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# Sick Populations and Sick Subpopulations: Reducing Disparities in Cardiovascular Disease Between Blacks and Whites in the United States 

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#### Abstract

Background-Cardiovascular disease (CVD) death rates are much higher in blacks than whites in the United States (US). It is unclear how CVD risk and events are distributed among blacks vs. whites and how interventions reduce racial disparities.

Methods-We developed risk models for fatal and for fatal-and-nonfatal CVD using 8 cohorts in the US. We used 6,154 adults aged 50-69 years in the National Health and Nutrition Examination Survey 1999-2012 to estimate the distributions of risk and events in blacks and whites. We estimated the total as well as disparity impacts of a range of population-wide, targeted and riskbased interventions on 10 -year CVD risks and event rates.

Results- $25 \%$ ( $95 \%$ confidence interval 22-28) of black men and $12 \%$ (10-14) of black women were at $\geq 6.67 \%$ risk of fatal CVD (almost equivalent to $20 \%$ risk of fatal or nonfatal CVD), compared with $10 \%(8-12)$ of white men and $3 \%(2-4)$ of white women. These high-risk individuals accounted for $55 \%$ (49-59) of CVD deaths among black men and $42 \%$ (35-46) in black women, compared with $30 \%$ (24-35) in white men and $18 \%$ (13-22) in white women. We


[^0]estimated that an intervention that treated multiple risk factors in high-risk individuals could reduce black-white difference in CVD death rate from 1,659 to 1,244 per 100,000 in men and from 1,320 to 897 in women. Rates of fatal-and-nonfatal CVD were generally similar between black and white men. In women, a larger proportion of women were at $\geq 7.5 \%$ risk of CVD ( $30 \%$ versus $19 \%$ in whites) and an intervention that targeted multiple risk factors among this group was estimated to reduce black-white differences in CVD rates from 1,688 to 1,197 per 100,000.

Conclusions-A substantially larger proportion of blacks have a high risk of fatal CVD and bear a large share of CVD deaths. A risk-based intervention that reduces multiple risk factors could substantially reduce overall CVD rates and racial disparities in CVD death rates.

## Keywords

coronary heart disease risk; disparities; prevention; risk factor

## Introduction

Cardiovascular diseases (CVD) are the leading causes of death in the United States (US), with substantially higher death rates among blacks than whites. ${ }^{1,2}$ Previous research has shown that up to three quarters of absolute disparities between blacks and whites in CVD mortality may be due to differences in classic risk factors (i.e. raised blood pressure and serum cholesterol, diabetes, obesity, and smoking). ${ }^{3,4}$ Therefore, interventions that reduce these risk factors are expected to reduce disparities in CVD mortality between blacks and whites but it is not clear which types of interventions, population-wide or targeted, can reduce racial disparities. Population-wide interventions can have large impacts on overall disease burden, ${ }^{5}$ but their impact on disparities depend on how they change risk factors in different subgroups of the population. For example, health education may reduce or widen disparities depending on how it is delivered. ${ }^{6-8}$ The disparity impact of interventions that target high-risk individuals (identified using a single risk factor or a combination of risk factors) will depend on whether the worse-off group has more or less high-risk individuals. Therefore, it is essential to have information on not only the average CVD risk and events, but also how CVD risk and events are distributed in better-off and worse-off subgroups of the population.

Some studies have qualitatively or quantitatively assessed the impacts of current risk factor exposures or scenarios of reducing risk factors on disparities in CVD or total mortality. ${ }^{3,4,9-14}$ Most of these studies have considered hypothetical risk factor reductions as opposed to interventions that could be implemented in practice. Other studies have used inconsistent or incomparable data and methods for calculating mortality effects across different risk factors, therefore reducing comparability. Furthermore, no study has assessed the disparity impact of risk-based prevention that is recommended by recent clinical guidelines, ${ }^{15,16}$ because information on distributions of absolute CVD risk by race was not available. In this paper, we analyzed the total as well as disparity impacts of a range of population-wide, targeted and risk-based interventions on 10-year CVD risks and rates using consistent methods and data. We hypothesized that a much larger proportion of blacks are at high risk of CVD than whites, and hence the disparity in high-risk subgroup is responsible for a large part of disparity in event rates between races.

## Methods

## Overview

We estimated the effects of three types of interventions on CVD risk and events, as well as their disparities between blacks and whites: (1) population-wide interventions (alone or in combination); (2) interventions to lower risk factor level among individuals with high levels for a single risk factor; and (3) a risk-based intervention that targeted individuals with high predicted 10-year CVD risk and treated several risk factors simultaneously (Table 1). We first estimated the 10-year risk and events of both fatal and fatal-and-nonfatal coronary heart disease (CHD) or stroke for a representative sample of blacks and whites in the US. Risks were predicted based on systolic blood pressure (SBP), serum total cholesterol (TC), diabetes and smoking, using risk prediction equations that were recalibrated for each age-sex-race group. ${ }^{29}$ We then assessed how each intervention changed the predicted risk as well as events for each age-sex-race group.

## Data on Risk Factors

We used data on risk factors from 7 rounds of the National Health and Nutrition Examination Survey (NHANES) 1999-2012 to have stable estimates for each age-sex-race subgroup. We included black or white participants who were 50 to 69 years old and did not have a history of CHD or stroke. We excluded participants older than 70 years of age to focus on the age range commonly considered for premature event and mortality.

We accounted for complex survey design to make estimates of risk factor, predicted risk, and events representative of the national population. We used TC as opposed to LDL-cholesterol because LDL-cholesterol was only measured in half of the participants. Diabetes was defined as having a fasting plasma glucose (FPG) $\geq 126 \mathrm{mg} / \mathrm{dL}$, hemoglobin A1c (HbA1c) $\geq$ $6.5 \%$, history of diagnosis by a health professional, or use of insulin or oral hypoglycemic agents.

## Data on CVD Deaths

In our primary analysis, we used fatal CVD as the primary outcome because data on nonfatal events, which is required for risk equation recalibration, is not available for the national US population (see below on methods to estimate fatal-and-nonfatal rates by race). We used mortality data from the National Center for Health Statistics (NCHS), to calculate death rates in 2011. We defined CVD death as death from CHD (ICD10 codes I20-I25) or stroke (ICD10 codes I60-I69).

## Effect Sizes for Interventions

We obtained the effects of interventions on risk factors from meta-analyses of randomized controlled trials, observational studies, or policy evaluation analyses, as detailed in Table 1. We used a larger effect size for the impact of salt reduction on blood pressure among blacks versus whites based, ${ }^{18}$ but used the same effect size between blacks and whites for all other interventions because proportional effects have been found to be generally similar by race. ${ }^{30-33}$ Under the risk-based intervention scenario, we used individuals' absolute CVD risks to determine whether they were affected by the interventions and assigned
interventions (e.g. antihypertensive and statins) only to individuals who were not already
receiving them. We applied smoking cessation to smokers irrespective of their absolute CVD risks. We note that the level of evidence supporting interventions varies: for example, the impact of population-wide interventions has only been estimated in observational studies, ${ }^{18,21,24}$ whereas the effect of statins on CVD has been consistently shown in many randomized trials. ${ }^{34}$ We also note that an individual may receive both population-wide and targeted interventions in practice, although these two types of interventions were analyzed separately here.

## Statistical Analysis

We used risk prediction equations (or risk scores) for fatal CVD and for total CVD developed from 8 prospective cohorts in the US, as described elsewhere. ${ }^{29}$ Briefly, the models use four inputs to estimate individual-level 10-year risk: (1) the participants' risk factor levels; (2) coefficients (i.e. log hazard ratios) for each risk factor estimated from the cohorts; (3) mean risk factor level for the same age-sex-race subgroup as the index participant; (4) average CVD event rate for the same age-sex-race subgroup as the index participant. The risk factors in the model were SBP, TC, diabetes and smoking. We used this new risk predication equation because it is based on data from multiple cohorts; it allows a straightforward recalibration by sex and race; it allows the age pattern of CVD risk to vary across race-sex subgroups; and it includes interactions between age and SBP, TC, diabetes and smoking and an interaction between sex and diabetes to account for the fact that the proportional effects of these risk factors on CVD vary by age and sex. ${ }^{35-39}$ For this application, we modified the risk scores to separate current from former smokers. The coefficients of the risk scores and the validation methods and results are presented in onlineonly Data Supplemental Table 1 and Supplemental Table 2. We assumed the same proportional associations between risk factors and fatal CVD risk for blacks and whites based on previous evidence. ${ }^{30-33}$ We relaxed this assumption by using race-specific coefficients in a sensitivity analysis (online-only Data Supplemental Table 3).

In the primary analysis, we first recalibrated the risk score by replacing the CVD event rate and mean risk factor levels with the observed age-sex-race-specific rates in the US population. We then used the recalibrated risk score and individual-level data from NHANES to estimate the 10-year risk of fatal CVD for each participant under the current risk factor levels. We report the mean predicted risk and number of events, as well as their relative or absolute differences between blacks and whites. We also present how the population and events were distributed by risk level in each sex-race subgroup. We further report the proportions of population and events at fatal CVD risk $\geq 2.5 \%$, hereafter referred to as 'moderate-risk' and $\geq 6.67 \%$, hereafter referred to as 'high-risk'. As almost one-third of CVD events are fatal in the US, ${ }^{40}$ these risk thresholds approximately correspond to $\geq$ $7.5 \%$ (the AHA/ACC threshold ${ }^{15}$ ) and $\geq 20 \%$ (the ATP-III threshold ${ }^{16}$ ) for fatal and nonfatal CVD. In our secondary analysis, we used fatal-and-nonfatal CVD events as outcome. We calculated the age-sex-race-specific event rate of fatal-and-nonfatal CVD (CHD and stroke) using the corresponding death rate multiplying by the race-specific total-to-fatal event ratio. We used the total-to-fatal event ratios for CHD and stroke as reported in the Reasons for Geographic And Racial Differences in Stroke (REGARDS) cohort ${ }^{40,} 41$ to
account for the higher case fatality rates in blacks. We reported the proportions of population and events at total CVD risk $\geq 7.5 \%$ and $\geq 20 \%$. In a sensitivity analysis, we calculated the 10-year CVD risk using the American College of Cardiology/American Heart Association (ACC/AHA) 2013 Pooled Cohort Equations. ${ }^{42}$

To estimate the effects of interventions on CVD risk and events, we first estimated their effects on risk factor(s), and then re-calculated the 10 -year risk and events using the postintervention risk factor levels. We chose this approach instead of directly applying the impact of interventions on CVD risk because for many of the interventions analyzed here, the outcome of epidemiological studies is risk factor level. For example, we estimated the impact of reducing incidence of diabetes from the Diabetes Prevention Program and combined that with evidence on the effect of diabetes on CVD from meta-analyses of observational studies. ${ }^{28,35}$ As there are no studies that show a direct impact of diabetes prevention on CVD mortality, we conducted a separate analysis by removing diabetes prevention from the risk-based multiple risk factor intervention.

We quantified uncertainty by sampling repeated draws of different inputs to analysis, as described in the online-only Data Supplemental Text. All analyses were conducted using Stata 12.0 (StataCorp, College Station, Texas) and R 3.02. The study was approved by the institutional review board of the Harvard School of Public Health (Boston, MA, USA).

## Results

We included 6,154 blacks and whites from 7 rounds of NHANES (online-only Data Supplemental Figure 1). About one-third of participants were black. TC levels were similar between blacks and whites, whereas other risk factor levels were higher in blacks (onlineonly Data Supplemental Table 4).

Mean 10-year risk of fatal CVD was $5.1 \%$ in black men versus $3.4 \%$ in white men (risk ratio (RR) of 1.49 ), and $3.0 \%$ in black women versus $1.7 \%$ in white women (RR of 1.79). This was equivalent to a 10-year CVD death rate of 5,052 per 100,000 in black men versus 3,393 in white men (rate difference (SD) of 1,659 per 100,000, and 2,989 in black women versus 1,669 in white women (RD of 1,320 ).

The distributions of both population and events by 10-year risk of fatal CVD were shifted to the right among blacks compared with whites; the distribution of events had a heavier tail than that of population as most of the events arise from the high-risk individuals (Figure 1). As a result, $25 \%$ ( $95 \%$ confidence interval $22-28$ ) of black men were at high-risk ( $\geq 6.67 \%$ risk of fatal CVD in 10 year) compared with only $10 \%$ (8-12) of white men (Table 2 and online-only Data Supplemental Figure 2). This high-risk subgroup accounted for 55\% (4959) of CVD deaths in black men compared with $30 \%$ (24-35) in white men. For women, $12 \%(10-14)$ of blacks versus $3 \%(2-4)$ of whites at high-risk accounted for $42 \%(35-46)$ of CVD deaths in blacks versus $18 \%$ (13-22) in whites. Compared with the results of fatal CVD, black-white disparities in total CVD were substantially smaller for men (Figure 2, Table 3 and online-only Data Supplemental Figure 3). In women, disparities were only noticeable for those with $\geq 7.5 \%$ CVD risk where $30 \%$ (27-33) of blacks versus $19 \%$ (18-
21) of white women accounted for $61 \%$ (57-63) of CVD events in blacks versus $46 \%$ (44-
49) in whites.

Population-wide interventions (i.e. salt reduction, improving diet, WHO EMPOWER tobacco control policies, increasing price of sugar-sweetened beverages) and targeted interventions on single risk factors (i.e. antihypertensive and statins treatment, referral for quitting smoking, diabetes prevention program) were estimated to reduce the 10-year CVD death rate by at most 440 per 100,000 in men and 290 in women. The risk-based multiple risk factor intervention was estimated to reduce the average CVD death rate by 1,086 per 100,000 in men and 669 in women for blacks, and 671 per 100,000 in men and 246 in women for whites (Table 4).

Population-wide interventions and targeted interventions on single risk factors did not substantially reduce the proportion of population at high risk of fatal CVD ( $\geq 6.67 \%$ risk of fatal CVD in 10 years) in either whites or blacks (Figure 3). In contrast, the risk-based multiple risk factor intervention was estimated to reduce the proportion of high-risk population by at most 12 percentage points for men and 6 percentage points for women. Results were similar for moderate-risk group ( $\geq 2.5 \%$ risk of fatal CVD in 10 years), where the risk-based multiple risk factor intervention was estimated to reduce the moderate-or-high-risk proportion by at most 13 percentage points for men and 9 percentage points for women compared with at most 6 and 4 percentage points in population-wide or targeted interventions (Figure 4). Our sensitivity analysis using separate fatal CVD risk scores for blacks and whites showed similar results (online-only Data Supplemental Figure 4 and Supplemental Figure 5).

None of the interventions analyzed here had a potential to reduce black-vs-white fatal CVD rate ratios (Table 4). When we considered disparities in absolute CVD rates, combining the four selected population-wide interventions was estimated to reduce black-white disparities by 198 per 100,000 ( $12 \%$ of total absolute disparity) in men and 141 ( $11 \%$ ) in women.

Among targeted single-risk interventions, the diabetes prevention program had the largest potential, with an estimated reduction in absolute disparity by 142 (9\%) in men and 173 $(13 \%)$ in women. The risk-based multiple risk factor intervention had much larger potential and could reduce absolute disparities by 415 per $100,000(25 \%)$ in men and $423(32 \%)$ in women. Removing diabetes prevention from the risk-based multiple risk factor intervention reduced the estimated impact of risk-based intervention by $41 \%$ to $50 \%$ but this intervention still had the largest potential for reducing absolute black-white disparities.

For fatal-and-nonfatal CVD rates, there were no significant disparities between blacks and whites in men. The estimated black-white disparities in women could be reduced by 217 per 100,000 ( $13 \%$ of total disparity in absolute risk) through a combination of four populationwide interventions. Implementing a diabetes prevention program alone was estimated to reduce disparities by 412 ( $24 \%$ ) and the risk-based multiple risk factor intervention by 491 (29\%) (Table 5). The sensitivity analysis using ACC/AHA 2013 Pooled Cohort Equations also showed consistent with the main analysis (online-only Data Supplemental Table 5 and Supplemental Table 6).

## Discussion

We found that a substantially larger proportion of blacks ( $25 \%$ of men and $12 \%$ of women) in the US had a high risk of fatal CVD than their white counterparts ( $10 \%$ of men and $3 \%$ of women). These high-risk individuals bore about half of the burden of fatal CVD events in the population. An intervention that could identify high-risk individuals and treat multiple risk factors could both deliver large total benefits and substantially reduce the absolute black-white disparities. Population-wide and targeted interventions on single risk factors had smaller potential on reducing racial disparities in CVD compared with a risk-based intervention on multiple risk factors. Total CVD risks were similar in black versus white men and the disparity between black and white women could be substantially reduced by a risk-based multiple risk intervention.

Our results on the disparities in risk factor exposure, and in their role as a cause of racial disparities in CVD, are consistent with those of previous analyses. ${ }^{1,3,4,9,11,13}$ A previous study proposed that population-wide interventions have a larger effect on health disparities than interventions that target high-risk individuals, but the two scenarios were only qualitatively compared. ${ }^{10}$ Other studies quantified the effects of hypothetical risk factor reductions on disparity in mortality without considering specific interventions. ${ }^{3,4,9,11,13}$ In addition, previous studies often used a single risk factor to identify high-risk individuals and considered interventions on one risk factor at a time. ${ }^{11,43}$

A key strength of our analysis is that we have assessed not only the aggregate risk and events within each group but also how risk and events were distributed, providing important information on who needs intervention and what the expected impact of intervention is. In addition, we compared the total and disparity impacts of a wide range of population-wide and targeted interventions using consistent methods and data. Risk factor distributions were from a nationally representative survey, mortality data were from vital registration system, and effect sizes for interventions were obtained from large meta-analyses of randomized trials or observational studies that had adjusted for important confounders. We also systematically quantified the uncertainty as a result of the sampling variability in the national surveys and the uncertainty of coefficients from the risk prediction equations. Finally, our primary model included age interaction between risk factors and CVD incorporating evidence from many prospective studies. ${ }^{35}$

Our study has some limitations. First, although we estimated the risk distributions for both fatal CVD and total CVD, reliable national data on total CVD incidence is not available for model recalibration, especially by race. Recent data from a large prospective cohort (REGARDS) shows that black men have higher incidence of fatal CHD and lower incidence of non-fatal CHD than white men, resulting in similar incidence of total CHD for black and white men. ${ }^{40}$ Using estimates of case fatality rates from REGARDS to recalibrate model for total CVD risk eliminated much of the racial disparity in total CVD risk and thus it is expected that the interventions evaluated here would have minimal impact on racial disparities in total CVD risk. Second, our analysis focused on primary prevention of CVD. However, patients with history of CVD have a high risk of subsequent cardiovascular events and should receive treatments for risk factors. In the US, $9 \%$ of blacks and $6 \%$ of whites
aged 50-69 in the 2011-2012 NHANES survey had a history of CVD. Were these
proportions to be added to our estimates of prevalence of high-risk status, disparities would be slightly larger than our estimates. Our analysis did not include patients with CVD because existing risk scores for these patients require data on predictors such as electrocardiography (ECG) results, coronary imaging and biomarkers that are not measured in NHANES. ${ }^{44,}{ }^{45}$ Third, we assumed that compliance with interventions would be similar to those observed in the randomized trials and observational studies used to generate the intervention effects, which may lead to overestimating the impact of interventions on blackwhite disparities. Although compliance may vary by race, prior work suggested that noncompliance is likely due to barriers of access to and poor quality of healthcare. ${ }^{46,47}$ If insurance coverage and healthcare quality were similar across races, it is unlikely that compliance would differ substantially, as has been observed for antiretroviral therapy. ${ }^{48}$ Fourth, smoking cessation interventions have been shown to affect disadvantaged populations more strongly. However, detailed data on the differential impacts of smoking cessation by race is not available. Therefore, our estimates for the impact of smoking cessation on black-white disparities in CVD risk should be considered conservative. Fifth, there is limited evidence on direct impact of diabetes prevention on CVD and it remains unclear whether the Diabetes Prevention Program prevents or delays the onset of diabetes. Our sensitivity analyses of removing diabetes prevention from the risk-based multiple risk factor intervention confirmed the largest impact on reducing the absolute back-white disparities still came from the risk-based intervention. Finally, the effects of some interventions (e.g. reducing salt in package food, WHO's MPOWER tobacco control policies) may be cumulative over decades. Our analyses did not incorporate the cumulative effects and hence may underestimate the long-term effect of these interventions.

In conclusion, although prevention and treatment have helped reduce CVD rates over the past few decades in the US, mortality rates remain higher in blacks than whites. ${ }^{1,2,40}$ Eliminating racial disparities in health is one of the overarching goals of the Healthy People 2020 agenda. ${ }^{49}$ As disparities in CVD are caused by disparities in broader social, economic and environmental determinants, policies and strategies are needed to address these factors and to facilitate healthy life-style and environment. Meanwhile, our findings suggest a much larger proportion of blacks are at high risk of fatal CVD than whites, and this high-risk subpopulation is bearing almost half of the deaths in the population. Therefore, by targeting this sick subpopulation with combination risk-based therapy, we can reduce a large share of events. While such approach has been advocated for the US population as a whole, ${ }^{15}$ achieving its potential as a means to reduce racial disparities will require increasing health insurance coverage and a strong primary care system that is equipped with well-trained health workers and appropriate infrastructure to provide low-cost essential drugs. Previous research has shown that universal health insurance over age 65 in the US is associated with lower racial differences in cardiovascular risk factors. ${ }^{50}$ An accessible and high-quality primary care program has also successfully reduced cardiovascular health inequality in other countries. ${ }^{51}$ The window of opportunity for addressing cardiovascular health disparity lies in the Affordable Care Act of $2010^{52}$ that has already shown promise in improving access to primary care services ${ }^{53}$ and commits to eliminate barriers to health for disadvantaged communities, along with the new guideline for risk-based multidrug treatment for CVD. ${ }^{15}$

Their intersection could help identify important opportunities to improve the access and affordability of risk-based treatment for CVD in underserved population, and finally improve cardiovascular health for all.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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## Clinical Perspective

1) What is new?

- We investigated how risk of fatal and fatal-plus-nonfatal cardiovascular disease (CVD), estimated using a risk prediction model, is distributed among whites and blacks in the US and how population-wide or targeted interventions on CVD risk factors would reduce these racial disparities.
- We used a nationally representative sample of adults aged 50 to 69 years in the US and a CVD risk prediction model that was recalibrated separately for blacks and whites.

2) What are the clinical implications?

- Our results indicated that there are substantial disparities in risk of fatal CVD.
- A large proportion of fatal CVD events among blacks were concentrated among a small proportion of the population; in contrast, racial disparities in risk of fatal-and-nonfatal CVD were only noticeable among women.
- Population-wide and targeted interventions on single risk factors did not reduce black-white disparities in fatal CVD risk substantially.
- An intervention that focused on high-risk individuals and reduced multiple risk factors simultaneously could reduce black-white disparities in fatal CVD risk by a quarter in men and a third in women.
- Focusing preventive interventions on the high-risk individuals has a large potential to improve overall CVD health and reduce racial disparities.


Figure 1.
Distributions of predicted 10-year risk of fatal CVD in the population (A) and among cases (B).


Figure 2.
Distributions of predicted 10-year risk of fatal-and-nonfatal CVD in the population (A) and among cases (B).


Figure 3.
Impact of population-wide, single raised risk factor, and risk-based interventions on proportion of population with $\geq 6.67 \%^{*} 10$-year risk of fatal CVD. * This threshold approximately equals to $\geq 20 \%$ for risk of fatal-and-nonfatal CVD given one third of CVD events are fatal in the US. ${ }^{40}$


Figure 4.
Impact of population-wide, single raised risk factor, and risk-based interventions on proportion of population with $\geq 2.5 \%^{*} 10$-year risk of fatal CVD. * This threshold approximately equals to $\geq 7.5 \%$ for fatal-and-nonfatal CVD given one third of CVD events are fatal in the US. ${ }^{40}$
Table 1
Selected risk factors, their exposure metrics, and examples of population-wide, single raised risk factor, and risk-based interventions.

| Risk factors | Exposure metric (unit) | Population-wide interventions | Single raised risk factor interventions | Multiple populationlevel interventions | Risk-based interventions |
| :---: | :---: | :---: | :---: | :---: | :---: |
| High blood pressure | Systolic blood pressure (SBP, mmHg ) | Reducing salt intake in packaged and prepared food | Two antihypertensive drugs at standard dose if diabetic, or SBP $\geq$ 140 mmHg for non-diabetic adults aged $<60$, or SBP $\geq 150 \mathrm{mmHg}$ for non-diabetic adults aged $\geq 60^{\prime \prime}$ | Multiple risk factor intervention at the population level, including reducing salt intake, dietary improvement, tobacco control, and increasing the price of sugar-sweetened beverages | Multiple risk factor intervention, including blood pressure and lipid lowering medications, smoking cessation and life-style modification if 10year risk of fatal CVD $\geq 2.5 \%$ (or total CVD risk $\geq$ $7.5 \%$ ) |
| High serum cholesterol | Serum total cholesterol (TC, $\mathrm{mmol} / \mathrm{L}$ ) | Community-based dietary improvement to reduce dairy fat and replace saturated with unsaturated fats, and increase vegetable and fruit consumption ${ }^{\text {t }}$ | High-intensity statin if 10 -year ASCVD risk $27.5 \%$, or LDL cholesterol $\geq 4.9 \mathrm{mmol} / \mathrm{L}$, moderate-intensity statin if diabetic aged 40-75 and 10-year ASCVD risk $<7.5 \%$ \# |  |  |
| Tobacco smoking | Current smoker prevalence (percentage) | Tobacco control package to ban smoking in indoor workplace, offer cessation treatment in general store, put warning on package, ban advertisements, and increase tobacco tax ${ }^{t}$ | Referral to smoking cessation intervention such as group behavioral therapy ** |  |  |
| Diabetes | Diabetes prevalence (percentage) | Increase in price of sugarsweetened beverages $\mathcal{\xi}$ | Intensive lifestyle modification intervention if diabetic ${ }^{\text {t/ }}$ |  |  |

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Table 2
Proportion of population and proportion of fatal CVD events occurring among high-risk individuals by sex and race.

|  | Fatal CVD risk $\geq \mathbf{2 . 5 \%} \%$ |  | Fatal CVD risk $\geq 6.67 \% *$ |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Proportion of <br> population (\%) | Proportion of <br> event (\%) | Proportion of <br> population (\%) | Proportion of <br> event (\%) |
|  |  |  |  |  |
| White | $50(47-52)$ | $77(75-78)$ | $10(8-12)$ | $30(24-35)$ |
| Black | $66(62-69)$ | $88(87-90)$ | $25(22-28)$ | $55(49-59)$ |
| Women |  |  |  |  |
| White | $17(15-18)$ | $49(46-52)$ | $3(2-4)$ | $18(13-22)$ |
| Black | $36(33-39)$ | $74(72-76)$ | $12(10-14)$ | $42(35-46)$ |

These risk thresholds are approximately equal to $\geq 7.5 \%$ (the AHA/ACC threshold ${ }^{15}$ ) and $\geq 20 \%$ (the ATP-III threshold ${ }^{16}$ ) for fatal-and-nonfatal CVD as almost one third of CVD events are fatal in the US. ${ }^{40}$

Table 3
Proportion of population and proportion of fatal-and-nonfatal CVD events occurring among high-risk individuals by sex and race.

|  | Fatal-and-nonfatal CVD <br> risk $\mathbf{7 . 5 \%}$ |  | Fatal-and-nonfatal CVD <br> risk $\geq 20 \%$ |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Proportion of <br> population (\%) | Proportion of <br> event (\%) | Proportion of <br> population (\%) | Proportion of <br> event (\%) |
|  |  |  |  |  |
| White | $68(66-71)$ | $86(85-87)$ | $15(13-16)$ | $32(29-34)$ |
| Black | $70(66-73)$ | $87(86-89)$ | $18(15-20)$ | $37(32-40)$ |
| Women |  |  |  |  |
| White | $19(18-21)$ | $46(44-49)$ | $2(1-3)$ | $10(6-13)$ |
| Black | $30(27-33)$ | $61(57-63)$ | $4(3-5)$ | $16(12-19)$ |

Impact of interventions on 10-year rate of fatal CVD (per 100,000) by sex and race

|  |  |  | Men |  | Absolute change in rate difference | Relative change in rate difference |  |  | Women |  | Absolute change in rate difference | Relative change in rate difference |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & \text { 10-year CVD } \\ & \text { death rate } \end{aligned}$ |  | Rate ratio (black vs. | Rate difference |  |  | 10-year CVD death rate |  | Rate ratio (black vs. | Rate difference |  |  |
|  | Black | White | - | - |  |  | Black | White |  |  |  |  |
| Current | 5,052 | 3,393 | 1.49 | 1,659 | NA | NA | 2,989 | 1,669 | 1.79 | 1,320 | NA | NA |
| Population-wide interventions |  |  |  |  |  |  |  |  |  |  |  |  |
| 1) Salt reduction | 4,939 | 3,329 | 1.48 | 1,610 | -49 | -3\% | 2,945 | 1,644 | 1.79 | 1,301 | -19 | -1\% |
| 2) Community-based dietary improvement to reduce serum cholesterol | 4,906 | 3,290 | 1.49 | 1,616 | -44 | -3\% | 2,893 | 1,618 | 1.79 | 1,276 | -45 | -3\% |
| 3) MPOWER package for smoking | 4,939 | 3,356 | 1.47 | 1,583 | -77 | -5\% | 2,934 | 1,647 | 1.78 | 1,287 | -34 | -3\% |
| 4) Increasing price of sugar-sweetened beverages for diabetes prevention | 4,972 | 3,362 | 1.48 | 1,609 | -50 | -3\% | 2,931 | 1,651 | 1.77 | 1,280 | -41 | -3\% |
| 5) Multiple interventions (1-4) | 4,612 | 3,151 | 1.46 | 1,461 | -198 | -12\% | 2,729 | 1,549 | 1.76 | 1,179 | -141 | -11\% |
| Single raised risk factor interventions |  |  |  |  |  |  |  |  |  |  |  |  |
| 6) Treatment for hypertension | 4,885 | 3,290 | 1.48 | 1,595 | -64 | -4\% | 2,906 | 1,629 | 1.78 | 1,277 | -43 | -3\% |
| 7) Treatment for dyslipidaemia | 4,687 | 3,143 | 1.49 | 1,544 | -116 | -7\% | 2,731 | 1,554 | 1.76 | 1,177 | -144 | -11\% |
| 8) Referral for quitting smoking | 4,875 | 3,342 | 1.46 | 1,533 | -126 | -8\% | 2,926 | 1,643 | 1.78 | 1,283 | -38 | -3\% |
| 9) Diabetes prevention program | 4,704 | 3,187 | 1.48 | 1,517 | -142 | -9\% | 2,699 | 1,551 | 1.74 | 1,148 | -173 | -13\% |
| Risk-based interventions |  |  |  |  |  |  |  |  |  |  |  |  |
| 10) Multiple risk factors (6-9 if risk of CVD death $22.5 \%$ ) * | 3,966 | 2,722 | 1.46 | 1,244 | -415 | -25\% | 2,320 | 1,423 | 1.63 | 897 | -423 | -32\% |
| 11) Multiple risk factors ( 10 without diabetes prevention) | 4,296 | 2,882 | 1.49 | 1,414 | -246 | -15\% | 2,601 | 1,495 | 1.74 | 1,106 | -214 | -16\% |

*The eligibility for each intervention was defined based on risk of CVD death $\geq 2.5 \%$ except for smoking cessation. Smoking cessation was provided to all current smokers irrespective of his/her predictive risk.
Impact of interventions on 10-year rate of fatal-and-nonfatal CVD (per 100,000) by sex and race

|  |  |  | Men |  | Absolute change in rate difference | Relative change in rate difference |  |  | Women |  | Absolute change in rate difference | Relative change in rate difference |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 10-year CVD event rate |  | Rate ratio (black vs. white) | Rate difference |  |  | $\begin{aligned} & \text { 10-year CVD } \\ & \text { event rate } \end{aligned}$ |  | Rate ratio (black vs. white) | Rate difference |  |  |
|  | Black | White | - | - |  |  | Black | White |  |  |  |  |
| Current | 13,082 | 12,343 | 1.06 | 739 | NA | NA | 6,868 | 5,179 | 1.33 | 1,688 | NA | NA |
| Population-wide interventions |  |  |  |  |  |  |  |  |  |  |  |  |
| 1) Salt reduction | 12,881 | 12,189 | 1.06 | 691 | -48 | -6\% | 6,791 | 5,123 | 1.33 | 1,668 | -20 | -1\% |
| 2) Community-based dietary improvement to reduce serum cholesterol | 12,703 | 11,965 | 1.06 | 739 | -1 | 0\% | 6,645 | 5,011 | 1.33 | 1,634 | -55 | -3\% |
| 3) MPOWER package for smoking | 12,852 | 12,224 | 1.05 | 629 | -110 | -15\% | 6,748 | 5,144 | 1.31 | 1,604 | -84 | -5\% |
| 4) Increasing price of sugarsweetened beverages for diabetes prevention | 12,962 | 12,264 | 1.06 | 697 | -42 | -6\% | 6,736 | 5,151 | 1.31 | 1,585 | -104 | -6\% |
| 5) Multiple interventions (1-4) | 12,204 | 11,613 | 1.05 | 590 | -149 | -20\% | 6,358 | 4,886 | 1.30 | 1,471 | -217 | -13\% |
| Single raised risk factor interventions |  |  |  |  |  |  |  |  |  |  |  |  |
| 6) Treatment for hypertension | 12,829 | 12,144 | 1.06 | 686 | -54 | -7\% | 6,739 | 5,106 | 1.32 | 1,633 | -56 | -3\% |
| 7) Treatment for dyslipidaemia | 12,147 | 11,347 | 1.07 | 800 | 61 | 8\% | 6,329 | 4,819 | 1.31 | 1,510 | -179 | -11\% |
| 8) Referral for quitting smoking | 12,816 | 12,190 | 1.05 | 626 | -113 | -15\% | 6,719 | 5,130 | 1.31 | 1,589 | -99 | -6\% |
| 9) Diabetes prevention program | 12,449 | 11,868 | 1.05 | 581 | -158 | -21\% | 6,290 | 5,013 | 1.25 | 1,277 | -412 | -24\% |
| Risk-based interventions |  |  |  |  |  |  |  |  |  |  |  |  |
| 10) Multiple risk factors (6-9 if risk of CVD death $22.5 \%$ ) * | 10,791 | 10,246 | 1.05 | 546 | -193 | -26\% | 5,740 | 4,542 | 1.26 | 1,197 | -491 | -29\% |
| 11) Multiple risk factors (10 without diabetes prevention) | 11,430 | 10,654 | 1.07 | 776 | 37 | 5\% | 6,135 | 4,683 | 1.31 | 1,452 | -236 | -14\% |

${ }^{*}$ The eligibility for each intervention was defined based on risk of total CVD $\geq 7.5 \%$ except for smoking cessation. Smoking cessation was provided to all current smokers irrespective of his/her predictive


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[^1]:    *We assumed a moderate salt reduction of about $1.4 \mathrm{~g} /$ day based on the United Kingdom's successful experience in reducing salt intake at the population level. 17 The effects of modest sodium reduction on blood pressure vary significantly by age, race and hypertensive status. We calculated the SBP reduction using the regression equation from a large meta-analysis of more than 100 randomized controlled trials. ${ }^{18}$
    ${ }^{*}$ The dietary improvement intervention is based on the success of the North Karelia Project in Finland, which includes substituting vegetable oil margarine for butter, non-fat or low-fat milk for fatty milk, lean meat for meat high in saturated fat, using vegetable oil for cooking, and increasing vegetable and fruit consumption. We assumed such an intervention would reduce the population mean TC by 0.5 $\mathrm{mmol} / \mathrm{L}$, the average decline over 5 years in Finland. ${ }^{19}$ Similar dietary intervention was also found to improve population mean TC level in New Zealand. ${ }^{20}$
    *The tobacco control package is based on the WHO MPOWER tobacco control policies, which include banning smoking in all indoor workplaces, providing NRT and bupropion in general store or pharmacy with prescription, putting bold and graphic warning to cover at least $50 \%$ of the package, banning all direct advertisements, and increasing $10 \%$ in the retail price of cigarette due to tax. We assumed a $11 \%$ reduction in smoking prevalence based on the previous policy evaluation reports. ${ }^{21}$
    $\xi_{\text {We modeled a }} 50 \%$ increase in price of sugar-sweetened beverages, which we estimated would reduce consumption by about $50 \%$ based on the experience in Mexico where raising prices by $10 \%$ decreased consumption by $12 \% .{ }^{22}$ Having consumption in the US would translate into a 0.5 serving/day reduction, as average sugar-sweetened beverage intake in adults is about one serving per day. 23

