3 response bias and MASQ GDD were more than twice as large in MZ than DZ twins, and (2) an ADE model provided a better fit than an ACE model (findings available upon request).

Since both hedonic capacity and perceived stress have been associated with depression severity (e.g., Candrian *et al.* 2007; Kasch *et al.* 2002), and stress diminishes reward responsiveness (e.g., Bogdan & Pizzagalli, 2006), three independent bivariate ADE Cholesky decomposition models were applied to evaluate shared and residual A, D and E contributions to depression, reward responsiveness (block 3 response bias), and perceived stress. The first Cholesky model specified three latent factors (A_1 , D_1 , and E_1) with pathways influencing both depression (MASQ GDD; a_{11} , d_{11} , and e_{11}) and reward responsiveness (a_{12} , d_{12} , and e_{12}), in

addition to three factors (A_2 , D_2 , and E_2) accounting for residual influences specific to reward responsiveness (Fig. 2). The second and third model were identical to the first, with the exception that perceived stress replaced reward responsiveness and perceived stress replaced GDD, respectively. These bivariate models yielded correlations between additive genetic (r_a), dominant genetic (r_d), and individual-specific environment factors (r_e) influencing the two phenotypes under investigation.

Full models were compared to nested submodels containing reduced parameters. Akaike's information criterion (AIC), which combines degrees of freedom with χ^2 goodness of fit, was used to evaluate model fit; the model with the lowest AIC value not significantly departing from the full model was chosen as the best fitting model, since it provides the best balance between parsimony and exploratory power. Model fitting analyses were performed with Mx (Neale *et al.* 1999) following established procedures (e.g., Kendler *et al.* 2007; Orstavik *et al.* 2007).

Analyses focused on response bias in block 3 because this variable fully captures overall reward responsiveness after contingencies have been learned.² The general depression distress (GDD) scale of the MASQ was used as a measure of depression because this subscale, unlike the BDI-II scale, is relatively unrelated to anhedonic symptoms (Watson *et al.*, 1995). This statistical non-overlap was important in light of prior findings linking decreased reward responsiveness to anhedonic symptoms (e.g., Pizzagalli *et al.* 2005).

Control analyses

An analysis of variance (ANOVA) with *Block* (1, 2, 3) was performed on response bias scores across all subjects. *Stimulus Type* (Rich, Lean) was added as a factor to hit rates and RT ANOVAs. Post-hoc Newman-Keuls tests evaluated significant ANOVA effects. Pearson

correlations were calculated to investigate relations between (1) response bias and (2) depressive/anxiety symptoms (MASQ and BDI) as well as perceived stress (PSS).

RESULTS

Twin analyses

All MZ twin correlations were higher than corresponding DZ correlations (Table 1). Bivariate Cholesky ADE decomposition model fitting results are shown in Table 2. The best fitting model for GDD and reward responsiveness was model IV, which dropped all dominant genetic pathways (d_{11} , d_{12} , d_{22}) and the common individual-specific environmental pathway (e_{12}) from the model. This model estimated that additive genetic influences explained 46% (95% CI =0.07-0.72) and 43% (95% CI =0.00-0.76) of the variance in reward responsiveness and GDD, respectively, while individual-specific environment/measurement error accounted for the remainder. According to this model, reward responsiveness and GDD are influenced by some of the same genes ($r_a=0.29$; 95% CI =-0.28-1.00). Thus, the overall heritability estimate of reward responsiveness can be subdivided into a small portion that was attributable to genetic effects also acting on GDD (0.04) as well as residual effects that were unique to reward responsiveness (0.42).³

The best fitting model for GDD and perceived stress was Model III, which dropped all dominant genetic pathways (d_{11} , d_{12} , d_{22}). According to this model, genetic contributions accounted for 40% (95% CI=0.00-0.75) and 44% (95% CI=0.05-0.70) of the variance in GDD and perceived stress, respectively; individual-specific environment/measurement error accounted for the remainder. The genetic correlation was estimated to be high ($r_a=0.67$) but the C.I. was wide (95% CI=-1.00-1.00). The individual-specific environment correlation was moderate ($r_e=0.33$; 95% CI=-0.10-0.66). Thus, the overall heritability estimate of perceived stress can be

subdivided into a large portion that was attributable to genetic effects acting on GDD (0.20) as well as a residual part that was unique to perceived stress (0.24). Similarly, the overall individual-specific environmental contribution can be subdivided into a small portion attributable to individual-specific environmental factors contributing to GDD (0.05) and unique contributions to stress perception (0.51).

The best fitting model for perceived stress and reward responsiveness was Model III. According to this model, additive genetic factors contributed to 45% (95% CI=0.12-0.70) and 48% (95% CI=0.14-0.73) of variance in stress perception and reward responsiveness, respectively; the majority of this genetic variance was shared between perceived stress and reward responsiveness ($r_a=0.72$; 95% CI=0.11-1.00) while individual-specific environment/measurement error factors were negatively correlated ($r_e=-0.43$; 95% CI=-0.69--0.04; Fig. 3). Thus, the overall heritability estimate of reward responsiveness can be subdivided into a large portion that overlaps with genetic effects acting on stress perception (0.25) as well as a residual component that was unique to reward responsiveness (0.23). Similarly, the overall individual-specific environmental contribution can be subdivided into a small portion attributable to factors contributing to perceived stress (0.10) and unique contributions to reward responsiveness (0.42).

Control analyses

Consistent with past research (e.g., Bogdan & Pizzagalli, 2006; Pizzagalli *et al.* 2005), analyses on response bias produced a main effect of *Block* [$F(2,136)=3.51$, $p<0.05$, partial $\eta^2=0.05$] due to increases from block 1 (0.10 ± 0.18) to block 2 (0.14 ± 0.19) and block 3 (0.16 ± 0.18 ; Newman-Keuls all $ps<0.04$). Control analyses on hit rates and RT data confirmed these results; rich hit rate increased over time and was greater than lean hit rate in each block and

RT decreased over time, more so for the rich stimulus type ($F_s > 3.50$, $p_s < 0.05$; all Newman-Keuls $p < 0.03$). Collectively, these findings suggest that the task elicited the intended effects; participants developed a behavioral preference toward the more frequently rewarded (rich) stimulus, as evident from the response bias, hit rate, and RT findings. Contrary to prior studies (Pizzagalli *et al.* 2005; Bogdan & Pizzagalli, 2006), however, no significant correlations emerged between self-report data and response bias (all $|r| < .19$, all $p_s > .12$).

DISCUSSION

The main goals of the present study were to (1) evaluate the feasibility of a twin approach to investigate genetic contributions to a laboratory-based anhedonic phenotype, (2) provide preliminary heritability estimates for reward responsiveness and perceived stress, and (3) assess the genetic and environmental correlation between perceived stress and reward responsiveness.

The present findings provide initial evidence that both reward responsiveness and perceived stress are heritable and influenced by individual-specific environmental factors. Consistent with previous literature assessing components of hedonic capacity (e.g., Loas, 1996) findings revealed that additive genetic factors and individual-specific environment/measurement error contributed to 46% and 54% of the variance in reward responsiveness, respectively. Moreover, replicating prior findings (Federenko *et al.* 2006), heritability estimates suggested that additive genetic factors contributed to 44% of the variance in stress perception while individual-specific environment contributed to the remainder. Of note, the genetic correlation between GDD and reward responsiveness was modest ($r_a = 0.29$). This finding is in line with conceptualizations suggesting that low positive affect and high negative affect are separate components of depression, with the former uniquely differentiating depression from anxiety and general negative affectivity being a nonspecific factor linked to both disorders (Watson *et al.* 1995).

More generally, this finding highlights the heterogeneity of depression and provides support for the endophenotypic research conceptualization (e.g. Hasler *et al.* 2004). In contrast to GDD and reward responsiveness, the genetic overlap between perceived stress and GDD was large ($r_a=0.67$); this overlap may be the result of robust associations between neuroticism, stress perception, and depression (e.g., Federenko, *et al.* 2006; Kendler *et al.* 2006).

Critically, this study suggests substantive overlap between genetic and individual-specific environmental factors influencing stress perception and reward responsiveness. Thus, genes that enhance perceived stress also increase reward responsiveness ($r_a=0.72$); conversely, individual-specific environmental factors that enhance perceived stress decrease reward responsiveness ($r_e=-0.43$). Genetic overlap between stress perception and reward responsiveness is intriguing, particularly when considering a large body of animal and human work emphasizing links between increased stress sensitivity and vulnerability to addiction, including evidence that stress can enhance the rewarding properties of addictive drugs (Hyman *et al.* 2006; Kreek *et al.* 2005). The negative correlations between environmental factors influencing perceived stress and reward responsiveness, on the other hand, raises the possibility that life stressors increasing stress perception might have deleterious consequences on the ability to modulate behavior as a function of reinforcers. Although speculative, this interpretation is consistent with prior findings of (1) a negative relationship between perceived stress and reward responsiveness (Pizzagalli *et al.* 2007) and (2) increased anhedonia when facing laboratory (Bogdan & Pizzagalli, 2006) and naturalistic (Berenbaum & Connelly, 1996) stressors. The positive genetic correlation and negative environment correlation between stress perception and reward responsiveness may account for the lack of a phenotypic correlation in the present study.

The limitations of this study warrant attention. First, although comparable to some prior twin studies (e.g., Berenbaum *et al.* 1990; Kendler *et al.* 1991; Mathews *et al.* 2007), the small sample size limited our statistical power; this is evidenced by large 95% confidence intervals. Second, data were collected outside controlled laboratory settings, which may have contributed to measurement error. However, MZ and DZ correlations were similar to those reported from other studies with larger samples (e.g., Federenko *et al.* 2005; Hay *et al.* 2001), and the general pattern of behavioral performance was comparable to prior independent samples tested with the same reward task in the laboratory.⁴ Unlike prior studies using this paradigm, however, no significant correlations emerged between the behavioral task and depressive measures, highlighting an important limitation of this study.

Despite these limitations, this is the first twin study, to our knowledge, that assesses: 1) hedonic capacity with an objective behavioral measure, and 2) genetic and environmental correlations between general depression, reward responsiveness and perceived stress. The findings of this study extend prior research using this probabilistic reward task in which reduced reward responsiveness has been associated with 1) elevated depressive (particularly anhedonic) symptoms (Pizzagalli *et al.* 2005) and a clinical diagnosis of depression (Pizzagalli *et al.* in press); 2) acute laboratory-induced stress (Bogdan & Pizzagalli, 2006) and elevated perceived stress (Pizzagalli *et al.* 2007); and 3) pharmacologically induced reduction of dopaminergic transmission (Pizzagalli *et al.* 2008). Collectively, these findings indicate that laboratory-based assessments of quantifiable aspects of depressive phenotypes might provide a powerful tool for parsing the heterogeneity characteristic of this complex and debilitating disease. In addition to replicating the present findings, molecular genetic approaches will be required to test the potential contributions of various candidate genes to hedonic capacity (e.g., Bogdan *et al.* 2006;

Noble, 2003) and perceived stress (e.g., Otte *et al.* 2007), which promise to provide critical insights into the etiology of depression.

Footnotes

¹ Sixteen pairs were excluded from analyses due to performance below chance level (n=4), unclear zygosity (n=3), incomplete reinforcement exposure (n=1), task noncompliance (n=1), pregnancy (n=1), outlier status (n=1) or the use of psychotropic medications (n=5).

² Structural equation modeling was not performed on overall reward learning (i.e., block 3 response bias – block 1 response bias) because MZ and DZ correlations for this variable were not significant.

³ Shared genetic contributions to reward responsiveness (0.04) were calculated as $[\text{sqrt}(0.46) \times 0.29]^2$, where 0.46 is the additive genetic influences to reward responsiveness and 0.29 is the genetic correlation (r_a) between GDD and reward responsiveness. Residual genetic contributions (0.42) were calculated as $0.46 - 0.04$.

⁴ To evaluate the psychometric properties of the present signal detection task, we compared the current data to data collapsed across three independent studies that were collected in the laboratory setting (Pizzagalli *et al.* 2005, 2007, in press). Results suggest no significant differences: *Study*, $F(1,241)=1.83$, $p>0.17$, *Study x Block*, $F(2,482)=0.09$, $p>0.90$. Together with the Cronbach's alpha reliability estimates for the questionnaires, these findings suggest that the subjective and objective data collected in this study had satisfactory psychometric properties.

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DECLARATION OF INTEREST

Mr. Bogdan reports no competing interests. Dr. Pizzagalli has received research support from GlaxoSmithKline and Merck & Co., Inc. for research unrelated to the present study.

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Table 1. Twin correlations for response bias and self-report measures

	MZ (n = 20 pairs)	DZ (n = 15 pairs)
Response Bias		
Block 3	.59***	-.05
Block 3 – Block 1	.05	-.18
Self Report measures		
MASQ GDA	.41*	.04
MASQ AA	.39*	-.11
MASQ GDD	.35	.06
MASQ AD	.68***	.39
BDI-II Total	.74***	.09
BDI-II Anhedonia ^a	.36	.01
BDI-II Melancholia ^b	.55**	.26
PSS	.35	.27

Notes: MASQ: Mood and Anxiety Symptom Questionnaire (Watson *et al.* 1995; GDA: General Distress-Anxious Symptoms; AA: Anxious Arousal; AD: Anhedonic Depression). BDI-II: Beck Depression Inventory-II (Beck *et al.* 1996). BDI-II Anhedonia (Pizzagalli *et al.* 2005): sum of BDI items associated with anhedonic symptoms (item 4: loss of pleasure; item 12: loss of interest; item 15: loss of energy; item 21: loss of interest in sex). BDI-II Melancholia (Pizzagalli *et al.* 2005): Sum of BDI items associated with melancholic symptoms (items 4, 12, 21; item 5: guilty feelings; item 11: agitation; item 6b: early morning awakening). PSS: Perceived Stress Scale (Cohen *et al.* 1983)

*** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$

Table 2. Bivariate Model Fitting for Block 3 Response Bias and Perceived Stress

Model	Common Pathways						Specific Pathways			Model Fit Parameters						
	Factor 1			Factor 2			Factor 2			GDDRB		GDDPSS		PSSRB		
	a ₁₁	d ₁₁	e ₁₁	a ₁₂	d ₁₂	e ₁₂	a ₂₂	d ₂₂	e ₂₂	df	AIC	p	AIC	p	AIC	p
I	+	+	+	+	+	+	+	+	+	129	178.58		708.87		176.23	
II	+	+	+	+		+	+		+	131	175.92	.51	705.25	.83	173.44	.55
III	+		+	+		+	+		+	132	175.06	.48	703.61	.86	171.44	.75
IV	+		+	+			+		+	133	173.14	.63	703.84	.56	174.07	.21
V	+		+			+	+		+	133	173.39	.59	704.13	.52	174.86	.16
VI	+		+			+			+	134	176.37	.17	704.60	.33	179.27	.02

Where a = additive genetic factors, d = dominant genetic factors, e = nonshared environment/measurement error, df = degrees of freedom, + = included pathway. GDDRB = bivariate Cholesky model with general distress depression (MASQ GDD) (Factor 1) and block 3 response bias (Factor 2). GDDPSS = bivariate Cholesky model with MASQ GDD (Factor 1) and perceived stress (Factor 2). PSSRB = bivariate Cholesky with perceived stress (Factor 1) and block 3 response bias (Factor 2). AIC = Akaike's Information Criterion. The lowest AIC value determined the best model fit. Best fitting models are bolded.

Figure legends

FIG. 1. Schematic representation of the task design and trial presentation.

FIG. 2. Path diagram for the bivariante Cholesky decomposition ADE model. The “A” and “D” components are correlated with $r = 1.00$ between MZ twins and 0.5 and 0.25 in DZ twins, respectively. Three independent bivariate models were run: Model 1: Factor 1 = GDD, Factor 2 = block 3 response bias; Model 2: Factor 1 = GDD, Factor 2 = perceived stress; Model 3: Factor 1 = perceives stress, Factor 2 = block 3 response bias.

FIG. 3. The best fitting bivariate Cholesky decomposition model for perceived stress and reward responsiveness. Numbers provided are % of variance.

FIG. 1

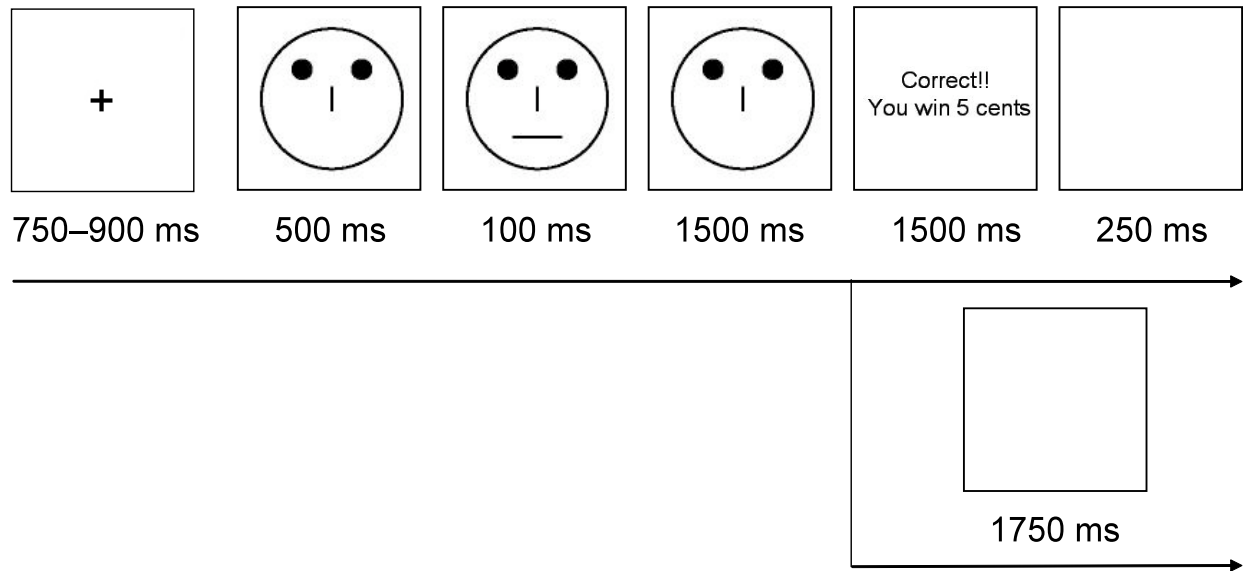


FIG. 2

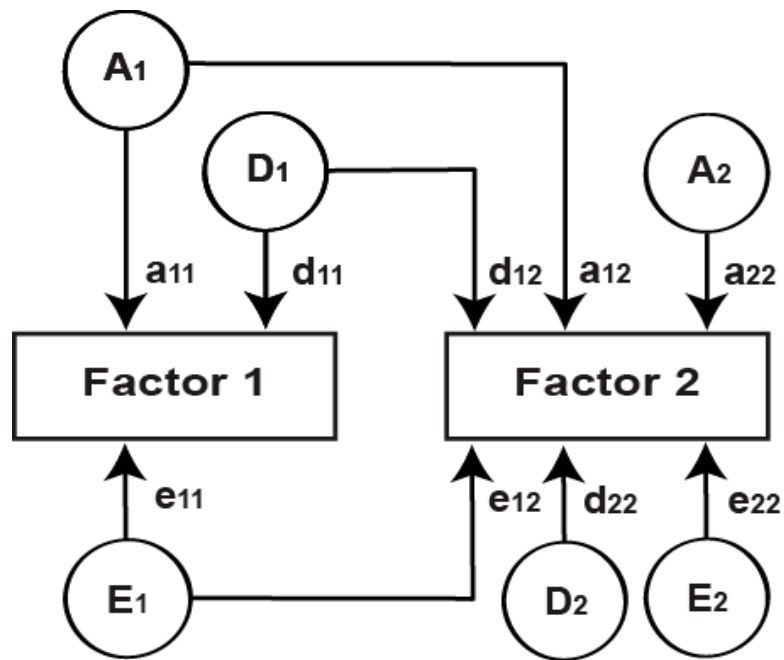


FIG. 3

