Visual Search for Feature and Conjunction Targets with an Attention Deficit

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Visual Search for Feature and Conjunction Targets with an Attention Deficit

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Abstract

Brain-damaged subjects who had previously been identified as suffering from a visual attention deficit for contralesional stimulation were tested on a series of visual search tasks. The experiments examined the hypothesis that the processing of single features is preattentive but that feature integration, necessary for the correct perception of conjunctions of features, requires attention (Treisman & Gelade, 1980; Treisman & Sato, 1990). Subjects searched for a feature target (orientation or color) or for a conjunction target (orientation and color) in unilateral displays in which the number of items presented was variable. Ocular fixation was controlled so that trials on which eye movements occurred were cancelled. While brain-damaged subjects with a visual attention disorder (VAD subjects) performed similarly to normal controls in feature search tasks, they showed a marked deficit in conjunction search. Specifically, VAD subjects exhibited an important reduction of their serial search rates for a conjunction target with contralesional displays. In support of Treisman’s feature integration theory, a visual attention deficit leads to a marked impairment in feature integration whereas it does not appear to affect feature encoding.

INTRODUCTION

One of the most influential proposals in vision research during the last decade is Treisman’s feature integration theory of attention (Treisman, 1988; Treisman & Gelade, 1980; Treisman & Sato, 1990; Treisman & Schmidt, 1982; Treisman & Souther, 1985). According to this model, the encoding of visual features (e.g., orientation or color) is performed through analyses that are preattentive and accomplished in a spatially parallel fashion across the visual field. Conversely, one task that is assumed to require attentional processing is the integration of visual features, which is a necessary condition for correctly perceiving feature conjunctions (e.g., colored shapes).

To now, most of the evidence for these two hypotheses comes from studies performed on normal, neurologically intact individuals. In the present paper, our main purpose is to examine more directly the role of attention in visual feature integration by studying the performance of left brain-damaged subjects in visual search tasks where the target is distinguished from the distractors by a single feature or by a conjunction of features. In particular, we are interested in the performance of a subgroup of left brain-damaged subjects who have been shown, in a previous series of experiments (Arguin, Cavanagh, & Joanette, 1992), to suffer from a disorder affecting the allocation of attention to stimuli presented in their contralesional visual hemifield.

VISUAL FEATURES AND THEIR INTEGRATION

Objects in our visual environment differ from one another on a number of perceptual dimensions. Empirical demonstrations have indicated, in the human as well as the animal, that the visual system comprises a number of functional modules, each specialized for the processing of visual features on a specific perceptual dimension (see, among others, Allman, Baker, Newsome, & Peterson, 1981; Cavanagh, 1988; De Yoe & Van Essen, 1988; Kaas, 1989; Zeki, 1978). However, the identification of visual objects not only requires that features along each of these dimensions be specified, but also demands an explicit representation of how they are combined to constitute each individual object (Treisman & Gelade,
At present, it appears that the most viable hypothesis regarding the process responsible for conjoining features is that it is performed by focusing visuospatial attention at the location occupied by the stimulus (Keele, Cohen, Ivry, Liotti, & Yee, 1988; Treisman & Sat0, 1990). Treisman and her collaborators have proposed a theory of visual feature integration that applies this principle (Treisman, 1982, 1983, 1985, 1988, 1991; Treisman & Gelade, 1980; Treisman & Paterson, 1984; Treisman & Schmidt, 1982; Treisman & Sat0, 1990; Treisman & Souther, 1985; Treisman et al., 1977). This theory proposes two distinct stages in visual processing. The first, called preattentive, allows the spatially parallel encoding of features that are present in the visual field. At the preattentive stage, information regarding the location of specific features is not made explicitly available. The second stage, called attentive, is responsible for the integration of features that have been encoded at the preattentive stage by focusing visuospatial attention at the location occupied by a particular stimulus. Without attention, the features that are encoded cannot be located properly—they are said to be free-floating—and may thus lead to the perception of illusory conjunctions (i.e., the erroneous integration of features that do not belong to the same object; Briand & Klein, 1987; Eglin, 1987; Prinzmetal, 1981; Prinzmetal, Presti, & Posner, 1986; Treisman & Paterson, 1984; Treisman & Schmidt, 1982; Treisman et al., 1977; Virzi & Egeth, 1984).

Empirical support for Treisman's feature integration theory has been reported with the use of various experimental paradigms in neurologically intact subjects (see Treisman & Sat0, 1990, for a review). One main source of evidence relative to the function of visuospatial attention in feature integration comes from the visual search paradigm. In the present paper, we will concentrate on this particular paradigm.

In the visual search task, subjects are asked to indicate, as rapidly as possible while avoiding errors, whether a prespecified stimulus, called target, is present in an array of other stimuli, called distractors, whose number is variable. When the target differs from distractors by a unique feature (e.g., red target, green distractors) a large number of experiments show that the time required by normal individuals to detect the presence of a target is independent of the number of stimuli that are presented (Arguin & Cavanagh, 1988; Bergen & Julesz, 1983; Cavanagh, Arguin, & Treisman, 1990; Dick, Ullman, & Sagi, 1987; Nakayama & Silverman, 1986; Quinlan & Humphreys, 1987; Sagi & Julesz, 1985; Treisman, 1982, 1983, 1985, 1988; Treisman & Gelade, 1980; Treisman & Gorman, 1988; Treisman & Souther, 1985; Treisman et al., 1977). Thus, the slope of the linear regression of response times (RTs) as a function of the number of items is about 0 msec/item. These observations indicate that detection of a feature target is performed by a spatially parallel process that does not involve the selection of individual stimuli by visuospatial attention (Snodgrass & Townsend, 1980), and thus support the hypothesis of a preattentive encoding of visual features. More specifically, this preattentive process is said to involve the monitoring of signals coming from a sensory map responsible for the encoding of the target's distinctive feature.

In contrast to feature targets, conjunction targets can be distinguished from distractors only by considering the combinations of features constituting the items presented (Treisman & Gelade, 1980). For example, if the target is a red horizontal bar, one subgroup of distractors is made of red vertical bars and another of green horizontal bars. With such displays, the target cannot be distinguished from distractors by a unique feature since it shares its color with some distractors and its orientation with others. In most of the experiments requiring normal subjects to search for a conjunction target, results show linearly increasing RTs with the number of stimuli displayed (Arguin & Cavanagh, 1988; 1990; Arguin, Jo-anette, & Cavanagh, 1990; Dehaene, 1989; Egeth, Virzi, & Garbart, 1984; Houck & Hoffman, 1986; Pashler, 1987; Quinlan & Humphreys, 1987; Treisman, 1982, 1983, 1985, 1988; Treisman & Gelade, 1980; Treisman & Sato, 1990; Treisman et al., 1977; Wolfe, Cave, & Franzel, 1989). This is taken as an indication that search is performed through the sequential examination of each item in turn (Snodgrass & Townsend, 1980). It thus appears that focusing visuospatial attention at the location occupied by a particular stimulus is necessary to correctly perceive the conjunction of its features, in support of Treisman's feature integration theory.

However, there exists another set of observations that pose a challenge to Treisman's feature integration theory. Indeed, experiments led by Nakayama and Silverman (1986), Steinman (1987), McLeod, Driver, and Crisp (1988), Sagi (1988), Wolfe et al. (1989), Dehaene (1989), Treisman and Sato (1990), and Arguin and Cavanagh (1990) show that visual search for some types of conjunction targets can be performed correctly with a very shallow increase in RTs as a function of the number of items displayed, suggesting a process that is spatially parallel instead of serial. On this basis, a number of authors have proposed that attention may not be required to integrate some feature combinations (McLeod et al., 1988; Nakayama & Silverman, 1986; Sagi, 1988), contrary to the basic hypothesis of Treisman's feature integration theory.

**BRAIN DAMAGE AND VISUAL SEARCH FOR FEATURE AND CONJUNCTION TARGETS**

Riddoch and Humphreys (1987) and Eglin, Robertson, and Knight (1989) examined the visual search performance of brain-damaged subjects who have been diagnosed as showing visual hemineglect, which is a
syndrome generally thought to reflect, at least in part, a
deficit of visual attention for contralesional stimulation
(Heilman, 1985; Kinsbourne, 1970; Mesulam, 1981; Mor-
row & Ratcliff, 1988; Roy, Reuter-Lorenz, Roy, Copland,
& Moscovitch, 1987). Specifically, Riddoch and Hum-
phreys (1987) and E格林 et al. (1989) have compared the
effect of the number of items displayed on RTs as a
function of the side (ipsilesional vs. contralesional)
of the display to which the target was presented. Perform-
ance was examined in search tasks involving either
feature or conjunction targets. Riddoch and Humphreys
(1987) tested subjects only with bilateral displays while
E格林 et al. (1989) used both unilateral and bilateral
displays. None of these studies performed any control
for the locus of eye fixation at the onset of trials or
checked for the occurrence of eye movements during
the search process. In fact, many of the subjects tested
by Riddoch and Humphreys (1987; two out of three
patients) and by E格林 et al. (1989; four out of seven
patients) suffered from either a verified or probable
visual field defect. Therefore, it is quite possible that a
number of neglect subjects in those experiments
searched the displays by examining them through their
ipsilesional hemifield. These are not ideal conditions to
test the performance of the damaged hemisphere on
visual search tasks and this may explain some of their
findings, as indicated below.

The results reported by Riddoch and Humphreys
(1987) suggest that neglect subjects performed a spatially
parallel search for a feature target, whatever side to
which the target was presented. In contrast, E格林 et al.
(1989) reported results suggesting that neglect subjects
performed the search for a feature target through a serial
process irrespective of the side of the display on which
the target appeared. This result differs from that observed
in normal individuals who, as indicated earlier, typically
search for a feature target through a spatially parallel
process. Both the studies reported by Riddoch and Hun-
phreys (1987) and by E格林 et al. (1989) showed main
effects indicating longer RTs for a contralesional target
than for an ipsilesional one. It is possible that this last
result may be partly attributable to the eye movements
that neglect subjects with a visual field deficit had to
perform in order to view the target in their ipsilesional
hemifield. An examination of the individual data reported
by Riddoch and Humphreys (1987) and by E格林 et al.
(1989) nevertheless suggests that this is not the only
factor involved in the RT asymmetry observed in neglect
subjects since it was apparent even in patients without
any visual field defect.1

With a conjunction target, the results reported in these
two studies indicate a regular increase of RTs with the
number of items displayed, thus suggesting a serial
search that involved focused attention. Again, RTs were
longer with contralesional targets. However, no signifi-
cant difference in the effect of the number of items on
RTs as a function of the side to which the target was
presented is reported, even though Riddoch and Hum-
phreys (1987) underlined a clear trend for a slower serial
search with contralesional targets. As mentioned previ-
ously, no control of ocular fixation was performed in
these experiments and this may have been a factor in
the lack of significant hemifield differences in the rate of
conjunction search.

In the present paper, we report a series of three visual
search experiments that were designed as a test of two of
Treisman's basic hypotheses about the encoding of
visual features and feature conjunctions, namely that sin-
gle feature encoding can be performed in a spatially
parallel manner without the aid of focused attention, and
that the correct perception of feature conjunctions re-
quires that attention be focused at the stimulus location.
Strict eye movement monitoring and brief exposure du-
rations were used to ensure control over the laterali-
ization of the stimuli presented.

The tasks used required subjects to search for a feature
target differing from distractors either by its orientation
(horizontal vs. vertical; Exp. 1) or its color (red vs. green;
Exp. 2), or to search for a conjunction target that differed
from distractors by a specific combination of orientation
and color (Exp. 3). The features used in the latter ex-
periment were the same as those used in Exps. 1 and 2.
Stimulus displays were made of one, two, three, or four
items and were shown unilaterally for a brief duration.
Ocular fixation was controlled in order to make sure of
the proper laterization of stimulation.

The experimental group was made up of left brain-
damaged2 subjects who had previously shown evidence
for an attention deficit for contralesional stimulation in
a separate set of visuospatial cuing experiments (VAD
group—for visual attention deficit; Arguin et al., 1995).
The attention deficit of the VAD group was revealed by
longer RTs to contralesional than ipsilesional targets at
short stimulus onset asynchronies (50 and 150 msec)
following a spatial cue that indicated, on 80% of the trials,
the hemifield in which the target would occur (invalid
cues were presented on the remaining 20% of trials).
The occurrence of this hemifield asymmetry was not
directly related to the validity of the spatial cue (i.e.,
whether it indicated the target location or not) and was
even seen in a task where no cue validity effect occurred
(see the subject description in the Method section for a
full description). The RT asymmetry between ipsilesional
and contralesional targets was much reduced or absent
at longer stimulus onset asynchronies (600 and 1000
msec) between cues and targets. This pattern of results
suggested that VAD subjects had a deficit in maintaining
a proper level of alertness for contralesional stimulation.
Of the control groups used in the present set of visual
search experiments, one was made of left brain-damaged
subjects who did not show any evidence for an attention
disorder (i.e., normal performance) in visuospatial cuing
tasks (nVAD group). The other control group was made
up of neurologically intact subjects whose age approxi-
mately matched that of the brain-damaged subjects (normal controls). None of the brain-damaged subjects showed any visual field impairment. The aspect of main interest in the visual search results reported below were differences between hemifields in the effect of the number of items presented on RTs. Following Treisman’s feature integration theory, an attention deficit should not affect performance in the search for a feature target since this task is assumed to be executed by preattentive processes. In contrast, since it is assumed that search for a conjunction target requires attentional processing, it was expected that subjects suffering from a visual attention deficit (VAD group) would show impaired performance with right-hemifield (contralateral) displays relative to control subjects. Specifically, if attention is required for visual feature integration, subjects with an attention deficit should take longer to process each individual item located in the part of space to which they have difficulty to attend, thus leading to a larger increase of RTs with the number of stimuli presented.

RESULTS AND DISCUSSION

Experiment 1—Orientation Search

Figure 1 shows the correct RTs observed in each group when the feature target differed from distractors displayed with it by its orientation. Table 1 depicts the error rates observed in this task. The correlation between RTs and error rates was null ($r = 0.00$), thus indicating the absence of a speed-accuracy trade-off. To briefly summarize the analyses described below, the results of Exp. 1 did not reveal any significant linear increase of RTs as a function of the number of stimuli presented in any condition or for any of the subject groups. Rather, on target-absent trials, all groups showed evidence for longer RTs if only a single item was shown than if two, three, or four were exposed. For the VAD group, data analyses also suggested a comparable effect of the number of stimuli on target-present trials. No evidence for an effect of the number of items on target-present trials was observed in the control and nVAD groups.

**Results**

A mixed-factor ANOVA was performed on the median correct RTs observed in each subject. Within-subject factors were the visual hemifield to which the stimuli were displayed (hemifield), the target’s presence or absence (target), and the number of items presented (number). The between-subject factor was made from the three groups (control, nVAD, and VAD). Similar analyses were also performed on the data of Exps. 2 and 3.

The analysis of variance applied on the correct RTs showed significant two-way interactions of group × number of items [$F(6, 45) = 2.5, p < 0.05$], and of target × number [$F(3, 45) = 4.2, p < 0.025$]. No other effect was significant.

Analysis of the simple effects of the group × number interaction showed that the effect of the number of items
was significant only in the VAD group \([F(3, 45) = 2.8, \text{n.s.}]\) for the normal controls, \([F(3, 45) = 2.0, \text{n.s.}]\) for the nVAD group, \([F(3, 45) = 3.5, p < 0.025]\) for the VAD group. A linear regression applied on the pooled results of the VAD group indicated a nonlinear decrease of RTs with the number of items presented (slope = -17.6 msec/item, \(r = 0.58\)). This effect mainly consisted in longer RTs when only one stimulus was presented than when two, three, or four were displayed.

Analysis of the simple effects of the target \(\times\) number interaction indicated that the effect of number of items was significant only on target-absent trials \([F(3, 45) = 2.7, \text{n.s.}]\) for target-present, \([F(3, 45) = 3.5, p < 0.025]\) for target-absent. A linear regression applied on the pooled results on target-absent trials showed that the variation of RTs as a function of the number of items was nonlinear \((r = 0.28)\) and was most apparent in that RTs were longer if only one stimulus was displayed than when two, three, or four items were presented.

**Discussion**

The main aspect of the results of Exp. 1 is that there was no significant increase of RTs with the number of stimuli presented in any condition or group. This is consistent with the existing literature indicating the spatially parallel search for feature targets. In fact, RTs often decreased as the number of items in the display increased. Thus, for all groups but only on target-absent trials, RTs were longer if only one item was presented than if two or more were in the display. An indication for a comparable effect of the number of items on target-present trials was observed in the VAD group. This latter effect was quite weak however, and only emerged because of shorter RTs with displays made of two stimuli. Thus, RTs of the VAD group on target-present trials were of 1028, 1002, 1024, and 1027 msec for display sizes of 1, 2, 3, and 4, respectively. Performance differences between target-present and target-absent trials have been reported on several occasions in the literature on feature search (e.g., Cavanagh et al., 1990; Farmer & Taylor, 1980; Quinlan & Humphreys, 1987; Treisman, 1982, 1991; Treisman & Gelade, 1980). These differences indicate that even though feature encoding proceeds through a spatially parallel process in both kinds of trials, the particular criteria on which the decisions as to the presence or the absence of the target may vary (see Treisman, 1991; Treisman & Gelade, 1980; Treisman & Gormican, 1988).

The effect of the number of items on RTs that occurred on target-absent trials suggests that subjects used a strategy in which target-absent responses were based on the homogeneity of the stimulus display. On this account, when two or more stimuli were shown on a target-absent trial, the homogeneity of the display (all distractors were vertical) was used by subjects as a signal that the target (i.e., a contrasting stimulus—horizontal bar) was absent. Obviously, when only one item was shown on target-absent trials, this homogeneity detection strategy could not apply. In that case, subjects may have had to rely on some explicit identification of the stimulus or on a comparison between it and an internal target template, thus leading to longer RTs. Two points should be mentioned about this homogeneity detection strategy on target-absent trials.

First, the application of such a process is apparently not uncommon, as it has previously been observed in several experiments with normal subjects (Duncan & Humphreys, 1989; Estes, 1972; Farmer & Taylor, 1980; Gordon, 1968; Gordon, Dulewicz, & Winwood, 1971; McIntyre, Fox, & Neale, 1970; Moraglia, Maloney, Fekete, & Al-Basi, 1989; Treisman, 1991; Treisman & Sato, 1990; Wolfe, Friedman-Hill, Stewart, & O'Connell, 1992). One example of this was reported by Farmer and Taylor (1980). In an experiment where subjects had to search for an achromatic target among chromatic distractors, these authors manipulated the level of homogeneity of the distractors. They showed that RTs to target-absent trials were greatly reduced by the use of homogeneous distractors, as opposed to a heterogeneous set.

Second, the application of the strategy of homogeneity detection indicates the capacity of subjects in encoding orientation through a spatially parallel process. Indeed, the successful application of a homogeneity detection process implies the parallel processing of orientation stimuli. That is, no benefit from display homogeneity

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**Table 1. Percentages of Errors Observed in the Orientation Search Experiment (Experiment 1)**

<table>
<thead>
<tr>
<th>Number of Items</th>
<th>Group</th>
<th>L-A</th>
<th>L-P</th>
<th>R-A</th>
<th>R-P</th>
</tr>
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<td>1.3</td>
<td>5.3</td>
<td>1.3</td>
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</tr>
<tr>
<td></td>
<td></td>
<td>2.7</td>
<td>1.3</td>
<td>2.0</td>
<td></td>
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<tr>
<td></td>
<td></td>
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<td>1.3</td>
<td>7.3</td>
<td>1.3</td>
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<td>6.6</td>
<td>1.3</td>
<td>10.6</td>
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<td>12.0</td>
<td>12.0</td>
</tr>
<tr>
<td></td>
<td></td>
<td>2.7</td>
<td>8.0</td>
<td>8.0</td>
<td>12.0</td>
</tr>
<tr>
<td></td>
<td></td>
<td>5.3</td>
<td>2.7</td>
<td>8.0</td>
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<td></td>
<td></td>
<td>5.3</td>
<td>9.3</td>
<td>12.0</td>
<td>8.0</td>
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<td>4.4</td>
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<td>4.4</td>
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</tbody>
</table>

*L: Left hemifield; R: right hemifield; A: target-absent trials; P: target-present trials.*
would have occurred in the present experiment if search had been based on the sequential processing of the stimuli.

The converse of homogeneity detection, which is the detection of heterogeneity (i.e., a discrepant item), did not occur on target-present trials in Exp. 1. This type of search process would predict longer RTs with displays of one item than with displays made of two or more stimuli. In contrast, the target-present results indicated no significant variation of RTs with the number of items. Such observations are consistent with the view that the presence of a feature target (here a horizontal bar) is detected on the basis of activity within the sensory map involved in the encoding of the distinctive feature of the target (e.g., Treisman, 1991; Treisman & Gormican, 1988).

The main purpose of Exp. 1 was to examine whether, as in neurologically intact individuals, the orientation of visual lines can be encoded by a spatially parallel process in subjects who suffer from an attention deficit (VAD group). Since the orientation search performance of the VAD subjects was nearly identical to that of the control and nVAD groups, the results of Exp. 1 support the hypothesis that an attention deficit does not prevent the spatially parallel processing of orientation.

**Experiment 2—Color Search**

Figure 2 shows the RTs observed in each group when the target differed from distractors by its color. Table 2 depicts the error rates observed in this task. The correlation between RTs and error rates was positive ($r = +0.22$), thus indicating the absence of a speed-accuracy trade-off.

In summary, the results of Exp. 2 indicated that the number of items did not affect RTs with left-hemifield exposures. With right-hemifield displays however, the number of stimuli shown affected performance, although differently on target-absent and target-present trials. Thus, with target-absent right-hemifield presentations, RTs were longer when only one item was shown than when two, three, or four were present. With right-hemifield displays in which the target was present, normal controls as well as VAD subjects showed a significant linear increase of RTs with the number of items presented. The effect of the number of stimuli did not differ between these two groups.

**Results**

The analysis of variance applied on the correct RTs in Exp. 2 revealed a significant two-way interaction of target × number of items [$F(3, 45) = 9.1, p < 0.001$], a three-way interaction of hemifield × target × number [$F(3, 45) = 4.1, p < 0.025$], and a four-way interaction of hemifield × target × group × number [$F(6, 45) = 3.3, p < 0.01$]. No other significant effect was found. Since the four-way interaction included all the factors examined, only the simple effects related to this interaction were analyzed.

For left-hemifield displays, analysis of the hemifield ×
Table 2. Percentages of Errors Observed in the Color Search Experiment (Experiment 2)*

<table>
<thead>
<tr>
<th>Number of Items</th>
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<th>L-P</th>
<th>R-A</th>
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<td>nVAD</td>
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<td>8.9</td>
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<td></td>
</tr>
</tbody>
</table>

*Conventions are the same as in Table 1.

target × group × number interaction showed no significant effect. In contrast, a similar analysis performed on the data observed with right-hemifield stimulation showed a target × number interaction \(F(3, 45) = 13.8, p < 0.001\) as well as a target × group × number interaction \(F(6, 45) = 3.8, p < 0.005\).

An examination of the group and number effects on target-absent right-hemifield trials indicated a significant effect only of the number of items \(F(3, 45) = 5.2, p < 0.005\). A linear regression of the data observed on these trials, pooled across groups, showed a reduction of RTs with increasing number of stimuli (slope = -12.7 msec/item, \(r = 0.63\)). This effect was mainly manifest by the longer RTs observed when only one item was presented relative to displays with two, three, or four stimuli. With target-present right-hemifield displays, the analysis showed an effect of number of items \(F(3, 45) = 6.9, p < 0.005\) and a marginally significant interaction of group × number of items \(F(6, 45) = 2.2, p < 0.06\). Decomposition of this interaction revealed a significant effect of number of items in both the normal controls \(F(3, 45) = 4.7, p < 0.025\) and the VAD group \(F(3, 45) = 5.2, p < 0.01\). In contrast, no effect of the number of stimuli \(F(3, 45) = 1.3, n.s.\) was observed in the nVAD group on target-present right-hemifield trials. Linear regressions on the target-present right-hemifield data for the control and VAD groups showed linear increases of RTs with the number of stimuli in both groups (controls: slope = 32.8 msec/item, \(r = 0.99\); VAD: slope = 58.6 msec/item, \(r = 0.92\)). A separate ANOVA was run in order to directly compare these effects of the number of items in the control and VAD group. It indicated that the magnitude with which RTs increased with additional items on target-present trials with right-hemifield stimulation did not differ between the two groups [interaction of group × number: \(F(3, 33) = 1.4, n.s.\)].

Discussion

The results seen in Exp. 2 with left-hemifield displays are congruent with the hypothesis of a spatially parallel search for a color target. Thus no group showed any effect of the number of items on RTs with left-hemifield stimulation. This result corresponds to what is generally reported in the literature on color search in a normal population and suggests that with left-hemifield displays, the task was performed through the monitoring of signals from a feature map responsible for encoding the target. The observations with right-hemifield stimulations are more complex however. Thus, even though there is evidence in all groups that color could be encoded by a spatially parallel process with right-hemifield displays, the actual operations by which target-present and target-absent decisions were reached were markedly different.

On right-hemifield target-absent trials, all groups showed longer RTs to displays containing only one item than to displays of two, three, or four stimuli. This result suggests a homogeneity detection strategy similar to that observed on target-absent trials in Exp. 1. Thus, with displays within which stimulus homogeneity could be detected (set sizes of two, three, or four), RTs were reduced relative to a condition in which no such signal was present (set size of one). As mentioned in the discussion of Exp. 1, this type of effect suggests that the color of the stimuli could be encoded by a spatially parallel process. Hence, had colors been treated through a serial process in this condition, no benefit from display homogeneity would have been observed.

In sharp contrast, however, normal controls and VAD subjects, but not nVAD subjects, showed significant linear increases of RTs with increasing number of items on right-hemifield target-present trials. This suggests that before responding that the target was present, subjects performed a sequential examination of individual items. Serial search for a color target has been observed before, but only in cases where the colors used for the distractors were very similar to that of the target (Treisman, 1991; Treisman & Gormican, 1988). Here, the target and distractors were highly discriminable (red vs. green) and such contrasting stimuli normally give flat RT functions with varying numbers of items. Finally, our results on right-hemifield target-present trials do not appear attributable to the use of equiluminant colors in Exp. 2 (see Method section). Previous visual search observations by Cavanagh et al. (1990) and by D'Zmura (1991) with equi-
The correlation between RTs and error rates was positive ($r = +0.39$), thus indicating the absence of a speed-accuracy trade-off.

The results of Exp. 3 indicated linear increases of RTs as a function of the number of items displayed and this effect did not vary as a function of target presence or absence. With left-hemifield stimulation, the only significant effect observed was one of the number of stimuli. Thus, RTs increased linearly with the number of items presented and this effect did not vary significantly as a function of group. In contrast, with right-hemifield displays, VAD subjects exhibited a larger increase of RTs with the number of stimuli than the control group. The nVAD group showed a performance that was comparable with that of the normal controls. Details of the data analyses are presented below.

### Results

The analysis of variance applied on the correct RTs revealed significant main effects of hemifield [$F(1, 15) = 28.5, p < 0.001$] and of number of items [$F(3, 45) = 42.5, p < 0.001$], two-way interactions of hemifield × group [$F(2, 15) = 20.8, p < 0.001$], of group × number [$F(6, 45) = 2.6, p < 0.05$], and of hemifield × number [$F(3, 45) = 7.6, p < 0.001$], and a three-way interaction of hemifield × group × number [$F(6, 45) = 3.2, p < 0.025$]. No other significant effect was found. Only the three-way interaction was analyzed any further since it included all the factors involved in the other significant effects.

For left-hemifield stimulation, simple effects analysis of the hemifield × group × number interaction only showed a main effect of the number of stimuli [$F(3, 45) = 23.11, p < 0.001$]. The linear regression of RTs as a function of the number of stimuli for left-hemifield displays, which was applied on the results pooled across groups and target-absent and target-present trials, indicated a fairly large linear increase of RTs with the number of items (slope = 58.1 msec/item, $r = 0.96$).

The analysis of the simple effects for right-hemifield exposures also indicated an effect of the number of stimuli [$F(3, 45) = 37.14, p < 0.001$], but also a significant interaction of group × number [$F(6, 45) = 3.46, p < 0.01$]. Analyses of this interaction revealed that the number of items presented had a significant effect on RTs in all three groups ($F(3, 45) = 11.3, p < 0.001$) for the normal controls, [$F(3, 45) = 6.0, p < 0.005$] for the nVAD group, [$F(3, 45) = 22.1, p < 0.001$] for the VAD group). Linear regressions applied to these data indicated, in the three groups, linear increases of RTs with increasing number of stimuli (controls: slope = 67.1 msec/item,
Figure 3. Mean RTs observed in Exp. 3 (conjunction search). (A) Control group. (B) nVAD group. (C) VAD group. Square, left hemifield display; circle, right-hemifield display; filled symbols, target-present trials; empty symbols, target-absent trials. Note that the scale of the vertical axis differs from that used in Figures 1 and 2.

Table 3. Percentages of Errors Observed in the Conjunction Search Experiment (Experiment 3)*

<table>
<thead>
<tr>
<th>Number of items</th>
<th>Group</th>
<th>L-A</th>
<th>L-P</th>
<th>R-A</th>
<th>R-P</th>
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</tbody>
</table>

*Conventions are the same as in Table 1.

$r = 1.0$; nVAD: slope = 68.9 msec/item, $r = 0.99$; VAD: slope = 168.8 msec/item, $r = 0.98$). Separate analyses designed to compare the magnitude of the effect of the number of items with right-hemifield stimulation between the normal controls and each of the brain-damaged groups were conducted. These showed that while the effect of the number of items did not differ significantly between the control and nVAD groups (interaction of group x number: $F(3, 39) < 1$), subjects in the VAD group exhibited a larger increase of RTs with additional stimuli than the normal controls (interaction of group x number: $F(3, 39) = 5.3, p < 0.005$).

Discussion

One first important aspect of the results observed during the search for a conjunction target is that there was a linear increase of RTs with the number of stimuli presented, as demonstrated by the linear regression analyses. This is consistent with the existing literature on conjunction search and indicates that the task was performed by the serial processing of the stimuli displayed.

A second point worth noting in the results of Exp. 3 is that the effect of number of items displayed on RTs did not vary significantly as a function of target presence or absence. This is at variance with most of the literature on conjunction search, which indicates slopes of RTs as
Thus, VAD patients, who have an attention deficit for contralateral displays in a task requiring the search for a conjunction target, showed an impaired search rate with contralateral conjunction displays. More specifically, their search performance suggests an increase in the time required to process each individual conjunction stimulus presented in their contralateral visual field.

GENERAL DISCUSSION

The purpose of the present study was to assess the effect of an attention disorder on the perceptual processes involved in the encoding of visual features and in their integration. Previous research, notably by Treisman and collaborators, suggested that the encoding of visual attributes is a preattentive process performed by modules specialized for the analysis of specific visual dimensions. In contrast, the integration of visual features, necessary for the correct perception of their conjunctions, is a process that requires attention to be focused at the stimulus location. According to these hypotheses, a disorder of visual attention should result in defective feature integration, while leaving intact the processes that are responsible for the encoding of these features.

The experiments reported here support these hypotheses. Thus, with respect to the encoding of visual features, Exps. 1 and 2 demonstrated that brain-damaged subjects with an attention deficit performed similarly to normal controls in visual search tasks in which the target differed from distractors by a single feature, namely orientation and color. More specifically, these observations provided evidence for the capacity of patients with an attention disorder to encode these visual features in a spatially parallel manner. Such a dissociation between an attention disorder and the encoding of visual features is consistent with the notion of preattentive processing of orientation and color features.

In contrast, in Exp. 3, left brain-damaged subjects with an attention disorder exhibited a marked deficit with contralateral displays in a task requiring the search for a conjunction target. Specifically, they showed a much steeper increase of RTs with the number of items with contralateral displays than that seen in normal controls with the same stimulation.

Among the alternative accounts that can be provided for the conjunction search deficit in the VAD group, some may readily be rejected. First, their impaired search rate for a conjunction target with contralateral stimulation may not be attributed to a nonspecific effect of brain damage since the nVAD group, who appeared to have intact attentional functions, did not show any impairment in this task relative to the control group. Moreover, the results of the VAD group cannot be accounted for by a deficit in encoding visual features. Thus, the results of Exps. 1 and 2, which involved the search for orientation and color targets, respectively, revealed no particular impairment in the VAD group and suggested a spatially parallel feature encoding in these subjects.
Therefore, the results of Exp. 3 indicate the importance of attentional processes for feature integration. That is, the slowing of serial search for a conjunction target with contralesional displays in the VAD group appears to be specifically related to their attention deficit for contralesional stimulation.

These observations from Exp. 3 are inconsistent with the hypothesis of preattentive feature integration that has been proposed by a number of investigators who reported very shallow slopes of RTs with the number of items displayed during the search for particular conjunction targets (McLeod et al., 1988; Nakayama & Silverman, 1986; Sagi, 1988). In fact, it appears that these rapid search rates for some features conjunctions may be accounted for while maintaining the hypothesis of attentional feature integration. Thus, Wolfe and his collaborators (Wolfe et al., 1989; see also Cave & Wolfe, 1990 and Wolfe, Stewart, Friedman-Hill, Yu, Shorter, & Cave, 1990) as well as Treisman and Sato (1990) reported evidence indicating that preattentive processes that serve for feature encoding may contribute to rapidly direct attention to the target location—what they termed "parallel guidance." When such guidance is highly effective, this is assumed to lead to very shallow slopes in the search for a conjunction target since attention may be rapidly directed to its location. In this context, variations in the search rates between types of conjunction targets in normal individuals are explained, at least partially, by differences in the amount of noise that is present in the guidance signal coming from preattentive processes. That is, a decreased effectiveness of the guidance signal implies that more items need to be examined before the target can be found. Even with preattentive guidance, however, attention is assumed to be critical for feature integration in all cases, which is congruent with the results reported here.

What are the subcomponents of the conjunction search process that may be affected by the attention disorder of the VAD group and be responsible for their slowed search rates for a contralesional conjunction target? Three candidates may be considered. The first is an increase in the time necessary to integrate the features of a contralesional item once attention is focused at the location it occupies. The second is the time required to perform a spatial shift of attention from one stimulus to the next, which may be increased for contralesional stimulation in the VAD group. Third, it may be that the effectiveness of the preattentive guidance of the focus of spatial attention to the target location is decreased for contralesional stimuli in the VAD group. The observations reported here provide some clues with respect to this question and suggest that an increase in the time required for the feature integration process may be responsible for the impaired search rate of the VAD group for a conjunction target.

First, comparisons of the average slopes shown by the normal controls studied here in Exp. 3 and those reported by Wolfe et al. (1989) in a similar task suggest that preattentive guidance was not a factor in our experiment. In Exp. 3, the features used to make the stimuli were the colors red and green and the orientations horizontal and vertical and the average search rate shown by the normal controls was of about 51 ms/item. In contrast, Wolfe et al. (1989), using the same features to construct their stimuli, observed much shallower slopes of RTs as a function of the number of items presented (5.6 ms/item on target-present trials and 13.1 ms/item on target-absent trials). Of course, age (Plude & Doussard-Roosevelt, 1989) may have been a factor in this large slope difference between the normal controls studied here and the subjects used by Wolfe et al. (1989). Still, under exposure conditions that were similar to those used in Exp. 3 and with identical stimuli, young normal subjects (mean age of 25 years) showed an average slope of about 25 ms/item (Arguin et al., 1990), which is substantially larger than what was reported by Wolfe et al. (1989). One point that distinguishes Exp. 3 and the experiment reported by Arguin et al. (1990) from that of Wolfe et al. (1989) was the use of tachistoscopic displays in the former two, while exposure durations were unlimited in the latter. It therefore appears that preattentive guidance of the focus of attention, which can lead to very shallow slopes in the search for a conjunction target, may be functional only when the search items are visible. This possibility could account for the slope discrepancies between our own results and those of Arguin et al. (1990), which were observed under brief exposure conditions, and those reported by Wolfe et al. (1989) with unlimited viewing of the displays.

One aspect of the results of Exp. 2 (color search) also suggests that the rate at which VAD subjects are able to shift their spatial attention from one item to another in their (contralesional) right visual field is comparable to that of the normal controls. It should be recalled that in Exp. 2, serial search was observed in both the VAD and control groups on right-hemifield target-present trials. In addition, the slopes shown by these two groups did not differ significantly, suggesting that the speed of attention shifts from one item to the next did not differ. Given that the control and VAD groups did not differ on the speed of spatial attention shifts in Exp. 2, it appears unlikely that a difference on this process occurred in Exp. 3 (conjunction search).

To conclude, we have shown here that brain-damaged subjects with a disorder affecting the allocation of attention to the contralesional visual hemifield are markedly impaired in a task requiring the search for a conjunction target presented in that hemifield. By elimination, it appears that this deficit results from an increase in the time required for the feature integration process once attention has been focused at the location occupied by an item. In contrast to their deficit in feature integration, subjects with an attention disorder have been shown to perform the search for single features as effectively as
neurologically intact individuals. These observations are congruent with the basic hypotheses of Treisman's feature integration theory, according to which feature encoding is preattentive while the integration of the separate features of an object requires attention.

**METHOD**

**Subjects**

Inclusion as well as exclusion criteria were applied for subject selection in order to control factors that otherwise might have rendered the results interpretation equivocal, and to make sure that subjects were able to perform the tasks adequately. Inclusion criteria were (1) right-handedness, as assessed with the Edinburgh Handedness Inventory (score between +80 and +100; Oldfield, 1971) and (2) brain lesion of vascular origin, a first occurrence of brain damage, with a unique lesion lateralized to one hemisphere, as assessed by a CT scan examination (applies only to brain-damaged subjects). Exclusion criteria were (1) reduced and uncorrected visual acuity; this was assessed by asking the subject to read a few sentences printed in small type and presented at a distance of about 50 cm; (2) ocular disease such as glaucoma or cataracts; the presence of either of these diseases was established either from the subject's self report or, in the case of patients, from their hospital charts; (3) color vision anomaly, as assessed by Ishihara's isochromatic plates (Ishihara, 1964); (4) visual field deficits, as assessed by campimetry (applies only to brain-damaged subjects); (5) visual hemineglect, as assessed with the bells test (Gauthier, Dehaut, & Joanette, 1989; applies only to brain-damaged subjects); extinction on double simultaneous stimulation was not assessed; (6) an aphasia that may impair the understanding of instructions (applies only to brain-damaged subjects). This was determined through consultation of the patients' hospital charts, who had all received a formal language investigation by a speech therapist.

Following the application of the selection criteria, eight left brain-damaged subjects whose mean age was 65 years were studied. The minimum time period between lesion onset and testing was of 2 months. Ten neurologically intact subjects whose mean age was 71 years served as controls.

All these subjects had previously been examined (Arguin et al., 1992) on a series of visuospatial cuing tasks (Eriksen & Hoffman, 1973, 1974; Jonides, 1981; Müller & Rabbit, 1989; Posner, 1980; Posner, Walker, Friedrich, & Rafal, 1984, 1987). Of the brain-damaged subjects, three (VAD group—visual attention deficit) showed results that differed from those of the normal controls. VAD subjects had longer RTs to contralesional than ipsilesional targets at short temporal intervals (50 and 150 msec SOAs) following a spatial cue that indicated the hemifield in which the target would occur on 80% of the trials (valid cues).

This result occurred in the following cuing conditions: (1) with a central cue (arrowhead at fixation) that was valid or invalid, (2) with the peripheral onset of a square at an invalid location, and (3) with bilateral peripheral vertical gratings in which the spatial cue was the direction of motion of the bars in the gratings. Note that with these motion cues, no effect involving the effect of cue validity was found. The hemifield asymmetries observed in each of the conditions just mentioned were much reduced or absent at longer SOAs (600 and 1000 msec) between cues and targets.

This pattern of results indicates a deficit in allocating visual attention to contralesional stimulation in the VAD subjects. Specifically, the impairment appeared to affect the capacity of VAD subjects to maintain a high level of alertness for contralesional stimulation in the intervening period between trials. Thus, the presentation of a warning signal (the spatial cues—valid or invalid) served to progressively increase the level of alertness for contralesional stimulation in VAD subjects. At short SOAs (50 and 150 msec), the level of alertness in these patients was higher for ipsilesional than for contralesional stimulation but this asymmetry was no longer present at late SOAs (600 and 1000 msec). A similar deficit in maintaining alertness for contralesional stimulation during the time interval separating one trial from the next has previously been revealed in right brain-damaged patients through the use of a spatial cuing paradigm by Posner, Inhoff, Friedrich, and Cohen (1987). Previous work on alerting in neurologically intact subjects has shown that alerting does not affect the rate of information encoding in automatic pathways. Rather, it appears to affect the rate at which a higher level attention system reacts—not necessarily through a physical response—to the build-up of perceptual information in those automatic pathways (Posner & Boies, 1971; Posner et al., 1987; Posner, Klein, Summers, & Buggie, 1973; Thomas, 1974; see Posner, 1978 and Posner & Peterson, 1990 for a detailed review of alerting).

In contrast to VAD subjects, the other brain-damaged subjects (nVAD group) showed performances in the visuospatial cuing tasks that were similar to those of the normal controls and which did not differ as a function of visual hemifield.

Figure 4 shows a reconstruction of the lesion sites observed in the nVAD subjects. The lesions suffered by the three VAD patients that have been tested in the visual search experiments reported in this paper are illustrated in Figure 5. Two other VAD patients who were examined on the spatial cuing experiments could not be tested in the visual search task since one showed a color vision deficit (deuteranopy) on the Ishihara plates and the other exhibited very important difficulties to perform a response choice through a button press (which was required for the visual search experiments) because of strong perseveration tendencies. The first of these patients suffered from a large cerebellar lesion while the
Figure 4. Reconstruction of the lesional sites observed in the nVAD subjects from their CT scans. The gray areas represent lesional sites observed in only one subject, and black areas represent lesional sites seen in two subjects. Numbers identify the patient who showed the lesion illustrated closest to it.

second had a lesion of the internal capsule. In spite of an apparent concentration of the brain damage shown in Figure 5 in the area of the parieto-occipitotemporal junction, it is not clear that the attention disorder of the VAD patients is specifically related to a lesion in this area. First, as indicated, two other VAD patients showed lesions that did not affect this region. Second, one of the nVAD patients (subject 5, Fig. 4) also showed a lesion in the parietotemporal area. The clearest difference between the brain damage showed by the VAD and nVAD subjects was the size of the lesions. Thus, the average volume of the lesions observed in the VAD patients was about four times larger than that seen in nVAD subjects.

Three VAD subjects were retested in the present experiments to determine the aspects of visual search that may be related to deficits of visual attention.

Materials

Experiments were controlled by an Amiga microcomputer and stimuli were displayed on an RGB monitor located at a distance of 57 cm from the subjects. Subjects rested their head on a chin rest. A two-button computer mouse was used for response production. The experiments were run in a dimly lit room.

An ocular fixation control was used to ensure that the stimuli were displayed to the proper visual field. This was achieved by the use of a pupil and corneal reflection tracking system (ISCAN, model RK-426) linked to the control computer. This tracking system analyzes video images obtained from a black and white camera equipped with an infrared filter. An infrared source served for illumination. Trials began only when the subject fixated his eyes on a central fixation point. If, during a trial, the subject shifted his ocular fixation toward the array of stimuli (displayed left or right of fixation) so that the distance between ocular fixation and the centermost stimulus was under 2.0° of visual angle, that trial was immediately terminated and run again later in the session. Across experiments, an average of 1.7% of trials had to be eliminated this way.

Stimuli

A white fixation stimulus (luminance of about 61 cd/m²; CIE coordinates, x = 0.29, y = 0.32) was shown at the center of the display screen between trials. It was made of a 0.2° dot surrounded by a 1.3° empty circle. Subjects were instructed to keep their eyes directed toward the fixation location as much as possible throughout the experiment.

In Exp. 1, stimuli were white horizontal (2.5° wide × 0.4° high) and vertical (0.5° wide × 2.4° high) bars. The
stimulus designated as the target was a horizontal bar and distractors were vertical bars.

In Exp. 2, stimuli were filled circles (1.3° in diameter) that were either red (luminance of about 11 cd/m²; CIE coordinates, \(x = 0.64, y = 0.33\)) or green (luminance of about 14 cd/m²; CIE coordinates, \(x = 0.32, y = 0.58\)). Prior to the experiment, these colors were adjusted to equiluminance individually for each subject with the minimum flicker technique (Troscianko & Low, 1985). The target the subjects had to report was a red circle and distractors were green circles.

In Exp. 3, stimuli were horizontal (2.5° wide × 0.4° high) and vertical (0.5° wide × 2.4° high) bars that were either red or green. As in Exp. 2, these colors were adjusted to equiluminance for each subject individually with the minimum flicker technique prior to the experiment. The target stimulus was a red horizontal bar. One subgroup of distractors was made of red vertical bars and another of green horizontal bars. In displays comprising an even number of distractors, half of them were red vertical and half were green horizontal. In other cases, the display contained one more distractor (selected randomly) from one subgroup than from the other. Thus, in Exp. 3, the target the subjects had to report differed from distractors by a unique conjunction of orientation and color.

In each experiment, stimulus duration was established individually for each subject during the 35 practice trials that immediately preceded the experimental session. This was done according to the following procedure. During the five first practice trials, stimulus duration was fixed at 400 msec. From the sixth practice trial, each trial on which the subject gave a correct response was followed by a reduction in exposure duration. Inversely, each trial on which the subject made an error was followed by an increase in display duration. Initially, the value by which stimulus duration was either increased or decreased was set at 134 msec. Afterward, each time a correct response was followed by an error and each time an error was followed by a correct response, this value was halved. The minimum and maximum possible exposure durations were 150 and 600 msec, respectively. This procedure rapidly established an exposure duration that remained quite stable during the last few practice trials and that yielded an adequate level of accuracy. The exposure duration in the experimental session was that used on the last practice trial. Table 4 shows the mean display durations used in each group in Exps. 1, 2, and 3.

Stimuli were always displayed unilaterally (either to the left or right of the fixation point). Each item in an array was located randomly within an eight-location grid that was organized in two columns and four rows. The distance separating the center of each of these locations...
was 3.9° horizontally and 3.6° vertically. To break the regularity of the display, each item was shifted between ±0.35° along the horizontal and ±0.28° along the vertical relative to the center of its assigned location in the matrix. The direction and size of this shift were determined randomly. The minimum distance between the centermost stimulus and the fixation location was 5.6°.

Procedure

The order in which the experiments were conducted was the same for all subjects (Exp. 1 first and Exp. 3 last). The three experiments required subjects to respond on each trial whether a prespecified target was present in the display. Subjects were instructed to respond as rapidly as possible while avoiding errors. All subjects used their left hand to respond since half the brain-damaged subjects suffered from a right-arm plegia or paresis. To indicate that the target was present, subjects pressed the right button on a two-button computer mouse with their index finger and pressed the left button with their middle finger to indicate that it was absent. All responses were immediately followed by an auditory feedback emitted by the control computer. Correct responses emitted between 150 and 3000 msec after the onset of the stimuli were immediately followed by a high-pitch sound and errors were followed by a low-pitch sound. Responses given less than 150 msec or more than 3000 msec after the onset of stimulation were also followed by a low-pitch sound. These latter trials were eliminated from data analyses and run again later in the experimental session. Across experiments, an average of 0.002% of trials were eliminated for RTs under 150 msec and an average of 0.22% of trials were eliminated for RTs above 3000 msec.

The number of stimuli displayed on a single trial was one, two, three, or four. The selection of these values was determined by two separate constraints: (1) A sufficient number of levels of the set size factor was required to obtain meaningful estimates of its effect on RTs. (2) The maximum number of items that could be presented had to be sufficiently low that accuracy remained high even in the more difficult task of conjunction search (Exp. 3). The target was present on half the trials and absent in the other half.

The effects of four factors were examined. Those were the subject group (control, nVAD, and VAD), the hemifield to which the stimuli were presented (left or right), target presence (present vs. absent), and the number of stimuli (1, 2, 3, or 4). Conditions were distributed randomly within a 240-trial experimental session, with the constraint that there were 15 trials per condition. This experimental session was immediately preceded by 35 practice trials. The main dependent variable was RT.

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Notes

1. We thank an anonymous reviewer for pointing this out.
2. Six right brain-damaged subjects were also examined. However, only one of them could be assessed satisfactorily since the others tended to fall asleep after a few minutes into the experiment. This is probably related to the vigilance problems that are frequently associated with right-brain lesions (De Renzi & Faglioni, 1965; Heilman, 1985) and the right-hemisphere superiority of normal individuals in vigilance tasks (Heilman & Van Den Abell, 1979, 1980). The facts that subjects were tested in a dimly lit room and that the search tasks involved no verbal interactions with the experimenter may also have contributed to the difficulty experienced in testing right brain-damaged patients. The subject who could be tested performed similarly to the left brain-damaged individuals who did not show evidence for an attention deficit. Since he is the only subject with a right brain lesion in our sample, his results will not be reported.

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REFERENCES


Plude, D. J., & Doussard-Roosevelt, J. A. (1989). *Aging, selec-


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3. Andrea Peru, Leonardo Chelazzi. 2008. Local (focused) and global (distributed) visual processing in hemispatial neglect. *Experimental Brain Research* **187:**3, 447-457. [CrossRef]


5. LEON Y. DEOUCELL, YARON SACHER, NACHUM SOROKER. 2005. Assessment of spatial attention after brain damage with a dynamic reaction time test. *Journal of the International Neuropsychological Society* **11:**06. [CrossRef]


13. Ramesh S. Bhatt, Carolyn Rovee-Collier. 1996. Infants' Forgetting of Correlated Attributes and Object Recognition. *Child Development* **67:**1, 172. [CrossRef]