



Disability weights for infectious diseases in four European countries: comparison between countries and across respondent characteristics

Citation

Maertens de Noordhout, C., B. Devleeschauwer, J. A. Salomon, H. Turner, A. Cassini, E. Colzani, N. Speybroeck, et al. 2017. "Disability weights for infectious diseases in four European countries: comparison between countries and across respondent characteristics." *The European Journal of Public Health* 28 (1): 124-133. doi:10.1093/eurpub/ckx090. <http://dx.doi.org/10.1093/eurpub/ckx090>.

Published Version

doi:10.1093/eurpub/ckx090

Permanent link

<http://nrs.harvard.edu/urn-3:HUL.InstRepos:37068243>

Terms of Use

This article was downloaded from Harvard University's DASH repository, and is made available under the terms and conditions applicable to Other Posted Material, as set forth at <http://nrs.harvard.edu/urn-3:HUL.InstRepos:dash.current.terms-of-use#LAA>

Share Your Story

The Harvard community has made this article openly available. Please share how this access benefits you. [Submit a story](#).

[Accessibility](#)

.....
The European Journal of Public Health, Vol. 28, No. 1, 124–133

© The Author 2017. Published by Oxford University Press on behalf of the European Public Health Association.

This is an Open Access article distributed under the terms of the Creative Commons Attribution License (<http://creativecommons.org/licenses/by/4.0/>), which permits unrestricted reuse, distribution, and reproduction in any medium, provided the original work is properly cited.

doi:10.1093/eurpub/ckx090 Advance Access published on 11 September 2017

Disability weights for infectious diseases in four European countries: comparison between countries and across respondent characteristics

Charline Maertens de Noordhout¹, Brecht Devleeschauwer², Joshua A. Salomon³, Heather Turner⁴, Alessandro Cassini^{5,6}, Edoardo Colzani⁵, Niko Speybroeck¹, Suzanne Polinder⁷, Mirjam E. Kretzschmar^{8,9}, Arie H. Havelaar^{10,11}, Juanita A. Haagsma^{6,12}

1 Institute of Health and Society, Université catholique de Louvain (Clos Chapelle-aux-Champs, 30) Brussels, Belgium

2 Department of Public Health and Surveillance, Scientific Institute of Public Health (WIV-ISP), Brussels, Belgium

3 Department of Global Health and Population, Harvard School of Public Health, Boston, MA, USA

4 Department of Statistics, University of Warwick, Coventry, UK

5 European Centre for Disease Prevention and Control, Stockholm, Sweden

6 Julius Center for Health Sciences and Primary Care, University Medical Center Utrecht, Utrecht, The Netherlands

7 Department of Public Health, Erasmus MC University Medical Center, Rotterdam, The Netherlands

8 National Institute for Public Health and the Environment, Centre for Infectious Disease Control, Bilthoven, The Netherlands

9 Julius Center for Health Sciences and Primary Care, University Medical Center Utrecht, Utrecht, The Netherlands

10 Department of Animal Health and Emerging Pathogens Institute, University of Florida, Gainesville, FL, USA

11 Institute for Risk Assessment Sciences, Faculty of Veterinary Medicine, Utrecht University, Utrecht, The Netherlands

12 Institute for Health Metrics and Evaluation, Seattle, WA, USA

Correspondence: Charline Maertens de Noordhout, Institute of Health and Society (IRSS), Université catholique de Louvain, Clos Chapelle-aux-Champs, 30 bte B1.30.15, Brussels–1200, Belgium, Tel: +32 2 764 31 79, Fax: +32 2 764 34 70, e-mail: charline.maertens@uclouvain.be

Background: In 2015, new disability weights (DWs) for infectious diseases were constructed based on data from four European countries. In this paper, we evaluated if country, age, sex, disease experience status, income and educational levels have an impact on these DWs. **Methods:** We analyzed paired comparison responses of the European DW study by participants' characteristics with separate probit regression models. To evaluate the effect of participants' characteristics, we performed correlation analyses between countries and within country by respondent characteristics and constructed seven probit regression models, including a null model and six models containing participants' characteristics. We compared these seven models using Akaike Information Criterion (AIC). **Results:** According to AIC, the probit model including country as covariate was the best model. We found a lower correlation of the probit coefficients between countries and income levels (range r_s : 0.97–0.99, $P < 0.01$) than between age groups (range r_s : 0.98–0.99, $P < 0.01$), educational level (range r_s : 0.98–0.99, $P < 0.01$), sex ($r_s = 0.99$, $P < 0.01$) and disease status ($r_s = 0.99$, $P < 0.01$). Within country the lowest correlations of the probit coefficients were between low and high income level (range $r_s = 0.89–0.94$, $P < 0.01$). **Conclusions:** We observed variations in health valuation across countries and within country between income levels. These observations should be further explored in a systematic way, also in non-European countries. We recommend future researches studying the effect of other characteristics of respondents on health assessment.

Introduction

The Global Burden of Disease (GBD) studies, which started in the early 1990s, presented an important set of comparative findings on the impact of different diseases, injuries and risk factors on population health.^{1,2} One of the most important results of these initiatives has been the introduction of the disability-adjusted life year (DALY), a new measure of overall disease burden, reflecting the number of healthy life years lost due to illness, disability or early death. This metric accounts for both premature mortality, expressed as years of life lost (YLLs) and time spent in states of reduced health, expressed as years lived with disability (YLDs). An essential parameter to calculate YLDs is the disability weight (DW), representing the severity of health loss associated with illness or disability. The DW is a value on a scale from zero to one: a DW of 0 means that a health state is equivalent to full health, whereas a DW equal to 1 means that a health state is equivalent to death. The value of a DW can be derived from the valuations of a panel of judges stated toward a set of

hypothetical health states expressing the relative undesirability of the health state.^{3,4} There are three main types of health state valuation methods: (1) trade-off methods such as the time trade-off, and the person trade-off; (2) discrete choice experiments and (3) scales such as the Rating or Visual Analogue Scale or a combination of these.⁵

In 2012, an update of the GBD study was performed, including a new set of DWs for 220 unique health states.⁶ To assess the DWs, the GBD 2010 study proposed a new ordinal health state valuation methodology based on paired comparisons questions, a type of discrete choice question and a new population health equivalence technique, a type of trade-off method, to anchor DWs to the scale (from zero to one) aiming to address criticisms of previous approaches.^{7–9} With paired comparisons two descriptions of hypothetical health states are presented to respondents who have to decide which they regard as being healthier.

In 2008, the European Centre for Disease Prevention and Control (ECDC) launched an initiative to estimate the current and future

burden of communicable disease in the European Union (EU) and European Economic Area (EEA) countries using DALYs.^{10,11}

For the purpose of estimating the burden of communicable disease in Europe, the DWs developed for the GBD 2010 study⁶ had a number of limitations. The set of DWs was derived from a household survey based on responses from the general population of five non-European countries and from an open-access web-survey study performed in 167 countries with most of the participants from North America and Australia. This web-survey study was available in English, Spanish and Mandarin, whereas German (16%) is the most spoken language in Europe.¹² Consequently the DWs were not based on the values representative of socio-demographic groups of the EU/EEA population.¹³ Moreover, DWs for 74 health states related to infectious diseases did not exist at all. Therefore ECDC, the Institute for Health Metrics and Evaluation and others started a collaboration to assess new DWs. They performed a web-based survey in Hungary, the Netherlands, Sweden and Italy¹⁴ because these four countries were considered geographically representative of the EU/EEA Member States.

The results of this study showed a high degree of agreement between the GBD 2010 and the European study. However, the putative differences in DWs between respondent characteristics with respect to their ranking of infectious diseases were not assessed. Previous studies already demonstrated that respondent characteristics such as age,^{15,16} income level,¹⁷ country,^{18,19} health status²⁰ or profession²¹ may have an impact on health evaluation, and consequently on derived utility or DW. However, none of these previous studies specifically looked into the implications for DWs derived from pairwise comparisons. If significant differences in health valuation would exist in the derivation of infectious disease DWs, these may have to be taken into account in the selection of study populations of future studies aiming to assess infectious disease DWs or in the translation of study results to the general population.

The objective of the current study is therefore to compare DWs between the four European countries and across respondents' characteristics.

Methods

Study design and participants

We performed a web-based survey between September and November 2013 in Hungary, the Netherlands, Sweden and Italy.¹⁴ The eligible study population was aged from 18 to 65 years. The participants were recruited through existing large panels of people available through the Growth From Knowledge (GFK) Company for participating in online surveys.

Outcome trees

In addition to acute illnesses, infections can also lead to long-term consequences. To estimate the full burden caused by an infection, all health outcomes caused by the infection have to be estimated.^{10,22} This process, which relates all outcomes to their infectious causes results in an 'outcome tree' or a 'disease model',²³ representing all health outcomes that are causally related to a hazard, their possible transition and transition probabilities in a flow chart.²⁴ In total, 32 infectious diseases and six healthcare-associated infections were included in the burden of communicable disease in Europe study. These 32 infectious diseases and six healthcare-associated infections had 74 unique health states for which DWs were evaluated in this study.²⁴

The short health state descriptions (35 words or fewer) were developed by Salomon et al. to emphasize the major functional consequences and symptoms associated with each health state with simple non-clinical vocabulary. The health state descriptions used in this study are available from Ref. 14.

Procedures

We adapted the questionnaire developed by Salomon et al.⁶ by including additional questions about disease experience and by changing income and educational categories to fit the situation in the four countries. For disease experience status we asked if the participant suffered from a disease or the consequences of an injury. We used existing health state descriptions developed by Salomon et al.⁶ for 74 health states related to infectious diseases.

Three different versions of the questionnaire were designed. Each version included questions about demographics of the participant (age, sex, educational and income level and disease experience), and each participant received only one version of the questionnaire. The first version was composed of 15 pairwise comparison (PWC) questions on chronic (life-long health) states and three population equivalence questions (PHE). The second version of the questionnaire included 15 paired comparison questions on temporary (lasting 1 week) health states and three PHE questions. The third version included five PWC questions with a chronic framing to accommodate PHE questions and three PHE questions. PHE questions, which allow to anchor the final DWs from 0 to 1 and to elicit trade-offs between mortality and nonfatal health states, were finally not used in the European DW study. For the PWC questions, participants were asked to imagine that two people have the same number of years left to live, and that they will experience the health problems described for the rest of their lives (or for 1 week for temporary health state). Then, the respondents were presented with two descriptions of hypothetical people, each living in a particular health state, and then asked which person they regarded as being healthier.

For instance:

- *'The first person has diarrhea three or more times a day with occasional discomfort in the belly.'*
- *'The second person has a low fever and mild discomfort, but no difficulty with daily activities.'*

Who do you think is healthier overall, the first person or the second person?'

The questionnaires and health states were translated from English to Hungarian, Dutch, Swedish and Italian and then checked by an independent native speaker.

The questionnaire was assigned to participants through a computer algorithm that first attributed the study version with the lowest percentage of participants at the time of assignment. Then after the version was allocated, the algorithm selected the health states to include in the questionnaire based on the minimum number of allocations that the health state had at that moment.

Statistical analysis

To assess the representativeness of our study population, we applied chi-squared tests to assess the comparability of the age, sex and educational distribution of the study population with those of the respective national populations. Because income levels were categorized, we assigned the midpoint of each income category to all respondents in that category, which we compared with mean household income values from Eurostat.²⁵

To investigate the differences between health states, we ran probit regression analyses on the choice responses. The first health state of the pair had a value of 1, the second health state of the pair had a value of -1 and those not in the pair had a value of 0. The resulting probit regression coefficients for each health state expressed the relative differences in valuation of the health state on an arbitrary scale. To project the results from the probit regression model on a DW scale ranging from zero to one, we ran a locally weighted scatterplot smoothing regression model (loess) of the probit regression coefficients vs. the logit transformed GBD2010 DWs.¹⁴

We then predicted logit transformed DWs for each of the probit coefficients from the loess fit. To change the scale and obtain values ranging between 0 and 1 we applied an inverse logistic transformation of these predicted DWs. We estimated uncertainty around the mean DWs through a bootstrap procedure. We first generated 200 samples of the pairwise coefficients using their probit estimated mean and standard deviation. We then used the generated samples to calculate 200 loess fits. We generated 200 samples for each of the DWs using loess fit, producing 40 000 bootstrap samples for each DW. We finally calculated uncertainty intervals (UI) as the 2.5th and 97.5th percentile of the corresponding distribution of bootstrapped samples intervals for the DW. These analyses were performed using the overall database of 255 health states available in Haagsma et al.¹⁴ We hereby extracted the 74 probit coefficients and DWs of the health states relating to infectious diseases, which were the primary focus of this study.

We performed the same analysis of the responses across countries, sex, age, disease experience status, income and education level and within countries by sex, age, disease experience status, income and education level.

The Spearman correlation coefficient was also used to evaluate the relation of both the probit coefficients between countries, age groups, sex, disease experience status, educational and income level.

To evaluate the effect of participants' characteristics, we also estimated difference between two health states in models with interactions that allow each country, or each population subgroups to have its own set of distinct values, assuming that the difference between the two state values was normally distributed. We then compared models using Akaike's Information Criterion (AIC).

We interpreted a *P* value < 0.05 as significant. Analyses were done with R (version 3.0.2) and using the speedglm package.²⁶

Results

Characteristics of respondents

Overall, 30 660 respondents completed the DW survey of which 6053 participants (20%) were from Hungary, 8054 (26%) from Italy, 8005 (26%) from the Netherlands and 8548 (28%) from Sweden.

The average age was 41.5 years (standard deviation [sd] 13.3) in Hungary, 41.7 years in Italy (sd 12.3), 44.1 years in the Netherlands (sd 12.9) and 41.9 years in Sweden (sd 13.7). Approximately half of the respondents were male in each of the four countries.

The majority of the respondents had a medium educational level in Hungary (49%), Italy (48%) and Sweden (51%) but while in the Netherlands the level of education of the participants was more equally distributed (low: 35%, medium: 35% and high: 30%).

Most of the respondents had a low income level in Italy (47%), medium income level in the Netherlands (56%) and Sweden (50%) and high income level in Hungary (47%) (table 1).

Approximately one in three respondents reported that they were suffering of a disease or of the consequences of an injury in Hungary (34%), the Netherlands (29%) and Sweden (35%) and one in five respondents in Italy (18%) (table 1).

Comparison of the study population with data from Eurostat showed that the study population was representative of the national populations in the four countries in terms of sex distribution, age distribution and educational levels.²⁷ For income levels we observed that, except in Sweden, mean income of the respondents was higher than in the general population table 1.

Probit regression models

We observed that all models, except the model including disease experience status (model 3) (AIC = 417 722), had lower AIC than the model without covariate (AIC = 417 572). We also observed that the model including country as covariate had the lowest AIC (411 142), the next lowest being the based on age (AIC = 416 940),

then on educational level (AIC = 417 178), on sex (AIC = 417 271) and finally on disease experience status (AIC = 417 722). That means that probit model including country as covariate was the best model because AIC represents the loss of information caused by using models to represent the process generating the actual data. We did not compare the AIC for the model including income level as covariate with the other models because it included less data because of the missing answers (table 2).

Probit regression coefficients–correlation coefficients

Between country

All correlation coefficients were significantly different from zero ($P < 0.001$).

Overall, we found a lower correlation of the probit coefficients between country than between age-groups, educational, income, sex and disease experience status (figure 1). The highest linear correlation of the probit coefficients was found between Italy and Hungary ($r_s = 0.95$) and the lowest correlation was observed between Hungary and the Netherlands ($r_s = 0.89$) (figure 1).

The correlation of the probit coefficients was high between age groups (range r_s : 0.98–0.99), educational levels (range r_s : 0.98–0.99), sex ($r_s = 0.99$) and disease experience status ($r_s = 0.99$) and lower for income levels (range r_s : 0.97–0.99).

Within-country

All correlation coefficients were significantly different from zero ($P < 0.001$).

We observed a higher correlation of the probit coefficients by age-groups in the Netherlands (range r_s : 0.97–0.98) than in Hungary (range r_s : 0.95–0.96), Italy (range r_s : 0.95–0.97) or Sweden (range r_s : 0.96–0.97). In all countries, the highest correlation was observed between 35–49 years and 50–65 years groups. In every country, except for the Netherlands, the lowest observed correlation was between 18–34 years and 50–64 years. In the Netherlands, the lowest observed correlation was between 18–34 years and 35–49 years groups ($r_s = 0.97$).

We observed roughly the same pattern for educational levels. The correlation of the coefficients was higher in the Netherlands (range r_s : 0.97–0.98) than in Hungary (range r_s : 0.94–0.98), Italy (range $r_s \geq 0.94$ –0.96) and Sweden (range r_s : 0.95–0.98). In every country, except Sweden, the highest correlation was observed between low and medium educational level. In Sweden the highest correlation was observed between medium and high educational level ($r_s = 0.98$). Except for the Netherlands, the lowest observed correlation was between low and high educational level. In the Netherlands, the lowest observed correlation was between medium and high educational groups ($r_s = 0.97$).

We observed relatively less concordance of the responses between income levels across country. The correlation of the coefficients was better in Hungary (range r_s : 0.94–0.95) than in the Netherlands (range r_s : 0.92–0.98), Italy (range r_s : 0.89–0.95) and Sweden (range r_s : 0.93–0.98). The lowest correlation was observed between medium and high income levels in the Netherlands ($r_s = 0.92$) and between low and high income levels in Hungary ($r_s = 0.94$), Italy ($r_s = 0.89$) and Sweden ($r_s = 0.93$). The highest correlation was observed between low and medium income levels in the Netherlands ($r = 0.98$) and Sweden ($r = 0.98$) and between medium and high levels in Hungary ($r_s = 0.94$) and Italy ($r_s = 0.93$).

We observed a higher correlation of the coefficients between disease experience status in Sweden ($r_s = 0.99$) than in the Netherlands ($r_s = 0.97$), Hungary ($r_s = 0.97$) and Italy ($r_s = 0.96$).

In all countries, we observed a high correlation of the probit coefficients between sex (range r_s : 0.97–0.98).

Overall, within country, we observed a lower correlation of the coefficients by income levels than by age groups, educational levels, sex and disease experience status.

Table 1 Description of the study population

	Total (n = 30660)	Hungary (n = 6053)	Hungarian population (%) ^a	P-value ^b	Italy (n = 8054)	Italian population (%) ^a	P-value ^b	Netherlands (n = 8005)	Dutch population (%) ^a	P-value ^b	Sweden (n = 8548)	Swedish population (%) ^a	P-value ^b
Sex													
Men	14 719 (48.0%)	2889 (47.7%)	47.6	0.999	3886 (48.2%)	48.4	0.999	3842 (48.0%)	49.5	0.943	4102 (48.0%)	49.9	0.902
Age													
18–34 years	9574 (31.2%)	2047 (33.8%)	34.1	0.993	2522 (31.3%)	30.2	0.982	2148 (26.8%)	33.3	0.584	2857 (33.4%)	36.5	0.869
35–49 years	10 753 (35.1%)	2026 (33.5%)	34		2983 (37.0%)	38.1		2883 (36.0%)	34.3		2861 (33.5%)	33.4	
50–65 years	10 333 (33.7%)	1980 (32.7%)	31.9		2549 (31.6%)	31.7		2974 (37.2%)	32.4		2830 (33.1%)	30.1	
Educational level													
Low	9125 (29.8%)	2344 (38.7%)	80.9	0.242	2668 (33.1%)	86.1		2801 (35.0%)	71.4		1312 (15.3%)	69.9	
Medium	14 000 (45.7%)	2976 (49.2%)	19.1		3896 (48.4%)	13.9		2804 (35.0%)	28.6		4324 (50.6%)	30.1	
High	7535 (24.6%)	733 (12.1%)			1490 (18.5%)			2400 (30.0%)			2912 (34.1%)		
Income level (n = 23950)													
Low	9458 (39.4%)	1301 (30.0%)		na	3177 (47.2%)		na	2053 (35.3%)		na	2927 (41.4%)		na
Medium	10 879 (45.4%)	1016 (23.4%)			3083 (45.8%)			3271 (56.3%)			3509 (49.7%)		
High	3613 (15.1%)	2026 (46.6%)			472 (7.0%)			484 (8.3%)			631 (8.9%)		
Mean [SD] (€)		6778 [3276]	5469	na	25764 [17723]	18997	na	39 839 [23850]	23 802	na	23 404 [12 541]	27 728	na
Disease experience													
Yes	8900 (29.0%)	2085 (34.4%)	NA		1448 (18.0%)			2337(29.2%)	NA		3026 (35.4%)	NA	

a: According to Eurostat: <http://ec.europa.eu/eurostat/data/database> (2013 population).

b: P values resulting from the chi-square test.

NA: Not Available.

na: not applicable.

Except for income level, Dutch participants showed higher correlation of the probit coefficients for all participants' characteristics.

Disability weights

'Distance vision, mild impairment' had the lowest DW in Hungary [H]: 0.007 [95% UI 0.004–0.013], Italy [I]: 0.004 [95% UI 0.002–0.008] and Sweden [S]: 0.005 [95% UI 0.003–0.009]. 'Hearing loss, mild' had the lowest DW in Netherlands [N]: 0.004 [95% UI 0.002–0.009]. 'Intensive care unit admission' had the highest DW for all the countries (H: 0.597 [95% UI 0.464–0.690]; I: 0.530 [95% UI 0.430–0.617]; N: 0.624 [95% UI 0.493–0.719] and S: 0.717 [95% UI 0.552–0.766]) (table 3).

The median DW of all health states based on the responses of the Hungarian respondents was 0.118 (range: 0.007–0.597), 0.101

(range: 0.004–0.530) in Italian respondents, 0.108 (range: 0.004–0.624) in Dutch respondents and 0.111 (range: 0.005–0.717) in Swedish respondents. The range of DWs was higher in Swedish participants (0.712) and lower in Italian participants (0.526) but the shape of the DW distributions were right-skewed in all countries (Supplementary Appendix) and highly correlated ($r_s \geq 0.99$, $P < 0.01$), including in the range below 0.2, where most of the European DWs reside ($r_s \geq 0.97$, $P < 0.01$).

The differences of DWs ranged from –0.135 ('Thrombocytopenic purpura') to 0.197 ('Lymphogranuloma Venereum–local infection') between Hungary and Italy, from –0.221 ('AIDS cases, not receiving ARV treatment') to 0.295 ('Motor plus cognitive impairments, severe') between Hungary and the Netherlands and from –0.165 ('Musculoskeletal problems, generalized, severe') to 0.250 ('Stoma') between Hungary and Sweden. The differences of DWs also ranked from –0.203 ('End-stage renal disease, on dialysis') to 0.226 ('Motor plus cognitive impairments, severe') between Italy and the Netherlands, from –0.187 ('Intensive care unit admission') to 0.241 ('Stoma') between Italy and Sweden and from –0.232 ('Motor plus cognitive impairments, severe') to 0.154 ('Tuberculosis, not HIV infected') between the Netherlands and Sweden. The biggest observed differences of DW were for 'motor plus cognitive impairments, severe' between Hungary and the Netherlands, between the Netherlands and Sweden and between Italy and the Netherlands (Hungary: 0.496; Italy: 0.426, The Netherlands: 0.200, Sweden: 0.432) and for 'Stoma', between Italy and Sweden and between Hungary and Sweden (Hungary: 0.301, Italy: 0.292, Sweden: 0.051). There were no differences in DW between Hungary and Netherlands for 'Near Vision Impairment', for 'Distance vision, mild impairment' and 'Infertility, primary' between Netherlands and Italy, for 'Unilateral hearing loss' between Netherlands and Sweden, for 'Hearing loss, mild',

Table 2 AIC of seven probit models, including a null model and six models containing one covariate, i.e. country, sex, age, disease experience status, income level, or educational level

Probit models	AIC
(1) Null model	417 572
(2) Country	411 142
(3) Disease experience status	417 722
(4) Sex	417 271
(5) Educational level	417 178
(6) Income level	NA ^a
(7) Age	416 940

^aNot available: AIC not calculated because it included less data than the other models.

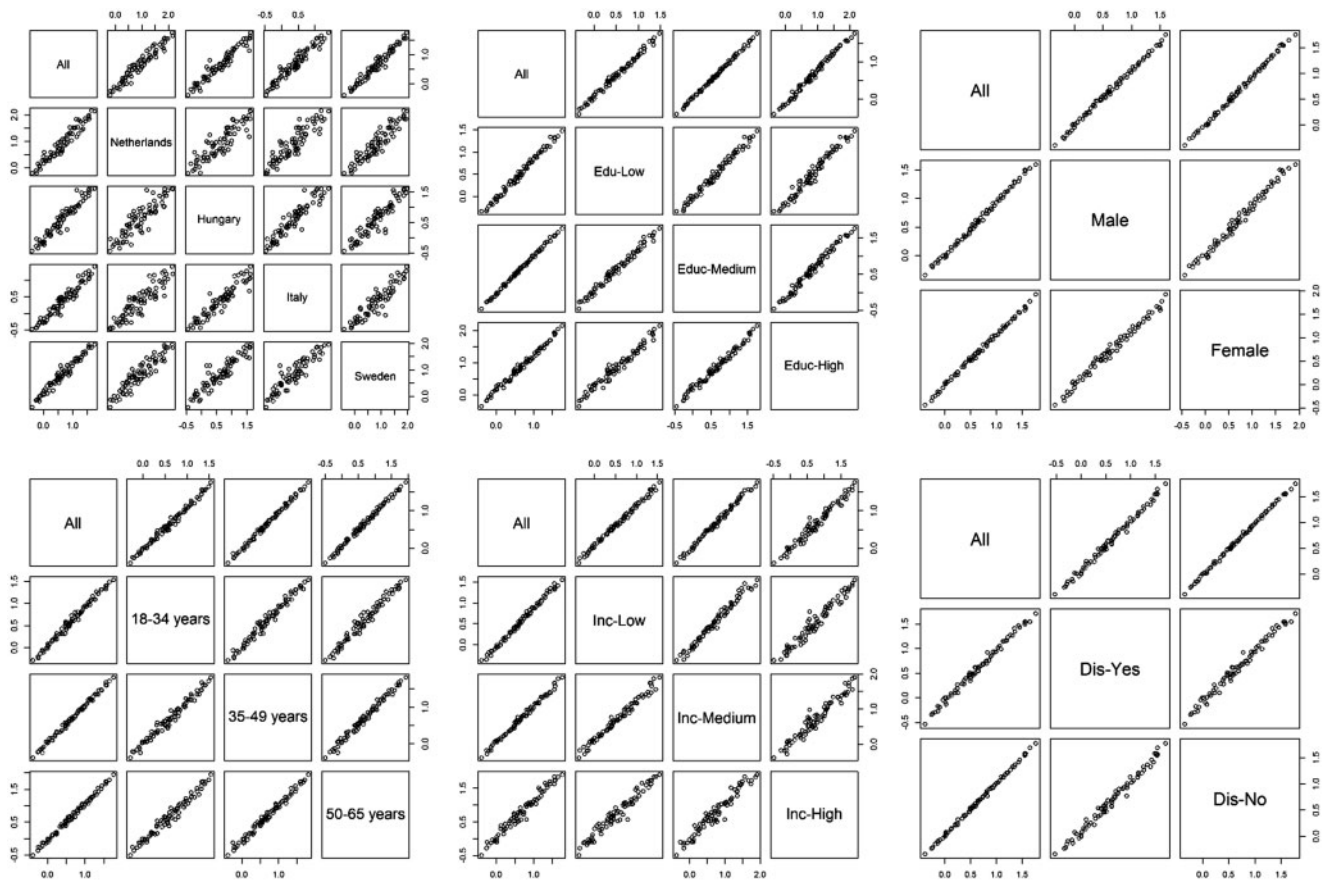


Figure 1 Correlation matrix of probit coefficients by country, educational level, sex, age, income level and disease experience status

Table 3 Comparison of the rank order (descending) of the health states between countries

Health states	Netherlands	Hungary	Italy	Sweden
Intensive care unit admission	1	1	1	1
AIDS cases, not receiving ARV treatment	2	12	3	4
Terminal phase, without medication (for cancers, end-stage kidney/liver disease)	3	3	2	3
End-stage renal disease, on dialysis	4	5	14	7
Musculoskeletal problems, generalized, severe	5	11	8	2
Terminal phase, with medication (for cancers, end-stage kidney/liver disease)	6	7	5	6
Subacute sclerosing panencephalitis – phase 3	7	2	4	5
COPD and other chronic respiratory problems, severe	8	6	13	13
Tuberculosis, HIV infected	9	13	9	16
Musculoskeletal problems, generalized, moderate	10	22	15	12
Tuberculosis, not HIV infected	11	18	16	21
Encephalopathy - severe	12	8	10	11
Encephalopathy - moderate	13	10	12	10
Abdominopelvic problem, severe	14	15	24	15
HIV cases, symptomatic, pre-AIDS	15	19	11	17
Tuberculosis of vertebrae	16	20	20	18
COPD and other chronic respiratory problems, moderate	17	14	19	22
Motor impairment, severe	18	9	6	8
Infectious disease, post-acute consequences (fatigue, emotional lability, insomnia)	19	31	28	25
Crohn's disease or ulcerative colitis	20	29	30	24
Epididymo-orchitis	21	27	29	38
Diabetic neuropathy	22	32	37	26
Infectious disease, acute episode, severe	23	41	47	33
Diarrhea, severe	24	24	27	14
Motor plus cognitive impairments, severe	25	4	7	9

(Continued)

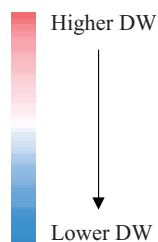
Table 3 Continued

Subacute sclerosing panencephalitis – phase 2	26	16	18	19
Decompensated cirrhosis of the liver	27	34	32	31
Generic uncomplicated disease: anxiety about diagnosis	28	44	31	48
Thrombocytopenic purpura	29	39	23	34
HIV/AIDS cases, receiving ARV treatment	30	53	39	36
Chronic kidney disease (stage IV)	31	40	46	40
Diarrhea, moderate	32	36	41	23
Musculoskeletal problems, upper limbs, moderate	33	37	44	32
Distance vision blindness	34	26	26	29
Musculoskeletal problems, lower limbs, severe	35	28	35	30
Subacute sclerosing panencephalitis – phase 1	36	43	43	47
Motor plus cognitive impairments, moderate	37	30	22	20
Diarrhea, mild	38	55	54	41
Distance vision, severe impairment	39	25	25	28
Generic uncomplicated disease: worry and daily medication	40	50	38	53
Invasive device/drain	41	21	21	43
Abdominopelvic problem, moderate	42	33	33	27
Musculoskeletal problems, lower limbs, moderate	43	35	40	46
Stoma	44	17	17	51
Infectious disease, acute episode, moderate	45	56	56	55
Hearing loss, complete, with ringing	46	38	34	35
Hearing loss, moderate, with ringing	47	45	51	45
Hearing loss, profound, with ringing	48	42	36	39
End-stage renal disease, with kidney transplant	49	71	57	63
Lymphogranuloma Venereum - local infection	50	23	53	52
Irritable bowel syndrome	51	47	59	37

(Continued)

Table 3 Continued

Hearing loss, mild, with ringing	52	54	58	50
COPD and other chronic respiratory problems, mild	53	61	68	60
Hearing loss, severe, with ringing	54	51	45	49
Motor impairment, moderate	55	46	52	42
Abdominopelvic problem, mild	56	69	65	64
Conjunctivitis without corneal scar	57	63	69	68
Hearing loss, profound	58	57	50	58
Musculoskeletal problems, upper limbs, mild	59	48	42	54
Distance vision, moderate impairment	60	49	55	62
Motor plus cognitive impairments, mild	61	52	48	44
Musculoskeletal problems, lower limbs, mild	62	60	60	56
Ear pain	63	65	64	65
Hearing loss, moderate	64	62	63	59
Motor impairment, mild	65	64	72	66
Hearing loss, severe	66	59	61	61
Infertility, secondary	67	68	73	73
Hearing loss, complete	68	58	49	57
Infectious disease, acute episode, mild	69	73	67	70
Near vision impairment	70	67	62	67
Infertility, primary	71	66	70	71
Unilateral hearing loss	72	70	66	69
Distance vision, mild impairment	73	74	74	74
Hearing loss, mild	74	72	71	72



'Infectious disease, acute episode, mild' and 'Chronic kidney disease (stage IV)' between Hungary and Sweden (Supplementary Appendix).

Discussion

In this study, we aimed to assess the impact of country, sex, age, disease or injury experience status, educational and income levels on the evaluation of infectious disease health states.

First, we found a lower correlation of the probit coefficients between countries than between the other respondent characteristics. It means that we observed more variation of the health states valuation between countries than between the other respondent characteristics. Especially the correlation of the probit coefficient was lower between Hungary and the Netherlands than between the other pairs. This translated in some major differences, as for 'motor plus cognitive impairments, severe', that was ranked as the 25th worse health state in the Netherlands while instead fourth, seventh and ninth in Hungary, Italy and Sweden. We also observed that the model including country as covariate had the lowest AIC. This is in line with the results of the correlation analysis. The observed variations of health state valuation across country could reflect a combination of different demographics and/or different attitudes but also expressed that the included characteristics did not fully explain the between country differences.

We found a high correlation of the probit coefficients between educational levels, both across and within countries. This means that educational levels had no influence on health valuation. In 2008, Haagsma et al. also demonstrated no significant effects of educational level on DWs for injury consequences in Dutch population.²⁸

In all countries, we observed a better correlation of the coefficients between 35–49 years and 50–65 years groups than between the other age groups. Dolan et al. also demonstrated that age group has an impact on health valuation.^{15,16}

Within country we observed a lower correlation of the probit coefficients between income levels, especially between low and high income levels, except for the Netherlands where the lowest correlation was found between medium and high income levels.

Our results also showed that the probit coefficients did not vary widely across disease experience status and sex.

Although this study highlighted the effect of participants' characteristic on health valuation using a sample of 30 660 answers, it also had some limitations.

First, we performed a web-based survey and did not include respondents aged 65 years or older. Elderly people are more often affected by some health states as hearing loss or distance vision loss, and may suffer of comorbidities.²⁹ However, we did not demonstrate a relationship between DWs and disease experience status. Therefore this limitation may have had only a limited impact on the derived DWs.

Second, as in Salomon et al. in 2010, our study elicited DW by presenting brief health state descriptions in lay language and some aspects of health states were inevitably omitted in the interest of brevity. For example, people suffering from chronic irritable bowel syndrome need to undergo a restrictive diet which is not included in the health state description and can lead to a higher long-term disability.³⁰ In addition, there may also have been differences between languages in the connotation of the words that were used in the health state descriptions. This could impact health state preferences of respondents and subsequently increase or decrease the value of the DW and impact the final ranking of the DWs.

Third, comparisons of DWs derived using a paired comparison approach across countries and participant characteristics are difficult and have to be interpreted with caution.³¹ Due to the nature of the paired comparison questions, the results hold only relative information. It is therefore not possible to compare the mean of all DWs across countries or between participants' characteristics and

for instance to conclude if one country has in average a higher or lower valuation of health. Instead, we explored different approaches for evaluating relative differences. First, we fitted different probit regression models to evaluate the interaction effects of country and participant characteristics. Comparison of models was only possible based on AIC values, which do not have an absolute interpretation and for which no null hypothesis distribution is available. AIC values thus do not allow to conclude on statistical significance of differences between models. Second, we used correlation coefficients to compare the rankings between different countries or participant characteristics. Because of the high number of observations, all correlation coefficients were significantly different from zero. Furthermore, the resulting correlation coefficients were all relatively high (>0.84). However, a definition of what could be considered sufficiently high depends on context and cannot be provided by statistical models. Although we observed a high agreement of the health states valuation for all sub-groups, we still observed variation of the ranking of DWs between some sub-groups. This has to be considered because that could have an impact on the final derived DWs and indirectly on burden of infectious diseases estimates and policy priorities.

Despite these limitations, we believe that this study brings an empirical basis for understanding the health state preferences related to infectious diseases of a diverse set of European respondents. We observed that health state evaluations are highly correlated across participants' characteristics but that there exists some variation in health assessment across country and within country between income levels. This means that country and income levels may have an impact on health state valuation. These variations should be further explored in a systematic way, also in non-European countries. There is a need of future studies aiming to be more representative of study populations in future DW elicitation exercises. We also recommend future researches studying the effect of other characteristics of respondents on health assessment.

Supplementary data

Supplementary data are available at *EURPUB* online.

Acknowledgements

The authors are grateful to Piotr Kramarz (European Centre for Disease Prevention and Control), Wouter Rottier (UMC Utrecht), Heidi van Ammerlaan (UMC Utrecht), Jolanda van Oirschot (GfK) and Marcel Temminghoff (GfK) who provided valuable contributions to this study.

Author contributions

CMdN: study design, data analysis, drafting the manuscript; BD: data analysis, drafting the manuscript; JS: study design, data analysis, critically revising the manuscript; HT: data analysis, critically revising the manuscript; AC: study design, critically revising the manuscript; EC: study design, critically revising the manuscript; NS: data analysis, critically revising the manuscript; SP: study design, critically revising the manuscript; MK: study design, critically revising the manuscript; AH: study design, critically revising the manuscript; JH: supervision data-collection, study design, data analysis, and drafting the manuscript.

Funding

This study was funded by the European Center for Disease Prevention and Control and the Institute for Health Metrics and Evaluation. The corresponding author had full access to all study data, and the analysis, interpretation, and the decision to publish were solely the responsibility of the authors.

Conflicts of interest: A Cassini and E Colzani are staff of the European Centre for Disease Prevention and Control, which provided partial financial support of the study, provided assistance with translation of health lay descriptions in the survey. Juanita A Haagsma is affiliated with the Institute for Health Metrics and Evaluation, which also provided part of the funding.

Key points

- Using the answers of more than 30 000 Europeans, we evaluated if disability weights (DWs) varied by country, age, sex, disease experience status, income and educational levels.
- We observed that there is some variation in health states valuation across countries and within country between income levels. Country and income levels may have an impact on health state valuation and on derived DWs.
- There is a need of future studies aiming to be more representative of study populations in future DW elicitation exercises.
- We also recommend future researches studying the effect of other characteristics of respondents on health assessment.

References

- Ezzati M, Lopez AD, Rodgers A, Comparative Risk Assessment Collaborating G, et al. Selected major risk factors and global and regional burden of disease. *Lancet* 2002;360:1347–60.
- Murray CJ, Lopez AD. Evidence-based health policy—lessons from the Global Burden of Disease Study. *Science* 1996;274:740–3.
- Drummond M, O'Brien B, Stoddart GL, et al. *Methods for the Economic Evaluation of Health Care Programmes*. USA, NY: Oxford Medical Publications.
- Murray CJ, Acharya AK. Understanding DALYs (disability-adjusted life years). *J Health Econ* 1997;16:703–30.
- Haagsma J, Polinder S, Cassini A, et al. Review of disability weight studies: comparison of methodological choices and values. *Population Health Metrics* 2014;12:20.
- Salomon JA, Vos T, Hogan DR, et al. Common values in assessing health outcomes from disease and injury: disability weights measurement study for the Global Burden of Disease Study 2010. *Lancet* 2012;380:2129–43.
- van Osch SM, Wakker PP, van den Hout WB, et al. Correcting biases in standard gamble and time tradeoff utilities. *Med Decis Making* 2004;24:511–7.
- Doctor JN, Bleichrodt H, Lin HJ. Health utility bias: a systematic review and meta-analytic evaluation. *Med Decis Making* 2010;30:58–67.
- Krabbe PF, Essink-Bot ML, Bonsel GJ. The comparability and reliability of five health-state valuation methods. *Soc Sci Med* 1997;45:1641–52.
- Kretzschmar M, Mangen MJ, Pinheiro P, et al. New methodology for estimating the burden of infectious diseases in Europe. *PLoS Med* 2012;9:e1001205.
- European Centre for Disease Prevention and Control. *Methodology Protocol for Estimating Burden of Communicable Diseases*. Stockholm: ECDC, 2010.
- Les Européens et leur langue, 2012. Available at: http://ec.europa.eu/public_opinion/archives/ebs/ebs_386_fr.pdf (31 August 2015, date last accessed).
- Field MJ, Gold MR. *Directions for the Development and Application of Population Health Metrics*. Washington, DC: Institute of Medicine: National Academy Press, 1998.
- Haagsma JA, Maertens de Noorhout C, Polinder S, et al. Assessing disability weights based on the responses of 30,660 people from four European countries. *Population Health Metrics* 2015;13: p. 10.
- Dolan P. Effect of age on health state valuations. *J Health Serv Res Policy* 2000;5:17–21.
- Hofman CS, Makai P, Boter H, et al. The influence of age on health valuations: the older olds prefer functional independence while the younger olds prefer less morbidity. *Clin Interv Aging* 2015;10:1131–9.
- Neethling I, Jelsma J, Ramma L, et al. Disability weights from a household survey in a low socio-economic setting: how does it compare to the global burden of disease 2010 study? *Glob Health Action* 2016;9:31754.
- Lien K, Tam VC, Ko YJ, et al. Impact of country-specific EQ-5D-3L tariffs on the economic value of systemic therapies used in the treatment of metastatic pancreatic cancer. *Curr Oncol* 2015;22:e443–452.
- Fang M, Oremus M, Tarride JE, et al. A comparison of health utility scores calculated using United Kingdom and Canadian preference weights in persons with alzheimer's disease and their caregivers. *Health Qual Life Outcomes* 2016;14:105.
- Dyer MT, Goldsmith KA, Sharples LS, et al. A review of health utilities using the EQ-5D in studies of cardiovascular disease. *Health Qual Life Outcomes* 2010;8:13.
- Korpi T. Is utility related to employment status? Employment, unemployment, labor market policies and subjective well-being among Swedish youth. *Labour Econ* 1997;4:125–47.
- Develeeschauwer B, Haagsma JA, Angulo FJ, et al. Methodological framework for world health organization estimates of the global burden of foodborne disease. *PLoS One* 2015;10:e0142498.
- Develeeschauwer B, Havelaar AH, Maertens de Noordhout C, et al. DALY calculation in practice: a stepwise approach. *Int J Public Health* 2014;59: 571–574.
- Mangen MJ, Plass D, Havelaar AH, et al. The pathogen- and incidence-based DALY approach: an appropriate [corrected] methodology for estimating the burden of infectious diseases. *PLoS One* 2013;8:e79740.
- Eurostat: Revenu moyen et médian par âge et sexe. 2015. Available at: <http://appsso.eurostat.ec.europa.eu/nui/show.do>.
- Enea M, RM, TK. *Fitting Linear and Generalized Linear Models to Large Data Sets*. Available at: <https://cran.r-project.org/web/packages/speedglm/speedglm.pdf>. 2015.
- Eurostat: Smarter, greener, more inclusive? *Indicators to Support the Europe 2020 Strategy*. Available at: <http://ec.europa.eu/eurostat/documents/3217494/6655013/KS-EZ-14-001-EN-N.pdf/a5452f6e-8190-4f30-8996-41b1306f7367> ed. pp. 200; 2015:200.
- Haagsma JA, van Beeck EF, Polinder S, et al. Novel empirical disability weights to assess the burden of non-fatal injury. *Inj Prev* 2008;14:5–10.
- Marengoni A, Angleman S, Melis R, et al. Aging with multimorbidity: a systematic review of the literature. *Ageing Res Rev* 2011;10:430–9.
- Mazzawi T, Hausken T, Gundersen D, et al. Effects of dietary guidance on the symptoms, quality of life and habitual dietary intake of patients with irritable bowel syndrome. *Mol Med Rep* 2013;8:845–52.
- Nord E. Disability weights in the Global Burden of Disease 2010: unclear meaning and overstatement of international agreement. *Health Policy* 2013;111:99–104.